# TRANSACTIONS OF SOCIETY OF ACTUARIES 

 1985 VOL. 37
## REPORT OF THE COMMITTEE TO RECOMMEND NEW DISABILITY TABLES FOR VALUATION*

The Committee's charge was to develop new disability tables for possible adoption by appropriate authorities for valuation of individual policy reserves (active lives) and individual and group claim reserves (disabled lives). The 1964 Commissioner's Disability Table ( 1964 CDT) has been recognized as being inadequate for claim reserves and is thought to be too conservative for active life reserves for policies sold to females in general and to males in the more popular occupation classes. A table was needed to better represent current products and experience.

The Committee believed that any new valuation table should be the simplest table that would embrace all of the factors that seem to affect policy liabilities. Its plan was to:

1. Develop an experience table involving all the variables we could statistically and rationally relate to either the incidence or termination of disability. The significant variables were determined to be age, sex, occupation class, elimination period, cause (sickness or accident), and duration from the date of disablement.
2. Eliminate any variable that does not significantly affect policy liabilities. Company and year of exposure were also found to be significant but were not kept isolated. The DTS Valuation Table was developed as an industry average table reflecting exposure periods of broad economic experience.
3. Develop appropriate margins to be added to such experience table to form a valuation table.

## DISABILITY TABLE STUDY (DTS) BASIC TABLE

Development of such a table required collection of data in sufficient volume, detail, and credibility as to warrant a sound analytical approach to determining the contingency factors affecting claim costs, to mathematically quantify those factors (which we will call variables) and to produce a simple, credible means to express those results in an experience table. Our approach was to build from a disability termination study that had been initiated by Mr. John Haynes Miller to collect and analyze data on policyholder terminations from claim. This study would tell us what variables appeared to be

[^0]significant in influencing claim continuance. Although far fewer variables were available, the same statistical methods were used to establish the variables that were significant for incidence rates.

The DTS Basic Table has been developed on this basis and appears in Appendix E of this report. Throughout this report, the letters DTS refer to the Disability Table Study. The report will refer to the tables as the DTS Basic Table and the DTS Valuation Table.

Appendix A provides a detailed description of the process of collecting and editing the DTS claim termination study data. Twenty companies participated, submitting usable experience data on 133,936 closed claims.

Appendix B describes the method used to determine the significant variables and to calculate numerical factors to reflect each variable's related significance in the rate of termination from claim. It describes further the practical application of this method to the determination of factors which would produce smooth termination rates for the first 2 years from disablement. Appendix B also describes the different methods used for determining rates of termination for the third through the tenth years and for ultimate years. Group long-term disability (LTD) experience was the primary influence for rates from the latter part of the second year through the tenth year. The method used to determine ultimate rates by attained age and by sex for durations 11 years and over was to evaluate ultimate data from several sources.

Appendix C describes the method used to determine disability incidence rates from data from several sources: DTS, Society of Actuaries (SOA), and the New York Study. Included are a large number of graphs, displayed here to provide the reader the means to evaluate the graduation process employed as well as for a quick assessment of the relationship of incidence rates to the parameters-age, sex, occupation class, elimination period, and cause (accident or sickness).

Appendix D contains illustrative values determined from the DTS Basic Table defined in Appendixes B and C. Comparisons are made to the SOA data as well as to the 1964 CDT.

Appendex E illustrates the methods for constructing a conventional continuance table from the DTS Basic Table. This appendix also shows two sample tables (males-class 3-e.p. 7 days, and males-class 1-e.p. 30 days). as well as the total DTS Basic Table. The DTS Basic Table is expressed in variable form.

Each variable found to be significant has been evaluated for the period of significance as indicated by Exhibit 1. The conventional continuance tables are readily constructed from these factors. It is expected, however, that most companies will work more directly from incidence rates and probabilities of termination from claim.

EXHIBIT 1
Disability Table Study
Determining Variables for Rates of Incidence and Termination*


[^1]The DTS table data base and the simple Fortran computer program the Committee used is available from the Society's office. Over three hundred requests for the diskette already have been filled. It should be noted that the diskette is essentially the working version of the Committee's report and is incorporated into the National Association of Insurance Commissioners' (NAIC) recommendation whereby the DTS Valuation Table is now known as the "1985 Commissioner's Individual Disability Tables A."

## SUITABILITY

There are several characteristics of the DTS table that will make it a suitable table as a basis for a valuation table for both active life reserves and claim reserves for individual disability income policies:

1. The DTS table was developed from exposures of the mid to late 1970s. At that time, the industry was going through a period of claims deterioration, to about 1976, and
the beginning of a claims improvement trend thereafter. The DTS table is, therefore, on the conservative side, relative to the good claims experience of the early 1980s.
2. The DTS table is sufficiently flexible as to lend itself to any company's particular mix of business by sex, elimination period, or occupation class.
3. The DTS table, although appearing complex, is very easy to use.
4. Each feature of the parametric approach is readily understandable.
5. This variable factor approach gives companies good detail with which to analyze the adequacy of their reserves over short periods of time and the tools to isolate any discriminating factor and adjust for it at the proper point.
6. The analytical approach and the subsequent method of determining termination rates should give the DTS table a high incidence of credibility and, of course, reliability.
7. Above all, the DTS table will promote the Society's intended position of prescribing sound principles of valuation, in contrast to specific minimal reserves from an aggregate table. Regulators can, with the DTS Valuation Table, enforce sound principles. Traditionally defined minimal reserves, though simple for regulation mechanics, have little reliability or credibility. Regulators, we believe, will be appreciative of a better way to set reserve standards tailored to the product and the company. This table offers an approach that considers the occupation class, elimination period, cause and sex, as well as age and claim duration.
8. The DTS table is easy to modify in order to add a contingency margin for a specific purpose.
9. During the second year of disablement, the termination rates were graduated from the level indicated in loss-of-time (LOT) experience to the level indicated by LTD. The rates are then based upon LTD experience for the third through the sixth years and graded into the ultimate rates of the eleventh year. LTD termination rates are considerably lower than LOT rates during the first 2 years of disablement and somewhat lower thereafter. Therefore, the table includes some implied margin beginning in the second year.

There are, of course, characteristics of any company's business which would make it not completely homogeneous to the underlying aggregate experience included in the DTS. Characteristics that would seem to inherently affect reserving requirements would include such items as:

1. the dimensions of the occupation classes;
2. the use of specialty classes;
3. either tight or liberal underwriting;
4. the relationship between benefits and earnings;
5. prudent claims-handling techniques, including rehabilitation activity;
6. geographic concentration of business;
7. definition of disability in the insuring clause;
8. other special features that might result in longer periods of claim, or reductions in the elimination period; or
9. even the quality of the field force.

There was definite evidence of antiselection by amount observed in our analysis, but we did not have information on the suspected underlying cause-
the relation of insured benefits to the household spendable income-and so we could not produce reliable relative numerical values for an amount variable.

We have no definite evidence of antiselection on the residual clause. Nevertheless, it seems logical that a person could be on claim for a longer period of time under residual, even though the aggregate amount paid may not be any greater than full benefits for the regular, shorter period.

## ALTERNATIVE MARGINS

The need for a small margin arises from the uncertainty in incidence (affecting the number of claims) and from the uncertainty of recovery (affecting the aggregate amount of claims payments). This need would be appropriately covered by a margin in the claims cost (affecting the active life reserves) and a margin in the claim reserves. It is not feasible for a valuation table to be so strong as to cover the worst possible experience of all companies. Nevertheless, there should be small margins to give some assurance of adequacy of reserves for the most likely unusual occurrences.

Minimal reserves could be prescribed as multiples of the DTS reserve factors on a scale graded by the size of a company's block of individual disability income business. Unfortunately, the approach would place too heavy a burden on a company that is growing conservatively or would not produce strong enough reserves for a company growing aggressively.

Adding a flat percentage margin of, say, 10 percent is very practical but is not objective, and when the margin is set high enough to adequately cover most cases, it would subject the more standard policies to unnecessary strain. Providing for a small margin by modifying a particular variable seems to be a better way to fulfill the purpose.

The adverse part of claims experience during the 1970s was caused mostly by the prolonging of early claims (short deferring of recovery), rather than by higher claim rates. Claims incidence on SOA data actually showed slight improvement during that period on policies with longer elimination periods and at the higher ages.

Increasing an incidence factor, while directly increasing active life policy reserves, would not affect claim reserves. Nor should it, because higher incidence could lead to higher termination from claims where there are more claimants less severely impaired.

Decreasing the termination rate by a percentage during the early influential months of a claim will add a margin to most active life reserves as well as increase all of the claim reserves in the early durations, where it is really needed.

Since the DTS terminations are highly influenced by group LTD experience by the end of the second year and through the tenth year, and since
terminations reflect, essentially, ultimate insured disabled life mortality experience thereafter, it would seem prudent to also allow for extra morbidity where it is likely to occur and would be most significant, during the first year of claim, and grading off during the second year of claim.

A possible 5 percent adverse deviation from normal claims terminations rates during the first year of disablement is well within the range of managerial judgment. Such a change could arise insidiously before the company actuary or the industry could recognize the trend or identify the cause. The DTS Valuation Table includes such a margin.

## RECOMMENDATION

The DTS Valuation Table consists of the DTS Basic Table allowing for 95 percent of standard termination from disablement rates during the first year of disablement, grading to 100 percent of the DTS standard termination rates in the eighteenth month. An illustration of the approximate reserve margins can be obtained by reviewing Exhibit D-8 in Appendix D of this report. Active life reserve margins would be from 5-10 percent, and claim reserve margins would be about 10 percent in the first 2 months of disablement. The claim reserve margin will decrease each month and disappear by the eighteenth month.

We recommend that the Society of Actuaries propose the DTS Valuation Table (Exhibits 2, 3a-3c, and 4) to the NAIC as the minimal valuation table for individual disability income active life and claim reserves. ${ }^{1}$ We recommend that this DTS Valuation Table be used with 1980 CSO ultimate mortality, sex distinct. Select mortality would be more precise than ultimate, but we believe it is more acceptable to be consistent with life insurance valuation standards.

The Committee did not have sufficient data to evaluate variables for policies with 6 -month elimination periods. The 90 -day elimination table would be used to calculate costs for policies with greater elimination periods even though this would interject some possible conservatism in active life reserves for such policies. For most insurers, the proportion of such policies would be minor.

The industry currently takes some comfort in the reserve margin being provided by low-valuation interest rates (3-3.5 percent) as an offset to current valuation morbidity deficiency. Greater confidence in valuation adequacy is obtained, of course, where reserving margins are more explicit with respect to each contingency. This is accomplished for morbidity by the

[^2]EXHIBIT 2
DTS Valuation Table
(Incidence of disability
rates per 1,000 lives exposed)

|  | Age | Male-Accident |  |  |  |  | AGF: | Mani-Sickness |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Elimination Period |  |  |  |  |  | Elminaton Period |  |  |  |  |
|  |  | 0 day | 7 l ay | 14 amy | 30 uar | 90 day |  | 0 day | 7 bar | 14 dar | 30 day | 90 dar |
| $\overline{\text { Class } 1}$ | 25 | 33.97 | 25.84 | 13.13 | 4.90 | 86 | 25 |  | 32.26 | 18.22 | 5.51 | 1.01 |
|  | 35 | 32.88 | 24.42 | 11.99 | 4.23 | 51 | 35 | $\cdots$ | 36.11 | 21.55 | 6.48 | 1.13 |
|  | 45 | 30.40 | 20.40 | 9.86 | 4.50 | 65 | 45 |  | 47.12 | 31.19 | 12.63 | 2.70 |
|  | 55 | 30.19 | 18.32 | 9.63 | 4.71 | . 80 | 55 |  | 69.48 | 52.75 | 25.11 | 7.78 |
|  | 62 | 33.45 | 16.11 | 10.39 | 5.47 | 1.18 | 62 | . | 91.52 | 74.06 | 41.24 | 15.20 |
| Class 2 | 25 | 59.96 | 47.98 | 30.01 | 10.48 | 2.07 | 25 |  | 46.61 | 27.01 | 12.17 | 2.23 |
|  | 35 | 59.96 | 44.62 | 28.83 | 10.14 | 2.09 | 35 |  | 52.79 | 33.37 | 14.47 | 2.56 |
|  | 45 | 56.74 | 38.49 | 25.67 | 9.86 | 2.14 | 45 |  | 65.97 | 46.91 | 25.40 | 6.21 |
|  | 55 | 51.66 | 31.31 | 20.50 | 10.03 | 2.20 | 55 |  | 92.99 | 71.27 | 41.37 | 15.74 |
|  | 62 | 52.84 | 29.85 | 19.86 | 10.92 | 2.57 | 62 |  | 116.81 | 93.05 | 58.54 | 25.94 |
| Class 3 | 25 | 75.80 | 62.68 | 42.87 | 23.69 | 7.04 | 25 |  | 46.83 | 32.22 | 14.75 |  |
|  | 35 | 74.78 | 58.37 | 39.59 | 22.57 | 6.48 | 35 |  | 52.72 | 38.32 | 18.70 | 3.52 |
|  | 45 | 69.76 |  |  |  | 5.97 | 45 |  | 67.05 | 51.53 | 29.45 | 7.83 |
|  | 55 | 66.37 | 44.27 | 30.51 | 18.49 | 5.46 | 55 |  | 92.60 | 76.39 | 52.66 | 20.07 |
|  | 62 | 65.04 | 39.98 | 27.96 | 18.56 | 5.30 | 62 |  | 116.23 | 98.78 | 78.56 | 36.04 |
| Class 4 | 25 | 89.42 | 77.60 | 52.59 | 27.03 | 8.73 | 25 |  | 48.20 | 33.28 | 15.07 | 3.04 |
|  | 35 | 91.59 | 73.24 | 50.53 | 26.93 | 8.17 | 35 |  | 53.75 | 39.27 | 19.33 | 3.59 |
|  | 45 | 84.64 | 62.13 | 42.61 | 24.78 | 7.68 | 45 |  | 70.03 | 52.71 | 30.13 | 7.97 |
|  | 55 | 79.77 | 52.03 | 37.34 | 22.78 | 7.27 | 55 |  | 95.01 | 77.91 | 55.87 | 20.45 |
|  | 62 | 79.95 | 49.76 | 36.11 | 22.96 | 7.20 | 62 |  | 119.16 | 101.41 | 81.62 | 36.63 |

EXHIBIT 2-Continued

|  | Agie | female-accident |  |  |  |  | Age | Fiemale-Sigkness |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Elimination Period |  |  |  |  |  | Elimination Period |  |  |  |  |
|  |  | 0 day | 7 day | 14 day | 30 day | 90 day |  | Opar | 7 bar | 14 day | 30 day | 90 uay |
| Class 1 | 25 | 23.06 | 19.92 | 12.96 | 6.00 | 1.14 | 25 |  | 61.10 | 39.29 | 14.03 | 2.55 |
|  | 35 | 26.28 | 20.87 | 13.39 | 6.21 | . 91 | 35 |  | 84.38 | 56.89 | 24.75 | 4.37 |
|  | 45 | 32.36 | 22.77 | 13.78 | 6.83 | 1.11 | 45 |  | 94.57 | 68.33 | 34.14 | 7.64 |
|  | 55 | 45.05 | 26.77 | 14.82 | 8.06 | 1.46 | 55 |  | 90.28 | 61.49 | 34.23 | 10.31 |
|  | 62 | 69.00 | 31.56 | 17.54 | 9.91 | 2.25 | 62 | $\cdots$ | 93.06 | 69.44 | 45.30 | 13.85 |
| Class 2 | 25 | 35.05 | 31.48 | 23.39 | 13.40 | 3.22 | 25 |  | 80.97 | 53.57 | 20.03 | 3.75 |
|  | 35 | 39.36 | 32.01 | 23.36 | 14.02 | 3.20 | 35 |  | 116.02 | 80.05 | 35.34 | 6.60 |
|  | 45 | 47.46 | 33.55 | 24.40 | 15.02 | 3.40 | 45 | $\cdots$ | 134.18 | 92.93 | 47.62 | 10.81 |
|  | 55 | 62.53 | 37.10 | 26.13 | 16.11 | 3.75 | 55 |  | 117.29 | 84.93 | 49.00 | 14.95 |
|  | 62 | 88.91 | 44.31 | 29.27 | 17.88 | 4.46 | 62 |  | 120.40 | 87.53 | 63.15 | 18.86 |
| Class 3 | 25 | 41.93 | 38.01 | 27.94 | 17.63 | 6.19 | 25 |  | 86.64 | 57.85 | 24.83 | 5.03 |
|  | 35 | 46.30 | 38.45 | 28.54 | 18.20 | 6.54 | 35 |  | 124.79 | 96.77 | 44.67 | 8.43 |
|  | 45 | 53.01 | 39.08 | 29.09 | 19.24 | 6.75 | 45 |  | 145.58 | 116.19 | 58.44 | 14.43 |
|  | 55 | 66.71 | 41.96 | 30.86 | 20.99 | 7.08 | 55 |  | 122.98 | 99.89 | 59.99 | 17.86 |
|  | 62 | 90.05 | 48.12 | 33.60 | 23.74 | 7.26 | 62 |  | 125.95 | 101.06 | 69.18 | 22.76 |
| CIASS 4 | 25 | 52.41 | 47.52 | 34.93 | 22.04 | 7.74 | 25 |  | 90.24 | 60.26 | 25.86 | 5.23 |
|  | 35 | 57.87 | 48.07 | 35.67 | 22.75 | 8.17 | 35 |  | 130.00 | 100.81 | 46.53 | 8.79 |
|  | 45 | 66.26 | 48.86 | 36.36 | 24.05 | 8.45 | 45 |  | 151.65 | 121.04 | 60.87 | 15.03 |
|  | 55 | 83.39 | 52.45 | 38.58 | 26.25 | 8.85 | 55 |  | 128.10 | 104.05 | 62.49 | 18.61 |
|  | 62 | 112.57 | 60.16 | 42.00 | 29.67 | 9.08 | 62 |  | 131.20 | 105.27 | 72.07 | 23.71 |

## EXHIBIT 3a

DTS Valuation Table
Factors for Calculation of Weekly Termination Rates*

| Factors** | We:k 1 |  |  | Week 2 |  |  | Week 3 |  |  | Wexk 4 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate | . 132 |  |  | 114 |  |  | . 111 |  |  | . 119 |  |  |
| Age 25: | 1.019 |  |  | 1.138 |  |  | 1.127 |  |  | 1.105 |  |  |
| EP-0,7,14,30 | 1.000 |  |  | $\begin{array}{rlllll}1.053 & .941 \\ 951 & 968 & 1.0121 .053\end{array}$ |  |  | $\begin{array}{llllll}1.131 & 1.066 & .788\end{array}$ |  |  |  |  |  |
| Class-1,2,3,4 | . 978.981 | .995 1.011 |  |  |  |  | $\begin{array}{ll}.963 & .9831 .0091 .036\end{array}$ |  |  | $\begin{array}{lllll}.983 & .9971 .0051 .009\end{array}$ |  |  |
| Sex-M,F | 1.154 .859 |  |  | 1.142 . 858 |  |  | 1.101 .897 |  |  | 1.079 .922 |  |  |
| Cause-A,S | $1.034 \quad .957$ |  |  | . 9561.018 |  |  | . 9121.074 |  |  | . 8941.098 |  |  |
| Age 35: | 1.014 |  |  | . 961 |  |  | . 959 |  |  | . 997 |  |  |
| EP-0,7,14,30 | 1.000 |  |  | $\begin{cases}1.062 & .934 \\ 1.046 & .999\end{cases}$ | . 977 | . 960 | $\left\lvert\, \begin{array}{rrr} 1.176 & 1.067 & .757 \\ 1.006 & .998 & .995 \end{array}\right.$ |  | .991 | $\begin{array}{\|ll} 1.1301 .049 & .815 \\ 1.0071 .001 & .996 \end{array}$ |  | .991 |
| Class-1,2,3,4 | 1.1111 .030 | . 957 | . 882 |  |  |  |  |  |  |  |  |  |
| Sex-M, F | 1.101 .901 |  |  | 1.190 . 824 |  |  | 1.146 . 862 |  |  | 1.090 . 913 |  |  |
| Cause-A,S | . 995 . 994 |  |  | $1.044 \quad .933$ |  |  | $\begin{array}{r}1.996 \quad .984 \\ \hline\end{array}$ |  |  | . 9601.023 |  |  |
| Age 45: | 1.027 |  |  | . 894 |  |  | . 898 |  |  | . 943 |  |  |
| EP-0,7,14,30 | 1.000 1.070 |  |  | $\begin{array}{lr}1.082 \\ 1.135 & 1.916 \\ 1.029\end{array}$ | . 951 | . 884 | $\left\{\begin{array}{lll} 1.218 & 1.053 & .741 \\ 1.061 & 1.017 & .977 \end{array}\right.$ |  | . 939 | $\begin{array}{lll}1.185 & 1.023 & .797 \\ 1.041 & 1.011 & .984\end{array}$ |  | . 960 |
| Class-1, $2,3,4$ | 1.2151 .070 | . 934 | . 796 |  |  |  |  |  |  |  |  |  |
| Sex-M,F | 1.038 . 955 |  |  | 1.146 . 856 |  |  | 1.110 .890 |  |  | $1.063 \quad .936$ |  |  |
| Cause-A,S | . 9771.013 |  |  | $1.132 \quad .860$ |  |  | $1.090 \quad .898$ |  |  | $1.046 \quad .939$ |  |  |
| Age 55: | 1.016 |  |  | . 949 |  |  | . 942 |  |  | . 948 |  |  |
| EP-0,7,14,30 | 1.000 006 |  |  | $\begin{array}{\|rr} 1.136 & .873 \\ 1.193 & 1.057 \\ 1.002 & .978 \\ 1.191 & .817 \\ \hline \end{array}$ | . 935 | . 832 | $\left[\begin{array}{lll} 1.263 & 1.001 & .751 \\ 1.120 & 1.039 & .959 \end{array}\right.$ |  | . 887 | $\left\lvert\, \begin{array}{lll} 1.228 & .988 & .797 \\ 1.086 & 1.028 & .970 \end{array}\right.$ |  | . 918 |
| Class-1,2,3,4 | 1.2431 .080 | . 936 | . 769 |  |  |  |  |  |  |  |  |  |
| Sex-M,F | . 9721.020 |  |  |  |  |  | 1.000 .988 |  |  | 1.000 .995 |  |  |
| Cause-A,S | 1.031 .960 |  |  |  |  |  | $1.171 \quad 836$ |  |  | 1.142 . 860 |  |  |
| Age 62: | . 924 |  |  | 1.058 |  |  | 1.072 |  |  | 1.007 |  |  |
| EP-0, $7,14,30$ | 1.000 |  |  | $\begin{cases}1.109 & .894 \\ 1.185 & 1.066\end{cases}$ | . 941 | . 825 | $\begin{array}{rr} 1.210 & .958 \\ 1.167 & 1.057 \\ .873 & 1.132 \\ 1.266 & .773 \\ \hline \end{array}$ | $\begin{array}{r} .819 \\ .949 \end{array}$ | . 847 | $\begin{array}{\|rr} 1.210 & .965 \\ 1.143 & 1.049 \\ .922 & 1.080 \\ 1.257 & .781 \\ \hline \end{array}$ | 827 |  |
| Class-1,2,3,4 | 1.2051 .072 |  | . 797 |  |  |  |  |  |  |  | . 955 | . 868 |
| Sex-M, F | . 9081.092 |  |  | . 8501.153 |  |  |  |  |  |  |  |  |
| Cause-A,S | $1.245 \quad 794$ |  |  | $1.300 \quad 749$ |  |  |  |  |  |  |  |  |

*The termination rate is the product of the Duration Rate and the corresponding variable factors for the respective Age, e.g., for Week 2 (.114), age 35 (.961), EP 7day (.934). class 2 (.999), male (1.190), accident (1.044), the termination rate is . 127 .
**Age is age at disablement.
Duration is from the date of disablement.
Class 1 includes the 2 lowest premium occupation classes of a 5 -class manual, or the lowest premium class of a 4 -class manual.

EXHIBIT 3a-Continued
dTS Vaiuation Table
Factors for Calculation of Weekly Termination Rates*

| Factors** | Week 5 |  |  | Weik 6 |  |  | Week 7 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate | . $112^{+}$ |  |  | . 117 |  |  | . 120 |  |  |
| Age 25: | 1.048 |  |  | 1.060 |  |  | 1.066 |  |  |
| EP-0, 7, 14,30 | $1.1561 .246 \quad 1.036$ |  | . 597 | 1.0761 .2101 .048 |  | . 689 | 1.0181 .1771 .053 |  | .760.993 |
| Class-1, 2, 3, 4 | $1.0061 .0061 .000 \quad .984$ |  |  | . $9921.0081 .007 \quad .990$ |  |  | . 9881.0101 .009 |  |  |
| Sex-M,F | $1.060 \quad .942$ |  |  | 1.036 . 965 |  |  | 1.022 .978 |  |  |
| Cause-A.S | . 8841.112 |  |  | . 8781.118 |  |  | . 8741.125 |  |  |
| Age 35: | . 985 |  |  | 1.019 |  |  | 1.043 |  |  |
| EP-0, 7, 14,30 | 1.2491 .191 | 985 | . 608 | 1.1641 .153 | . 998 | . 701 | 1.1191 .121 | . 006 | . 759 |
| Class-1, 2, 3,4 | 1.0071 .003 | . 997 | . 988 | . 9991.003 | . 002 | . 994 | . 9961.001 | . 003 | . 998 |
| Sex-M, F | 1.055 .946 |  |  | 1.019 .981 |  |  | . 9941.005 |  |  |
| Cause-A,S | . 9371.050 |  |  | . 9251.062 |  |  | . 9161.073 |  |  |
| Age 45: | . 962 |  |  | . 988 |  |  | 1.007 |  |  |
| EP-0,7,14,30 | 1.2981 .123 | 938 | . 652 | 1.2061 .096 | . 962 | . 738 | 1.1721 .073 | . 974 | . 783 |
| Class-1, 2, 3,4 | 1.0251 .009 | 990 | . 972 | 1.0151 .006 | . 995 | . 983 | 1.0101 .003 | . 996 | . 990 |
| Sex-M,F | 1.033 . 966 |  |  | 1.005 .995 |  |  | . 9841.016 |  |  |
| Cause-A.S | $1.014 \quad .970$ |  |  | $1.002 \quad .981$ |  |  | . $989 \quad .994$ |  |  |
| Age 55: | . 977 |  |  | . 969 |  |  | . 964 |  |  |
| EP-0,7,14,30 | 1.2981 .056 | . 897 | . 725 | 1.2201 .052 | . 930 | . 786 | 1.1961 .041 | . 946 | . 814 |
| Class-1,2,3,4 | 1.0601 .023 | . 979 | . 938 | 1.0411 .018 | . 985 | . 956 | 1.0301 .013 | . 989 | . 968 |
| Sex-M,F | . 9971.001 |  |  | . 99951.005 |  |  | . 9901.010 |  |  |
| Cause--A,S | 1.118 .879 |  |  | 1.111 . 884 |  |  | 1.098 .895 |  |  |
| Age 62: | 1.028 |  |  | . 965 |  |  | . 920 |  |  |
| EP-0,7,14,30 | 1.2571 .004 | . 867 | . 815 | 1.1961 .031 | . 896 | . 849 | 1.1911 .031 | 910 | . 857 |
| Class-1,2,3,4 | 1.1201 .044 | 962 | . 885 | 1.0901 .040 | . 971 | . 906 | 1.0711 .037 | 977 | . 921 |
| Sex-M,F | . 95511.045 |  |  | 1.988 1.012 |  |  | 1.011 |  |  |
| Cause-A,S | $1.245 \quad .790$ |  |  | $1.260 \quad .780$ |  |  | $1.253 \quad .785$ |  |  |

*The termination rate is the product of the Duration Rate and the corresponding variable factors for the respective Age, e.g., for Weck 2 (.114), age 35 (.961), EP 7day (.934). class 2 (.999), male (1.190), accident (1.044), the termination rate is .127 .
**Age is age at disablement.
Duration is from the date of disablement.
Class 1 includes the 2 lowest premium occupation classes of a 5 -class manual, or the lowest premium class of a 4 -class manual.
${ }^{\dagger}$ Use .080 for 30 -day elimination periods to allow for the short week from 30 to 35 days.

