Adjustment Coefficient in the Sparre Anderson Model with Reinsurance

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Abstract:

Assumptions:

1. Reinsurance treaty: quota-share retention level a, retention limit M.

2. A diffusion process is added to the claim process.

Objective:

Choose the a and M to maximize the adjustment coefficient.

1. Reinsurance Forms

• Proportional (Pro Rata): premiums and losses are shared proportionately between the ceding company and the reinsurer.

Quota share: same percentage applies to all policies.

Surplus share: share varies for each risk.

• Non-proportional

Catastrophe: big losses involving more than one insured.

Stop loss: pays losses in excess of a predetermined amount.

Spread loss: ceding company's losses for a certain year are "spread" over a period of years.

2. Definitions

 $X_{a,M} = min(aX, M)$: the insurer retains payment when a claim of size X occurs.

 $\{T_i\}_{i=1}^{\infty}$: inter-claim time, i.i.d. non-negative r.v. exponential distribution with $E(T_i) = \frac{1}{\lambda}$

 $S_n = T_1 + T_2 + \cdots + T_n$: the claim process with $S_0 = 0$.

 $\{N(t)\}_{t\geq 0}$: the number of claims, an ordinary renewal process. $N(t) = \sup\{n : S_n \leq t\}.$

 $\{W(t)\}$: a diffusion process affects the surplus process, a Wiener process with infinitesimal drift 0 and infinitesimal variance 2D > 0. Independent of S(t) and $W(t) \sim N(0, 2Dt)$

 $\{X_i\}_{i=1}^{\infty}$: claim amount, i.i.d. non-negative r.v.s, independent of $\{T_i\}_{i=1}^{\infty}$. With distribution function F and $E(X_i) = \mu$. Satisfies :

- F(0) = 0; 0 < F(x) < 1 for $0 < x < +\infty;$
- $\frac{dF}{dx}$ exist and continuous;

• $M_X(r)$ (moment generating function) exists for $r \in (-\infty, \tau)$ for some $0 < \tau \leq +\infty$ and $\lim_{r \to \tau} M_X(r) = \lim_{r \to \tau} E[e^{rX}] = +\infty$.

P: the insurer's gross premium income per unit of time.

- $\alpha :$ positive loading coefficient.
- e: insurer's expenses rate.
- c: commission payment rate.
- *u*: non-negative initial surplus.

2. Assumptions

1. The reinsurer will pays the commission back to the insurer according to the business volume, which is c(1-a)P.

2. The insurer cannot reinsure the whole risk with a certain profit, i.e., e > c and $(1 - e)P - (1 + \alpha)\gamma\mu < 0$. The reinsurance premium:

$$P_{a,M} = (1-c)(1-a)P + (1+\alpha)\gamma \int_{M/a}^{\infty} (ax - M)dF(x).$$

3. The surplus process can be expressed as

$$U_{a,M}(t) = u + [(1-e)P - P_{a,M}]t - \sum_{i=1}^{N(t)} \min(aX_i, M) + W(t).$$

Note that for $X \sim N(\mu, \sigma^2)$, $M(t) = E[e^{tX}] = e^{\mu t + \sigma^2 t^2/2}$

4. The insurer's net(of expenses and reinsurance) risk at time t:

$$Y_{a,M}(t) = [(1-e)P - P_{a,M}]t - \sum_{i=1}^{N(t)} \min(aX_i, M) + W(t).$$

$$Y_{a,M}^i = -[Y_{a,M}(S_i) - Y_{a,M}(S_{i-1})].$$

Note that for $X \sim \exp(\lambda)$, $E[e^{tX}] = \int_0^\infty e^{tx} \lambda e^{-\lambda x} dx = \frac{\lambda}{\lambda - t}$.

5. The insurer's expected net profit per period of time (after reinsurance and expenses):

$$E[W(a,M)] = (1-e)P - P_{a,M} - \lambda E[X_{a,M}].$$

3. Preliminaries

Centeno(1985) has proved the following lemma

LEMMA 1: Let $A = \{a : 0 < a \leq 1 \text{ and there exists an } M$ such that $E[W(a, M)] = 0\}$ and $a_0 = \frac{(e-c)P}{(1-c)P-\lambda E(X)}$. Under our assumption on the reinsurance premium $P_{a,M}$, we have

1. $A = (a_0, 1]$.

