



“On the Laplace Transform of the Aggregate Discounted Claims with Markovian Arrivals,” Jiandong Ren, April 2008

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In this paper Dr. Ren has obtained some very elegant matrix expressions for the Laplace transform and the first two moments of the aggregate discounted claims in a risk process in a Markovian environment (MAP). The purpose of this discussion is to show how to calculate the n th moment of the aggregate discounted claims recursively when the force of interest is constant $\delta > 0$, that is, $v(t) = e^{\delta t}$.

We slightly change the notation and use $D_0(i, j)$ and $D_1(i, j)$ to denote the (i, j) -th element of D_0 and D_1 , respectively. We extend the assumption in the paper under discussion by assuming that claim amounts that occur at a transition of the underlying Markov process from state i to state j depend on both i and j . We denote by $p_{i,j}$ and $\hat{p}_{i,j}(s) = \int_0^\infty e^{-sx} p_{i,j}(x) dx$ the density function and the Laplace transform, respectively, of claim amounts that occur at a transition.

First, we shall show that $L(\xi, t)$ satisfies the following first-order linear partial differential equation:

$$\delta \xi \frac{\partial L(\xi, t)}{\partial \xi} + \frac{\partial L(\xi, t)}{\partial t} = D_0 L(\xi, t) + R(\xi) L(\xi, t), \tag{D.1}$$

where $R(\xi) = (D_1(i, j) \hat{p}_{i,j}(\xi))_{i,j=1}^m$.

The idea for deriving this partial differential equation is similar to that used in proving the integro-differential equation for the Gerber-Shiu function in a MAP risk model (see Badescu 2008). For an infinitesimal $h > 0$, conditioning on three possible events that can occur in $[0, h]$ —no change in the MAP phase (state), a change in the MAP phase accompanied by no claims, and a change in the MAP phase and a claim instant—we have

$$\begin{aligned} L_{i,j}(\xi, t) &= [1 + D_0(i, i)h]L_{i,j}(\xi e^{-\delta h}, t - h) + \sum_{k=1, k \neq i}^m D_0(i, k)hL_{k,j}(\xi e^{-\delta h}, t - h) \\ &+ \sum_{k=1}^m D_1(i, k)h\hat{p}_{i,k}(\xi e^{-\delta h})L_{k,j}(\xi e^{-\delta h}, t - h). \end{aligned} \tag{D.2}$$

Taylor’s expansion gives

$$L_{i,j}(\xi e^{-\delta h}, t - h) = L_{i,j}(\xi, t) - \delta \xi h \frac{\partial L_{i,j}(\xi, t)}{\partial \xi} - h \frac{\partial L_{i,j}(\xi, t)}{\partial t} + o(h), \tag{D.3}$$

where $\lim_{h \rightarrow 0} (o(h)/h) = 0$. Substituting (D.3) into (D.2), dividing both sides by h , and letting $h \rightarrow 0$, we have

$$\begin{aligned} \delta \xi \frac{\partial L_{i,j}(\xi, t)}{\partial \xi} + \frac{\partial L_{i,j}(\xi, t)}{\partial t} &= \sum_{k=1}^m D_0(i, k)L_{k,j}(\xi, t) \\ &+ \sum_{k=1}^m D_1(i, k)\hat{p}_{i,k}(\xi)L_{k,j}(\xi, t). \end{aligned} \tag{D.4}$$

Rewriting (D.4) in matrix form gives (D.1).

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Let $M_n(t)$ denote the matrix with the (i, j) element being

$$M_{ij}^{(n)}(t) = \mathbb{E}_i[S^n(t)I(J(t) = j)].$$

$M_{ij}^{(n)}(t)$ is the n th moment of the discounted aggregate claims by time t , and the phase at t is j given that the initial phase is i . Using Taylor's series expansion, we have

$$L(\xi, t) = \sum_{n=0}^{\infty} \frac{(-1)^n \xi^n}{n!} M_n(t). \tag{D.5}$$

Furthermore,

$$\hat{p}_{ij}(\xi) = \sum_{n=0}^{\infty} \frac{(-1)^n \xi^n}{n!} \mu_{ij}^{(n)},$$

where $\mu_{ij}^{(n)}$ is the n th moment of p_{ij} , and hence,

$$D_1(i, j)\hat{p}_{ij}(\xi) = \sum_{n=0}^{\infty} \frac{(-1)^n \xi^n}{n!} D_1(i, j)\mu_{ij}^{(n)}.$$