EXHIBIT 3a-Continued
DTS Valuation Table
Factors for Calculation of Weekly Termination Rates*

|  | WEEK 8 |  |  | Week 9 |  |  | Wixk 10 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\frac{\text { Factors** }}{\text { DURATION Rate }}$ | . 119 |  |  | . 116 |  |  | . 111 |  |  |
| Age 25: | 1.073 |  |  | 1.079 |  |  | 1.086 |  |  |
| EP-0, $0,14,30$ | $\begin{array}{\|cccc} .9801 .147 & 1.054 & .820 \\ .983 & 1.009 & 1.010 & .997 \end{array}$ |  |  | $\begin{array}{\|lllll} .958 & 1.118 & 1.049 & .873 \\ . .978 & 1.007 & 1.012 & 1.004 \\ \hline \end{array}$ |  |  | $\begin{array}{llll} .951 & 1.087 & 1.038 & .921 \\ .972 & 1.002 & 1.013 & 1.013 \end{array}$ |  |  |
| Class-1,2,3,4 |  |  |  |  |  |  |  |  |  |
| Sex-M,F | $\begin{array}{rrr} .983 & 1.009 & 1.010 \\ 1.012 & .988 \end{array}$ |  |  | $1.004 \quad .995$ |  |  | . 9971.002 |  |  |
| Cause-A,S | . 8711.129 |  |  | . 8701.131 |  |  | . 8711.131 |  |  |
| Age 35: | 1.058 |  |  | 1.066 |  |  | 1.068 |  |  |
| EP-0, $7,14,30$ | $\begin{array}{rrrrr}1.082 & 1.099 & 1.013 & .807 \\ .993 & 1.000 & 1.004 & 1.003\end{array}$ |  |  | 1.0511 .0821 .017 .848 |  |  | $1.0251 .0691 .019 \quad .885$ |  |  |
| Class-1,2,3,4 |  |  |  | . 990 . 9991.0051 .006 |  |  | .961 1.040 |  |  |
| Sex-M,F | .9931 .0001 .0041 .003.9781 .022 |  |  | $.9671 .033$ |  |  |  |  |  |
| Cause-A,S | . 9121.078 |  |  | . 9131.078 |  |  | . 9191.072 |  |  |
| Age 45: | 1.019 |  |  | 1.024 |  |  | 1.022 |  |  |
| EP-0, $7,14,30$ | 1.1431 .057 | . 983 | . 818 | 1.1131 .046 | . 990 | . 851 | 1.0831 .040 | . 995 | . 882 |
| Class-1,2,3,4 | 1.0061 .000 | . 997 | . 995 | 1.004 .999 | . 998 | . 998 | $1.002 \quad .999$ | . 999 | . 000 |
| Sex-M,F | . 9691.031 |  |  | . 9591.042 |  |  | . 9511.050 |  |  |
| Cause-A,S | . 9821.001 |  |  | . 9811.003 |  |  | . 986.999 |  |  |
| Age 55: | . 961 |  |  | . 957 |  |  | . 953 |  |  |
| EP-0, $7,14,30$ | 1.1711 .031 | . 957 | . 841 | 1.1471 .021 | . 964 | . 869 | 1.1211 .013 | . 967 | . 900 |
| Class-1,2,3,4 | 1.0231 .009 | . 991 | . 976 | 1.0201 .007 | . 993 |  | 1.0191 .005 | . 993 | . 982 |
| Sex-M,F | . 9841.016 |  |  | . 9761.024 |  |  | . 9661.034 |  |  |
| Cause-A,S | 1.089 .902 |  |  | $1.084 \quad .908$ |  |  | $1.082 \quad .910$ |  |  |
| Age 62: | . 890 |  |  | . 874 |  |  | .87! |  |  |
| EP-0,7.14.30 | 1.1801 .024 | . 917 | . 876 | 1.1661 .010 | . 919 | . 907 | 1.147 . 987 | . 917 | . 951 |
| Class-1,2,3,4 | 1.0581 .033 | . 980 | . 933 | 1.0481 .028 | . 982 | . 944 | 1.0431 .022 | . 984 | . 953 |
| Sex-M,F | $\begin{array}{lll}1.025 & .975\end{array}$ |  |  | 1.024 .976 |  |  | 1.008 .991 |  |  |
| Cause--A,S | $11.245 \quad .790$ |  |  | $1.236 \quad 796$ |  |  | $1.223 \quad .806$ |  |  |

*The termination rate is the product of the Duration Rate and the corresponding variable factors for the respective Age, e.g., for Week 2 (.114), age 35 (.961), EP 7day (.934), class 2 (.999), male (1.190), accident ( 1.044 ), the termination rate is .127 .
**Age is age at disablement.
Duration is from the date of disablement.
Class 1 includes the 2 lowest premium occupation classes of a 5 -class manual, or the lowest premium class of a 4 -class manual.

EXHIBIT 3a-Continued
DTS Valuation Table
Factors for Calculation of Weekly Termination Rates*

| Factoms** | Week 11 | Week 12 | Weik 13 |
| :---: | :---: | :---: | :---: |
| Duration Rate | . 104 | . 094 | 082 |
| Age 25: | 1.096 | 1.110 | 1.133 |
| EP--0,7,14,30 | . 9631.0511 .018 . 964 | . 9961.008 . 9851.007 | $\begin{array}{lllll}1.059 & .949 & .9351 .050\end{array}$ |
| Class-1, 2, 3,4 | 966 . 9941.0151 .026 | . 957 . 9821.0171 .045 | . 944 . 9641.0211 .074 |
| Sex-M,F | .9901 .008 | . 9841.013 | . 9751.018 |
| Cause-A,S | . 8761.127 | . 8841.118 | . 8971.104 |
| Age 35: | 1.062 | 1.049 | 1.027 |
| EP-0.7,14,30 | $1.0031 .0581 .017 \quad .920$ | . 9851.0491 .008 . 955 | . 9711.038 . 9898 |
| Class-1,2,3,4 | . 981.9961 .0071 .015 | . 974 . 9941.0091 .002 | . 962 . 9931.0121 .032 |
| Sex-M,F | . 9581.042 | . 9591.039 | 9671.026 |
| Cause-A.S | .9301 .060 | . 9501.040 | 9841.006 |
| Age 45: | 1.012 | . 993 | . 962 |
| EP-0.7,14,30 | $1.048 \quad 1.039 \quad .998 \quad .914$ | 1.0071 .043 . 9997 | .952 1.054 . 989 . 995 |
| Class-1,2,3,4 | 1.0011 .0001 .000 .999 | $1.0001 .0031 .000 \quad .995$ | 1.0001 .0081 .001 .989 |
| Sex-M,F | . 9461.055 | . 9431.057 | . 9421.053 |
| Cause-A,S | . 998 . 989 | $1.020 \quad .969$ | $1.058 \quad .935$ |
| Age 55: | . 948 | . 941 | . 932 |
| EP-0,7,14,30 | $1.0901 .005 \quad .966$. 938 | $\begin{array}{llll}1.052 & .997 & .959 & .989\end{array}$ | . 999 . 988 . 9431.062 |
| Class-1, 2, 3,4 | $\begin{array}{lllll}1.022 & 1.006 & .992 \quad .980\end{array}$ | $\begin{array}{lllll}1.031 & 1.009 & .989 & .971\end{array}$ | 1.0481 .015 . 984 . 953 |
| Sex-M,F | . 9531.048 | . 9351.066 | . 9081.092 |
| Cause-A,S | 1.086 .909 | $1.094 \quad .904$ | $1.110 \quad .891$ |
| Age 62: | . 881 | . 907 | . 946 |
| EP- $0.7,14.30$ | $\begin{array}{llll}1.119 & .956 & .9131 .017\end{array}$ | $\begin{array}{llll}1.079 & .914 & .906 & 1.114\end{array}$ | $1.024 \quad .853 \quad .8941 .265$ |
| Class-1, 2, 3,4 | 1.0411 .016 . 984 . 961 | $\begin{array}{lllll}1.043 & 1.009 & .982 & .967\end{array}$ | $\begin{array}{lllll}1.052 & .998 & .978 & .972\end{array}$ |
| Sex-M,F | . 9751.024 | . 9201.083 | . 8441.175 |
| Cause-A.S | $1.210 \quad 816$ | $11.193 \quad 829$ | $11.166 \quad .849$ |

*The termination rate is the product of the Duration Rate and the corresponding variable factors for the respective Age, e.g., for Week 2 (.114), age 35 (.961), EP 7day (.934), class 2 (.999), male (1.190), accident (1.044), the termination rate is .127 .
${ }^{* *}$ Age is age at disablement
Duration is from the date of disablement.
Class 1 includes the 2 lowest premium occupation classes of a 5 -class manual, or the lowest premium class of a 4 -class manual.
$190$


[^3]
## EXHIBIT 3c

DTS VALUATION TABLE
factors for Calcllation of annual Termination Rates
years 3 through 10

| Year | 3 | 4 | 5 | 6 |
| :---: | :---: | :---: | :---: | :---: |
| Duration Rate | 123 | . 084 | . 062 | . 050 |
| Male: | 1.080 | 1.129 | 1.179 | 1.200 |
| Female: | . 920 | . 871 | . 821 | . 800 |
| Age: 25 | 2.085 | 1.832 | 1.554 | 1.262 |
| 35 | 1.164 | 1.103 | 1.017 | . 909 |
| 45 | . 727 | . 757 | . 767 | . 754 |
| 55 | . 536 | . 616 | . 697 | . 832 |
| 62 | . 489 | . 691 | 965 | 1.244 |
| Year | 7 | 8 | 9 | 10 |
| Duration Rate | . 045 | . 042 | . 042 | . 043 |
| Male: | 1.212 | 1.210 | 1.204 | 1.200 |
| Female: | . 788 | . 790 | . 796 | . 800 |
| Age: 25 | . 994 | . 776 | . 617 | . 524 |
| 35 | . 792 | . 696 | . 631 | 582 |
| 45 | . 741 | . 737 | . 739 | . 751 |
| 55 | . 984 | 1.103 | 1.182 | 1.226 |
| 62 | 1.489 | 1.688 | 1.830 | 1.918 |

## EXHIBIT 4

## DTS Valuation Table Ultimate Termination Rates for Duration 11 Years and Over <br> by Attained Age

| Altained Age | Male | Female | Attained Age | Male | Female |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 30 | . 0238 | . 0160 | 65 | . 0665 | . 0446 |
| 31 | . 0240 | . 0161 | 66 | . 0707 | . 0474 |
| 32 | . 0242 | . 0162 | 67 | . 0753 | . 0504 |
| 33 | . 0244 | . 0163 | 68 | . 0802 | . 0538 |
| 34 | . 0246 | . 0165 | 69 | . 0857 | . 0574 |
| 35 | . 0249 | . 0167 | 70 | . 0916 | . 0614 |
| 36 | . 0251 | . 0168 | 71 | . 0986 | . 0657 |
| 37 | . 0254 | . 0170 | 72 | . 1051 | . 0704 |
| 38 | . 0258 | . 0173 | 73 | . 1127 | . 0755 |
| 39 | . 0261 | . 0175 | 74 | . 1210 | . 0811 |
| 40 | . 0265 | . 0178 | 75 | . 1301 | . 0871 |
| 41 | . 0270 | . 0181 | 76 | . 1398 | . 0937 |
| 42 | . 0275 | . 0184 | 77 | . 1504 | . 1008 |
| 43 | . 0280 | . 0188 | 78 | . 1619 | . 1085 |
| 44 | . 0286 | . 0192 | 79 | . 1743 | . 1168 |
| 45 | . 0292 | . 0196 | 80 | . 1878 | . 1258 |
| 46 | . 0299 | . 0200 | 81 | . 2022 | . 1355 |
| 47 | . 0306 | . 0205 | 82 | . 2178 | . 1459 |
| 48 | . 0315 | . 0211 | 83 | . 2345 | . 1571 |
| 49 | . 0324 | . 0217 | 84 | . 2525 | . 1691 |
| 50 | . 0334 | . 0224 | 85 | . 2717 | . 1820 |
| 51 | . 0345 | . 0231 | 86 | . 2922 | . 1958 |
| 52 | . 0357 | . 0239 | 87 | . 3140 | . 2104 |
| 53 | . 0370 | . 0248 | 88 | . 3372 | 2259 |
| 54 | . 0384 | . 0257 | 89 | . 3618 | . 2424 |
| 55 | . 0400 | . 0268 | 90 | . 3877 | . 2598 |
| 56 | . 0417 | . 0279 | 91 | . 4149 | . 2780 |
| 57 | . 0436 | . 0292 | 92 | . 4435 | . 2971 |
| 58 | . 0456 | . 0306 | 93 | . 4732 | . 3171 |
| 59 | . 0479 | . 0321 | 94 | . 5041 | . 3378 |
| 60 | . 0503 | . 0337 | 95 | . 5360 | . 3591 |
| 61 | . 0530 | . 0355 | 96 | . 5686 | . 3801 |
| 62 | . 0559 | . 0375 | 97 | . 6020 | . 4033 |
| 63 | . 0592 | . 0397 | 98 | . 6357 | . 4259 |
| 64 | . 0627 | . 0420 | 99 | . 6695 | . 4486 |

implicit margin inherent in the method of construction, the parametric approach of the DTS Valuation Table, and the explicit front-end margin in rates of termination from claim. Valuation interest rates then need to be updated as well, and a change in interest rates is being reviewed by an Academy of Actuaries' committee that will recommend model reserve standards to the NAIC.

Although the DTS Valuation Table is reasonably representative of group LTD experience after the first 2 years of claim, a consensus was reached during the exposure period that additional margin, varying by age, would be needed for a group LTD valuation standard. Accordingly, the Committee is not proposing a valuation table for group LTD at this time.

Although the proposed DTS Valuation Table would be the individual disability income valuation standard, there may be blocks of business where the resulting reserves are inadequate under prudent management. In such a case, as has been the situation historically, the company must hold adequate liabilities. Termination rates on the proposed DTS Valuation Table are easily modified to handle that contingency on a very flexible, practical, and even seriatim, basis.

Each company actuary should be required periodically, in any event, to demonstrate to the Commissioner's satisfaction that the reserves held are reasonable and adequate for each unique policy form.

The termination rate for any duration is the product of the duration rate and the appropriate factor from each set of factors for that duration.

The values for the individual ages were obtained by the Lagrange interpolation formula shown in Exhibit 5.

Age is age at disablement.
Duration is measured from the date of disablement.
Occupation Class 1 includes the 2 lowest premium classes of a 5 -class manual or the lowest premium class of a 4 -class manual.

Although many people contributed substantial amounts of time, the Committee wishes to express its special appreciation to Mr. Frank Knorr and Mr. John Haynes Miller for their very capable and most extensive effort in support of the work of this assignment.

The Committee also wishes to formally acknowledge with appreciation, the assistance received from Mr. Frank O’Grady, from Mr. Tappan Roy, and from the New York Insurance Department for making the results of their study available to us.

## EXHIBIT 5

5-POINT LAGRANGE INTERPOLATION FORMULA
Used for incidence rates and termination rates.
Given points $F(a), F(b), F(c), F(d)$, and $F(e)$, then:

$$
\begin{aligned}
F(x) & =\frac{(x-b)}{(a-b)} \frac{(x-c)}{(a-c)} \frac{(x-d)}{(a-d)} \frac{(x-e)}{(a-e)} F(a) \\
& +\frac{(x-a)}{(b-a)} \frac{(x-c)}{(b-c)} \frac{(x-d)}{(b-d)} \frac{(x-e)}{(b-e)} F(b) \\
& + \\
& \frac{(x-a)}{(e-a)} \frac{(x-b)}{(e-b)} \frac{(x-c)}{(e-c)} \frac{(x-d)}{(e-d)} F(e)
\end{aligned}
$$

for $a<x<e$,
$a, b, c, d$, and $e$ are ages $25,35,45,55$, and 62 , respectively.
When $x \leq 25$ :
for incidence rates, $\quad F(x)=F(25)$
for termination rates, $F(x)=F(25)+(25-x)[F(25)-F(26)]$
When $x \geq 62$ :
$F(x)=F(62)+(x-62)[F(62)-F(61)]$.

## APPENDIX A

## COLLECTION AND EDITING OF THE DATA USED IN DEVELOPING TERMINATION RATES

The original solicitation of data to be used in developing a table of disability termination rates was made in 1977 by John H. Miller through his Disability Newsletter. The data requested were records for each disability claim which either terminated in 1975 or 1976 or was outstanding at the end of 1976. A number of companies contributed data in response to this solicitation.

Subsequent to the formation in 1978 of the Committee to Recommend New Disability Tables for Valuation, John Miller obtained approval from all contributing companies to tum the data he had collected from them over to the Committee. During the next several years, the Committee solicited contributions from additional companies, as well as requested and received additional years of experience from many of the original contributors. The extent of the data contributed is shown in the following table.

| Contributing |  | Number | of Cla | TER | ated | CA | R YE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Company | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 | Toals |
| American Mutual | 0 | 0 | 245 | 280 | 0 | 0 | 0 | 525 |
| Durham Life | 0 | 0 | 0 | 78 | 0 | 0 | 0 | 78 |
| Franklin Life | 0 | 0 | 816 | 880 | 0 | 0 | 0 | 1,696 |
| IDS Life | 0 | 0 | 1,066 | 1.050 | 840 | 781 | 0 | 3.737 |
| John Hancock | 3,746 | 3,312 | 3,313 | 2,758 | 2,558 | 2.344 | 2,089 | 20,120 |
| Life of Virginia | 0 | 0 | 129 | 116 | 157 | 296 | 0 | 698 |
| Mass. Casualty | 0 | 0 | 0 | 2,449 | 0 | 0 | 0 | 2.449 |
| Mass. Indemnity | 0 | 0 | 4,137 | 3,629 | 0 | 0 | 0 | 7,766 |
| Mass. Mutual | 0 | 0 | 263 | 254 | 322 | 322 | 0 | 1,161 |
| Metropolitan | 0 | 0 | 5,847 | 6,175 | 5,694 | 0 | 0 | 17,716 |
| Monarch | 0 | 0 | 8,978 | 8,230 | 7,908 | 7,185 | 6,642 | 38.943 |
| Mutual of Omaha* | 1.261 | 1.334 | 1,379 | 1,333 | 1.256 | 0 | 0 | 6,563 |
| Northwestern Mut. | 0 | 0 | 260 | 295 | 0 | 0 | 0 | 555 |
| Provident L \& A | 0 | 0 | 1,726 | 2,207 | 2,319 | 2.151 |  | 8.403 |
| Provident Mutual | 0 | 0 | 0 | 450 | 0 | 0 | 0 | 450 |
| Prudential | 0 | 0 | 9,842 | 3,806 | 0 | 0 | 0 | 13,648 |
| State Mutual | 0 | 0 | 1.139 | 1.072 | 0 | 0 | 0 | 2.211 |
| Travelers* | 0 | 0 | 187 | 85 | 0 | 0 | 0 | 272 |
| Washington Nat. | 0 | 0 | 2,906 | 3.013 | 0 | 0 | 0 | 5,919 |
| Woodman A \& L | 0 | 0 | 0 | 1.026 | 0 | 0 | 0 | 1,026 |
| Totals: | 5,007 | 4.646 | 42.233 | 39.186 | 21,054 | 13.079 | 8.731 | 133.936 |
| *Second year and later data only |  |  |  |  |  |  |  |  |

The specifications of the data to be included and the format of the data are contained in the instructions furnished to each contributing company. These instructions are reproduced below.

INTERCOMPANY DISABILITY TERMINATION TABLE INSTRUCTIONS (Original Instructions of May 1977, as edited December 1979)

Data Specifications. For the development of disability termination rates the following data are required:

1. A claim record for each claim for total disability under coverage through an individual policy providing both accident and sickness total disability benefits. (The original instructions indicated that accident-only business could be submitted at the option of the contributor, but such contributions as were received have been excluded from the processing of the data.)
2. A separate record for:
a. Each claim terminated by death, recovery, or expiration of benefit period or of coverage in each year of observation.
b. Each claim outstanding at the end of each period of observation.
3. Partial Disability will be excluded.

## FORM OF DATA

The data should be submitted on 80 -column punch cards or on magnetic tape, using the following outlined format. The following format specifications are very similar to those outlined in the TSA 1959 Reports pages 15663 , as well as to those required for the 1975 submission of disability experience to the New York Insurance Department for its analysis of Disability Income Insurance Cost Differentials between Men and Women. However, in this study each claim requires only one record-neither summary cards nor exposure cards are necessary.

## CLAIM RECORD FORMAT

| $\frac{\text { Field }}{1}$ | $\frac{\text { Columns }}{1}$ | $\frac{\text { Description of Field }}{\text { skip }}$ |
| :--- | :---: | :--- |
| 2 | 2 | Last digit of calendar year of observation |
| (see field 30) |  |  |

3 3-5 Company Code Number.
46 Type of coverage or cause of disability: 3 for Accident, 4 for Sickness. ( 1 was used for Accident Only and 2 designates Sickness under a Sickness Only policynow a rarity.)

5 7-9 (optional) Contributing company's policy form code.

6
10

| Code |
| :---: |
| 1 |
| 2 |
| 3 |
| 4 |
| 5 |

6

7

Type of Renewal Provision
Renewal Type
Noncancelable
Guaranteed renewable
Nonrenewable for stated reasons only.
Collectively renewable
Level premium policies not included in 1-4 above.
Step rate policies not included in 1-4 above.
Other policies. Please explain the renewal conditions.

7 11-12 Age at expiration of coverage. Record the limiting age of coverage specified in the policy contract, even though it may be continued by company policy, to some more advanced age. If there is no expiry age specified in the policy, punch 99.
13 Sex: Men =1; Women = 2; Combined $=5$.
$9 \quad 14$ Occupational class. Please code from 0 to 7 according to the following table.

| Code | 4-class manual | 5-class manual | "Bureau Manual' | "NY" class code* |
| :---: | :---: | :---: | :---: | :---: |
| 0 |  | 4A |  | 1 |
| 1 |  | 3A |  | 1 |
| 2 | 3A |  | A \& B | , |
| 3 | 2A | 2A | C \& D | 2 |
| 4 | A | A | D* \& E | 3 |
| 5 | B | B | F \& G | 4 |
| 6 | C | C | H \& I | 4 |
| 7 | $\mathrm{D}+$ over | + over | J \& over | 4 |

9 Separation by class not available
If your classification system does not approximate one of the four above groupings, please send an explanation which will enable us to determine appropriate codes.
*Codes used for the New York Study, indicating our impression of the typical correspondence with other designations displayed above.
Definition of Occupation Classes:
Class 1: the lowest premium class, includes professional, technical and managerial occupations that are generally office duties only.

Class 2: includes supervisory and other skilled clerical and skilled technical people.
Class 3: nonhazardous light manual workers.
Class 4: hazardous work with heavy manual labor or using heavy equipment.
10 15-17 Elimination period in days for sickness benefits. This may be left blank if the cause is accident.
11 18-20 Elimination period in days for accident benefits. This may be left blank if same as sickness or if the cause is sickness.

13 24-26 Benefit period, accident. Enter maximum number of months for which total disability accident benefits are payable under the terms of the policy. Code 965 for benefits to age 65; 999 for lifetime benefits. This may be left blank if the same as sickness or if the cause is sickness.
27 skip
15 28-29 Attained age at disablement. Age nearest birthday at date of disablement.
16 30-31 (optional) Year of issue. Enter last two digits of the year of issue.
17
21-23 Benefit period, sickness. Enter the maximum number of months for which total disability sickness benefits are payable under the terms of the policy. Code 965 for benefits to age 65; 999 for lifetime benefits. This may be left blank if the cause is accident.

32-34 skip
35-38 Amount of monthly indemnity or $4.35 \times$ weekly indemnity, to nearer dollar.
39 Status of claim
Code
0 -claim open at end of observation period
1-claim terminated by recovery
2-claim terminated by death
3-claim terminated because benefit period was exhausted.

Note re: Status of claim, cause of termination. If data source does not distinguish between deaths and recoveries, terminations by either of
these causes may be coded: 4-claim terminated by recovery or death.

2040 (optional Month incurred. Enter calendar month of in-
(see field 31) curral (Jan. $=1$, Feb $=2, \ldots$, Sept $=9$, Oct $=0$, Nov. $=x$, Dec. $=\mathrm{y}$ ).
21 41-42 Enter the last two digits of the calendar year of incurral.

46-47 skip
24 48-52 Duration of disability. Show the duration in days for which total disability benefits were incurred (i.e. measured from the end of the elimination period). On open claims, show the duration for which total disability payments were incurred up to December 31 of the year of observation. (Estimates of future durations on open claims should be excluded.) For recurrent disabilities, follow the policy contract and report the total number of days for which benefits were incurred under the claim. In cases settled by legal judgment, arbitration, or compromise, compute an "adjusted duration of disability" which, when multiplied by the rate of disability indemnity, will reproduce the amount of the settlement.

Note re: Duration of disability. If the record for open claims shows only "last transaction date," or similar information, the days of duration to the year-end should be: Duration in days = December 31 , calendar year of observation - (date incurred) - (elimination period). If the record includes the number of days of disability up to the last transaction date, just increase it by the number of days between the last transaction and December 31 in order to obtain the duration.

53-57 (optional) Diagnosis of disability.
26 58-62 (optional) Impairment code. Companies that maintain records as to the types of impairment riders added to policies are encouraged to furnish this information. Companies coding this information would supply a copy of their code. $01,02,05,08$ ( 100 months), 10 . If to age 60 or 65 , code 60 or 65 , respectively. If no limit, enter 99 . If no provision for disability from 'his occupation,' enter 00.

Column 65: Indemnity provision
$0=$ no benefits payable if insured has earnings from a new occupation for which he is reasonably fitted by education, training, and experience.
1 = pro rata for earnings in his new occupation.
$2=$ pro rata for earnings in his regular occupation.
$3=1$ for 1 offset in earnings.
$4=2$ for 3 offset in earnings.
$5=1$ for 2 offset in earnings.
$6=$ claimant's option of either his occupation to 65 without reduction, or residual benefit.
7 = other-please define.
$9=$ no reduction in indemnity payable and no offset by reason of earnings in new occupation.

Note re: Definition of disability. If the determination of this information for each claim poses a major problem, a code appropriate for at least 90 percent of all claims, or 90 percent of all claims in each major category, may be used for all such cases. However, this alternative should not be employed if the actuary responsible for the submission believes it might result in a significant error in termination rates applicable to a particular definition, for which there is a credible volume of experience.

66-73 Claim identification number. This number will provide a means of reference to follow up inconsistencies and correct errors.
2974 Indicate whether policy is standard or substandard $0=$ standard $1=$ substandard
$2=$ substandard cases included but not identifiable $3=$ no substandard policies issued
75-77 Month and day claim was closed. Together with field 2 , this will provide the complete date of disability termination, or

31 78-79 Day claim was incurred. Together with fields 20 and 21 this will provide the complete date of incurral. As an alternative, field 31 together with fields 20 and 21 will suffice as a substitute for field 30 . Note: The separation of fields 30 and 31 from fields 2 and 20 , respectively, results from omitting the information for these fields in the original specifications.

Claims to be excluded: These are all cases where a claim has not been admitted by insurer but include, as outstanding, cases on which one or more payments have been made if policyholder is now contesting termination.

Successive or recurrent periods of disability: The following illustrates the entry for claim duration: Policyholder with three-month elimination period became totally disabled $1 / 1 / 74$. Disability terminated $1 / 1 / 75$. The duration of disability would be 9 months (expressed in days, field 24 ). If disability recurs $5 / 1 / 75$, original claim is reopened and then terminated $11 / 1 / 75$, the duration would be 15 months (original 9 plus additional 6). If claim is reopened again on 5/1/76 and remains open on $12 / 31 / 76$, the elapsed duration of disability (field 24) will be 23 months. If the insurer treated these three disability periods as three different claims, they should be so reported for this study.