2. For each $a \in A$, there is a unique M such that E[W(a, M)] = 0, i.e. there is a function Φ mapping A into $(0, \infty)$, such that $M = \Phi(a)$ is equivalent to E[W(a, M)] = 0.

3. $\Phi(a)$ is convex.

4. $\lim_{a\to a_0} \Phi(a) = +\infty$.

4. Two Models

The adjustment coefficient, $R_{a,M}$ is the positive solution of:

1. Modified Centeno's model

$$g_{a,M}(r) = M_{Y_{a,M}^{i}}(r) = E[e^{rY_{a,M}^{i}}]$$

= $E[e^{rX_{a,M}}]E[e^{-[(1-e)P - P_{a,M}]rT}]E[e^{-rW(1)}] = 1$

2. Modified Dufresne's model

$$\lambda \int_0^\infty e^{rx_{a,M}} dF(x) + Dr^2 = \lambda + [(1-e)P - P_{a,M}]r.$$

We want to figure out M and a to maximize $R_{a,M}$. We will discuss both models.

5. Modified Centeno's model

LEMMA 2.1: $R_{a,M}$ is the one and the only one positive solution of

$$E[e^{rX_{a,M}}]E[e^{-[(1-e)P - P_{a,M}]rT}]E[e^{-rW(1)}] = 1.$$

Proof:

1. $E[e^{rX_{a,M}}] = \int_0^\infty e^{rx_{a,M}} dF(x)$ is a non-decreasing convex function;

2. For
$$T \sim \exp(\lambda)$$
, $E[e^{-[(1-e)P - P_{a,M}]rT}] = \frac{\lambda}{\lambda + [(1-e)P - P_{a,M}]r};$

3. For $W(1) \sim N(0, 2D)$, $E[e^{-rW(1)}] = e^{Dr^2}$ is a non-decreasing convex function.

So the above equation is equivalent to

$$\lambda e^{Dr^2} \int_0^\infty e^{rx_{a,M}} dF(x) = \lambda + \left[(1-e)P - P_{a,M} \right] r \tag{1}$$

LEMMA 2.2: The adjustment coefficient is positive if and only if $(a, M) \in L$. *L* is the set of points for which the insurer's net expected profit is positive, i.e.

$$L = \{(a, M) : 0 \le a \le 1, M \ge 0 \text{ and } E[W(a, M)] > 0\}.$$

And for any $(a, M) \in L$, $H'_{a,M}(r)$ is positive at $r = R_{a,M}$.

Proof:

$$H_{a,M}(r) = \lambda e^{Dr^2 \int_0^\infty e^{rx_{a,M}} dF(x)} - [(1-e)P - P_{a,M}]r - \lambda = 0,$$

$$H_{a,M}'(0) = \frac{\partial H_{a,M}(r)}{\partial r}|_{r=0} = \lambda E[x_{a,M}] - [(1-e)P - P_{a,M}].$$

$$\xi = \begin{cases} +\infty & \text{if } M < +\infty \\ \tau & \text{for } M = +\infty \end{cases}, \longrightarrow \begin{cases} H_{a,M}(0) = 0 \\ \lim H_{a,M}(r)_{r \to \xi} = +\infty \end{cases}$$

Adjustment coefficient is positive $\iff H'_{a,M}(0) < 0.$

RESULT 1.1: For a fixed value of $a \in (a_0, 1]$, $R_{a,M}$ is a unimodal function of M, attaining its maximum value at the only point satisfying $M = \frac{1}{R_{a,M}} \ln[(1 + \alpha)] - DR_{a,M}$, where $R_{a,M}$ is the only positive solution of (1). Let $\hat{R}_a = \max(R_{a,M})$.