In matrix notation,

$$R(\xi) = \sum_{n=0}^{\infty} \frac{(-1)^n \xi^n}{n!} Q_n, \tag{D.6}$$

where $Q_n = (D_1(i, j)\mu_{ij}^{(n)})_{ij=1}^m$ with $Q_0 = D_1$. Substituting (D.6) and (D.5) into (D.1), and equating the coefficients of ξ^n gives the following first-order linear matrix differential equation for $M_n(t)$:

$$M'_n(t) + (n\delta I - D)M_n(t) = \sum_{k=1}^n \binom{n}{k} Q_k M_{n-k}(t), \quad n \in \mathbb{N}^+, \tag{D.7}$$

with the initial condition $M_n(0) = \mathbf{0}$. In particular, $M_1(t)$ satisfies

$$M'_1(t) + (\delta I - D)M_1(t) = Q_1 e^{Dt}, \quad t \geq 0.$$

Solving the differential equation (D.7) with $M_n(0) = \mathbf{0}$, we obtain the following recursive formula for $M_n(t)$:

$$M_n(t) = \sum_{k=1}^n \binom{n}{k} \int_0^t e^{-n\delta x} e^{Dx} Q_k M_{n-k}(t-x) dx, \quad t \geq 0, n \in \mathbb{N}^+. \tag{D.8}$$

In particular,

$$M_1(t) = \int_0^t e^{-\delta x} e^{Dx} Q_1 e^{D(t-x)} dx, \quad t \geq 0. \tag{D.9}$$

Denote $\mathbf{m}_n(t)$ as an $m \times 1$ column vector with the i th entry being $m_i^{(n)}(t) = \mathbb{E}_i[S^n(t)]$; the n th moment of $S(t)$ given that the initial phase is i . Clearly $m_{\gamma, n} = \mathbb{E}_{\gamma}[S^n(t)] = \boldsymbol{\gamma} \mathbf{m}_n(t)$ and $\mathbf{m}_n(t) = M_n(t) \mathbf{e}^{\top}$. Then it follows from (D.8) and (D.9) that

$$\mathbf{m}_n(t) = \sum_{k=1}^n \binom{n}{k} \int_0^t e^{-n\delta x} e^{Dx} Q_k \mathbf{m}_{n-k}(t-x) dx, \quad n \in \mathbb{N}^+.$$

In particular,

$$\begin{aligned} \mathbf{m}_1(t) &= \int_0^t e^{-\delta x} e^{Dx} Q_1 e^{D(t-x)} \mathbf{e}^{\top} dx = \int_0^t e^{-\delta x} e^{Dx} dx Q_1 \mathbf{e}^{\top} \\ &= [\delta I - D]^{-1} [I - e^{-t(\delta I - D)}] Q_1 \mathbf{e}^{\top}. \end{aligned}$$

When $t \rightarrow \infty$, we have the following asymptotic results for the moments of the discounted aggregate claims:

$$\begin{aligned} \mathbf{m}_n(\infty) &= \int_0^\infty e^{-n\delta x} e^{Dx} dx \sum_{k=1}^n \binom{n}{k} \mathbf{Q}_k \mathbf{m}_{n-k}(\infty), \quad n \in \mathbb{N}^+ \\ &= [n\delta \mathbf{I} - \mathbf{D}]^{-1} \sum_{k=1}^n \binom{n}{k} \mathbf{Q}_k \mathbf{m}_{n-k}(\infty). \end{aligned}$$

In particular,

$$\mathbf{m}_1(\infty) = [\delta \mathbf{I} - \mathbf{D}]^{-1} \mathbf{Q}_1 \mathbf{e}^\top.$$

Finally we remark that the recursive formula for the n th moment of the discounted aggregate claims can be obtained as special cases for the renewal risk process with phase-type interclaim times, the semi-Markovian risk process studied by Albrecher and Boxma (2005), and the Markov-modulated risk process studied in Kim and Kim (2007) who give the first two moments of the discounted aggregate claims.

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