Editing the Data: All of the claim records for each contributing company were processed through an edit program that tested the various fields of the record for valid data. A record that contained invalid data was rejected and printed on an error list. The error list was referred to the contributing company for review and correction. Most of the records that originally contained errors have been corrected and passed successfully through this edit program.

Exhibit A-1 shows the format of the output record containing the edited data. Acceptable values for each edited field are indicated. It will be noted that in a few cases the acceptable values were translated into a simple code to facilitate further processing.

Some special routines were used in creating certain of the output fields. A detailed description of these routines is given below.

## 1. Elimination Period and Benefit Period

The appropriate accident or sickness periods were selected depending upon whether the type was coded accident or sickness. If the type was coded "unknown" the sickness periods were used. However, if a further test showed that the duration of disability exceeded the sickness benefit period and also the accident benefit period exceeded the sickness benefit period, then the accident benefit period was used.

```
EXHIBIT A-1
FORMAT OF EDITED DATA
(Logical Record Length of 51, all fields in Integer form)
```



1 Sex 1-male 2-female
3 Benefit period in Months or 965-to age 65
999-Lifetime
1 Type 3-accident, 4-sickness, 5-unknown
1 Renewal Provision 1-Non Can
2-G.R.
3-Non Ren. for stated reasons only
4-C.R.
5-all other
0,9-Unknown
1 Impairment 0-standard, 1-substandard, 3-unknown
2 His. Occ. Period in years or
55-to age 55
65-to age 65
99-Lifetime
1 Indemnity Provision
0 -complete reduction
1-pro rata for new oce.
2-pro rata for regular occ.
5-1 for 2 offset
7-others
9-no reduction

$$
\begin{aligned}
& 1-1975,1976 \text { combined } \\
& 3-1973 \\
& 4-1974 \\
& 5-1975 \\
& 6-1976 \\
& 7-1977 \\
& 8-1978 \\
& C-1975,1976,1977,1978 \text { combined }
\end{aligned}
$$

| 23 | 6 | Termination Date |
| :--- | :--- | :--- |
| 29 | 1 | Status on Termination Date | | 0-open |
| :--- |
| 1-recovery |
| 2-death |
| 3-exhausted |
| 4-death or recovery |

## 2. Termination Coding

Several companies were able to code the cause of claim terminations as deaths or recoveries. In order to be able to use this information in subsequent processing, a special field was created to indicate the extent of termination coding.

## 3. Year Claim Closed

The original input coding instructions requested that the year closed be given in column 2 of the record and that, for claims open on December 31, this field be left blank. This proved unnecessary since the open claims could be detected by a claim status of 0 . Accordingly, the appropriate year code was included in each output record whether the claim was closed or open.

In the original submission, several companies included records for claims closed in both 1975 and 1976 but only for those open on December 31, 1976. In subsequent processing, it would be necessary to treat these claims differently from those of companies that handled each year separately. A special code was used to indicate the claims where the exposure period could extend over the two years.

## 4. Days of Disability

For other than lifetime benefit periods, the duration of disability was tested against the length of the benefit period. The record was rejected if the duration of disability exceeded the benefit period (except for those coded "exhausted" that fell within the permissible range as explained below).

The benefit period in months was converted to days using either a $360-$ day year or a 365 -day year. Some companies appeared to use one measure and some the other, so the appropriate one was selected.

## 5. Status of Claim

Several tests were performed on the relationship between the claim status
code, the benefit period, and the duration of disability. The record was rejected if the claim status showed a death or recovery, but the duration of disability equaled the benefit period. The record also was rejected if the claim was coded a limit claim, but the duration did not equal the benefit period. A small leeway was permitted here, from 29.8 to 30.7 times the number of months in the benefit period. If the duration fell within this range, it was arbitrarily set equal to the days in the benefit period.

## 6. Termination Date and Date of Disability

In order to calculate the exposure period for each claim, it is necessary to know these two dates. Most records did not contain either date, although some did contain one or the other or possibly the month and year incurred. The following routine was used to develop both dates based on the information available:

1. If the claim status indicated an open claim, the termination date was set to $12 / 31$ of the exposure year. The date of disability was calculated as the termination date minus the duration of disability.
2. For closed claims, if the month closed was given, the termination date was established based on the date incurred, if given. If only the month and year incurred were given, the day was arbitrarily made 1 . The date of disability was calculated as the date incurred plus the elimination period, and the termination date was calculated as the date of disability plus the duration of disability. The resulting termination date was checked against the year-claim-closed field since these years should be the same. If the termination date exceeded the year closed, the dates were adjusted as follows. The termination date was set to $12 / 30$ of the exposure year, and the date of disability recalculated to equal the termination date minus the duration of disability.
3. When the termination year was prior to the year closed, the following additional procedures were performed. If the day of the month incurred was not given but had been arbitrarily made 1 , it was changed to 30 and the comparison made again. This may have advanced the calculated termination year to be equal to the year-claimclosed field. If not, this claim was rejected.
4. If neither the termination date nor the date incurred were given, the termination date was arbitrarily set to $12 / 30$ of the exposure year, and the date of disability calculated as the termination date minus the duration of disability.
5. As a further validation of each company's data, the termination rates for each company were compared to the termination rates for all companies combined. A preliminary table of termination rates was constructed from the data for all companies. This table was used as the basis for expected terminations. Ratios of actual to expected terminations were calculated for every value of the input categories being analyzed. Each contributing company was asked to review the actual to expected ratios based on its data and compare them to the ratios for all companies combined. As a result of this review, several obvious errors and inconsistencies were detected and corrected.

## APPENDIX B

## DEVELOPMENT OF TERMINATION RATES

## Introduction

After termination experience data were solicited, collected, and edited, the data were analyzed, and termination rates were developed. The purpose of this Appendix B is to explain the analysis and development of the smooth weekly, monthly, and annual termination rates.

The following is an outline of the major steps in this process:
I. Summarization of Edited Data into Usable Form
a. Reformat File
b. Summarize Reformatted File
II. Analysis of Data
a. Approach to Problem
b. Identify Significant Variables
c. Identify Significant Interactions
d. Determine the Best Model
e. Analyze Contingency Table
III. Graduated Termination Rates
IV. Weekly Termination Rates
V. Ultimate Termination Rates

Steps I, II, and III were completed using monthly and annual data with the intention of replacing the first three months with weekly data for the first 13 weeks in Step IV. Step V was done in parallel with the other work and actually was completed first.

## Summarization of Edited Data into Usable Form

The overall objective of this summarization was to produce the number of (a) terminations from disability and (b) exposures to termination in contingency table form so that they could be analyzed using Everyman's Contingency Table Analysis (ECTA). An explanation of the ECTA Program is given on page 484. The collection and editing of the data for analysis are described in Appendix A. These edited data with generally one record per claim were reformatted to have one record per exposure month. Then, the data were summarized into a number of different arrays, i.e., contingency tables. The values in the arrays represent either terminations or exposures. The positions of the values in the arrays represent the variables that may
affect the termination from disability. The ECTA program then reads the arrays and analyzes the effects of the variables.

## REFORMAT FILE

The input data are in the form described in Exhibit A-1 Appendix A. This is the coded output from the editing procedures described in Appendix A. The data represent all the disability claims that terminated during a particular experience year (or observation period) plus all the claims that were still outstanding at the end of the period. There is one record per claim with the following exceptions:

1. If a company contributed data separately for different years of experience and a single claim was observed in two or more of the years, the result would be more than one record for that claim. The exposure months, however, do not overlap.
2. If a claimant has more than one policy with different contributing companies (or with the same contributing company, and the company contributes data separately for different policies), then there would be more than one record for that claimant. In this case, exposure months would overlap.
In addition to the fields in Exhibit A-1, there was a field for Age to Expiration of Coverage. However, there was widespread confusion over its meaning. No attempt was made to correct the data in this field because it was agreed that such an age would have little or no effect on termination rates.

The editing procedures described in Appendix A did not detect all errors in the input. In a few cases, the edit program had assigned incorrect codes. Before creating an output record, the following situations had to be correct:

1. One company contributed data observed in 1975 and 1976 combined only for those claims with dates of disablement in 1975 and data observed in 1976 for all other claims. However, one "Experience Year Code" (1-1975. 1976 combined) was found on all the records for that company.
2. Another company had "His Occ. Period" and "Indemnity Provision" miscoded.
3. Another company had "His Occ. Period" and "Indemnity Provision," and "Renewal Provision" miscoded.
4. The editing program miscoded some termination dates.

One company's data were in error and could not be corrected. Thus, all of that company's data were ignored.

Any record that had a date of termination before or the same day as the date of disability was ignored. In addition, there was only one record with the sex coded as unknown. This record was ignored.

For every input record used, the number of complete months of exposure was calculated, and one output record was written for every month of exposure. The format of the output records is shown in Exhibit B-1.

## EXHIBIT B-1

REFORMATTED DATA FOR PRODUCING MONTHLY TERMINATION RATES (Logical Record Length of 50, all fields in Binary form)

| Beginning Position | Field Length |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Description |  |  |  |
| 1 | 2 | Elimination Period Code | 1-0 day | 4-30 day | 7-180 |
|  |  |  | 2-7 day | 5-60 | 8-360 |
|  |  |  | 3-14 day | 6-90 |  |
| 3 | 2 | Age at Disablement | 1-20-24 | 5-40-44 | 9-60-64 |
|  |  |  | 2-25-29 | 6-45-49 | 10-65-69 |
|  |  |  | 3-30-34 | 7-50-54 | 11-70-74 |
|  |  |  | 4-35-39 | 8-55-59 | 12-75-79 |

2 Sex l-male 2-female

9

2 Impairment 1-standard, 2-substandard, 3-unknown
2 Benefit Period

$$
1-1-12
$$ months 3-25-60 months 2-13-24 4-to age $65 \quad 6$-Other months

2
Days to Expiration of Benefit Period

2 Type 1-accident, 2-sickness, 3-unknown
2 Renewal Provision 1-unknown 2-Non Can 5-C.R. 6-All Other
$1-0 \mathrm{yrs}$.
$2-1 \mathrm{yr}$.
$3-2 \mathrm{yrs}$.
Etc.

12-to age 55
14-to age 65
15--Lifetime

8-others 9-no reduction

2 Occ. Class Code
$\left.\begin{array}{cccccc}\text { ade } & \begin{array}{c}4 \text { Class } \\ \text { Manual }\end{array} & \begin{array}{c}5 \text { Class } \\ \text { Manual }\end{array} & \begin{array}{c}\text { Bureau } \\ \text { Manual }\end{array} & \begin{array}{c}\text { NY Class } \\ \text { Code }\end{array} \\ 2 & & & 3 \mathrm{~A}\end{array}\right)$

1-30-59 days 2-60-89 3-90-119 4-120-149 5-150(157.164)
\& greater

4-Nonrenewable for stated reasons only
reduction
2-pro rata for new oce.
3-pro rata for reg.
occ.
6-1 for 2 offset

| 25 | 2 | $\begin{array}{ll}\text { Experience } & \text { 1-1973 } \\ \text { Year } & 2-1974\end{array}$ | $\begin{array}{ll} 3-1975 & 5-1977 \\ 4-1976 & 6-1978 \end{array}$ |
| :---: | :---: | :---: | :---: |
| 27 | 2 | Duration from Disablement | 1-month I ( 0 day EP) 27-month 24 <br> 2 -month 1 ( 7 day EP) 28 -year 3 <br> 3 -month 1 ( 14 day EP) 29 -year 4 <br> 5 -month 2 <br> 30 -year 5, etc. <br> 6 -month 3 , etc. |
| 29 31 | 2 2 | Company Code (1 through Monthly Indemnity Amount |  |
| 33 | 2 | Termination Indicator | 0-exposure month only <br> 1- exposure month and termination from death or recovery during month. |
| 35 | 2 | $\begin{aligned} \text { Status Code } & 0 \text {-open } \\ & 1 \text {-recovery } \\ & 2 \text {-death }\end{aligned}$ | 3-exhausted benefit period 4-death or recovery |
| 37 | 2 | Benefit period in months or | 965-to age 65, 999-Lifetime |

There was much concern over the terminations that occurred near the end of the benefit period. Some of the data indicated that there was some sort of "reverse selection period" which began a few months before benefits were scheduled to run out. Therefore, exposure months within 5 months from the end of the benefit period were specifically identified by the field "Time to Expiration of Benefit Period." This field may be considered the compliment of "Exposure Month" which measures from the date of disablement.

A dilemma arose with 7- and 14-day elimination periods. If exposure months were measured from date of disablement, then the months at the end of the benefit period would not be accurate. On the other hand, measuring exposure months from the date of disability was not desirable in the first few months of the benefit period. The solution was a combinatin of the two definitions and can best be understood with an illustration.

14-DAY ELIMINATION PERIOD, 12-MONTH BENEFIT PERIOD


If a termination occurred between exposure months 7 and 8 in this illustration, it was ignored and other codes were changed so that it would be treated as "exhausted" without creating any records for months $8,9,10$, and 11 (the last month of the benefit period was always ignored).

A point was defined for each claim, (and for 7 - and 14-day elimination periods, each exposure month), as the date from which the exposure was to be measured.

Exposure months were then measured in two ways:

1. From the date of disablement. This was the way almost every exposure month was defined. This way of measuring exposure month is used when either the elimination period was $0,1,2,3,6$, or 12 months or the end of the exposure month is more than 5 months from the end of the benefit period. In this case. DMD is defined to be the "Day of the Month of Disablement."
2. From the date of disability. This was used only when both the elimination period was 7 or 14 days, and the end of the exposure month is within 5 months of the end of the benefit period. In this case, DMD is defined to be the "Day of the Month of Disability."

All calculations assume 30 days in every month. The calculations result in a month being included as an exposure month if:

1. the beginning of the month was in the experience period (the month begins on the first day after DMD):
2. the end of the month (occurring on DMD) was within the experience period;
3. the termination date had not occurred before the beginning of the month:
4. the end of the month was not the end of the expiration of the benefit period or later:
5. the end of the month was after the end of the elimination period; and
6. for two companies that contributed data only for experience observed after the first year of disablement, the end of the month was more than 12 months after disablement.
Similarly, the calculation results in a month being included as a termination if it is included as an exposure month, and the termination occurs during that month.

Some output fields will be explained here in more detail:
Days to Expiration of Benefit Period. For the values 1, 2, 3, and 4. the value indicates the number of months between the end of the month of disablement (occurring on DMD) and the end of the benefit period. The end of the benefit period was calculated using the "Benefit Period in Months" from the input record and was determined as exactly that number of months after the date of disability. The value of 5 indicates that the end of the month of disablement was 5 or more months before the end of the benefit period. "To Age 65 and Lifetime" benefits were all coded with a 5 in this field. It should be pointed out that if the elimination period was 7 or 14 days, the definition of DMD is different when this value is 5 than when it is 4 . Therefore, there may be 7 or 14 days between the end of the month when this value is 5 and the beginning of the month when it is 4 . If a termination occurred during the 7 or 14 days that are between months, then the termination was ignored by changing the Status Code to 3 (exhausted), and no exposure record was written for
month 4 . Any termination due to recovery or death after the month when this value is 1 was also ignored by changing the Status Code to 3 . That is, no records were written for experience within 30 days of the end of the "Benefit Period."

Experience Year. This indicates the year of the day at the beginning of the month of disablement.

Duration from Disablement. During the first month following disablement (month 1), only claims with zero-day elimination periods can be exposed to terminate for a complete month. These exposure months were identified with a 1 in this field. Claims with 7 - and 14 -day elimination periods can only be exposed to terminate for 23 or 16 days of the first month of disablement. However, these incomplete exposure months were still of some interest, so they were identified with 2 (month 1, 7 -day elimination period) and 3 (month 1, 14-day elimination period) in this field. There were no output records with 4 in this field. For the second through the twenty-fourth months, this field contained 3 plus the number of months from the date of disablement to the day (DMD) at the end of the month of exposure, which was almost always an integral number. When it was not an integral number, then any fractional parts of a month were ignored. For any month of exposure greater than twenty-four, this field contained 26 plus the number of years from the date of disablement to the day at the beginning of the month of exposure, where fractional parts of a year were ignored. If a claim was observed throughout its third year of disablement without a termination, then the output file would contain twelve records (one for each exposure month), all of them having a 28 in this field.

Termination Indicator. There was a value of zero in this field unless a termination due to death or recovery occurred during the month of disablement represented by the output record, in which case there was a value of one in this field. Therefore, the number of terminations could be determined by adding the values in this field, and the number of exposure months could be determined by counting the records.

Status Code. All of the output records that were generated from a single claim record (input record) had the same value in this field. The value was the same as that read from the input file or 3 (exhausted), if it had been changed (see the description of Days to Expiration of Benefit Period).

The output file contained about 870,000 records. It was split into four separate tape files to make further processing more efficient:

1. Records with duration for months 1,2 , and 3 .
2. Records with duration for months 4 through 24 .
3. Records with duration for years 3 through 10 .
4. Records with duration for years greater than 10 .

## SUMMARIZE REFORMATTED FILE

The next step in processing data into a contingency table required reading the proper exposure month tape, which was in the form described in Exhibit B-1. Note that again the output file of the prior step becomes the input file for the current step.

A number of different contingency tables were created because different factors needed to be analyzed. Construction of the contingency tables differed by:

1. The selection process
a. All contingency tables selected only data for ages 20 through 64 . Older ages were not used in any analysis.
b. Most analysis was done for a particular duration, so the records with the proper duration code had to be selected.
c. When company was a variable of interest in the anlaysis, the proper companies had to be selected.
2. Variables of interest
a. Initially all possible variables were of interest.
b. After the initial analysis, only the variables judged to be significant were of interest.
c. The variables of interest changed from one duration to another.

A contingency table is defined as a set of counts or frequencies obtained by classifying observations in two or more different ways. To illustrate this, a fairly simple contingency table is shown in Table B-1. A brief description of that contingency table is as follows:

1. Dimensions
2. Classifications (or variables)
3. Categories (or levels)
4. Counts (or values)

4
Elimination period, age, type, status
$0,7,14$, and 30 days (for elimination period), 20-39 and 40-64 (for age), accident and sickness (for type), and on (claim not terminated at the end of the month) and off (claim terminated before the end of the month)
Number of male claims observed during exposure month 2.

The contingency table is then a four dimensional array with 32 cells. The first dimension represents elimination period (which has 4 possible values), the second dimension represents age (with 2 possible values), the third di-

TABLE B-1
Illustrative 4-Dimensional Contingency Table

| elimination period |  |  |  | agit | rype | status |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0-day | 7-day | 14.day | 30-day |  |  |  |
| 50 | 65 | 55 | 50 | $<40$ \} | accident | (on) |
| 40 | 60 | 50 | 50 | $\geq 40$ ) |  |  |
| 3 | 50 | 40 | 60 | $\left.\begin{array}{l}<40 \\ \geq 40\end{array}\right\}$ | sickness, | terminated |
| 3 | 70 | 45 | 75 | $\geq 40$ ) |  |  |
| 40 | 45 | 40 | 30 | $<40$ ) | accident | (off) |
| 30 | 20 | 22 | 15 | $\geq 40$ f |  | terminated |
| 2 4 | 20 34 | 30 25 | 30 | $\left.\begin{array}{l}<40 \\ \geq 40\end{array}\right\}$ | sickness |  |

mension represents type (with 2 possible values), and the fourth dimension represents status. The cells represent all possible combinations of the values of the variables $(4 \times 2 \times 2 \times 2=32)$. The value 50 in the first cell represents the observed number of exposure months for 0 -day elimination period, age less than 40 , accident claims that did not terminate. The cell containing 22 in Table B-1 shows that there were 22 exposure months observed for a 14-day elimination period, age greater than or equal to 40 , accident claims in which terminations occurred. The observed data in our example do not include all the data because our selection process excluded (1) female data, (2) durations other than month 2 , (3) ages greater than 64 , and (4) data where the type was unknown.

The data in Table B-1 may also be used to calculate crude termination rates. For example, the crude termination rate for 30 -day elimination period, age less than 40 , accident claims is .375 or $30 \div(50+30)$. However, when termination rates were to be calculated, the form of the contingency tables was altered slightly by having the final variable be "exposed versus terminated" instead of "not terminated versus terminated." Contingency tables for the calculation of termination rates do not have the same requirements as those used to analyze the data. The latter requires that the cell counts be mutually exclusive. That is, an exposure month with a termination can only be assigned to one cell when the data are being analyzed. A violation of this approach would result in unreliable conclusions.

The actual creation of the contingency tables was a simple programming task. There were different programs for the different contingency tables. The IF statements determined the selection process, and an array set up in programs was defined by the variables of interest. Each contingency table was written into a disk file where it could be read and analyzed by the ECTA program.

## Analysis of Data

The objective of the data analysis was to determine the best form of a table of monthly and annual termination rates. This involved:

1. identification of significant variables.
2. identification of significant interactions among the variables,
3. determination of the best model, and
4. analyzation of the contingency table.

## APPROACH TO PROBLEM

The Committee spent a great deal of time experimenting with approaches to the first step, resulting in the selection of Contingency Table analysis. We then spent an even greater amount of time learning to use and modify the tool. The result was an approach which was used with slight variations for each of the next three steps. Consequently, most of the description will be of the approach and its application to identify the significant variables. With this as a foundation, the variations for identifying the most significant interactions and the best model are described.

Our objective was accomplished by analyzing a large number of runs of the ECTA program purchased from the University of Chicago (Department of Statistics) where it was developed under the direction of Professor Leo Goodman. Not every analysis technique incorporated in ECTA was used in this analysis; the ones that were used will be described here. The concept of modeling was used a great deal and also will be described.

The ECTA program reads a contingency table and develops another contingency table (array of numbers) with the same dimensions. The new contingency table contains the expected values of the cell counts under the proposed model. Unlike the original contingency table, the counts in the new table need not be integers. The new contingency table is similar to the original table in other ways, too. The similarities are defined by the model. The model tells which totals and subtotals in the new array must be the same as those in the original array. The totals and subtotals are identified by the number of its dimension. For example, if the original array has four dimensions, a 4 by 2 by 2 by 2 array, then ECTA will produce a new fourdimensional array with the same four dimensions. Since the last dimension has 2 possible values, we can divide all the cells into those that have a fourth-dimension value of "not terminated" or "terminated." Totals for each group of cells can be calculated by adding all the numbers in all the "on" cells and adding all the numbers in the "off" cells. For the array in Table B-1, these two totals are 768 and 424 , respectively. To have ECTA produce an array with the same totals for "on" and "off," a model of 4 must be specified, signifying that the totals for all the levels of the fourth

TABLE B-2
Two Sample Models

| Pari a-Moutio 4 |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ei.mination Period |  |  |  | Als | TYP1: | Stalls |
| O-day | 7 -day | 14-day | 30-day |  |  |  |
| 48 | 48 | 48 | 48 | $<40$ | accident | on |
| 48 | 48 | 48 | 48 | $\geq 40$ |  | (not |
| 48 | 48 | 48 | 48 | $<40$ | sickness | terminated) |
| 48 | 48 | 48 | 48 | $\geq 40$ |  |  |
| 26.5 | 26.5 | 26.5 | 26.5 |  |  |  |
| 26.5 | 26.5 | 26.5 | 26.5 | $\geq 40$ |  |  |
| 26.5 | 26.5 | 26.5 | 26.5 | $<40$ | sickness | (terminated) |
| 26.5 | 26.5 | 26.5 | 26.5 | $\geq 40$ |  |  |

Part b-Model:2

| Eimination Piriod |  |  |  | Agit | TYpe | Status |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| O-day | 7 -day | 14-day | 30-day |  |  |  |
| 38.25 | 38.25 | 38.25 | 38.25 | $<40$ | accident | on |
| 36.25 | 36.25 | 36.25 | 36.25 | $\geq 40$ |  |  |
| 38.25 | 38.25 | 38.25 | 38.25 | $<40$ | sickness | terminated) |
| 36.25 | 36.25 | 36.25 | 36.25 | $\geq 40$ |  |  |
| 38.25 | 38.25 | 38.25 | 38.25 | $<40$ | accident |  |
| 36.25 | 36.25 | 36.25 | 36.25 | $\geq 40$ |  |  |
| 38.25 | 38.25 | 38.25 | 38.25 | $<40$ | sickness | (terminated) |
| 36.25 | 36.25 | 36.25 | 36.25 | $\geq 40$ |  |  |

dimension must be preserved in the new array. When this is done with the given array, the result is the array in Table B-2 Part a. If the model of 2 had been specified, where we are concerned about preserving the totals for all the levels of the second dimension, then the result would be the array in Table B-2 Part b. In the new array, the totals of 612 for ages less than 40 and 580 for ages greater than 39 are the same as those in the original array.
In our analysis, we were interested in preserving termination rates. In the examples in Table B-2, the first model does not preserve termination rates for any of the individual cells although the overall termination rate of the entire set of data is preserved ( $424 \div 1,192$ ). The second model in Table B-2 does not preserve any termination rates at all; in fact, all the termination rates of the new contingency table in Table B-2 Part b are equal to .5. To preserve the termination rates of the two age groups, we must specify a model of $(2,4)$. This model will preserve the overall totals of the age groups ( 612 for young ages and 580 for old ages) and the overall totals of the ons and offs ( 768 for ons and 424 for offs) and will also preserve all the subtotals involving the age groups and on/off ( 375 for young ons, 393 for old ons, 237 for young offs, and 187 for old offs). Therefore, the termination rates for young ages ( $237 \div 612$ ) and for old ages ( $187 \div 580$ ) are preserved.

In the same manner, the model $(1,4)$ will preserve termination rates for each of the four elimination periods (dimension 1), and the model ( 3,4 ) will preserve termination rates for accident and sickness (dimension 3).

All of these models can also be combined in one model $(1,4),(2,4),(3$, 4). In basic terms this model preserves:

1. the overall totals of each elimination period,
2. the overall totals of the ons and offs,
3. the subtotals involving all the combinations of elimination period and on/off,
4. the overall totals of each age group,
5. the subtotals involving all combinations of ages and on/off.
6. the overall totals of accident and sickness, and
7. the subtotals of all combinations of accident/sickness and on/off.

Notice that this model does not preserve the values in each individual cell, only certain totals and subtotals. Likewise, every individual termination rate is not preserved, but the termination rates for each of the four elimination periods, for each age group, and for accident/sickness are preserved. However, termination rates for any combination of elimination period and age (or elimination period and accident/sickness or age and accident/sickness) are not necessarily preserved. An example of this is shown in Table B-3.

The model that preserves the value in each individual cell is called the

TABLE B-3
Two Models for Analysis of Monthly Termination Rates

| Part a-Model: (1,4), (2,4), (3,4), (1,2,3) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Elimination Perion |  |  |  | AGE | TYPE | states |
| 0-day | 7-day | 14-day | 30-day |  |  |  |
| 48.11 | 71.19 | 56.40 | 52.17 | $<40$ | $\left.\begin{array}{c}\text { accident } \\ \text { sickness }\end{array}\right\}$ <br> $\left.\begin{array}{l}\text { accident } \\ \text { sickness }\end{array}\right\}$ | on |
| 42.08 | 56.52 | 47.32 | 46.21 | $\geq 40$, |  |  |
| 3.67 | 44.67 | 40.89 | 57.89 | $<40$ |  |  |
| 4.14 | 72.62 | 45.38 | 78.72 | $\geq 40$, |  |  |
| 41.89 | 38.81 | 38.60 | 27.83 | $<40$ |  | off |
| 27.92 | 23.48 | 24.68 | 18.79 | $\geq 40$ \} |  |  |
| 3.33 | 25.33 | 29.11 | 32.11 | $<40$ |  |  |
| 2.86 | 31.38 | 24.62 | 33.28 | $\geq 40$, |  |  |
| Part b-Monel: (1.2.4), (3,4), (1.2.3) |  |  |  |  |  |  |
| Eimination perioin |  |  |  | Age |  | starus |
| 0 -day | 7-day | 14-day | 30-day |  |  |  |
| 51.10 | 70.71 | 55.13 | 52.19 | $<40$ | $\left.\begin{array}{l} \text { accident } \\ \text { sickness } \end{array}\right\}$ | on |
| 39.16 | 56.93 | 48.51 | 46.28 | $\geq 40$ |  |  |
| 3.90 | 44.29 | 39.87 | 57.81 | $<40$ |  |  |
| 3.84 | 73.07 | 46.49 | 78.72 | $\geq 40$ |  |  |
| 38.90 | 39.29 | 39.87 | 27.81 | $<40$ |  | off |
| 30.84 | 23.07 | 23.49 | 18.72 | $\geq 40$, |  |  |
| 3.10 | 25.71 | 30.13 | 32.19 | $<40$ |  |  |
| 3.16 | 30.93 | 23.51 | 33.28 | $\geq 40$ |  |  |

saturated model; in this example, it is ( $1,2,3,4$ ). This preserves every total and subtotal possible including the subtotals involving all combinations of all values of the four variables. For example, the total number of 0 -day EP, young age, accident, ons will remain the same. In fact, each cell will remain the same, yielding the identical array as was used for input. This makes the saturated model uninteresting.