Proof:

From the implicit function theorem, we have,

$$\frac{\partial R_{a,M}}{\partial M} = -\frac{(\partial/\partial M)H_{a,M}(r)}{(\partial/\partial r)H_{a,M}(r)}|_{r=R_{a,M}}$$
$$\frac{\partial^2 R_{a,M}}{\partial M^2}|_{\frac{\partial R_{a,M}}{\partial M}=0} = -\frac{(\partial^2/\partial M^2)H_{a,M}(r)}{(\partial/\partial r)H_{a,M}(r)}|_{r=R_{a,M},\frac{\partial R_{a,M}}{\partial M}=0}$$

From lemma 2.2,
$$\frac{\partial H_{a,M}(r)}{\partial r}|_{r=R_{a,M}} > 0.$$

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Target: find out *M*, s.t.
$$\frac{\partial R_{a,M}}{\partial M} = 0$$
 where $\frac{\partial^2 R_{a,M}}{\partial M^2}|_{\frac{\partial R_{a,M}}{\partial M} = 0} < 0$.

 $\iff M$ satisfies

$$\frac{\partial H_{a,M}(r)}{\partial M}|_{r=R_{a,M}} = 0$$
$$\frac{\partial^2 H_{a,M}(r)}{\partial M^2}|_{r=R_{a,M},\frac{\partial R_{a,M}}{\partial M} = 0} > 0.$$

$$\frac{\partial H_{a,M}(r)}{\partial M} = \lambda r [1 - F(\frac{M}{a})] [e^{Dr^2 + rM} - (1 + \alpha)],$$

implies $\frac{\partial R_{a,M}}{\partial M} = 0$ at least has one solution and one solution is $M = \frac{1}{R_{a,M}} \ln[(1+\alpha)] - DR_{a,M}$ (2)

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Since

$$\frac{\partial^2 H_{a,M}(r)}{\partial M^2} = \lambda r^2 e^{rM + Dr^2} \left[1 - F(\frac{M}{a})\right] + \frac{\lambda r f(\frac{M}{a})}{a} (1 + \alpha - e^{Dr^2 + rM}),$$

When M satisfies (2), we have

$$\frac{\partial^2 H_{a,M}(r)}{\partial M^2}\Big|_{r=R_{a,M},\frac{\partial R_{a,M}}{\partial M}=0} = \lambda r^2 (1+\alpha) [1-F(\frac{M}{a})] > 0$$

Hence the second derivative with respect to M of $R_{a,M}$ is negative whenever the first derivative is zero, which implies that for fixed $a \in (a_0, 1]$, $R_{a,M}$ has at most one turning point, and that when such a point exists it is a maximum. The maximum will exist and be finite at the only point satisfying $M = \frac{1}{R_{a,M}} \ln[(1+\alpha)] - DR_{a,M}$.

Compared to the original model, the excess of loss retention limit M decreased due to the diffusion process, which means that for increasing uncertainty, the insurance company should cede more business to the reinsurer.

RESULT 1.2: \hat{R}_a is a unimodal function of a, for $a \in (a_0, 1]$, attaining it's maximum at a = 1, if and only if $\lim_{a \to 1^-} \frac{d}{da} \hat{R}_a \ge 0$

Proof:

At equation(2), defines M as a function of $a: \Upsilon(a)$. Let $\hat{R}_a = R_{a,\Upsilon(a)}$.

First we have
$$\frac{d\hat{R}_a}{da} = -\frac{(\partial/\partial a)H_{a,M}(r)}{(\partial/\partial r)H_{a,M}(r)}|_{M=\Upsilon(a),r=\hat{R}_a}$$

$$\frac{d^2\hat{R}_a}{da^2}|_{\frac{d\hat{R}_a}{da}=0} = -\frac{(\partial^2/\partial a^2)H_{a,M}(r)\times(\partial^2/\partial M^2)H_{a,M}(r)-[(\partial^2/\partial a\partial M)H_{a,M}(r)]^2}{(\partial/\partial r)H_{a,M}(r)\times(\partial^2/\partial M^2)H_{a,M}(r)}|_{M=\Upsilon(a),r=\hat{R}_a,\frac{d\hat{R}_a}{da}=0}$$

Note here we already proved that

$$\frac{\partial^2 H_{a,M}(r)}{\partial M^2}|_{M=\Upsilon(a),r=\widehat{R}_a} > 0 \text{ and } \frac{\partial H_{a,M}(r)}{\partial r}|_{M=\Upsilon(a),r=\widehat{R}_a} > 0.$$

After some calculations, we have

 $(\partial^{2}/\partial a^{2})H_{a,M}(r) \times (\partial^{2}/\partial M^{2})H_{a,M}(r) - [(\partial^{2}/\partial a\partial M)H_{a,M}(r)]^{2}|_{r=R_{a,M};\frac{\partial}{\partial a}H_{a,M}(r)=0;\frac{\partial}{\partial a}H_{a,M}(r)=0}$ $= \lambda^{2}r^{4}e^{rM+2Dr^{2}}[1-F(\frac{M}{a})]\int_{0}^{M/a}x^{2}e^{rax}dF(x) > 0,$ which means $\frac{d^{2}\widehat{R}_{a}}{da^{2}}|_{\frac{d\widehat{R}_{a}}{da}=0} > 0.$

On the other hand, when $a \to a_0$, \hat{R}_a goes to zero and we can say that the maximum of \hat{R}_a is 1, if and only of if $\lim_{a\to 1^-} \frac{d}{da} \hat{R}_a \ge 0$. Hence the result is proved. It is the same as the original result.

Example 1. Exponential Distribution

Claim amount distribution: $F(x) = 1 - e^{-\beta x}$ and $f(x) = \beta e^{-\beta x}$.

Let $\alpha = 0.9, c = 0.2, e = 0.3, P = 1.7, \beta = 0.5, \lambda = 1, D = 0.004.$ It satisfies e > c and $(1 - e)P - (1 + \alpha)\lambda\mu = -2.61 < 0$

In this case, we have:

$$\int_0^\infty e^{rx_{a,M}} dF(x) = \frac{ra}{ra-\beta} e^{rM-\beta\frac{M}{a}} - \frac{\beta}{ra-\beta},$$
$$(1-e)P - P_{a,M} = (a+c-ca-e)P - (1+\alpha)\lambda\frac{a}{\beta}e^{-\beta\frac{M}{a}}$$

Hence (1) becomes

$$\lambda e^{Dr^2} \left(\frac{ra}{ra-\beta} e^{rM-\beta\frac{M}{a}} - \frac{\beta}{ra-\beta}\right) = \lambda + \left[\left(a+c-ca-e\right)P - (1+\alpha)\lambda\frac{a}{\beta}e^{-\beta\frac{M}{a}}\right]r$$

After some simplification, it becomes:

$$\frac{\lambda r^2 a^2}{\beta} (1+\alpha)^{1-\frac{\beta}{ar}} e^{D\beta r/a} = \lambda \beta e^{Dr^2} + (ra-\beta)(a+c-ca-e)rP + \lambda(ra-\beta).$$

The solution is $r = \beta/a$.

When a = 1, the adjustment coefficient $R_{a,M}$ attains its maximum value which satisfies:

$$\left\{\frac{\lambda r^2}{\beta(r-\beta)}(1+\alpha)^{1-\beta/r}e^{D\beta r} = \frac{\lambda\beta}{r-\beta}e^{Dr^2} + (1-e)rP + \lambda\right\}|_{r=R_{a,M}}$$

 $R_{a,M} = 0.5$ and M = 1.2817.

6. Modified Dufresne's model

LEMMA 3.1: $R_{a,M}$ is the one and the only one positive solution of

$$\lambda \int_0^\infty e^{rx_{a,M}} dF(x) + Dr^2 = \lambda + [(1-e)P - P_{a,M}]r.$$

The function can be rewritten as

$$\lambda \int_0^\infty e^{rx_{a,M}} dF(x) = -Dr^2 + [(1-e)P - P_{a,M}]r + \lambda.$$
 (3)

LEMMA 3.2: The adjustment coefficient is positive if and only if $(a, M) \in L$. L is the set of points for which the insurer's net expected profit is positive, i.e.

$$L = \{(a, M) : 0 \le a \le 1, M \ge 0 \text{ and } E[W(a, M)] > 0\}.$$

And for any $(a, M) \in L$, $H'_{a,M}(r)$ is positive at $r = R_{a,M}$.