One model that is of interest, though, is similar to the saturated model. That is, a model preserving the subtotals of all the combinations of all the values of all except one variable. In this model the status variable is not included. In the example, this model would be $(1,2,3)$ and would preserve the total number of 0 -day EP, young age, accidents, the total number of 0 -day EP, young age, sicknesses, and so on until all combinations of the four elimination periods, two age groups, and two types are exhausted. Since status is the only variable not mentioned in the model, the total number of 0 -day EP, young age, accidents is merely the ons plus the offs for that combination of EP, age, and type. Now, the ons plus offs are merely the total exposures, so the model ( $1,2,3$ ) preserves the exposures for each combination of the four elimination periods, two age groups, and two types.

By specifying a model such as $(1,4),(2,4),(3,4),(1,2,3)$, we will be assured that the new array is similar to the original array in the following ways:

1. The termination rates for 0-day, 7 -day, 14-day, and 30 -day elimination periods are the same as in the original array.
2. The termination rates for younger ages and older ages are the same as in the original array.
3. The termination rates for accident and sickness are the same as in the original array.
4. The exposures for each combination of elimination period, age group. and type are the same as in the original array.
By specifying the model $(1,2,4),(3,4),(1,2,3)$ we will be assured that the new array will be similar to the original array in the following ways:
5. The termination rates for each combination of the four elimination periods and two age groups will be the same as in the original array.
6. The termination rates for each type (accident and sickness) is the same as in the original array.
7. The exposures for each combination of elimination period, age group, and type are the same as in the original array.
These two models are typical of the ones used in the analysis of monthly termination rates and are shown in Table B-3. The second one differs from the first because it not only preserves the termination rates of the younger ages and older ages but also preserves the termination rates of each elimination period within the younger ages and each elimination period within
the older ages. We refer to this as the interaction of elimination period and age group. Since two variables are involved, it is called a two-way interaction. Terms signifying the other possible two-way interactions in our example are (1, 3, 4) and (2, 3, 4).

Once the ECTA program creates a new array, it also compares it with the original array by calculating a $\chi^{2}$ value. To calculate a $\chi^{2}$ value, the arrays are compared cell by cell; an amount equal to:

$$
\begin{gathered}
\frac{(a-e)^{2}}{e} \text { for Pearson } \chi^{2} \\
\text { or } \\
a \times \ln (a \div e) \text { for Likelihood Ratio } \chi^{2}
\end{gathered}
$$

is calculated for each cell, where $a$ is the number in the cell of the original array, $e$ is the number in the same cell of the new array, and $\ln$ is the natural $\log$ function. The $\chi^{2}$ value is merely the sum of the amounts for all the cells. In our analysis the Likelihood Ratio $x^{2}$ was used.

A $x^{2}$ value is small if the numbers in the cells of the new array are close to the numbers in the corresponding cells of the original array. Conversely, the $\chi^{2}$ is large if the numbers of the new array are not close to or different from the corresponding numbers of the original array. Therefore, the $\chi^{2}$ value can be thought of as a measure of how different the two arrays are, or as a measure of fit between the two arrays. These $\chi^{2}$ values have a $\chi^{2}$ distribution, so it can be determined if the two arrays are statistically significantly different with a certain level of confidence. For example, the $\chi^{2}$ value for the array in Table B-2 Part a (compared to the original array) is 282.68. Based on the $\chi^{2}$ distribution with 30 degrees of freedom, one can say that the new array is significantly different than the original at the 99 percent confidence level.

Comparing the array in Table B-3 Part a with the original array produces a $\chi^{2}$ value of 9.36 with 10 degrees of freedom. Basing a conclusion on this result, one cannot say that the two arrays are different at a 95 percent confidence level - not even with 80 percent confidence. Such a result may be sufficient to say that the model produces a satisfactory fit to the raw data. Since the model that produced the array in Table B-3 Part a only preserved the termination rates for the levels of each variable, it was not necessary to preserve interactions among variables to obtain a satisfactory fit.

The desired model in our analysis was "the simplest model with a satisfactory fit." In the array used in our examples, other models may be tested:

| Model | $\chi^{2}$ | Degrees of Freedom | Levels of <br> Significance |
| :--- | ---: | :---: | :---: |
| $(4),(1,2,3)$ | 22.98 | 15 | 92 |
| $(1,4),(1,2,3)$ | 14.27 | 12 | 72 |
| $(2,4),(1,2,3)$ | 17.51 | 14 | 77 |
| $(3,4),(1,2,3)$ | 22.35 | 14 | 93 |
| $(1,4),(2,4),(1,2,3)$ | 9.45 | 11 | 43 |
| $(1,4),(3,4),(1,2,3)$ | 14.27 | 11 | 78 |
| $(2,4),(3,4),(1,2,3)$ | 17.23 | 13 | 81 |

"Level of Significance" is defined here to be the largest confidence level for which it can be said that the two arrays are different. That is, if it can be said that the arrays are significantly different at a 92 percent confidence level but not with 93 percent confidence, then the level of significance is 92. If a satisfactory fit were defined by a level of significance of 75 or less, then the model $(1,4),(1,2,3)$ would be the simplest model that produces a satisfactory fit.

This approach is somewhat in contrast to the usual statistical analysis whereby the null hypothesis is attempted to be disproven using large confidence intervals. In this case, we are trying to show similarities between the sets of arrays rather than differences. From a statistical perspective, it is as if we were attempting to not reject the null hypothesis rather than the classical rejection of the null hypothesis as our proof.

An interpretation of our example would be that to satisfactorily predict a termination, one need only know the elimination period. Therefore, only four termination rates need to be derived-one for each elimination period. In this example, further observations could be made:

1. Knowledge of whether it is an accident or sickness case adds very little to predicting the number of terminations.
2. The best fit to the raw data can be achieved if the termination rate of elimination periods and age groups are preserved. (A very close fit may not be desirable because the random errors inherent in the raw data will also be reproduced.)
The actual data in our analysis were considerably more complex than the example we have been using here, so the $\chi^{2}$ values are not used directly. Instead, the difference between $\chi^{2}$ values was used. This difference has a $\chi^{2}$ distribution since likelihood ratio $\chi^{2}$ values had been used. Having the difference of two $\chi^{2}$ values equal another $\chi^{2}$ value is the direct result of using the Likelihood Ratio $x^{2}$.

First, the $\chi^{2}$ value and degrees of freedom for a basic model were recorded. Then, the model was changed slightly and the resulting $\chi^{2}$ and degrees of freedom were recorded. Using the difference between the $\chi^{2}$ values as the $\chi^{2}$ value for the change and the difference between the degrees
of freedom as the degrees of freedom for the change, it can be determined if the change in the model produced a significantly different array to represent the raw data. A high level of confidence, resulting from a large change in the $\chi^{2}$ values, would indicate that the change in the model had a large impact on trying to reproduce the original array.
To determine the important variables, a basic model was agreed upon which produced one $\chi^{2}$ value. Then this model was altered to eliminate one of the variables without changing any other part of the model. The resulting $\chi^{2}$ value was compared with the value for the basic model to determine if the elimination of that variable made significant difference. If the level of confidence is great, then the variable is important for maintaining a close fit to the original data. In other words, if we wanted to come up with a model that would reflect the termination rates of the experience data, then this variable needed to be included in the model.

## IDENTIFY SIGNIFICANT VARIABLES

Determining which variables were the most important became the first priority in the analysis because recognizing all variables would have required an array with 15 dimensions: duration, elimination period, age, class, sex, benefit period, time to expiration of benefit period, type, renewal provision, impairment, his occ., indemnity provision, observation period, company, and status.

If all levels of all variables were used, the contingency table would require $6.4 \times 10^{11}$ cells. Even if only two levels of all variables were used, there would be 32,768 cells. To reduce the number of variables, data for the second month of disablement were tested to determine the least important variables, which were then eliminated from any further study.

1. Month 2 was chosen because it contains the most exposure months and therefore more cells would contain data in them.
2. Elimination period was variable 1 and had 2 levels: less than 30 days and 30 days.
3. Age was variable 2 and had 2 levels: under 40 , and 40 through 64.
4. Class was variable 3 and had 2 levels: white collar (Occ. class $=1,2,3,4$-see Exhibit B-1) and blue collar (Occ. class $=5,6,7$ ).
5. Sex was variable 4 and had 2 levels: male and female.
6. Benefit period was variable 5 and had 2 levels: 2 years or less and greater than 2 ycars (to age 65 is assumed to be greater than 2 years).
7. Time to expiration of benefit period was not a variable in this test because virtually all the data are in the category of "more than 5 months to expiry," since the test concerns month 2 data only. This is to say that there was virtually no data with a benefit period less than 7 months.
8. Type was variable 6 and had 2 levels: accident and sickness.
9. Renewal provision was variable 7 and had 2 levels: Noncancelable and others.
10. Impairment was variable 8 and had 2 levels: standard and substandard.
11. His occ. period was variable 9 and had 2 levels: one year and others.
12. Indemnity provision was variable 10 and had 2 levels: some reduction and no reduction.
13. Observation period was variable 11 and had 2 levels: 1975-76 and 1973-74-7778.
14. Company was variable 12 and had 5 levels.
15. Status was variable 13 and had two levels-on/off.

The basic model was chosen to preserve the termination rates of all levels of all the variables as well as all 2-way interactions, all 3-way interactions that include sex or company, and all 4-way interactions that include sex and company. This can be written as:
$(1,2,4,12,13)$
$(1,3,4,12,13)$
$(1,5,4,12,13)$
$(1,6,4,12,13)$
$(1,7,4,12,13)$
$(1,8,4,12,13)$
$(1,9,4,12,13)$
$(1,10,4,12,13)$
$(1,11,4,12,13)$
$(2,3,4,12,13)$
$(2,5,4,12,13)$
$(2,6,4,12,13)$
$(2,7,4,12,13)$
$(2,8,4,12,13)$
$(2,9,4,12,13)$
$(2,10,4,12,13)$
$(2,11,4,12,13)$
$(3,5,4,12,13)$
$(3,6,4,12,13)$
$(3,7,4,12,13)$
$(3,8,4,12,13)$
$(3,9,4,12,13)$
$(3,10,4,12,13)$
$(3,11,4,12,13)$
$(5,6,4,12,13)$
$(5,7,4,12,13)$
$(5,8,4,12,13)$
$(5,9,4,12,13)$
$(5,10,4,12,13)$
$(5,11,4,12,13)$
$(6,7,4,12,13)$
$(6,8,4,12,13)$
$(6,9,4,12,13)$
$(6,10,4,12,13)$
$(6,11,4,12,13)$
$(7,8,4,12,13)$
$(7,9,4,12,13)$
$(7,10,4,12,13)$
$(7,11,4,12,13)$
$(8,9,4,12,13)$
$(8,10,4,12,13)$
$(8,11,4,12,13)$
$(9,10,4,12,13)$
$(9,11,4,12,13)$
$(10,11,4,12,13)$
(1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12)
This was chosen as the basic model because it was decided that 3-way interactions would be too complex except that, within each company, termination rates may differ completely between male rates and female rates. This model produced $\chi^{2}=782$, with 9,905 degrees of freedom.

The model was then changed to ignore elimination period (variable 1). This was done by removing 1 from all the terms of the model except the final term, which insured that the total exposures were preserved for each cell. The $\chi^{2}$ value for this model is 1,020 with 9,960 degrees of freedom. The difference between the two models has a $\chi^{2}$ value of 238 with 55 degrees of freedom. This means that by ignoring elimination period, we have produced an array that is statistically different, beyond the 99 percent level of confidence, from the basic array. Therefore, the inclusion of elimination period as a variable is necessary if we wish to produce an array that fits the basic model closely. In other words, elimination period is a statistically significant variable. The effects of ignoring the other variables are shown in Exhibit B-2.

When indemnity provision was ignored by eliminating 10 from all the terms of the basic model except the last term, the $\chi^{2}$ value changed to 798 with 9.960 degrees of freedom. The difference of 16 in the $\chi^{2}$ value with 55 degrees of freedom translates into no change in the array because of the elimination of indemnity provision ( 0 percent confidence that the two arrays are different). Therefore, indemnity provision can be eliminated as a variable without affecting fit. In other words, it was not found to be a statistically significant variable.

EXHIBIT B-2
Month 2
Effects of Variable Elimination in Contingency Table Analysis

| Variabif Missingi | Cm-Sulared Vailie | Degreis of Friemom | Levei of Significance |
| :--- | :---: | :---: | :---: |
| Basic Model | 782 | 9.905 | - |
| Elimination Period | 1.020 | 9,960 | 100 |
| Age | 990 | 9,960 | 100 |
| Class | 846 | 9,960 | 81 |
| Sex | 871 | 9,960 | 100 |
| Benefit Period | 877 | 9.960 | 100 |
| AccidenuSickness | 891 | 9,960 | 100 |
| Renewal Provision | 812 | 9,960 | 0 |
| Impairment | 835 | 9,960 | 44 |
| His Occ. | 802 | 9,960 | 0 |
| Indemnity Provision | 798 | 9,960 | 0 |
| Observation Year | 810 | 9,960 | 0 |
| Company | 1,316 | 10,173 | 100 |

Because of this analysis of month 2 data, five variables were eliminated because they were statistically unimportant: renewal provision, impairment, his occ., indemnity provision, and observation period.

Twelve new arrays were produced for analysis of the remaining variables.

1. The first six arrays contained data for the first six months of disablement.
2. The last six arrays contained data for the last six quarters of the first two years of disablement. That is, the seventh array contained data for months 7,8 , and 9 ; the eighth array contained data for months 10,11 , and 12 ; and so on until the last array for months 22, 23, and 24.

The following variables were included in the arrays:

1. Elimination period was in the first seven arrays and had 3 levels in the first array ( 0 , 7, and 14 days); 4 levels in the next two arrays ( $0,7,14$, and 30 days); 5 levels in the next three arrays $(0,7,14,30$, and 90 days); and 2 levels in the seventh array ( $0.7,14,30$, and 90 days combined and 180 days).
2. Age was a variable with 5 levels in all twelve arrays. The levels were 20-29, 30-39, 40-49, 50-59, and 60-64.
3. Class was a variable with 4 levels in all arrays. The levels were the 4 classes defined, in the New York Study.
4. Sex was a variable with 2 levels in all arrays. The levels were male and female.
5. Benefit period was in the first six arrays and had 6 levels (1-12 months, 13-24 months, 25-60 months, to age 65 , lifetime, and other).
6. Time to expiration of benefit period was a variable in the last 5 arrays. It had 2 levels (within 5 months from the end of the benefit period and more than 5 months).
7. Type was a variable with 2 levels in all arrays. The levels were accident and sickness.
8. Duration was a variable in the last 6 arrays. The 3 levels were the first, second, and third months of the quarter.
9. Company was a variable in all arrays. In the first eight arrays there were 5 companies, and in the last four arrays there were 6.

The basic model used was essentially the same as that used for the month 2 data, namely, the model that preserves the termination rates of all levels of all variables as well as all 2-way interactions, all 3-way interactions that include sex or company, and all 4-way interactions that include sex and company.

The tedious process of eliminating each variable from the basic model and calculating the level of significance was done for each array to determine the important variables at each duration.

## IDENTIFY SIGNIFICANT INTERACTIONS

To determine the important interactions, the basic model was also changed to exclude interactions. One-by-one every interaction (that did not involve company) was eliminated while being careful to still keep other interactions in the model. For example, to determine the importance of the age-sex interaction, the terms involving this interaction had to be changed so that the model no longer preserved the male termination rates by age and female termination rates by age while not changing the interactions of age with other variables or sex with other variables.

## DETERMINE THE BEST MODEL

Once we had the levels of significance for each variable and interaction, the simplest model with a satisfactory fit needed to be determined. This required model was produced by simplifying the basic model through the elimination of unimportant variables and interactions. The rules for determining which variables and interactions should be eliminated were:

1. It does not have a high level of significance, and it also has no hope of being significant in later durations, or
2. there is no logic to support its inclusion.
(Other considerations were also made, such as reducing the number of variables to a manageable number and the reasonableness of the factors that would be produced, and so on.)

Exhibit B-3 shows the levels of significance of the variables and interactions tested in the twelve arrays. The rules for eliminating variables and interactions left room for judgment concerning what a high level of confi-

EXHIBIT B-3
Level of Confidence of Variables and Interactions

| Monch | 1 | 2 | 3 | 4 | 5 | 6 | 7.8.9 | 10.11.12 | 13.14.15 | 16.17.18 | 19,20,21 | 22,23,24 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EP | 100 | 100 | 80 | 0 | 0 | 0 | 0 |  |  |  |  |  |
| AGE | 100 | 100 | 100 | 100 | 14 | 63 | 91 | 22 | 0 | 0 | 0 | 0 |
| CLASS | 83 | 45 | 40 | -1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| SEX | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 9 | 0 | 0 | 0 | 0 |
| BP | 0 | 16 | 0 | 0 | 0 | 0 |  |  |  |  |  |  |
| EXP. | 0 |  |  |  |  |  |  | 0 | 0 | 0 | 0 | 30 |
| A/S | 100 | 100 | 98 | 24 | 0 | 0 | 100 | 66 | 55 | 0 | 0 | 0 |
| DUR. |  |  |  |  |  |  | 55 | 1 | 0 | 0 | 0 | 0 |
| CO . | 100 | 100 | 1 | 0 | 0 | 0 | 0 | 2 | 0 | 0 | 0 | 0 |
| EP-AGE | 99 | 55 | 100 | 83 | 74 | 61 | 53 |  |  |  |  |  |
| EP-CLASS | 85 | 97 | 92 | 100 | 52 | 62 | 54 |  |  |  |  |  |
| EP-SEX | 95 | 39 | 83 | 95 | 70 | 80 | 56 |  |  |  |  |  |
| EP-A/S | 19 | 54 | 98 | 53 | 13 | 17 | 0 |  |  |  |  |  |
| AGE-CLASS. | 99 | 91 | 98 | 92 | 66 | 59 | 66 | 48 | 98 | 45 | 35 | 89 |
| AGE-SEX | 100 | 65 | 75 | 4 | 53 | 95 | 83 | 89 | 90 | 78 | 69 | 24 |
| AGE-EXP |  |  |  |  |  |  |  | 47 | 2 | 0 | 24 | 87 |
| AGE-A/S | 100 | 100 | 100 | 100 | 65 | 99 | 100 | 40 | 96 | 100 | 53 | 30 |
| CLASS-SEX | 24 | 86 | 94 | 46 | 67 | 24 | 39 | 18 | 15 | 27 | 40 | 27 |
| CLASS-EXP |  |  |  |  |  |  |  | 18 | 0 | 15 | 47 | 15 |
| CLASS-A/S. | 31 | 31 | 46 | 31 | 67 | 67 | 83 | 86 | 88 | 94 | 27 | 100 |
| SEX-EXP |  |  |  |  |  |  |  | 29 | 31 | 0 | 3 | 100 |
| SEX-A/S | 95 | 84 | 100 | 43 | 84 | 16 | 89 | 16 | $88$ | $98$ | $31$ | 31 |
| EXP-A/S |  |  |  |  |  |  |  | 84 | $19$ | $66$ | $31$ | 66 |
| SEX-EP-AGE | 2 | 9 | 41 | 22 | 0 | 3 | 53 |  |  |  |  |  |
| SEX-EP-CLASS | 6 | 0 | 44 | 4 | 0 | 11 | 8 |  |  |  |  |  |
| SEX-EP-A/S . | 12 | 13 | 31 | 83 | 17 | 2 | 29 |  |  |  |  |  |
| SEX-AGE-CLASS | 0 | 6 | 4 | 21 | 1 | 1 | 16 | 2 | 0 | 0 | 0 | 0 |
| SEX-AGE-EXP. |  |  |  |  |  |  |  | 40 | 0 | 0 | 53 | 1 |
| SEX-AGE-A/S ... | 87 | 13 | 6 | 59 | 4 | 6 | 28 | 59 | 41 | 0 | 0 | 2 |
| SEX-CLASS-EXP |  |  |  |  |  |  |  | 3 | 0 | 3 | 0 | 0 |
| SEX-CLASS-A/S. | 24 | 13 | 5 | 18 | 31 | 3 | 54 | 13 | 5 | 1 | 11 | 0 |
| SEX-EXP-A/S. |  |  |  |  |  |  |  | 56 | 3 | 19 | 3 | 56 |

dence was. The variables and interactions that remained in the simplified model are enclosed in boxes in Exhibit B-3. Some comments on the selection of this model are:

1. Elimination period was included as a variable in months 4,5 , and 6 so that the difference between 90 days and less than 90 days could be quantified.
2. Age and sex were included as variables for all durations because of their significance in later durations.
3. His Occ. period also had some hope of being significant at the end of one year or two years of disability. However, further tests were conducted, but no evidence of significance was found.
4. Benefit period was not found to be significant, and no logic was found to support its inclusion.
5. Time to expiration of benefit period was not found to be significant although the two levels showed very different termination rates. To avoid distortion, only data more than 5 months from the end of the benefit period were included.
6. Company was included in the steps to identify significant interactions to avoid distortions. Once these determinations were made, extensive studies of all the companies demonstrate that little distortion of the termination rates result from combining the data of all companies.
7. No 3-way interactions were significant enough to be included in the model.

The model can be stated as follows:
I. In months 1, 2, and 3 the terms of the model are:
A. (elimination period, age, status)
B. (age, class, status)
C. (age, sex, status)
D. (age, type, status)
E. (elimination period, age, class, sex, type)
II. In months 4,5 , and 6 the terms are:
A. (eliminatin period, status)
B. (age, type, status)
C. (sex, status)
D. (elimination period, age, sex, type)
III. In months 7 through 12 the terms are:
A. (age, type, status)
B. (sex, status)
C. (age, sex, type)
IV. In months 13 through 24 the terms are:
A. (age, status)
B. (sex, status)
C. (age, sex)

The determination of this model marks the completion of the analysis of the variables. Some analysis included studying some factors produced by
different models but the ultimate determination of the factors is a completely different step in the process of developing an experience table.

## ANALYZE CONTINGENCY TABLE

By this time, you should have some feel for what contingency table analysis is and its potential to the actuary. At this point, we will (1) give a description of both the analysis and the modifications made to it so that the remainder of this section will be clearer and (2) give a source for further reference.

Brief Description of Application-You already have noted that a contingency table is a multidimensional array of mutually exclusive counts or frequencies. If one of those dimensions has only two values depicting change of status, we can model the odds of change of status. Specifically, we have shown how the probability of terminating claims status can be computed by dividing a cell for off claims by the sum of that cell and the corresponding cell for on claims. The odds for terminating claims status is a simpler calculation; namely, the quotient of an "off"' cell and its corresponding "on" cell.

Thus, once we have a model that produces acceptable cell counts, we can divide the mathematical expression for that model for the "off" cells by the mathematical expression for the "on'" cells, simplify the algebra, and have a mathematical model for an array with one less dimension which contains the odds of terminating claims for each combination of variable values. Since we are not interested in the cell counts per se, but only interested in being able to produce a reasonable fit, this model is more interesting. The only drawback is that it is in terms of odds, and we are used to dealing in probabilities.

Our approach was to use the contingency tables with mutually exclusive counts to perform all of the analysis. This maintains the validity of the statistical tests used to decide upon the best model. Once the model was chosen, we reran ECTA with exposures instead of "ons,' and the resulting model produced probabilities that are exactly equivalent to the odds already produced.

We made one other change in the form of the model. The model used by ECTA is called a log-linear model. The name comes from the fact that the model works with the logarithm of the values, rather than the values themselves, and limits itself to linear relationships. The resulting model is translated back to antilog values for output. This results in the model being a multiplicative model.

Specifically, the model for the "odds" (or "probabilities") is an overall average "odds" (or "probabilities'’) and a set of factors for each variable
and interaction defined in the model. The set of factors for each variable is a vector whose length is equal to the number of values that that variable assumes in the model. The set of factors for each interaction defined in the model is an array whose rank is equal to the number of variables in the interaction and whose shape is defined by the number of values each such variable assumes in the model.

As an illustration, consider the variables sex, cause, and occupational class. The factors for each of these variables would be a vector length 2 representing male and female, a vector length 2 for accident and sickness, and a vector length 4 for the four occupational classes. If the model had an interaction between cause and occupational class, it would be a 2 by 4 matrix with a row for each cause and a column for each class. If the model had a three-way interaction among these three variables, the set of factors would be a 2 by 2 by 4 array with a plane for each sex, a row for each cause, and a column for each occupational class.

The mathematical form of the model is that the "odds" (or "probability") for each cell is equal to the product of the overall average "odds" (or "probability") and the factor from each set of factors corresponding to the value of each variable that identifies the cell. As an aside, we found this form to be simple conceptually and, thus, appealing-if the data could be represented without involving high-level interactions. This was the case for termination rates. Incidence rates, however, required so many high-level interactions that use of the method was limited to establishing the statistical significance of the variables.

The form of the model did have one aspect with which we were uncomfortable. Because the model dealt with logarithms, rather than values, the overall average was a geometric mean. Similarly, if you multiplied each factor along any dimension in a set of factors, the product was one. We modified the ECTA result so that the overall average is an arithmetic mean and the mean of each set of factors is equal to 1 . This modified model produces exactly the same results, and it is easier to interpret the model's parameters. The overall average rate is a simple average of all the rates defined by the model.

For a better understanding of how the model reproduces the crude data, consider the following:

1. Think of the original contingency table as two arrays, one of "offs" and one of "ons."
2. Replace the "ons" with exposures.
3. Expand any model into its full array of probabilities. It will be the same size and shape as either the "ons" or "offs."
4. The element by element multiplication of points 2 and 3 yields an array of expected "offs" based upon the model.
5. Summarize the array of actual "offs" and the array of expected "offs" as many times as you have sets of factors in the model with the summarizations conforming to the shape of each set of factors.
6. Divide, element by element, each summarization of actual "offs"' by the corresponding summarization of expected "offs." Each quotient will be equal to one.
To the extent that the model is reasonably simple, it does produce some smoothing. However, the precise fit, particularly, when it applies to relatively sparse areas of the contingency table, does retain some of the anomalies usually found in crude data. Hence, we still have a need for graduation, which is considered in the next step.

Source for Additional Information-Contingency Table Analysis is a relatively new field developed by statisticians. Our Committee learned of this area through its request for technical assistance from Mr. Tappin Roy, then of the Travelers Research Corporation. Mr. Roy recommended the technique as most appropriate for our problem and with assistance from a Committee member, Mr. William Daniels, produced an APL program implementing the tool. The limited APL workspace size limited our application of the tool. A statistical expert, Mr. Edward Seligman, was added to the Committee, and he guided us in the learning process of adapting the tool to our problem. In addition, Mr. Seligman presented a paper at the Fourteenth Actuarial Research Conference at the University of Iowa in 1979 entitled "Applications of Multi-Dimensional Contingency Tables to the Analysis of Termination Counts in Disability Income Claim Data." A more extensive unpublished paper, together with a bibliography may be obtained by contacting Mr . Seligman. Mr. Frank Knorr, who actually applied the technique to our problem, has also presented a paper to the American Statistical Association which is published in the ASA 1983 Proceedings of the Social Statistics Section, August 15-18, 1983, Toronto.

## Graduated Termination Rates

The goal of this step was to build an experience table of termination rates for all durations (month 1 to the final age of the table). The termination rates were to be smooth and vary according to the variables and interactions defined by the model.

A modification of the ECTA program was used to produce a termination rate and modification factors for each of the twelve arrays, representing the first six months and the last six quarters of the first two years of disablement. The ultimate table, developed by Mr. John H. Miller was used for termination rates after 10 years of disablement. The latest Group LTD experience was used to get smooth rates between the first and eleventh years. The technique that was used to smooth the termination rates in the 10 -year select

## EXHIBIT B-4

## EXAMPLE OF OUTPUT FROM ECTA PROGRAM FOR UNGRADUATED DATA FOR THIRD QUARTER

Estimated lambda effects, their standard errors, and the standardized values are followed by the corresponding Tau Parameter of the multiplicative model.
Variables where each level of the effect is the difference between that level of the variable, and the average effect: !-age
Variables of only two levels where the single effect shown is the difference of the first level, and the average effect: 2-sex, 3-type, 4-exposure/"off."

| Grand Mean Ehfect | Ertect | STANDARD Error | STANDARDLZED Valle | tau Parm |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 5.78603 |  |  | 325.71802 |  |
| Effect for Variables | 1.06619 | 0.01740 | 61.29123 | Overall Rate |  |
| ${ }_{1,4}$ |  |  |  | 0.11855 |  |
| For Levels of Var 1 (by age group) |  |  |  | Age Factors (For age only) |  |
| 20s | -0.14010 | 0.03232 | -4.33482 | 1.32340 |  |
| 30 s | -0.08111 | 0.02665 | -3.04393 | 1.17612 |  |
| 40s | $-0.01008$ | 0.02696 | -0.37400 | 1.02037 |  |
| 50 s | 0.08002 | 0.02668 | 2.99884 | 0.852100.73893 |  |
| 60 s | 0.15127 | 0.05347 | 2.82925 |  |  |
|  |  |  |  | Sex Factors |  |
|  |  |  |  | Male | Female |
| 2,4 | 0.01720 | 0.01740 | 0.98862 | 0.96619 | 1.03499 |
|  |  |  | $\begin{aligned} \text { Average Factor } & =1.00059 \\ \text { Adjusted Factors } & =0.966-\mathrm{Male} \\ & =1.034-\mathrm{Female} \\ \text { Adjusted Rate } & =0.11862 \end{aligned}$ |  |  |
|  |  |  |  | Type Factors (for type only) |  |
|  |  |  |  | Accident | Sickness |
| 3,4 | -0.06524 | 0.01740 | $-3.75050$ | 1.13938 | 0.87767 |
| 1,3.4 <br> For Levels of Var 1 (by age group) |  |  |  | Age/Type Factors (For Interactions) |  |
|  |  |  |  | Accident | Sickness |
| 20 s | 0.03848 | 0.03232 | 1.19049 | 0.92593 | 1.07999 |
| 30s | 0.06462 | 0.02665 | 2.42498 | 0.87877 | 1.13796 |
| 40 s | 0.00160 | 0.02696 | 0.05953 | 0.99680 | 1.00321 |
| 50s | -0.03914 | 0.02668 | -1.46683 | 1.08143 | 0.92470 |
| 60 s | -0.06556 | 0.05347 | -1.22614 | 1.14010 | 0.87711 |
| Age Group | Age Only $\times$ Type Only $\times$ Age Type Interaction |  |  |  |  |
|  | Accident | Sickness | Adjusted Factors |  |  |
|  |  |  | Accident | Sickness |  |
| 20 s | 1.39617 | 1.254 | 1.352 | 1.214 |  |
| 30s | 1.178 | 1.175 | 1.140 | 1137 |  |
| 40s | 1.159 | 0.89842 | 1.122 | $\begin{aligned} & 0.870 \\ & 0669 \end{aligned}$ |  |
| S0s | 1.050 | 0.69155 | 1.016 |  |  |
| 60 s | 0.960 | 0.569 | 0.929 | $\begin{aligned} & 0.669 \\ & 0.551 \\ & \hline \end{aligned}$ |  |
|  |  |  | $\begin{aligned} \text { Average Factor } & =1.033 \\ \text { Adjusted Rate } & =0.123 \end{aligned}$ |  |  |

period was the Whittaker-Henderson Type B Multidimensional Graduation Method. Although some smoothing was achieved through grouping and use of the ECTA program to produce termination rates and factors, the ultimate termination rates did not need to be smoothed because they had been defined by a formula.

In the analysis of the termination experience in the first 2 years of disablement, twelve arrays were used to represent the termination experience in the first six months and the last six quarters of the 2 years. Twelve new arrays needed to be produced to derive the termination rates and factors for the same twelve periods. The new arrays differ from the other arrays in three ways:

1. Only the variables defined in the model for that period were used.
2. The "on" cells of the variable status were changed to include "ons" plus "offs." This changed the variable from "on/off" to "exposures/off."
3. To eliminate cells with zeros, the value of .01 was added to each "off" cell and another value added to each "exposure" cell so that the ratio of .01 to that value was approximately the same as the ratio of the total number of "offs" to the total number of exposures. This assigned a termination rate, which was equal to the overall termination rate, to any cell that has no exposures.
Once the twelve new arrays were set up, the ECTA program was used again. This time the part designed for log-linear analysis was used. There were 2 slight changes made to the ECTA program that simplified the output (an annotated copy of which is included as Exhibit B-4) for our application. These were:
4. The printing of the log-linear data was suppressed unless it involved the variable exposure/off.
5. The formula for the Tau Parameter was changed from $e^{u}$ to $e^{-2 u}$. This value and its reciprocal were printed under the heading of Tau Parm.
We also converted the geometric output of ECTA to the more easily understood arithmetic output described earlier. Both sets of factors for the average monthly termination rate in the third quarter (applicable to the eighth month of disablement) are as follows:

Arithmetic

| Arithmetic |  |  | Geometric |  |
| :---: | :---: | :---: | :---: | :---: |
| Rate |  | 0.123 | Rate | 0.119 |
| Male |  | 0.966 | Male | 0.966 |
| Female |  | 1.034 | Female | 1.035 |
| Age | Accident | Sickness | Accident | Sickness |
| $20-29$ | 1.352 | 1.214 | 1.396 | 1.254 |
| $30-39$ | 1.140 | 1.137 | 1.178 | 1.175 |
| $40-49$ | 1.122 | 0.870 | 1.159 | 0.898 |
| $50-59$ | 1.016 | 0.669 | 1.050 | 0.692 |
| $60-64$ | 0.929 | 0.551 | 0.960 | 0.569 |

The routine, which makes this conversion, takes each set of geometric factors and performs these 3 steps:

1. determines the average of the set of geometric factors,
2. divides every geometric factor of the set by the average from step 1 to produce the corresponding arithmetic set, and
3. multiplies the overall rate by the average from step 1 to produce a partial conversion to an arithmetic mean.
This will not affect the specific termination rates at all.
To smooth the termination rates, every specific termination rate was calculated. For month 2 there were 320 specific rates that depend on age, sex, elimination period, class, and type. These form a five-dimensional array of termination rates. Similar arrays were formed for the other durations in the first two years of disablement.

Then all arrays were combined into one six-dimensional array of termination rates. Similarly a six-dimensional array of weights were formed from the exposures of the corresponding termination rates. The six dimensions are:

1. Elimination period- 5 levels: $0.7,14,30$, and 90 days. In month 1 , the only elimination periods with any exposures were 0,7 , and 14 days; the termination rates for 7 and 14 days did not represent a full month termination rate, so they were adjusted by dividing them by .75 and .5 , respectively, for graduation purposes. The exposures were not adjusted in month 1; they were merely the denominator of the termination rate calculation and not complete exposure months for 7 and 14 days. In months 2 and 3 , the only elimination periods with any exposures were $0,7,14$, and 30 days. In months 4, 5, and 6, all elimination periods had exposures; however, 0, 7, 14, and 30 days all had the same termination rates. Durations greater than 6 months had the same set of termination rates for all climination periods.
2. Age at disablement-5 levels: $20-29,30-39,40-49,50-59,60-64$.
3. Class-A levels: The four New York Study classifications. Durations greater than three months had the same termination rates for all classes.
4. Sex- 2 levels: male and female.
5. Type- 2 levels: accident and sickness. Durations greater than 12 months had the same termination rates for both types.
6. Duration of disablement-24 levels: one for each month in the first 2 years of disablement. For durations greater than 6 months, termination rates had only been calculated for the third, fourth, fifth, sixth, seventh, and eighth quarters. They were used to represent the termination rates of months $8,11,14,17,20$, and 23 , respectively. The weights for these months were set equal to one-third of the exposure months of the entire quarter, while the weights for months $7,9,10,12,13,15,16$, $18,19,21,22$, and 24 were set equal to zero.
Using the Whittaker-Henderson Type B graduation method, smooth termination rates were created for all months, including those months that had

## weights of zero. The graduation was done in four different parts:

1. Months 1 through 6-graduations of three dimensions at the same time: age (minimizing fourth differences), duration (minimizing fourth differences) and class (minimizing second differences). This was done for each elimination period, sex, and type. These had the effect of forcing the class factors to converge as duration increases and within each duration, the class factors would be near the least-squares straight line.
2. Months 1 through 24 -graduations of two dimensions at the same time: age (minimizing fourth differences) and duration (minimizing fourth differences). This was done for each elimination period, class, sex, and type. The termination rates that were graduated in this part were made up of graduated rates from the first part and the ungraduated termination rates for each quarter after month 6. Interpolated and extrapolated termination rates were created where the weights were equal to zero.
3. Years 3 through 10 -graduation of two dimensions at the same time: age and duration for male and female separately. The values that were graduated were the logarithms of the coefficients of selection. Using large smoothness factors insured that the coefficients of selection could be written as an exponentially decreasing function of age and duration. Coefficients of selection are defined as the ratio of the select termination rate to the ultimate termination for the same attained age. Large coefficients of selection represent large differences from the ultimate rates which occur in the early part of the select period. The ungraduated coefficients of selection were the ratio of the latest Group LTD termination rates (for years 1974-78) ${ }^{2}$ to the ultimate termination rates developed by Mr. John H. Miller. For year 11 these were fixed at 1, and for year 2 these were based on the graduated rates from the second part of this graduation.
4. Months 13 through 24 -graduation of two dimensions at the same time: age (minimizing fourth differences) and duration (minimizing third differences). This was done for male and female separately. These used termination rates from the third and fourth quarters as well as from the third year of disablement. This was done to produce termination rates that graded smoothly from the first year to the third year.
Once the four parts of the graduation were completed, the graduated termination rates were multiplied by the actual exposures for each cell of the six-dimensional array (For month 1, elimination periods 7 and 14 days, the termination rates were first adjusted by multiplying them by .75 and .5 , respectively). The new set of terminations and exposures were summarized the same as before so that it could be used as input for the ECTA program to produce new termination rates and factors for the first six months and the last six quarters of the first two years of disablement. Termination rates and factors were also produced by the ECTA program for years 3 through 10 .

The duration rates and factors for the graduated termination rates are shown in Exhibits B-5a to B-5c. They represent the results of the gathering, analyzing, and processing of the termination rate data. The quarterly termination rates shown in this exhibit are actually monthly rates to be used

[^4]
## EXHIBIT B-5a

Factors for Calculation of Monthly Termination Rates
Based on Graduated Data
(average $=1$ )


EXHIBIT B-5b
Factors for Calculation of Quarterly Termination Rates
Based on Graduated Data
(average $=1$ )

| Quarter: | Third | Fourth | Fifth | Sixth | Siventh | Eighth |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | . 124 | 066 | . 048 | . 032 | . 021 | . 016 |
| Sex: |  |  |  |  |  |  |
| Male | . 943 | . 931 | . 975 | . 984 | . 997 | 1.009 |
| Female | 1.057 | 1.069 | 1.025 | 1.016 | 1.003 | . 991 |
| Age: |  |  |  |  |  |  |
| 20-29 A, S | 1.2591 .262 | 1.5341 .344 | 1.625 | 1.825 | 2.042 | 2.098 |
| 30-39 A.S | 1.1271 .240 | 1.2471 .248 | 1.292 | 1.303 | 1.289 | 1.217 |
| 40-49 A.S | 1.0191 .048 | 1.054 .966 | . 937 | . 835 | 720 | . 679 |
| 50-59 A,S | . 869.820 | . 831.676 | . 629 | . 542 | 463 | 487 |
| 60-64 A,S | . 706.651 | . 602.499 | . 517 | 496 | . 486 | 519 |

EXHIBIT B-5c
Factors for Determination of Annual Termination R ites
Based on Graduated Data
(average $=1$ )

| Ycar: | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | 123 | . 084 | . 062 | . 050 | . 045 | . 042 | . 042 | . 043 |
| Sex: <br> Male Femate | 1.080 .920 | 1.129 .871 | 1.179 .821 | $\begin{array}{r} 1.200 \\ .800 \end{array}$ | 1.212 .788 | 1.210 .790 | 1.204 .796 | 1.200 .800 |
| Age: |  |  |  |  |  |  |  |  |
| 20-29 | 2.085 | 1.832 | 1.554 | 1.262 | 994 | . 776 | . 617 | . 524 |
| 30-39 | 1.164 | 1.103 | 1.017 | . 909 | 792 | . 696 | . 631 | . 582 |
| 40-49 | . 727 | . 757 | 767 | . 754 | 741 | . 737 | . 739 | . 751 |
| 50-59 | 536 | 616 | 697 | 832 | 984 | 1.103 | 1.182 | 1.226 |
| 60-64 | 489 | 691 | 965 | 1.244 | 1.489 | 1.688 | 1.830 | 1.918 |

for each month of the quarter. These were subsequently changed to produce a different rate for each month of the quarter.

One last note should be pointed out. Although time to expration of benefit period was not determined to be an important variable, the actors produced by data within 5 months from the end of the benefit period were dramatically different. Therefore, the data used in the graduation process represented only data (exposures and terminations) that were not near the end of the benefit period, that is, more than 5 months from the end. The actual termination rates for claims exposed near the end of the benefit period have been measured to be 16 percent and 236 percent greater than the termination in Exhibits B-5b for the fourth and eighth quarters, respectively.

## Weekly Termination Rates

In order to study weekly termination rates in the first three months of disablement, the data had to be changed to a more manageable form. It was to be in a form that was flexible enough so that it would allow the easy study of weekly as well as daily or monthly termination rates. The form of the data is very similar to that used in the analysis of monthly termination rates in the first two years of disablement.

The input data for this reformatting were the same as the input data used for the summarization of data to produce monthly termination rates. The format of the records of the input file can be found in Exhibit A-1. The file contains one record per disability claim, except when the claim was observed in more than one experience period or if data were submitted separately for different policies owned by the same claimant.

The selection process required calculating the duration in days from the date of disablement to the first day of exposure for each input record. The first day of exposure is the later of the first day after the elimination period expires or January 1 of the first experience year for that record. Each month is assumed to have 30 days. If the duration is greater than 90 days, the input record is not used. This eliminates all records with elimination periods greater than or equal to 90 days. Since most of the data are for experience years 1975 and 1976, most of the selected records have dates of disablement in 1975 and 1976.

The same corrections to certain fields were made to this data as were made to the data used for monthly termination rates.

There was one output record for each input record selected. The format is shown in Exhibit B-6. This format is similar to the file used in the analysis of the termination rates of the first two years (Exhibit B-1). However, instead of having a duration field, this output required two duration fields: (a) the duration from date of disablement to the first day of exposure to termination (which was used in the selection process) and (b) the duration to the last day of exposure. The duration $b$ is measured as the number of days from the date of disablement to the earlier of the termination date or the date 90 days after disablement or the date at the end of the experience period. If the termination date was more than 90 days after disablement, then the status code was made equal to 3 (exhausted). Also if the status code is 1,2 , or 4 , then a 1 appears as the termination indicator signifying that the disability terminated at duration $b$ because of death or recovery. The output file containing almost 150,000 records was used to generate terminations and exposures for weekly termination rates.

Exposures were calculated differently than for monthly termination rates.

## EXHIBIT B-6

Reformatted Data for Producing Weekly Termination Rates (Logical Record Length of 50, all fields in Binary form)


## EXHIBIT B-6-Continued

2 Indemnity Provisio
1- complete reduction
2 -. pro rata for new occ.
3 - pro rata for regular occ.
6-1 for 2 offset
8 -- others
9 - no reduction

2 Experience 1-1973
Year

3-1975 5-1977
$4-1976 \quad 6-1978$

2 Company Code (1 through 21)
2 Monthly Indemnity Amount
2 Termination Indicator
0 -exposure days only
1-the last day is an exposure
and a termination
2 Status Code

| 0-open <br> 1-recovery | 3-reached 90 days without <br> termination <br> 2-death |
| :--- | :--- |
| 4-death or recovery |  |

2 Benefit Period in Months or 965-to age 65 999-Lifetime

2 Duration at First Day of Exposure ( 1 through 90)
2 Duration at Last Day of Exposure (1 through 90)

For each 7-day period, the number of days of exposure was calculated and then divided by 7 . That is:

## Exposure $=$ A. One plus

B. the difference between

1. the latest of
a. the first day in the 7-day period,
b. the first day after the end of the elimination period,
c. the first day of the experience period, and
2. the earliest of
a. the last day in the 7 -day period,
b. the termination date,
c. the last day of the experience period,
d. 90 days after the date of disablement.
C. divided by seven
D. zero if ( B 2 ) minus ( B 1 ) is negative.

This yields an exposure value of one (one week's exposure) if a claim was exposed to termination for the entire week and did not terminate. This also yields a value of 0.7143 for 30 -day EP claims exposed for the entire fifth week. Terminations in the middle of the week also resulted in fractional exposure weeks.

If a termination occurred during a 7 -day period, it was counted as one termination for that week. These exposures and terminations were passed to the ECTA program to produce rates and factors. The model used was the same as the model used for the first three months of the monthly termination rates.

The graduation process was also similar to that used for monthly termination rates; however, no adjustment was needed for 7 -day and 14-day EP since all the rates are true weekly termination rates. Even week 5 for 30 day EP claims needed no adjustment before graduation since 0.7143 of a week's worth of terminations were divided by 0.7143 of a week's worth of exposure resulting in a weekly termination rate. The graduated rates were multiplied by the exposures to get smooth terminations. The smooth terminations and exposures were passed to the ECTA program to produce the smooth rates and factors found in Exhibit B-7.

EXHIBIT B-7
Factors for Calculation of Weekly Termination Rates Based on Graduated Data
( AVERAGE $=1$ )

| Week: | 1 |  |  | 2 |  |  |  | 3 |  |  | 4 |  |  | 5 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | . 139 |  |  | . 120 |  |  |  | . 117 |  |  | . 125 |  |  | 118* |  |  |
| Age: 20-29 | 1.019 |  |  | 1.138 |  |  |  | 1.127 |  |  | 1.105 |  |  | 1.048 |  |  |
| EP: 0,7,14,30 | 1.0008080 |  |  | 1.053 .941 |  |  |  | $\begin{array}{\|rrrr} 1.131 & 1.066 & .788 \\ .963 & .983 & 1.009 & 1.036 \end{array}$ |  |  | $\begin{array}{\|ccc} 1.061 & 1.074 & .849 \\ .983 & .997 & 1.005 \quad 1.009 \end{array}$ |  |  | 1.1561 .2461 .036 |  | . 597 |
| Class: 1,2,3,4 | . 978 . 981 | . 995 | 1.011 | . 951 | . 968 | . 012 | 1.053 |  |  |  | 1.0061 .006 | 1.000 | . 984 |
| Sex: M,F | $1.154 \quad .859$ |  |  | 1.142 | . 858 |  |  | 1.101 .897 |  |  |  |  |  | $1.079 \quad .922$ |  |  | $1.060 \quad .942$ |  |  |
| Cause: A,S | $1.034 \quad .957$ |  |  | . 956 | 1.018 |  |  | . 9121.074 |  |  | . 8941.098 |  |  | . 8841.112 |  |  |
| Age: 30-39 | 1.014 |  |  | . 961 |  |  |  | . 959 |  |  | . 997 |  |  | . 985 |  |  |
| EP: 0,7,14,30 | 1.000 |  |  | 1.062 .934 |  |  |  |  |  |  | $\left[\begin{array}{llll} 1.130 & 1.049 & .815 & \\ 1.007 & 1.001 & .996 & .991 \end{array}\right.$ |  |  | $1.2491 .191 \quad .985 \quad .608$ |  |  |
| Class: 1,2,3,4 | $\begin{array}{lllll}1.111 & 1.030 & .957 & .882\end{array}$ |  |  | $\begin{array}{lllll}1.046 & .999 & .977 & .960\end{array}$ |  |  |  |  |  |  | $\begin{array}{llll}1.007 & 1.003 & .997 & .988\end{array}$ |
| Sex: M,F | 1.101 .901 |  |  | $\begin{aligned} & 1.190 \\ & 1.044 \end{aligned}$ | . 824 |  |  | $\begin{array}{llll} 1.006 & .998 & .995 & .991 \\ 1.146 & .862 & & \end{array}$ |  |  |  |  |  | $\begin{array}{llll} 1.007 & 1.001 & .996 & .991 \\ 1.090 & 913 \end{array}$ |  | $1.090 \quad .913$ | 1.055 . 946 |  |  |
| Cause: A,S | . 999 . 994 |  |  |  | . 933 |  |  | . 996.984 |  |  | . 9601.023 |  |  | . 9371.050 |  |  |
| Age: 40-49 | 1.027 |  |  | . 894 |  |  |  | . 898 |  |  | . 943 |  |  | . 962 |  |  |
| EP: 0.7,14,30 | 1.000 lll |  |  | 1.082 .916 |  |  |  | $\left\lvert\, \begin{array}{llll} 1.218 & 1.053 & .741 & \\ 1.061 & 1.017 & .977 & .939 \end{array}\right.$ |  |  | $\left(\begin{array}{llll} 1.185 & 1.023 & .797 & \\ 1.041 & 1.011 & .984 & .960 \end{array}\right.$ |  |  | $1.298 \quad 1.123 \quad .938 \quad .652$ |  |  |
| Class: 1,2,3,4 | $\begin{array}{llll}1.215 & 1.070 & .934 & .796 \\ 1.038 & .955 & & \end{array}$ |  |  | $\begin{array}{llll}1.135 & 1.029 & .951 & .884\end{array}$ |  |  |  |  |  |  | $1.0251 .009 \quad .990 \quad .972$ |
| Sex: M,F |  |  |  | 1.146 |  |  |  | $1.110 \quad .890$ |  |  |  |  |  | $\left(\begin{array}{rrrr} 1.041 & 1.011 & .984 & .960 \\ 1.063 & .936 \end{array}\right)$ |  |  | 1.033 .966 |  |  |
| Cause: A,S | $\begin{array}{\|rr} 1.038 & .955 \\ .977 & 1.013 \end{array}$ |  |  | $1.132$ | . 860 |  |  | $1.090 \quad .898$ |  |  | $1.046 \quad .939$ |  |  | $1.014 \quad .970$ |  |  |
| Age: 50-59 | 1.016 |  |  | . 949 |  |  |  | . 942 |  |  | . 948 |  |  | . 977 |  |  |
| EP: 0,7,14,30 | 1.000 |  |  | $\begin{array}{llll} 1.136 & .873 \\ 1.193 & 057 & 935 & 832 \end{array}$ |  |  |  | $\left\{\begin{array}{llll} 1.263 & 1.001 & .751 \\ 1.120 & 1.039 & .959 & .887 \end{array}\right.$ |  |  | $\begin{array}{\|rrrr} 1.228 & .988 & .797 & \\ 1.086 & 1.028 & .970 & .918 \end{array}$ |  |  | $\begin{array}{llll} 1.298 & 1.056 & .897 & .725 \\ 1.060 & 1072 \end{array}$ |  |  |
| Class: 1, 2, 3,4 | $\begin{array}{rrrr}1.2431 .080 & .936 & .769 \\ .972 & 1.020 & & \end{array}$ |  |  |  |  |  |  | $\begin{array}{llll} 1.060 & 1.023 & .979 & .938 \end{array}$ |  |  |  |
| Sex: M,F |  |  |  | $\begin{aligned} & 1.002 \\ & 1.191 \end{aligned}$ | $.978$ |  |  |  |  |  | 1.120 1.039 .959 .887 <br> 1.000 988   |  |  | $\begin{array}{ll} 1.000 & .995 \\ 1.142 & 860 \end{array}$ |  |  | . 9971.001 |  |  |
| Cause: A,S | $\begin{array}{r} .9721 .020 \\ 1.031 \\ \hline \end{array}$ |  |  |  | . 817 |  |  | $\left\lvert\, \begin{array}{ll} 1.000 & .988 \\ 1 & 171 \\ 1 \end{array}\right.$ |  |  | $1.118 \quad 879$ |  |  |  |  |  |
| Age: 60-64 | . 924 |  |  | 1.058 |  |  |  | 1.072 |  |  | 1.007 |  |  | 1.028 |  |  |
| EP: 0,7,14,30 | 1.000 |  |  | $1.109 \quad .894$ |  |  |  | $\left\lvert\, \begin{array}{llll} 1.210 & .958 & .819 & \\ 1.167 & 1.057 & .949 & .847 \end{array}\right.$ |  |  | $\begin{array}{ll} 1.210 & .965 \\ 1.143 & 1.049 \end{array}$ | .827 |  | 1.2571 .004 | .867.962 | .815.885 |
| Class: 1,2,3,4 | 1.2051 .072 | . 938 | 797 | 1.185 | 1.066 | .941 | 825 |  |  |  | . 955 | . 868 |  |  |  |
| Sex: M,F | $\begin{array}{rr} .908 & 1.092 \\ 1.245 & .794 \\ \hline \end{array}$ |  |  | $\begin{array}{r} .850 \\ 1.153 \\ 1.300 \quad .749 \\ \hline \end{array}$ |  |  |  | $\begin{array}{rr} .873 & 1.132 \\ 1.266 \quad .773 \\ \hline \end{array}$ |  |  |  | $\begin{array}{r} .9221 .080 \\ 1.257 \\ \hline \end{array}$ |  |  |  |  | $\begin{array}{r}.9551 .045 \\ 1.245 \\ \hline\end{array}$ |  |  |
| Cause: A,S |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

*Use .084 for 30 -day elimination periods to allow for the short week from 30 to 35 days.