RESULT 2.1: For a fixed value of $a \in (a_0, 1]$, $R_{a,M}$ is a unimodal function of M, attaining its maximum value at the only point satisfying $M = \frac{1}{R_{a,M}} \ln[(1 + \alpha)]$, where $R_{a,M}$ is the only positive solution of (2.2). Let $\hat{R}_a = \max(R_{a,M})$.

Proof:

Similar to result 1.1,

$$\frac{\partial H_{a,M}(r)}{\partial M} = \lambda r [1 - F(\frac{M}{a})] [e^{rM} - (1 + \alpha)]$$
$$\frac{\partial^2 H_{a,M}(r)}{\partial M^2} = \lambda r [-\frac{1}{a} f(\frac{M}{a})] [e^{rM} - (1 + \alpha)] + \lambda r^2 [1 - F(\frac{M}{a})] e^{rM}.$$

which implies,

$$\frac{\partial R_{a,M}}{\partial M} = -\frac{(\partial/\partial M)H_{a,M}(r)}{(\partial/\partial r)H_{a,M}(r)}|_{r=R_{a,M}} = 0$$

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at lease has one solution and one of solution is

$$M = \frac{1}{R_{a,M}} \ln[(1+\alpha)]. \tag{4}$$

When the above relationship satisfied,

$$\frac{\partial^2 H_{a,M}(r)}{\partial M^2}\Big|_{r=R_{a,M},\frac{\partial R_{a,M}}{\partial M}=0} = \lambda r^2 [1 - F(\frac{M}{a})] e^{rM} > 0$$

which means.

$$\frac{\partial^2 R_{a,M}}{\partial M^2}\Big|_{\frac{\partial R_{a,M}}{\partial M}=0} = -\frac{(\partial^2/\partial M^2)H_{a,M}(r)}{(\partial/\partial r)H_{a,M}(r)}\Big|_{r=R_{a,M},\frac{\partial R_{a,M}}{\partial M}=0} < 0$$

Hence the second derivative with respect to M of $R_{a,M}$ is negative whenever the first derivative is zero, which implies that for fixed $a \in (a_0, 1]$, $R_{a,M}$ has at most one turning point, and that when such a point exists it is a maximum. The maximum will exist and be finite at the only point satisfying $M = \frac{1}{R_{a,M}} \ln[(1 + \alpha)]$.

RESULT 2.2:

 \hat{R}_a is a unimodal function of a, for $a \in (a_0, 1]$, attaining it's maximum at a = 1, if and only of $\lim_{a \to 1^-} \frac{d}{da} \hat{R}_a \ge 0$

Proof:

Here
$$\hat{R}_a = R_{a,\Upsilon(a)}$$
, where $\Upsilon(a) = M = \frac{1}{R_{a,M}} \ln[(1+\alpha)]$.

Similar to result 1.2, we first have $\frac{d^2 \hat{R}_a}{da^2} |_{\frac{d \hat{R}_a}{da} = 0} > 0.$

On the other hand, when $a \to a_0$, \hat{R}_a goes to zero and we can say that the maximum of \hat{R}_a is 1, if and only of if $\lim_{a\to 1^-} \frac{d}{da} \hat{R}_a \ge 0$. Hence the result is proved. It is the same as the original result.

Example 2. Exponential Distribution

Use the same assumptions as example 1, (3) becomes

$$\lambda(\frac{ra}{ra-\beta}e^{rM-\beta\frac{M}{a}}-\frac{\beta}{ra-\beta})+Dr^{2}=\lambda+(a+c-ca-e)rP-(1+\alpha)\lambda r\frac{a}{\beta}e^{-\beta\frac{M}{a}}$$

Consider M satisfying (4), after some simplification, we have :

$$\lambda a(1+\alpha)^{1-\frac{\beta}{ra}}\frac{ra}{\beta} + (ra-\beta)Dr = \lambda a + (a+c-ca-e)(ra-\beta)P.$$

The solution is $r = \beta/a$.

When a = 1, the adjustment coefficient $R_{a,M}$ attains its maximum value which satisfies:

$$(1+\alpha)^{1-\frac{\beta}{r}}\frac{\lambda r}{\beta(r-\beta)} + Dr = \frac{\lambda}{r-\beta} + (1-e)P.$$

 $R_{a,M} = 0.5$ and M = 1.2837.

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Any Questions?

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