EXHIBIT B-7-Continued
Factors for Calculation of Weekly Termination Rates Based on Graduated Data
( average = 1)

| Week: | 6 |  |  | 7 |  |  | 8 |  |  | 9 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | . 123 |  |  | 126 |  |  | . 125 |  |  | . 122 |  |  |
| Age:20-29 | 1.060 |  |  | 1.066 |  |  | 1.073 |  |  | 1.079 |  |  |
| EP: 0,7,14,30 | $\left\lvert\, \begin{array}{rrr} 1.076 & 1.210 & 1.048 \\ .992 & 1.008 & 1.007 \end{array}\right.$ |  | $\begin{aligned} & .689 \\ & .990 \end{aligned}$ | 1.0181 .1771 .053 |  | . 760 | . 98011.1471 .054 |  | . 820 | . $9581.1181 .049 \quad .873$ |  |  |
| Class: 1,2,3,4 |  |  | . 9861.0101 .009 . 993 | . $9831.0091 .010 \quad .997$ |  |  | . 9781.0071 .0121 .004 |  |  |
| Sex: M,F | $1.036 \quad .965$ |  |  | $1 \begin{array}{ll}1.022 \quad .978\end{array}$ |  |  | 1.012 .988 |  |  | $\begin{array}{ll}1.004 & .995\end{array}$ |  |  |
| Cause: A,S | . 8781.118 |  |  | . 8741.125 |  |  | . 8711.129 |  |  | . 8701.131 |  |  |
| Age: 30-39 | 1.019 |  |  | 1.043 |  |  | 1.058 |  |  | 1.066 |  |  |
| EP: $0,7,14,30$ | 1.1641 .153 |  |  | . 701 | $1.1191 .1211 .006 .759$ |  |  | $\begin{array}{\|rrrr} 1.082 & 1.099 & 1.013 & .807 \\ .993 & 1.000 & 1.004 & 1.003 \end{array}$ |  |  |  |  |  |
| Class: 1,2,3,4 | . 9991.003 | 1.002 | . 994 | $\begin{array}{llll}.9961 .001 & 1.003 & .998 \\ .994 & 1.005 & & \end{array}$ |  |  |  |  |  |  |  |  |
| Sex: M,F | 1.019 .981 |  |  |  |  |  | .9931 .0001 .0041 .003.9781 .022 |  |  | $.990 \quad .9991 .0051 .006$ |  |  |
| Cause: A.S | . 9251.062 |  |  | $\begin{array}{r} .9941 .005 \\ .9161 .073 \\ \hline \end{array}$ |  |  | . 9121.078 |  |  | . 9131.078 |  |  |
| Age: 40-49 | . 988 |  |  | 1.007 |  |  | 1.019 |  |  | 1.024 |  |  |
| EP: 0,7,14,30 | 1.2061 .096 | . 962 | . 738 | $\begin{array}{rr} 1.1721 .073 \\ 1.010 & 1.003 \\ .984 & 1.016 \\ .989 & .994 \\ \hline \end{array}$ | $\begin{array}{r} .974 \\ .996 \end{array}$ | $\begin{aligned} & .783 \\ & .990 \end{aligned}$ | 1.1431 .0571.0061 .000 | .983.997 | . 818 | $\begin{array}{rr} 1.113 & 1.046 \\ 1.004 & .999 \end{array}$ | 990.851 |  |
| Class: 1,2,3,4 | 1.0151 .006 | . 995 | . 983 |  |  |  |  |  | 1.0061 .000 . 9978 |  | $\begin{array}{llll} 1.004 & .999 & .998 & .998 \end{array}$ |  |
| Sex: M,F | 1.005 .995 |  |  |  |  |  | $\begin{aligned} & .9691 .031 \\ & 9821.001 \end{aligned}$ |  |  | $9591.042$ |  |  |
| Cause: A,S | 1.002 . 981 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Age: 50-59 | . 969 |  |  | . 964 |  |  | . 961 |  |  | . 957 |  |  |
| EP: $0,7,14,30$ | 1.2201 .052 | .930 | . 786 | $\left(\begin{array}{llll} 1.196 & 1.041 & .946 & .814 \\ 1.030 & 1.013 & .989 & .968 \end{array}\right.$ |  |  | $\left(\begin{array}{lll} 1.171 & 1.031 & .957 \\ 1.023 & 1.009 & .991 \end{array}\right.$ |  | $\begin{aligned} & .841 \\ & .976 \end{aligned}$ | $\begin{array}{llll} 1.147 & 1.021 & .964 \quad .869 \end{array}$ |  |  |
| Class: 1,2,3,4 | 1.0411 .018 | . 985 | . 956 |  |  |  | $\begin{array}{\|r\|rr\|} 1.020 & 1.007 & .993 \\ .976 & 1.024 & \\ \hline \end{array}$ |  |  |  |  |  |  |  |  |
| Sex: M,F | . 9951.005 |  |  | 1.030 1.013 1.010 .989 . 968 |  |  |  |  |  | $\left[\begin{array}{rrrr}1.023 & 1.009 & .991 & .976 \\ .984 & 1.016 & & \end{array}\right.$ |  |  |
| Cause: A,S | 1.111 . 884 |  |  | 1.098 . 895 |  |  | $1.089 \quad .902$ |  |  | $1.084 \quad .908$ |  |  |
| Age: 60-64 | . 965 |  |  | . 920 |  |  | . 890 |  |  | . 874 |  |  |
| EP: $0,7,14,30$ | 1.1961 .031 | . 896 | . 849 | 1.191 1.031 .910 .857 <br> 1.071 1.037 .977 .921 |  |  | 1.1801 .024 .917 .876  <br> 1.058 1.033 .980 .933 |  |  | 1.166 1.010 .919 .907 <br> 1.048 1.028 .982 .944 |  |  |
| Class: 1,2,3,4 | 1.0901 .040 | . 971 | . 906 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Sex: M,F | . 9881.012 |  |  | $1.011 \quad .988$ |  |  | $1.025 \quad .975$ |  |  | $\begin{array}{lll}1.024 & .976\end{array}$ |  |  |
| Cause: A,S | $1.260 \quad .780$ |  |  | $1.253 \quad .785$ |  |  | $11.245 \quad .790$ |  |  | $1.236 \quad .796$ |  |  |

EXHIBIT B-7-Continued
Factors for Calculation of Weekly Termination Rates Based on Graduated Data
( AVERAGE $=1$ )

| Week: | -10 |  |  | 11 |  |  | 12 |  |  | 13 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | . 117 |  |  | . 109 |  |  | . 099 |  |  | . 086 |  |  |  |
| Age: 20-29 | 1.086 |  |  | 1.096 |  |  | 1.110 |  |  | 1.133 |  |  |  |
| EP: 0,7.14,30 | $\begin{array}{llll} .951 & 1.087 & 1.038 & .921 \\ .972 & 1.002 & 1.013 & 1.013 \end{array}$ |  |  | . 9631.0511 .018 . 964 |  |  | . $9961.008 \quad .9851 .007$ |  |  | 1.059 .949 |  | . 9351.050 |  |
| Class: 1,2,3,4 |  |  |  | . 966 . 9941.0151 .026 |  |  | . 9571.9821 .0171 .045 |  |  | . 944 . 9641.0211 .074 |  |  |  |
| Sex: M,F | .9721 .0021 .0131 .013.9971 .002 |  |  | . 9901.008 |  |  | . 9841.013 |  |  | . 975 | 1.018 |  |  |
| Cause: A.S | . 8711.131 |  |  | . 8761.127 |  |  | . 8841.118 |  |  | . 8971.104 |  |  |  |
| Age: 30-39 | 1.068 |  |  | 1.062 |  |  | 1.049 |  |  | 1.027 |  |  |  |
| EP: 0,7.14,30 | $\begin{array}{ccccc}1.025 & 1.069 & 1.019 & .885 \\ .986 & .997 & 1.006 & 1.010\end{array}$ |  |  | $1.0031 .058 \quad 1.017 \quad .920$ |  |  | . $9851.0491 .008 \quad .955$ |  |  | $\begin{array}{rrrr} .971 & 1.038 & .989 & .992 \\ .962 & .993 & 1.012 & 1.032 \end{array}$ |  |  |  |
| Class: 1,2,3,4 |  |  |  | . $981 \quad .9961 .0071 .015$ |  |  | . 974 . 9941.0091 .022 |  |  |  |  |  |  |
| Sex: M,F | $\begin{aligned} & 9611.040 \\ & .9191 .072 \end{aligned}$ |  |  | . 9581.042 |  |  | . 9591.039 |  |  | $\begin{array}{llll} .962 & .993 & 1.0121 .032 \end{array}$ |  |  |  |
| Cause: A,S |  |  |  | . 9301.060 |  |  | . 9501.040 |  |  | $.9841 .006$ |  |  |  |
| Age: 40-49 | 1.022 |  |  | 1.012 |  |  | . 993 |  |  | . 962 |  |  |  |
| EP: 0,7,14,30 | 1.083 1.040 .995 .882 <br> 1.002 .999 .999 1.000 <br> .951 1.050   <br> .986 .999   |  |  | $\left[\begin{array}{ccc} 1.048 & 1.039 & .998 \\ 1.001 & 1.000 & 1.000 \\ .946 & 1.055 \\ .998 & .989 \end{array}\right.$ |  |  | $\begin{array}{ll} 1.0071 .043 \\ 1.000 & 1.003 \end{array}$ | .9971.000 |  | $\begin{array}{lll} .9521 .054 & .989 & .995 \end{array}$ |  |  |  |
| Class: 1,2,3,4 |  |  |  | . 995 | 1.000 |  |  |  | $1.008$ | $1.001$ | . 989 |
| Sex: M.F |  |  |  |  | . 9431.057 |  |  | . 942 | 1.053 |  |  |
| Cause: A,S |  |  |  |  | $1.020 \quad .969$ |  |  | 1.058 | . 935 |  |  |
| Age: 50-59 | . 953 |  |  |  |  | . 948 |  |  | . 941 |  |  | . 932 |  |  |  |
| EP: 0,7,14.30 | 1.1211 .013 | . 967 | . 900 |  |  | 1.0901 .005 | . 966 | . 938 | $1.052 \quad .997$ | . 959 | . 989 | . 999 | . 988 | . 943 | 1.062 |
| Class: 1,2,3.4 | 1.0191 .005 | . 993 | . 982 |  |  | 1.0221 .006 | . 992 | . 980 | 1.0311 .009 | . 989 | .971 | 1.048 | 1.015 | . 984 | . 953 |
| Sex: M,F | . 9661.034 |  |  | . 9531.048 |  |  | . 9351.066 |  |  | . 908 | 1.092 |  |  |
| Cause: A,S | $1.082 \quad .910$ |  |  | $1.086 \quad .909$ |  |  | $1.094 \quad .904$ |  |  | 1.110 | . 891 |  |  |
| Age: $60-64$ | . 871 |  |  | . 881 |  |  | . 907 |  |  | . 946 |  |  |  |
| EP: 0,7,14,30 | $\begin{array}{rr} 1.147 & .987 \\ 1.043 & 1.022 \\ 1.008 & .991 \\ 1.223 & .806 \end{array}$ |  | . 951 | $\left\{\begin{array}{lr} 1.119 & .956 \\ 1.041 & 1.016 \\ .975 & 1.024 \\ 1.210 & .816 \end{array}\right.$ | $\begin{array}{rr} .913 & 1.017 \\ .984 & .961 \end{array}$ |  | 1.079 .914 .906 1.114 <br> 1.043 1.009 .982 .967 |  |  | $\left\lvert\, \begin{array}{rr} 1.024 & .853 \\ 1.052 & .998 \\ .844 & 1.175 \end{array}\right.$ |  | . 8941.265 |  |
| Class: 1,2,3,4 |  |  | . 953 |  |  |  | . 978.972 |  |  |  |
| Sex: M,F |  |  |  |  |  |  |  |  | $\begin{array}{r} 1.0431 .009 \\ .9201 .083 \end{array}$ |  |  | $.982$ | $.967$ |
| Cause: A.S |  | $11.223 \quad .806$ |  |  | $1.210 \quad 816$ |  | $1.193 \quad 829$ |  |  | 1.166 | $.849$ |  |  |

## Ultimate Termination Rates-Durations 11 and Higher

For this purpose the new termination data collected from writers of individual disability insurance did not include sufficient exposures at the longer durations to be of any value. It was therefore necessary to rely upon the published data with respect to Group LTD insurance, the most recent study of experience under individual Waiver of Premium benefits, the experience under Social Security, and the study by Mutual of Omaha of its termination experience. From an analysis of all of these data, the target values at quinquennial ages were developed and from these the following formulas were devised for the development of graduated values.

For all terminations (death and recovery):
the termination rate $q_{x}^{t}=1.022-p_{x}^{t}$

$$
\text { where } 10^{4} \operatorname{colog} p_{x}^{t}=10^{.040(x-7.6)}
$$

For terminations due to death only:

$$
\begin{gathered}
q_{x}^{d}=1.007-p_{x}^{d} \\
\text { where } 10^{4} \operatorname{colog} p_{x}^{d}=10^{035(x+4.0)}
\end{gathered}
$$

The resulting values and comparisons with other data are presented in the accompanying table, Exhibit B-8, and the full set of Ultimate Termination Rates in Exhibit B-9.

With respect to the comparisons with experience under Benefit 2 and the Group Waiver of Premium benefits, it should be noted that there has been a considerable passage of time between the experience years involved and the present, the experience of which we are endeavoring to reflect.

Ultimate rates for females are .67 times the corresponding male rates. The factor of .67 was set empirically by reviewing the relationship of female to male death rates from several mortality tables and the ratios of recovery rates and total termination rates for Group LTD experience. Using the flat ratio leaves a practical working formula for generating rates at all ages with a minimum amount of conservatism.

## EXHIBIT B-8

Comparison of Disability Termination Rates per 1,000

| AcF. <br> ( $x$ ) | Cause* | DTS Ulitimate Rate (Malcs) | Ordinary <br> Waiver <br> 1969-74 | $\begin{aligned} & \text { OASDI } \\ & 1973-76 \end{aligned}$ | OASDI 1975-78 | OASDI 1973-77 | Group <br> Waiver <br> 1955-64 | $\begin{gathered} \text { Ben. } 2 \& 3 \\ (x+1 / 2) \\ 1930-50 \end{gathered}$ | Mutual of Omaha 1970-77 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | ars) | $\begin{aligned} & 11+ \\ & (1) \\ & \hline \end{aligned}$ | $11+$ (2) | $6+$ <br> (3) | $6+$ (4) | $\begin{aligned} & 11+ \\ & (5) \end{aligned}$ | (6) | (7) | (8) |
| 22 | D R T | $\begin{array}{r} 8.9 \\ 14.0 \\ 22.9 \end{array}$ |  |  |  |  |  |  |  |
| 27 | D R T | $\begin{array}{r} 9.8 \\ 13.6 \\ 23.4 \end{array}$ |  | $\begin{array}{r} 9.9 \\ 34.0 \\ 43.9 \end{array}$ | $\begin{array}{r} 9.7 \\ 33.9 \\ 43.6 \end{array}$ |  | $\begin{aligned} & 16 \\ & 20 \\ & 36 \end{aligned}$ |  |  |
| 32 | D R T | $\begin{aligned} & 11.2 \\ & 12.9 \\ & 24.1 \end{aligned}$ | $\begin{aligned} & 19.7 \\ & 19.7 \\ & 39.4 \end{aligned}$ | $\begin{aligned} & 12.6 \\ & 19.6 \\ & 32.2 \end{aligned}$ | $\begin{aligned} & 12.7 \\ & 20.4 \\ & 33.1 \end{aligned}$ | 63.1 | $\begin{aligned} & 17 \\ & 19 \\ & 36 \end{aligned}$ | $\begin{aligned} & 12.6 \\ & 52.8 \\ & 65.4 \end{aligned}$ |  |
| 37 | D R T | $\begin{aligned} & 13.2 \\ & 12.2 \\ & 25.4 \end{aligned}$ | $\begin{array}{r} 13.6 \\ 6.8 \\ 20.4 \end{array}$ | $\begin{aligned} & 16.0 \\ & 11.3 \\ & 27.3 \end{aligned}$ | $\begin{aligned} & 15.9 \\ & 12.4 \\ & 28.3 \end{aligned}$ | 21.3 | $\begin{aligned} & 18 \\ & 18 \\ & 36 \end{aligned}$ | $\begin{aligned} & 15.5 \\ & 46.3 \\ & 61.8 \end{aligned}$ | 40 |
| 42 | D R T | $\begin{aligned} & 16.3 \\ & 11.2 \\ & 27.5 \end{aligned}$ | $\begin{aligned} & 12.8 \\ & 12.8 \\ & 25.6 \end{aligned}$ | $\begin{array}{r} 21.9 \\ 8.0 \\ 29.9 \end{array}$ | $\begin{array}{r} 21.0 \\ 8.8 \\ 29.8 \end{array}$ | 25.1 | $\begin{aligned} & 26 \\ & 16 \\ & 42 \end{aligned}$ | $\begin{aligned} & 19.1 \\ & 39.8 \\ & 58.9 \end{aligned}$ | 49 |
| 47 | D R T | $\begin{array}{r} 20.9 \\ 9.7 \\ 30.6 \end{array}$ | $\begin{array}{r} 18.1 \\ 9.8 \\ 27.9 \end{array}$ | $\begin{array}{r} 28.8 \\ 5.4 \\ 34.2 \end{array}$ | $\begin{array}{r} 27.9 \\ 6.3 \\ 34.2 \end{array}$ | 29.9 | $\begin{aligned} & 33 \\ & 14 \\ & 47 \end{aligned}$ | $\begin{aligned} & 22.2 \\ & 33.3 \\ & 55.5 \end{aligned}$ | 42 |
| 52 | D R T | $\begin{array}{r} 27.8 \\ 7.8 \\ 35.6 \end{array}$ | $\begin{array}{r} 37.7 \\ 6.6 \\ 44.3 \end{array}$ | $\begin{array}{r} 39.0 \\ 3.1 \\ 42.1 \end{array}$ | $\begin{array}{r} 37.6 \\ 3.7 \\ 41.3 \end{array}$ | 38.5 | $\begin{aligned} & 39 \\ & 12 \\ & 51 \end{aligned}$ | $\begin{aligned} & 25.8 \\ & 26.8 \\ & 52.6 \end{aligned}$ | 64 |
| 57 | D R T | $\begin{array}{r} 37.9 \\ 5.7 \\ 43.6 \end{array}$ | $\begin{array}{r} 37.1 \\ 4.6 \\ 41.7 \end{array}$ | $\begin{array}{r} 51.6 \\ 1.6 \\ 53.2 \end{array}$ | $\begin{array}{r} 48.1 \\ 1.9 \\ 50.0 \end{array}$ | 49.5 | $\begin{array}{r} 46 \\ 7 \\ 53 \end{array}$ | $\begin{aligned} & 33.4 \\ & 20.2 \\ & 53.6 \end{aligned}$ | 64 |
| 62 | D R T | $\begin{array}{r} 52.9 \\ 3.0 \\ 55.9 \end{array}$ | $\begin{array}{r} 67.2 \\ 3.7 \\ 70.9 \end{array}$ | $\begin{array}{r} 54.2 \\ 1.4 \\ 55.6 \end{array}$ | $\begin{array}{r} 60.8 \\ .8 \\ 61.6 \end{array}$ | 61.7 | $\begin{array}{r} 58 \\ 5 \\ 63 \\ \hline \end{array}$ | $\begin{aligned} & 47.7 \\ & 13.7 \\ & 61.4 \end{aligned}$ | 52 |

* $\mathrm{D}=$ death; $\mathrm{R}=$ recovery; $\mathrm{T}=$ death and recovery.
(1) Rates based on the DTS formula for graduating the ultimate rates.
(2) Data provided by Mr. John H. Cook, from contributions to the intercompany Disability Waiver of Premium study.
(3) Actuarial Study No. 75 (Social Security).
(4) Actuarial Study No. 81 (Social Security).
(5) Data supplied by Mr. Francisco R. Bayo for ultimate experience after first 10 years of disablement.
(6) TSA 1968 Reports, page 194.
(7) TSA 1952 Reports, page 106.
(8) Derived from recent termination study by Mutual of Omaha,


## EXHIBIT B-9

Ultimate Termination Rates for
Duration 11 Years and Over
by Attained Age

| $\begin{aligned} & \text { Amtained } \\ & \text { Age } \end{aligned}$ | Mall: | Female | $\begin{gathered} \text { Atiantid } \\ \text { Age } \\ \hline \end{gathered}$ | Male | Femaie |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 30 | . 0238 | . 0160 | 65 | . 0665 | . 0446 |
| 31 | . 0240 | . 0161 | 66 | . 0707 | . 0474 |
| 32 | . 0242 | . 0162 | 67 | . 0753 | . 0504 |
| 33 | . 0244 | . 0163 | 68 | . 0802 | . 0538 |
| 34 | . 0246 | . 0165 | 69 | . 0857 | . 0574 |
| 35 | . 0249 | . 0167 | 70 | . 0916 | . 0614 |
| 36 | . 0251 | . 0168 | 71 | . 0986 | . 0657 |
| 37 | . 0254 | . 0170 | 72 | . 1051 | . 0704 |
| 38 | . 0258 | . 0173 | 73 | . 1127 | . 0755 |
| 39 | . 0261 | . 0175 | 74 | . 1210 | . 0811 |
| 40 | . 0265 | . 0178 | 75 | . 1301 | . 0871 |
| 41 | . 0270 | . 0181 | 76 | . 1398 | . 0937 |
| 42 | . 0275 | . 0184 | 77 | . 1504 | . 1008 |
| 43 | . 0280 | . 0188 | 78 | . 1619 | . 1085 |
| 44 | . 0286 | . 0192 | 79 | . 1743 | . 1168 |
| 45 | . 0292 | . 0196 | 80 | . 1878 | . 1258 |
| 46 | . 0299 | . 0200 | 81 | . 2022 | 1355 |
| 47 | . 0306 | . 0205 | 82 | . 2178 | . 1459 |
| 48 | . 0315 | . 0211 | 83 | . 2345 | . 1571 |
| 49 | . 0324 | . 0217 | 84 | 2525 | 1691 |
| 50 | 0334 | . 0224 | 85 | . 2717 | . 1820 |
| 51 | . 0345 | . 0231 | 86 | 2922 | . 1958 |
| 52 | . 0357 | . 0239 | 87 | . 3140 | . 2104 |
| 53 | . 0370 | . 0248 | 88 | . 3372 | . 2259 |
| 54 | . 0384 | . 0257 | 89 | . 3618 | . 2424 |
| 55 | . 0400 | . 0268 | 90 | . 3877 | . 2598 |
| 56 | . 0417 | . 0279 | 91 | . 4149 | . 2780 |
| 57 | . 0436 | . 0292 | 92 | . 4435 | . 2971 |
| 58 | . 0456 | . 0306 | 93 | . 4732 | . 3171 |
| 59 | . 0479 | . 0321 | 94 | . 5041 | . 3378 |
| 60 | 0503 | . 0337 | 95 | . 5360 | .3591 |
| 61 | . 0530 | . 0355 | 96 | . 5686 | . 3801 |
| 62 | . 0559 | . 0375 | 97 | . 6020 | . 4033 |
| 63 | . 0592 | . 0397 | 98 | . 6357 | 4259 |
| 64 | . 0627 | . 0420 | 99 | . 6695 | 4486 |

## APPENDIX C

## DEVELOPMENT OF INCIDENCE RATES

Only five companies furnished data in the age and class detail that the Committee requested, and because of the inadequate amount of data, many of the cells were deemed too small for our purpose. Accordingly, we approached the construction of incidence rates by using the 1976-79 industry data as our underlying base, and the results of the 1976 New York study (data base 1967-73) as a source of relationship among occupation classes. The SOA data are by class groups I (N.Y. classes 1 and 2 ) and II (N.Y. classes 3 and 4).

The basic assumption was that the ratio of the incidence rate for class 1 (or 2) to the incidence rate for classes 1 and 2 combined is the same for the SOA data as for the New York study. Ratios of the New York study incidence rates for each decennial age group by sex, cause, and elimination period were calculated for class I and class 2 and multiplied by the SOA corresponding incidence rate for class group I to obtain the generated SOA rates for class 1 and for class 2.

The same process was followed to determine rates for classes 3 and 4 from SOA class group II. Incidence rates were calculated separately, by identical methods, for accident and sickness.

Crude incidence rates determined by this method appeared to be in good shape for males, but not quite as good for females, as evidenced by the "crude rate" graphs. We tried several different graduating approaches on these crude rates, without success. Mechanical graduating methods did not seem to improve them, so a multidimensional graphic method was used to obtain graduated rates for both males and females. These graduated rates were then applied against SOA exposure distribution for each cell (based upon the New York exposure distribution) and modified very slightly so that the final graduated rates produce the same number of claims as the SOA data for each class group (I and II) and each cause (accident and sickness). This method worked well for the male rates and classes 1, 2, and 3, for females, but the volume of data was so small for female class 4 , that we could not produce class 4 results by this method. Accordingly, by studying the male results and the results for females at classes 1,2 , and 3 , we concluded that our best estimate for female class 4 was to generate class 4 incidence rates by dividing the class 3 accident rates by .80 and the class 3 sickness rates by 96 .

Summaries of the data for the five responding companies, the SOA and the New York study are included here as Exhibit C-1. The generated crude
incidence rates are in the outer column, labeled Generated Four Occ.
Graphs of the crude data and graduated data are included in Exhibit C-2. Associated with each set of crude and graduated rates is a corresponding set of implied 90 -day rates. The implied 90 -day rate is the probability of becoming disabled and remaining disabled through the ninetieth day. This $90-$ day point was a reference point that was used to provide another dimension to the graduation. It was used to evaluate the logical consistency between tables at that point in time. Graphs of the imputed rates also are included.

Incidence rates were determined from exposures in 10-year age groupings. The rates for each age group were not assigned to the central age of the age group. Rather, they were assigned to a more precise weighted age determined from the exposure distribution of the DTS quinquennial age data. These weighted ages are $25.5,34.5,44.5,54.0$, and 62.2 . Rates for ages $25,35,45,55$, and 62 were determined by interpolating with a fourth degree polynomial. Final graduated rates are shown in Exhibit C-3.

EXHIBIT C-I
(TWO OCC. DATA)


EXHIBIT C-1-Continued
(TWO OCC. DATA)


EXHIBIT C-1-Continued
(TWO OCC. DATA)


EXHIBIT C-1--Continued
(TWO OCC. DATA)


EXHIBIT C-1-Continued
(TWO OCC. DATA)

| S | COC | EP | X | SOA Two Occ. Data |  |  | New York Stuid |  |  | Four Company Decennial |  |  | Five Company Orioinal |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Expusure | Claims | Rate | Exposure | Claims | Rate |
| F | A I | 0 | 25 | 3,445 | 107 | 31.06 | 12,002 | 352 | 29.33 | 337 | 8 | 23.74 | 421 | 9 | 21.38 |
| F | A I | 0 | 35 | 4,830 | 144 | 29.81 | 14,709 | 457 | 31.07 | 388 | 14 | 36.08 | 464 | 20 | 43.10 |
| F | A I | 0 | 45 | 5,687 | 144 | 25.32 | 23.031 | 797 | 34.61 | 506 | 16 | 31.62 | 554 | 18 | 32.49 |
| F | AI | 0 | 55 | 8.809 | 258 | 29.29 | 23,595 | 971 | 41.15 | 869 | 40 | 46.03 | 890 | 40 | 44.94 |
| F | A I | 0 | 65 | 4,095 | 162 | 39.56 | 7,627 | 273 | 35.79 | 502 | 27 | 53.78 | 502 | 27 | 53.78 |
| F | AI | 7 | 25 | 4,090 | 97 | 23.72 | 10,603 | 248 | 23.39 | 1,216 | 31 | 25.49 | 2,375 | 69 | 29.05 |
| F | A I | 7 | 35 | 9,778 | 222 | 22.70 | 14,852 | 367 | 24.71 | 2,973 | 61 | 20.52 | 5,662 | 129 | 22.78 |
| F | A I | 7 | 45 | 10,957 | 276 | 25.19 | 24,078 | 528 | 21.93 | 3,805 | 62 | 16.29 | 8,008 | 168 | 20.98 |
| F | A I | 7 | 55 | 14,846 | 359 | 24.18 | 26,782 | 740 | 27.63 | 4,983 | 126 | 25.29 | 11,592 | 290 | 25.02 |
| F | AI | 7 | 65 | 1,665 | 39 | 23.42 | 2,512 | 66 | 26.27 | 601 | 15 | 24.96 | . 957 | 24 | 25.08 |
| F | A I | 14 | 25 | 11.772 | 208 | 17.67 | 8,597 | 101 | 11.75 | 5,056 | 77 | 15.23 | 5,751 | 89 | 15.48 |
| F | A I | 14 | 35 | 16,940 | 338 | 19.95 | 10,350 | 137 | 13.24 | 5,353 | 114 | 21.30 | 6,826 | 137 | 20.07 |
| F | AI | 14 | 45 | 12.817 | 222 | 17.32 | 13,508 | 165 | 12.21 | 4,302 | 68 | 15.81 | 6,096 | 96 | 15.75 |
| F | A I | 14 | 55 | 11,788 | 231 | 19.60 | 11,044 | 191 | 17.29 | 3,383 | 62 | 18.33 | 5,202 | 90 | 17.30 |
| F | A 1 | 14 | 65 | 3,028 | 82 | 27.08 | 2,673 | 40 | 14.96 | 687 | 14 | 20.38 | $\bigcirc 732$ | 15 | 20.49 |
| F | AI | 30 | 25 | 20.955 | 130 | 6.20 | 14,351 | 54 | 3.76 | 6,567 | 32 | 4.87 | 6.759 | 32 | 4.73 |
| F | A I | 30 | 35 | 29.841 | 231 | 7.74 | 20,106 | 98 | 4.87 | 8,760 | 61 | 6.96 | 9.076 | 64 | 7.05 |
| F | A I | 30 | 45 | 21,428 | 161 | 7.51 | 26,844 | 159 | 5.92 | 7,905 | 66 | 8.35 | 8,192 | 68 | 8.30 |
| F | A ! | 30 | 55 | 16,839 | 164 | 9.74 | 21,188 | 158 | 7.46 | 6,845 | 54 | 7.89 | 6,985 | 55 | 7.87 |
| F | Al | 30 | 65 | 3,015 | 31 | 10.28 | 4.024 | 36 | 8.95 | 899 | 6 | 6.67 | 9.901 | 6 | 6.66 |
| F | A I | 90 | 25 | 3,370 | 6 | 1.78 | 2.581 | 1 | 0.39 | 852 | 1 | 1.17 | 911 | 2 | 2.20 |
| F | A I | 90 | 35 | 6,862 | 4 | 0.58 | 4,195 | 11 | 2.62 | 1,574 |  |  | 1.717 |  |  |
| F | AI | 90 | 45 | 5,193 | 7 | 1.35 | 6,578 | 9 | 1.37 | 1,376 | 2 | 1.45 | 1.710 | 2 | 1.17 |
| F | A I | 90 | 55 | 4,007 | 9 | 2.25 | 6,196 | 15 | 2.42 | 1,189 | 6 | 5.05 | 1,616 | 8 | 4.95 |
| F | A I | 90 | 65 | 556 | 2 | 3.60 | 1.026 | 3 | 2.92 | 135 |  |  | 146 |  |  |
| SUBTOTAL |  |  |  | 236,613 | 3,634 |  | 313,052 | 5.977 |  | 71,063 | 963 |  | 94,045 | 1,458 |  |

EXHIBIT C-1-Continued
(TWO OCC. DATA)

| S | COC | EP | X | SOA Twu Occ. Data |  |  | New Yokk Studr |  |  | Four Company Decennial |  |  | Five Company Original. |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Ratt: |
| F | A II | 0 | 25 | 207 | 12 | 57.97 | 2,013 | 115 | 57.13 | 16 | 2 | 125.00 | 16 | 2 | 125.00 |
| F | A II | 0 | 35 | 706 | 38 | 53.82 | 2,875 | 174 | 60.52 | 38 |  |  | 38 |  |  |
| F | A 11 | 0 | 45 | 1,127 | 44 | 39.04 | 4,236 | 292 | 68.93 | 55 | 2 | 36.36 | 55 | 2 | 36.36 |
| F | A 11 | 0 | 55 | 1.818 | 73 | 40.15 | 3,688 | 223 | 60.47 | 109 | 2 | 18.35 | 109 | 2 | 18.35 |
| F | A II | 0 | 65 | 995 | 33 | 33.17 | 900 | 49 | 54.44 | 41 |  |  | 41 |  |  |
| F | A II | 7 | 25 | 689 | 12 | 17.42 | 2,822 | 126 | 44.65 | 345 | 8 | 23.19 | 374 | 9 | 24.06 |
| F | A II | 7 | 35 | 2,095 | 70 | 33.41 | 3,792 | 174 | 45.89 | 1,056 | 46 | 43.56 | 1,116 | 51 | 45.70 |
| F | A II | 7 | 45 | 2,312 | 80 | 34.60 | 4,414 | 252 | 57.09 | 1,206 | 32 | 26.53 | 1,245 | 36 | 28.92 |
| F | A II | 7 | 55 | 2,414 | 107 | 44.32 | 2,806 | 127 | 45.26 | 1,231 | 42 | 34.12 | 1,262 | 44 | 34.87 |
| F | A II | 7 | 65 | 258 | 15 | 58.14 | 123 |  |  | . 55 | 4 | 72.73 | 1.262 | 4 | 70.18 |
| F | A III | 14 | 25 | 2,033 | 59 | 29.02 | 1.362 | 23 | 16.89 | 1,435 | 34 | 23.69 | 1,438 | 34 | 23.64 |
| F | A II | 14 | 35 | 3,106 | 94 | 30.26 | 1,411 | 34 | 24.10 | 1,537 | 56 | 36.43 | 1,546 | 57 | 36.87 |
| F | A II | 14 | 45 | 2,638 | 103 | 39.04 | 1,650 | 54 | 32.73 | 1,409 | 51 | 36.20 | 1,423 | 51 | 35.84 |
| F | A II | 14 | 55 | 1,833 | 59 | 32.19 | 1,161 | 38 | 32.73 | 891 | 29 | 32.55 | 902 | 29 | 32.15 |
| F | A II | 14 | 65 | 440 | 8 | 18.18 | 222 | 12 | 54.05 | 128 | 3 | 23.44 | 128 | 3 | 23.44 |
| F | A II | 30 | 25 | 1,931 | 27 | 13.98 | 1,507 | 9 | 5.97 | 871 | 14 | 16.07 | 883 | 14 | 15.86 |
| F | A II | 30 | 35 | 2,993 | 59 | 19.71 | 1,649 | 23 | 13.95 | 1,168 | 27 | 23.12 | 1,182 | 27 | 22.84 |
| F | A II | 30 | 45 | 2,229 | 45 | 20.19 | 2,038 | 60 | 29.44 | 1,094 | 29 | 26.51 | 1.116 | 29 | 25.99 |
| F | A II | 30 | 55 | 1,799 | 34 | 18.90 | 1,332 | 37 | 27.78 | 843 | 23 | 27.28 | 852 | 23 | 27.00 |
| F | A II | 30 | 65 | 196 | 5 | 25.51 | 174 | 2 | 11.49 | 51 | 2 | 39.22 | 53 | 2 | 37.74 |
| F | A II | 90 | 25 | 169 |  |  | 89 |  |  | 44 |  |  | 44 |  |  |
| F | A II | 90 | 35 | 200 | 1 | 5.00 | 94 | 3 | 31.91 | 52 |  |  | 52 |  |  |
| F | A II | 90 | 45 | 113 | 1 | 8.85 | 183 | 4 | 21.86 | 32 | 1 | 31.25 | 32 | 1 | 31.25 |
| F | A II | 90 | 55 | 97 |  |  | 188 | 3 | 15.96 | 37 |  |  | 37 |  |  |
| F | A II | 90 | 65 | 13 |  |  | 45 |  |  | 8 |  |  | 8 |  |  |
|  | SUBT | AL |  | 32,411 | 979 |  | 40,774 | 1,834 |  | 13,752 | 407 |  | 14,009 | 420 |  |
|  | PAGE T | TAL |  | 269,024 | 4,613 |  | 353,826 | 7,811 |  | 84,815 | 1,370 |  | 108,054 | 1,878 |  |

EXHIBIT C-1-Continued
(TWO OCC. DATA)


EXHIBIT C-1-Continued
(TWO OCC. DATA)

|  |  |  |  |  | SOA | OCC. Da |  | New | York Stud |  | Four Co | Pany Di | cennial. | Five Com | Any Or | Inal |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| S | C | OC | EP | X | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |
| F | S | II | 0 | 25 |  |  |  |  |  |  |  |  |  |  |  |  |
| F | S | II | 0 | 35 | 12 |  |  |  |  |  |  |  |  |  |  |  |
| F | S | 11 | 0 | 45 | 33 | 4 | 121.21 |  |  |  |  |  |  |  |  |  |
| F | S | II | 0 | 55 | 133 | 20 | 150.38 |  |  |  | 2 |  |  | 2 |  |  |
| F | S | 11 | 0 | 65 | 45 | 14 | 311.11 |  |  |  |  |  |  |  |  |  |
| F | S | II | 7 | 25 | 817 | 63 | 77.11 | 4,765 | 552 | 115.84 | 350 | 34 | 97.14 | 379 | 40 | 105.54 |
| F | S | II | 7 | 35 | 2,727 | 319 | 116.98 | 6,159 | 991 | 160.90 | 1,090 | 141 | 129.36 | 1,150 | 151 | 131.30 |
| F | S | II | 7 | 45 | 3,224 | 412 | 127.79 | 7,842 | 1,418 | 180.82 | 1,240 | 193 | 155.65 | 1,279 | 201 | 157.15 |
| F | S | II | 7 | 55 | 3,498 | 456 | 130.36 | 5,670 | 973 | 171.60 | 1,328 | 172 | 129.52 | 1,359 | 179 | 131.71 |
| F | S | II | 7 | 65 | 468 | 54 | 115.38 | 801 | 128 | 159.80 | 87 | 10 | 114.94 | - 89 | 10 | 112.36 |
| F | S | II | 14 | 25 | 2,052 | 125 | 60.92 | 1,466 | 70 | 47.75 | 1,438 | 114 | 79.28 | 1,441 | 115 | 79.81 |
| F | S | II | 14 | 35 | 3,122 | 305 | 97.69 | 1,597 | 166 | 103.94 | 1,542 | 181 | 117.38 | 1,551 | 183 | 117.99 |
| F | S | II | 14 | 45 | 2,669 | 356 | 133.38 | 1,972 | 235 | 119.17 | 1,422 | 219 | 154.01 | 1,436 | 221 | 153.90 |
| F | S | II | 14 | 55 | 1,888 | 198 | 104.87 | 1,551 | 164 | 105.74 | 912 | 122 | 133.77 | +923 | 122 | 132.18 |
| F | S | II | 14 | 65 | - 452 | 43 | 95.13 | - 320 | 41 | 128.13 | 136 | 13 | 95.59 | 136 | 13 | 95.59 |
| F | S | II | 30 | 25 | 1,948 | 51 | 26.18 | 1,620 | 34 | 20.99 | 878 | 34 | 38.72 | 890 | 35 | 39.33 |
| F | S | II | 30 | 35 | 2,993 | 146 | 48.78 | 1,859 | 87 | 46.80 | 1,169 | 67 | 57.31 | 1,183 | 67 | 56.64 |
| F | S | II | 30 | 45 | 2,250 | 122 | 54.22 | 2,366 | 167 | 70.58 | 1,106 | 75 | 67.81 | 1,128 | 75 | 66.49 |
| F | S | II | 30 | 55 | 1,872 | 115 | 61.43 | 1,663 | 98 | 58.93 | 843 | 61 | 72.36 | - 853 | 61 | 71.51 |
| F | S | II | 30 | 65 | 214 | 16 | 74.77 | 253 | 16 | 63.24 | 52 | 1 | 19.23 | 54 | 1 | 18.52 |
| F | S | II | 90 | 25 | 167 |  |  | - 92 | 3 | 32.61 | 44 |  |  | 44 |  |  |
| F | S | II | 90 | 35 | 199 |  |  | 94 | 4 | 42.55 | 52 | 1 | 19.23 | 52 | 1 | 19.23 |
| F | S | II | 90 | 45 | 113 | 1 | 8.85 | 183 | 5 | 27.32 | 32 |  |  | 32 |  | 1.23 |
| F | S | II | 90 | 55 | 97 | 1 | 10.31 | 186 | 6 | 32.26 | 37 |  |  | 37 |  |  |
| F | S | II | 90 | 65 | 13 |  |  | 44 | 5 | 113.64 | 8 |  |  | 8 |  |  |
| SUBTOTAL |  |  |  |  | 31,006 | 2,821 |  | 40,503 | 5,163 |  | 13,768 | 1,438 |  | 14,026 | 1,475 |  |
| PAGE TOTAL |  |  |  |  | 262,718 | 14,700 |  | 345,896 | 25,090 |  | 84,683 | 5,272 |  | 107,923 | 7,450 |  |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Foul Company Dricennial |  |  | Five Company Original |  |  | Niw York Ratios | $\begin{array}{\|c\|} \hline \text { SOA } \\ \hline \text { Two } \\ \text { Oce. } \end{array}$ | $\begin{aligned} & \text { Generated } \\ & \text { Four Oce. } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| M | A 1 | 0 | 25 | 35.614 | 1,186 | 33.30 | 1,344 | 57 | 42.41 | 1.875 | 78 | 41.60 | 0.7429 | 41.93 | 31.15 |
| M | A 1 | 0 | 35 | 128,713 | 3,308 | 25.70 | 7,185 | 202 | 28.11 | 8,471 | 275 | 32.46 | 0.7258 | 40.72 | 29.55 |
| M | A 1 | 0 | 45 | 211.797 | 4,833 | 22.82 | 10,119 | 226 | 22.33 | 12,099 | 289 | 23.89 | 0.7789 | 31.04 | 24.18 |
| M | A 1 | 0 | 55 | 209,773 | 4,757 | 22.68 | 11,264 | 215 | 19.09 | 16,101 | 365 | 22.67 | 0.8007 | 23.78 | 19.04 |
| M | A 1 | 0 | 65 | 154,607 | 3,736 | 24.16 | 4,461 | 111 | 24.88 | 6,120 | 168 | 27.45 | 0.9251 | 24.30 | 22.48 |
| M | A 1 | 7 | 25 | 11,888 | 309 | 25.99 | 1,013 | 23 | 22.70 | 2,452 | 83 | 33.85 | 0.7662 | 34.97 | 26.80 |
| M | A 1 | 7 | 35 | 24,775 | 465 | 18.77 | 3,258 | 64 | 19.64 | 6,529 | 185 | 28.34 | 0.7244 | 35.09 | 25.42 |
| M | A 1 | 7 | 45 | 32,992 | 544 | 16.49 | 3,577 | 61 | 17.05 | 8,909 | 189 | 21.21 | 0.7516 | 27.56 | 20.71 |
| M | A 1 | 7 | 55 | 30,868 | 528 | 17.11 | 2,992 | 69 | 23.06 | 12,698 | 271 | 21.34 | 0.8043 | 21.79 | 17.53 |
| M | A I | 7 | 65 | 10,662 | 184 | 17.26 | 938 | 17 | 18.12 | 3,147 | 57 | 18.11 | 0.7951 | 23.82 | 18.94 |
| M | A 1 | 14 | 25 | 18,468 | 161 | 8.72 | 5,061 | 68 | 13.44 | 5,923 | 81 | 13.68 | 0.5935 | 22.38 | 13.28 |
| M | A 1 | 14 | 35 | 59,394 | 443 | 7.46 | 10,094 | 148 | 14.66 | 13,077 | 178 | 13.61 | 0.6052 | 20.16 | 12.20 |
| M | A 1 | 14 | 45 | 84,604 | 566 | 6.69 | 9.074 | 86 | 9.48 | 13,947 | 130 | 9.32 | 0.6396 | 15.63 | 10.00 |
| M | A 1 | 14 | 55 | 49,249 | 389 | 7.90 | 7,705 | 87 | 11.29 | 12,413 | 157 | 12.65 | 0.6636 | 14.51 | 9.63 |
| M | A 1 | 14 | 65 | 12,139 | 99 | 8.16 | 1,937 | 26 | 13.42 | 3,142 | 38 | 12.09 | 0.6979 | 15.89 | 11.09 |
| M | A 1 | 30 | 25 | 62,194 | 216 | 3.47 | 14,501 | 58 | 4.00 | 15,607 | 64 | 4.10 | 0.7692 | 6.29 | 4.84 |
| M | A 1 | 30 | 35 | 249,604 | 769 | 3.08 | 47,429 | 172 | 3.63 | 51,288 | 195 | 3.80 | 0.7831 | 5.43 | 4.25 |
| M | A 1 | 30 | 45 | 260,512 | 899 | 3.45 | 40,329 | 160 | 3.97 | 45,885 | 188 | 4.10 | 0.8021 | 5.65 | 4.53 |
| M | A 1 | 30 | 55 | 119,037 | 490 | 4.12 | 25,056 | 134 | 5.35 | 29,209 | 160 | 5.48 | 0.8009 | 5.88 | 4.71 |
| M | A 1 | 30 | 65 | 28,444 | 124 | 4.36 | 5,314 | 23 | 4.33 | 9,095 | 52 | 5.72 | 0.8026 | 6.92 | 5.56 |
| M | A 1 | 90 | 25 | 16.883 | 8 | 0.47 | 3,989 | 1 | 0.25 | 4,382 | 1 | 0.23 | 0.7240 | 1.12 | 0.81 |
| M | A 1 | 90 | 35 | 109,295 | 55 | 0.50 | 20,667 | 4 | 0.19 | 22,537 | 6 | 0.27 | 0.7819 | 0.65 | 0.51 |
| M | A 1 | 90 | 45 | 134,900 | 104 | 0.77 | 17,837 | 13 | 0.73 | 20,989 | 18 | 0.86 | 0.8116 | 0.79 | 0.64 |
| M | A 1 | 90 | 55 | 64,408 | 71 | 1.10 | 9,207 | 8 | 0.87 | 11,856 | 9 | 0.76 | 0.7181 | 1.09 | 0.78 |
| M | A 1 | 90 | 65 | 14,487 | 33 | 2.28 | 1,905 | 2 | 1.05 | 2,457 | 2 | 0.81 | 0.7807 | 1.54 | 1.20 |

EXHIBIT C-1--Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Stuoy |  |  | Four Company Decennial |  |  | Five Company original |  |  | Niw York Ratios | $\begin{aligned} & \text { SOA } \\ & \hline \text { Two } \\ & \text { Oce. } \end{aligned}$ | GeneratedFour Occ. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| M | A 2 | 0 | 25 | 21,316 | 1,366 | 64.08 | 1,221 | 96 | 78.62 | 1,435 | 115 | 80.14 | 1.4296 | 41.93 | 59.94 |
| M | A 2 | 0 | 35 | 58,188 | 3,310 | 56.88 | 6,529 | 304 | 46.56 | 6,964 | 346 | 49.68 | 1.6065 | 40.72 | 65.41 |
| M | A 2 | 0 | 45 | 106,486 | 4,492 | 42.18 | 12,097 | 414 | 34.22 | 12,666 | 443 | 34.98 | 1.4398 | 31.04 | 44.70 |
| M | A 2 | 0 | 55 | 138,960 | 5,120 | 36.85 | 21,781 | 523 | 24.01 | 23,972 | 612 | 25.53 | 1. 3009 | 23.78 | 30.94 |
| M | A 2 | 0 | 65 | 67,515 | 2,066 | 30.60 | 11,126 | 284 | 25.53 | 11,906 | 330 | 27.72 | 1.1715 | 24.30 | 28.47 |
| M | A 2 | 7 | 25 | 12,521 | 519 | 41.45 | 1,498 | 78 | 52.07 | 3,562 | 213 | 59.80 | 1.2219 | 34.97 | 42.74 |
| M | A 2 | 7 | 35 | 23,507 | 786 | 33.44 | 4,808 | 182 | 37.85 | 8,672 | 375 | 43.24 | 1.2905 | 35.09 | 45.28 |
| M | A 2 | 7 | 45 | 35,150 | 951 | 27.06 | 5,954 | 183 | 30.74 | 11,722 | 365 | 31.14 | 1.2332 | 27.56 | 33.98 |
| M | A 2 | 7 | 55 | 39,194 | 962 | 24.54 | 6,439 | 181 | 28.11 | 17,468 | 438 | 25.07 | 1.1541 | 21.79 | 25.15 |
| M | A 2 | 7 | 65 | 12,420 | 317 | 25.52 | 2,258 | 68 | 30.12 | 4,871 | 134 | 27.51 | 1.1759 | 23.82 | 28.01 |
| M | A 2 | 14 | 25 | 16,049 | 346 | 21.56 | 8,824 | 251 | 28.45 | 9,562 | 271 | 28.34 | 1.4678 | 22.38 | 32.85 |
| M | A 2 | 14 | 35 | 28,807 | 644 | 22.36 | 12,516 | 364 | 29.08 | 14,948 | 420 | 28.10 | 1.8140 | 20.16 | 36.56 |
| M | A 2 | 14 | 45 | 33,272 | 667 | 20.05 | 10,057 | 234 | 23.27 | 13.693 | 318 | 23.22 | 1.9165 | 15.63 | 29.95 |
| M | A 2 | 14 | 55 | 26,367 | 511 | 19.38 | 9,073 | 168 | 18.52 | 12,969 | 241 | 18.58 | 1.6283 | 14.51 | 23.62 |
| M | A 2 | 14 | 65 | 6,260 | 116 | 18.53 | 2,810 | 56 | 19.93 | 3,735 | 77 | 20.62 | 1.5858 | 15.89 | 25.21 |
| M | A 2 | 30 | 25 | 24,407 | 175 | 7.17 | 10,102 | 106 | 10.49 | 10,729 | 111 | 10.35 | 1.5881 | 6.29 | 10.00 |
| M | A 2 | 30 | 35 | 67,629 | 479 | 7.08 | 23,292 | 182 | 7.81 | 24,840 | 200 | 8.05 | 1.8004 | 5.43 | 9.77 |
| M | A 2 | 30 | 45 | 78,371 | 559 | 7.13 | 22,639 | 186 | 8.22 | 24,716 | 215 | 8.70 | 1.6579 | 5.65 | 9.36 |
| M | A 2 | 30 | 55 | 44,589 | 351 | 7.87 | 17,686 | 130 | 7.35 | 19,508 | 143 | 7.33 | 1.5316 | 5.88 | 9.01 |
| M | A 2 | 30 | 65 | 10,773 | 89 | 8.26 | 4,117 | 36 | 8.74 | 4,489 | 39 | 8.69 | 1.5211 | 6.92 | 10.53 |
| M | A 2 | 90 | 25 | 4,509 | 6 | 1.33 | 982 | 5 | 5.09 | 1,107 | 5 | 4.52 | 2.0333 | 1.12 | 2.27 |
| M | A 2 | 90 | 35 | 18,121 | 27 | 1.49 | 3,611 | 3 | 0.83 | 3,950 | 3 | 0.76 | 2.3152 | 0.65 | 1.50 |
| M | A 2 | 90 | 45 | 26,178 | 49 | 1.87 | 5,154 | 5 | 0.97 | 5,744 | 5 | 0.87 | 1.9706 | 0.79 | 1.56 |
| M | A 2 | 90 | 55 | 16,366 | 53 | 3.24 | 4,160 | 9 | 2.16 | 5,058 | 10 | 1.98 | 2.1095 | 1.09 | 2.30 |
| M | A 2 | 90 | 65 | 3,678 | 20 | 5.44 | 839 | 2 | 2.38 | 1,059 | 2 | 1.89 | 1.8637 | 1.54 | 2.88 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

|  | S | C OC | EP | X | New York Study |  |  | Four Company decennial |  |  | Five Company Original. |  |  | New York Ratios | $\begin{aligned} & \text { SOA } \\ & \text { Two } \\ & \text { Occ. } \end{aligned}$ | Generated <br> Four Oce. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
|  | M | A 3 | 0 | 25 | 62,163 | 6,927 | 111.43 | 3,829 | 354 | 92.45 | 3,829 | 354 | 92.45 | 0.9811 | 98.71 | 96.84 |
|  | M | A 3 | 0 | 35 | 126,406 | 11,221 | 88.77 | 21,313 | 1,270 | 59.59 | 21,313 | 1,270 | 59.59 | 0.9922 | 89.15 | 88.45 |
|  | M | A 3 | 0 | 45 | 164,959 | 11,417 | 69.21 | 34,297 | 1,569 | 45.75 | 34,297 | 1,569 | 45.75 | 0.9994 | 57.76 | 57.73 |
|  | M | A 3 | 0 | 55 | 144,940 | 8.201 | 56.58 | 44,472 | 1,657 | 37.26 | 44,472 | 1,657 | 37.26 | 1.0060 | 46.74 | 47.02 |
|  | M | A 3 | 0 | 65 | 46,918 | 2,357 | 50.24 | 19,053 | 596 | 31.28 | 19,053 | 596 | 31.28 | 1.0077 | 39.46 | 39.76 |
| $\underset{\infty}{n}$ | M | A 3 | 7 | 25 | 37,802 | 2,395 | 63.36 | 4,908 | 330 | 67.24 | 8,113 | 545 | 67.18 | 0.9648 | 58.26 | 56.21 |
|  | M | A 3 | 7 | 35 | 65,195 | 3,482 | 53.41 | 15,075 | 868 | 57.58 | 22,980 | 1,341 | 58.36 | 0.9424 | 58.52 | 55.15 |
|  | M | A 3 | 7 | 45 | 82,038 | 3,527 | 42.99 | 14,938 | 759 | 50.81 | 29,208 | 1,314 | 44.99 | 0.9589 | 46.69 | 44.77 |
|  | M | A 3 | 7 | 55 | 84,901 | 3,019 | 35.56 | 13,403 | 651 | 48.57 | 38,907 | 1,515 | 38.94 | 0.9589 | 39.99 | 38.35 |
|  | M | A 3 | 7 | 65 | 24.230 | 839 | 34.63 | 4,373 | 199 | 45.51 | 10,011 | 372 | 37.16 | 0.9843 | 35.92 | 35.36 |
|  | M | A 3 | 14 | 25 | 34,096 | 1.268 | 37.19 | 23,782 | 1,031 | 43.35 | 25,302 | 1,102 | 43.55 | 0.9860 | 48.38 | 47.70 |
|  | M | A 3 | 14 | 35 | 57,321 | 2,011 | 35.08 | 30,978 | 1,368 | 44.16 | 37,220 | 1,571 | 42.21 | 0.9701 | 47.23 | 45.82 |
|  | M | A 3 | 14 | 45 | 51,112 | 1,827 | 35.75 | 20,448 | 802 | 39.22 | 27,908 | 1,032 | 36.98 | 0.9830 | 39.81 | 39.14 |
|  | M | A 3 | 14 | 55 | 32,650 | 1.178 | 36.08 | 13,387 | 451 | 33.69 | 19,272 | 618 | 32.07 | 0.9922 | 34.09 | 33.82 |
|  | M | A 3 | 14 | 65 | 6,471 | 266 | 41.11 | 3,153 | 85 | 26.96 | 4.354 | 123 | 28.25 | 1.0552 | 27.67 | 29.20 |
|  | M | A 3 | 30 | 25 | 33,568 | 573 | 17.07 | 20,600 | 488 | 23.69 | 23,748 | 562 | 23.67 | 1.0036 | 24.07 | 24.16 |
|  | M | A 3 | 30 | 35 | 66,115 | 1.090 | 16.49 | 33,536 | 716 | 21.35 | 38,186 | 834 | 21.84 | 0.9414 | 24.24 | 22.82 |
|  | M | A 3 | 30 | 45 | 58,505 | 1.004 | 17.16 | 22,982 | 440 | 19.15 | 27,114 | 534 | 19.69 | 0.9765 | 22.93 | 22.39 |
|  | M | A 3 | 30 | 55 | 33,147 | 562 | 16.95 | 13,815 | 211 | 15.27 | 16,637 | 257 | 15.45 | 0.9793 | 18.67 | 18.28 |
|  | M | A 3 | 30 | 65 | 6.311 | 87 | 13.79 | 2,910 | 34 | 11.68 | 4,458 | 65 | 14.58 | 0.9008 | 14.92 | 13.44 |
|  | M | A 3 | 90 | 25 | 4,678 | 19 | 4.06 | 1,250 | 6 | 4.80 | 1,334 | 7 | 5.25 | 1.0453 | 5.11 | 5.34 |
|  | M | A 3 | 90 | 35 | 9,610 | 41 | 4.27 | 2,389 | 11 | 4.60 | 2,726 | 14 | 5.14 | 0.9304 | 4.34 | 4.03 |
|  | M | A 3 | 90 | 45 | 10,906 | 40 | 3.67 | 2,456 | 9 | 3.66 | 2,993 | 11 | 3.68 | 0.9147 | 4.86 | 4.44 |
|  | M | A 3 | 90 | 55 | 7,311 | 27 | 3.69 | 2,012 | 13 | 6.46 | 2.531 | 14 | 5.53 | 1.0568 | 5.43 | 5.74 |
|  | M | A 3 | 90 | 65 | 1.740 | 7 | 4.02 | 573 | 4 | 6.98 | 720 | 5 | 6.94 | 0.7413 | 5.28 | 3.92 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company decennial. |  |  | Five Company Original |  |  | New York Ratios | $\begin{array}{\|c\|} \hline \text { SOA } \\ \hline \text { Two } \\ \text { Occ. } \\ \hline \end{array}$ | Generated <br> Four Occ. <br> 105.67 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| M | A 4 | 0 | 25 | 16,679 | 2,028 | 121.59 | 1,129 | 105 | 93.00 | 1,129 | 105 | 93.00 | 1.0705 | 98.71 | 105.67 |
| M | A 4 | 0 | 35 | 33,177 | 3,057 | 92.14 | 5,917 | 410 | 69.29 | 5,917 | 410 | 69.29 | 1.0299 | 89.15 | 91.82 |
| M | A 4 | 0 | 45 | 41,141 | 2,856 | 69.42 | 9,341 | 508 | 54.38 | 9,341 | 508 | 54.38 | 1.0024 | 57.76 | 57.90 |
| M | A 4 | 0 | 55 | 35,501 | 1,948 | 54.87 | 11,800 | 451 | 38.22 | 11,800 | 451 | 38.22 | 0.9756 | 46.74 | 45.60 |
| M | A 4 | 0 | 65 | 10,609 | 511 | 48.17 | 4,583 | 135 | 29.46 | 4,583 | 135 | 29.46 | 0.9661 | 39.46 | 38.12 |
| M | A 4 | 7 | 25 | 11,173 | 821 | 73.48 | 1,746 | 120 | 68.73 | 1,943 | 128 | 65.88 | 1.1190 | 58.26 | 65.19 |
| M | A 4 | 7 | 35 | 16,751 | 1,162 | 69.37 | 5,029 | 345 | 68.60 | 5,449 | 390 | 71.57 | 1.2241 | 58.52 | 71.63 |
| M | A 4 | 7 | 45 | 13,441 | 754 | 56.10 | 4,886 | 293 | 59.97 | 5,289 | 315 | 59.56 | 1.2511 | 46.69 | 58.41 |
| M | A 4 | 7 | 55 | 7,643 | 413 | 54.04 | 3,370 | 202 | 59.94 | 3,645 | 217 | 59.53 | 1.4571 | 39.99 | 58.27 |
| M | A 4 | 7 | 65 | 1,268 | 58 | 45.74 | 874 | 34 | 38.90 | 946 | 39 | 41.23 | 1.3002 | 35.92 | 46.71 |
| M | A 4 | 14 | 25 | 9,810 | 388 | 39.55 | 7,974 | 392 | 49.16 | 8,054 | 400 | 49.66 | 1.0486 | 48.38 | 50.73 |
| M | A 4 | 14 | 35 | 13,492 | 550 | 40.76 | 10,008 | 542 | 54.16 | 10,166 | 550 | 54.10 | 1.1272 | 47.23 | 53.24 |
| M | A 4 | 14 | 45 | 9,912 | 392 | 39.55 | 5,608 | 262 | 46.72 | 5,758 | 267 | 46.37 | 1.0876 | 39.81 | 43.30 |
| M | A 4 | 14 | 55 | 4,722 | 181 | 38.33 | 2,661 | 128 | 48.10 | 2,749 | 131 | 47.65 | 1.0541 | 34.09 | 35.93 |
| M | A 4 | 14 | 65 | 742 | 15 | 20.22 | 353 | 8 | 22.66 | 376 | 8 | 21.28 | 0.5189 | 27.67 | 14.36 |
| M | A 4 | 30 | 25 | 11,352 | 191 | 16.83 | 8,178 | 235 | 28.74 | 8,248 | 236 | 28.61 | 0.9893 | 24.07 | 23.81 |
| M | A 4 | 30 | 35 | 18,338 | 389 | 21.21 | 11,413 | 317 | 27.78 | 11,487 | 319 | 27.77 | 1.2113 | 24.24 | 29.37 |
| M | A 4 | 30 | 45 | 14,498 | 279 | 19.24 | 7,316 | 192 | 26.24 | 7,371 | 192 | 26.05 | 1.0950 | 22.93 | 25.11 |
| M | A 4 | 30 | 55 | 7,631 | 144 | 18.87 | 3,763 | 83 | 22.06 | 3,785 | 83 | 21.93 | 1.0899 | 18.67 | 20.35 |
| M | A 4 | 30 | 65 | 1,204 | 28 | 23.26 | 661 | 11 | 16.64 | 661 | 11 | 16.64 | 1.5197 | 14.92 | 22.67 |
| M | A 4 | 90 | 25 | 984 | 3 | 3.05 | 387 | , | 2.58 | 389 | 1 | 2.57 | 0.7846 | 5.11 | 4.01 |
| M | A 4 | 90 | 35 | 1,948 | 12 | 6.16 | 766 | 5 | 6.53 | 768 | 5 | 6.51 | 1.3434 | 4.34 | 5.82 |
| M | A 4 | 90 | 45 | 2,062 | 12 | 5.82 | 679 | 3 | 4.42 | 680 | 3 | 4.41 | 1.4513 | 4.86 | 7.05 |
| M | A 4 | 90 | 55 | 1,274 | 3 | 2.35 | 527 | 2 | 3.80 | 539 | 2 | 3.71 | 0.6739 | 5.43 | 3.66 |
| M | A 4 | 90 | 65 | 287 | 4 | 13.94 | 126 | 0 | 0.00 | 126 | 0 | 0.00 | 2.5683 | 5.28 | 13.57 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company Decennial. |  |  | Five Company Original. |  |  | Nitw York Retios | $\begin{gathered} \text { SOA } \\ \text { Two } \\ \text { Occ. } \end{gathered}$ | Generated <br> Four Oce |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Ratc | Exposure | Claims | Rate |  |  |  |
| M | S 1 | 0 | 25 | 0 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| M | S 1 | 0 | 35 | 0 | 0 | 0.00 | 1 | 0 | 0.00 | 11 | 3 | 272.73 | 0.0000 | 100.00 | 0.00 |
| M | S 1 | 0 | 45 | 0 | 0 | 0.00 | 4 | 0 | 0.00 | 557 | 50 | 89.77 | 0.0000 | 122.32 | 0.00 |
| M | S 1 | 0 | 55 | 0 | 0 | 0.00 | 35 | 5 | 142.86 | 4,332 | 618 | 142.66 | 0.0000 | 124.79 | 0.00 |
| M | S 1 | 0 | 65 | 0 | 0 | 0.00 | 34 | 3 | 88.24 | 1,698 | 297 | 174.91 | 0.0000 | 138.16 | 0.00 |
| M | S 1 | 7 | 25 | 30,351 | 1,078 | 35.52 | 1,841 | 54 | 29.33 | 3,661 | 148 | 40.43 | 0.8536 | 49.02 | 41.84 |
| M | S 1 | 7 | 35 | 70,920 | 2,684 | 37.85 | 7,733 | 294 | 38.02 | 11,867 | 552 | 46.52 | 0.8223 | 42.77 | 35.17 |
| M | S 1 | 7 | 45 | 107,525 | 4,787 | 44.52 | 10,509 | 484 | 46.06 | 16,716 | 895 | 53.54 | 0.8333 | 54.59 | 45.49 |
| M | S 1 | 7 | 55 | 96,031 | 6,544 | 68.14 | 11,830 | 790 | 66.78 | 22,179 | 1.718 | 77.46 | 0.8490 | 77.19 | 65.53 |
| M | S 1 | 7 | 65 | 36,326 | 3,084 | 84.90 | 4,494 | 495 | 110.15 | 2, 6,846 | -799 | 116.71 | 0.8253 | 109.75 | 90.58 |
| M | S 1 | 14 | 25 | 20,675 | 298 | 14.41 | 5,166 | 98 | 18.97 | 6,094 | 121 | 19.86 | 0.8277 | 21.84 | 18.08 |
| M | S 1 | 14 | 35 | 57,833 | 1.054 | 18.22 | 10,429 | 226 | 21.67 | 13,573 | 302 | 22.25 | 0.8364 | 25.03 | 20.93 |
| M | S 1 | 14 | 45 | 72,747 | 2,094 | 28.78 | 9,542 | 308 | 32.28 | 14,627 | 492 | 33.64 | 0.8364 | 25.03 35.39 | 29.931 |
| M | S 1 | 14 | 55 | 47,795 | 2,334 | 48.83 | 8,270 | 441 | 53.33 | 13,120 | 739 | 56.33 | 0.8655 | 56.81 | 49.17 |
| M | S 1 | 14 | 65 | 14,981 | 1,028 | 68.62 | 2,142 | 171 | 79.83 | 3,377 | 289 | 85.58 | 0.8904 | 90.29 | 80.40 |
| M | S 1 | 30 | 25 | 75,220 | 314 | 4.17 | 14,851 | 86 | 5.79 | 16,041 | 95 | 5.92 | 0.7885 | 7.34 | 5.79 |
| M | S 1 | 30 | 35 | 327,115 | 1.762 | 5.39 | 49,519 | 393 | 7.94 | 53,618 | 433 | 8.08 | 0.8304 | 7.44 | 6.18 |
| M | S 1 | 30 | 45 | 389,990 | 3,838 | 9.84 | 42,595 | 657 | 15.42 | 48,470 | 753 | 15.54 | 0.8242 | 14.56 | 12.00 |
| M | S 1 | 30 | 55 | 208,467 | 4,282 | 20.54 | 26,566 | 718 | 27.03 | 31,017 | 866 | 27.92 | 0.8213 | 28.01 | 23.00 |
| M | S 1 | 30 | 65 | 54,481 | 1,943 | 35.66 | 5,850 | 279 | 47.69 | 9,669 | 538 | 55.64 | 0.9158 | 48.15 | 44.10 |
| M | S 1 | 90 | 25 | 17,406 | 14 | 0.80 | 4,009 | 0 | 0.00 | 4,402 | 0 | 0.00 | 0.8861 | 1.54 | 1.37 |
| M | S 1 | 90 | 35 | 112,542 | 148 | 1.32 | 20,741 | 40 | 1.93 | 22,623 | 47 | 2.08 | 0.9051 | 1.46 | 1.32 |
| M | S 1 | 90 | 45 | 140,529 | 446 | 3.17 | 18,040 | 56 | 3.10 | 21,265 | 76 | 3.57 | 0.8249 | 3.06 | 2.52 |
| $\mathbf{M}$ | S ! | 90 | 55 | 70,191 | 570 | 8.12 | 9,353 | 83 | 8.87 | 12,067 | 110 | 9.12 | 0.7656 | 9.03 | 6.91 |
| $\mathbf{M}$ | S 1 | 90 | 65 | 17,268 | 257 | 14.88 | 1,969 | 33 | 16.76 | 2,533 | 44 | 17.37 | 0.8262 | 15.61 | 12.90 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company decennial |  |  | Five Company Original |  |  | New York Ratios | $\begin{aligned} & \hline \text { SOA } \\ & \hline \text { Two } \\ & \text { Occ. } \end{aligned}$ | Generated <br> Four Occ. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| M | S 2 | 0 | 25 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 3 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| M | S 2 | 0 | 35 | 0 | 0 | 0.00 | 3 | 0 | 0.00 | 1.2 | 2 | 166.67 | 0.0000 | 100.00 | 0.00 |
| M | S 2 | 0 | 45 | 0 | 0 | 0.00 | 10 | 1 | 100.00 | 235 | 33 | 140.43 | 0.0000 | 122.32 | 0.00 |
| M | S 2 | 0 | 55 | 0 | 0 | 0.00 | 148 | 21 | 141.89 | 2,420 | 353 | 145.87 | 0.0000 | 124.79 | 0.00 |
| M | S 2 | 0 | 65 | 0 | 0 | 0.00 | 212 | 31 | 146.23 | 1,097 | 170 | 154.97 | 0.0000 | 138.16 | 0.00 |
| M | S 2 | 7 | 25 | 28,792 | 1,383 | 48.03 | 2,551 | 155 | 60.76 | 4,785 | 308 | 64.37 | 1.1544 | 49.02 | 56.59 |
| M | S 2 | 7 | 35 | 64,619 | 3,554 | 55.00 | 10,830 | 574 | 53.00 | 14,997 | 860 | 57.34 | 1.1950 | 42.77 | 51.11 |
| M | S 2 | 7 | 45 | 105,941 | 6,618 | 62.47 | 17,207 | 1,037 | 60.27 | 23,226 | 1,492 | 64.24 | 1.1692 | 54.59 | 63.83 |
| M | S 2 | 7 | 55 | 116,804 | 10,539 | 90.23 | 26,508 | 2,164 | 81.64 | 37,718 | 3,232 | 85.69 | 1.1241 | 77.19 | 86.77 |
| M | S 2 | 7 | 65 | 40,570 | 4,826 | 118.95 | 12,322 | 1,373 | 111.43 | 14,955 | 1,702 | 113.81 | 1.1564 | 109.75 | 126.91 |
| M | S 2 | 14 | 25 | 17,396 | 365 | 20.98 | 8,918 | 208 | 23.32 | 9,671 | 233 | 24.09 | 1.2048 | 21.84 | 26.32 |
| M | S 2 | 14 | 35 | 32,302 | 910 | 28.17 | 12.797 | 402 | 31.41 | 15,262 | 492 | 32.24 | 1.2929 | 25.03 | 32.36 |
| M | S 2 | 14 | 45 | 40,804 | 1,774 | 43.48 | 10,440 | 483 | 46.26 | 14,114 | 658 | 46.62 | 1.2763 | 35.39 | 45.17 |
| M | S 2 | 14 | 55 | 35,190 | 2,348 | 66.72 | 9,717 | 645 | 66.38 | 13,632 | 916 | 67.19 | 1.1826 | 56.81 | 67.19 |
| M | S 2 | 14 | 65 | 10,050 | 901 | 89.65 | 3,179 | 310 | 97.51 | 4,106 | 408 | 99.37 | 1.1633 | 90.29 | 105.04 |
| M | S 2 | 30 | 25 | 26.781 | 226 | 8.44 | 10.170 | 104 | 10.23 | 10,826 | 113 | 10.44 | 1.5940 | 7.34 | 11.70 |
| M | S 2 | 30 | 35 | 79,575 | 876 | 11.01 | 23,495 | 304 | 12.94 | 25,133 | 335 | 13.33 | 1.6971 | 7.44 | 12.63 |
| M | S 2 | 30 | 45 | 102,393 | 2,041 | 19.93 | 23,039 | 513 | 22.27 | 25,183 | 573 | 22.75 | 1.6694 | 14.56 | 24.31 |
| M | S 2 | 30 | 55 | 78,893 | 2,905 | 36.82 | 18,450 | 691 | 37.45 | 20,312 | 769 | 37.86 | 1.4723 | 28.01 | 41.23 |
| M | S 2 | 30 | 65 | 22,195 | 1.043 | 46.99 | 4,478 | 270 | 60.29 | 4,857 | 295 | 60.74 | 1.2067 | 48.15 | 58.11 |
| M | S 2 | 90 | 25 | 4,627 | 6 | 1.30 | 982 | 1 | 1.02 | 1,107 | 1 | 0.90 | 1.4285 | 1.54 | 2.20 |
| M | S 2 | 90 | 35 | 18,225 | 42 | 2.30 | 3,615 | 15 | 4.15 | 3,954 | 18 | 4.55 | 1.5861 | 1.46 | 2.32 |
| M | S 2 | 90 | 45 | 27,374 | 200 | 7.31 | 5,174 | 25 | 4.83 | 5,764 | 31 | 5.38 | 1.8990 | 3.06 | 5.80 |
| M | S 2 | 90 | 55 | 19,187 | 378 | 19.70 | 4,254 | 48 | 11.28 | 5,152 | 62 | 12.03 | 1.8574 | 9.03 | 16.77 |
| M | S 2 | 90 | 65 | 5,327 | 150 | 28.16 | 882 | 15 | 17.01 | 1,102 | 27 | 24.50 | 1.5632 | 15.61 | 24.40 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company decennial. |  |  | Five Company Original |  |  | New York Ratios | $\begin{aligned} & \text { SGA } \\ & \text { Two } \\ & \text { Occ. } \end{aligned}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| M | S 3 | 0 | 25 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 111.11 | 0.00 |
| M | S 3 | 0 | 35 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 60.11 | 0.00 |
| M | S 3 | 0 | 45 | 0 | 0 | 0.00 | 24 | 1 | 41.67 | 24 | 1 | 41.67 | 0.0000 | 119.74 | 0.00 |
| M | S 3 | 0 | 55 | 0 | 0 | 0.00 | 36 | 5 | 138.89 | 36 | 5 | 138.89 | 0.0000 | 139.80 | 0.00 |
| M | S 3 | 0 | 65 | 0 | 0 | 0.00 | 30 | 7 | 233.33 | 30 | 7 | 233.33 | 0.0000 | 140.20 | 0.00 |
| M | S 3 | 7 | 25 | 87,874 | 5,149 | 58.60 | 8,532 | 477 | 55.91 | 11,737 | 685 | 58.36 | 1.0125 | 48.30 | 48.91 |
| M | S 3 | 7 | 35 | 171,912 | 10,660 | 62.01 | 35,795 | 1,854 | 51.79 | 43,700 | 2,435 | 55.72 | 0.9962 | 52.89 | 52.69 |
| M | S 3 | 7 | 45 | 225,869 | 16,827 | 74.50 | 48,413 | 3,540 | 73.12 | 62,683 | 4,635 | 73.94 | 0.9861 | 67.04 | 66.11 |
| M | S 3 | 7 | 55 | 211,043 | 21,269 | 100.78 | 56,622 | 5,309 | 93.76 | 82,126 | 8,029 | 97.76 | 0.9865 | 91.53 | 90.29 |
| M | S 3 | 7 | 65 | 65,120 | 8,205 | 126.00 | 22,919 | 2,667 | 116.37 | 28,557 | 3,432 | 120.18 | 0.9953 | 113.75 | 113.21 |
| M | S 3 | 14 | 25 | 37.526 | 1,022 | 27.23 | 23,936 | 780 | 32.59 | 25.456 | 840 | 33.00 | 1.0165 | 31.27 | 31.79 |
| M | S 3 | 14 | 35 | 64,112 | 2,310 | 36.03 | 31,406 | 1,353 | 43.08 | 37,648 | 1,596 | 42.39 | 1.0050 | 38.04 | 38.23 |
| M | S 3 | 14 | 45 | 58,519 | 3,359 | 57.40 | 20,963 | 1,224 | 58.39 | 28,423 | 1,641 | 57.73 | 0.9979 | 51.11 | 51.00 |
| M | S 3 | 14 | 55 | 38,691 | 3,262 | 84.31 | 13,983 | 1,107 | 79.17 | 19,868 | 1,573 | 79.17 | 1.0126 | 74.60 | 75.54 |
| M | S 3 | 14 | 65 | 8.237 | 1,011 | 122.74 | 3,343 | 339 | 101.41 | 4,544 | 487 | 107.17 | 1.0176 | 105.89 | 107.76 |
| M | S 3 | 30 | 25 | 39,222 | 449 | 11.45 | 20,648 | 374 | 18.11 | 23,796 | 444 | 18.66 | 1.0281 | 15.12 | 15.54 |
| M | S 3 | 30 | 35 | 77.435 | 1,354 | 17.49 | 33,688 | 634 | 18.82 | 38,338 | 756 | 19.72 | 0.9719 | 17.33 | 16.84 |
| M | S 3 | 30 | 45 | 74.195 | 1,963 | 26.46 | 23.221 | 788 | 33.93 | 27,353 | 923 | 33.74 | 0.9822 | 29.42 | 28.89 |
| M | S 3 | 30 | 55 | 47,887 | 1,973 | 41.20 | 14,334 | 756 | 52.74 | 17,156 | 905 | 52.75 | 0.9702 | 49.06 | 47.59 |
| M | S 3 | 30 | 65 | 10.783 | 615 | 57.03 | 3,151 | 212 | 67.28 | 4,699 | 352 | 74.91 | 0.9674 | 69.08 | 66.83 |
| M | S 3 | 90 | 25 | 4.807 | 15 | 3.12 | 1,250 | 3 | 2.40 | 1,334 | 4 | 3.00 | 1.1337 | 3.42 | 3.88 |
| M | S 3 | 90 | 35 | 9,580 | 49 | 5.11 | 2,389 | 11 | 4.60 | 2,726 | 13 | 4.77 | 1.0007 | 3.65 | 3.65 |
| M | S 3 | 90 | 45 | 11.488 | 98 | 8.53 | 2,479 | 24 | 9.68 | 3,016 | 29 | 9.62 | 1.0087 | 8.44 | 8.51 |
| M | S 3 | 90 | 55 | 8,317 | 159 | 19.12 | 2,062 | 55 | 26.67 | 2,581 | 65 | 25.18 | 0.9717 | 18.73 | 18.20 |
| M | S 3 | 90 | 65 | 1,971 | 72 | 36.53 | 593 | 34 | 57.34 | 740 | 38 | 51.35 | 1.0061 | 40.46 | 40.71 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | C OC | EP | X | Niw Yorx Study |  |  | Four Company Ducennial. |  |  | Five Company Original. |  |  | New York Ratios | SOA <br> Two Occ. | GeneratedFour Occ. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Clains | Rate | Exposure | Claims | Rate |  |  |  |
| M | S 4 | 0 | 25 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 111.11 | 0.00 |
| M | S 4 | 0 | 35 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 60.11 | 0.00 |
| M | S 4 | 0 | 45 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 119.74 | 0.00 |
| M | S 4 | 0 | 55 | 0 | 0 | 0.00 | 3 | 2 | 666.67 | 3 | 2 | 666.67 | 0.0000 | 139.80 | 0.00 |
| M | S 4 | 0 | 65 | 0 | 0 | 0.00 | 5 | 0 | 0.00 | 5 | 0 | 0.00 | 0.0000 | 140.20 | 0.00 |
| M | S 4 | 7 | 25 | 26,797 | 1,487 | 55.49 | 2,842 | 127 | 44.69 | 3,039 | 130 | 42.78 | 0.9589 | 48.30 | 46.32 |
| M | S 4 | 7 | 35 | 47,676 | 3,008 | 63.09 | 10,818 | 569 | 52.60 | 11,238 | 614 | 54.64 | 1.0136 | 52.89 | 53.61 |
| M | S 4 | 7 | 45 | 52,349 | 4,193 | 80.10 | 14,027 | 990 | 70.58 | 14,430 | 1,033 | 71.59 | 1.0602 | 67.04 | 71.07 |
| M | S 4 | 7 | 55 | 41,469 | 4,528 | 109.19 | 14,931 | 1.408 | 94.30 | 15,206 | 1,448 | 95.23 | 1.0688 | 91.53 | 97.83 |
| M | S 4 | 7 | 65 | 11,358 | 1,477 | 130.04 | 5,363 | 618 | 115.23 | 5,435 | 632 | 116.28 | 1.0272 | 113.75 | 116.84 |
| M | S 4 | 14 | 25 | 10,473 | 264 | 25.21 | 7.989 | 273 | 34.17 | 8,069 | 275 | 34.08 | 0.9409 | 31.27 | 29.42 |
| M | S 4 | 14 | 35 | 14,242 | 499 | 35.04 | 10,069 | 407 | 40.42 | 10,227 | 412 | 40.29 | 0.9773 | 38.04 | 37.18 |
| M | S 4 | 14 | 45 | 10,777 | 627 | 58.18 | 5,713 | 336 | 58.81 | 5,863 | 346 | 59.01 | 1.0114 | 51.11 | 51.69 |
| M | S 4 | 14 | 55 | 5,326 | 403 | 75.67 | 2,751 | 238 | 86.51 | 2,839 | 245 | 86.30 | 0.9088 | 74.60 | 67.79 |
| M | S 4 | 14 | 65 | 908 | 92 | 101.32 | 376 | 34 | 90.43 | 399 | 39 | 97.74 | 0.8401 | 105.89 | 88.95 |
| M | S 4 | 30 | 25 | 11,877 | 120 | 10.10 | 8,195 | 131 | 15.99 | 8,265 | 133 | 16.09 | 0.9073 | 15.12 | 13.72 |
| M | S 4 | 30 | 35 | 19,776 | 395 | 19.97 | 11,481 | 269 | 23.43 | 11,555 | 271 | 23.45 | 1.1102 | 17.33 | 19.24 |
| M | S 4 | 30 | 45 | 15,680 | 458 | 29.21 | 7.401 | 247 | 33.37 | 7,456 | 251 | 33.66 | 1.0843 | 29.42 | 31.90 |
| M | S 4 | 30 | 55 | 8.791 | 434 | 49.37 | 3,882 | 235 | 60.54 | 3,904 | 237 | 60.71 | 1.1625 | 49.06 | 57.03 |
| M | S 4 | 30 | 65 | 1,464 | 107 | 73.09 | 716 | 47 | 65.64 | 716 | 47 | 65.64 | 1.2398 | 69.08 | 85.64 |
| M | S 4 | 90 | 25 | 1,006 | 1 | 0.99 | 386 | 0 | 0.00 | 388 | 0 | 0.00 | 0.3611 | 3.42 | 1.24 |
| M | S 4 | 90 | 35 | 1,963 | 10 | 5.09 | 766 | 5 | 6.53 | 768 | 5 | 6.51 | 0.9967 | 3.65 | 3.64 |
| M | S 4 | 90 | 45 | 2,110 | 17 | 8.06 | 683 | 5 | 7.32 | 684 | 5 | 7.31 | 0.9527 | 8.44 | 8.04 |
| M | S 4 | 90 | 55 | 1,340 | 31 | 23.13 | 550 | 17 | 30.91 | 562 | 17 | 30.25 | 1.1758 | 18.73 | 22.02 |
| M | S 4 | 90 | 65 | 315 | 11 | 34.92 | 131 | 4 | 30.53 | 131 | 4 | 30.53 | 0.9618 | 40.46 | 38.91 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company Decennial |  |  | Fine: Company Original. |  |  | NEW YORK Ratios | $\begin{aligned} & \text { SOA } \\ & \text { Two } \\ & \text { Ocs. } \end{aligned}$ | Generated <br> Four Occ. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| F | A 1 | 0 | 25 | 9,563 | 239 | 24.99 | 262 | 5 | 19.08 | 328 | 6 | 18.29 | 0.8521 | 31.06 | 26.47 |
| F | A 1 | 0 | 35 | 11,455 | 295 | 25.75 | 314 | 10 | 31.85 | 376 | 15 | 39.89 | 0.8289 | 29.81 | 24.71 |
| F | A 1 | 0 | 45 | 17,635 | 516 | 29.26 | 393 | 9 | 22.90 | 430 | 11 | 25.58 | 0.8455 | 25.32 | 21.41 |
| F | A 1 | 0 | 55 | 17,666 | 659 | 37.30 | 662 | 24 | 36.25 | 678 | 24 | 35.40 | 0.9065 | 29.29 | 26.55 |
| F | A 1 | 0 | 65 | 5.632 | 179 | 31.78 | 330 | 12 | 36.36 | 330 | 12 | 36.36 | 0.8879 | 39.56 | 35.13 |
| F | A 1 | 7 | 25 | 8,046 | 176 | 21.87 | 752 | 23 | 30.59 | 1,628 | 50 | 30.71 | 0.9352 | 23.72 | 22.18 |
| F | A 1 | 7 | 35 | 11,010 | 250 | 22.71 | 1,446 | 31 | 21.44 | 3,533 | 78 | 22.08 | 0.9189 | 22.70 | 20.86 |
| F | A 1 | 7 | 45 | 18,381 | 363 | 19.75 | 1,299 | 27 | 20.79 | 4,637 | 105 | 22.64 | 0.9006 | 25.19 | 22.69 |
| F | A 1 | 7 | 55 | 20,698 | 559 | 27.01 | 1.270 | 37 | 29.13 | 6,967 | 176 | 25.26 | 0.9775 | 24.18 | 23.64 |
| F | A 1 | 7 | 65 | 1.837 | 43 | 23.41 | 195 | 7 | 35.90 | 488 | 15 | 30.74 | 0.8909 | 23.42 | 20.87 |
| F | A 1 | 14 | 25 | 6,384 | 59 | 9.24 | 2,817 | 32 | 11.36 | 3,177 | 37 | 11.65 | 0.7867 | 17.67 | 13.90 |
| F | A 1 | 14 | 35 | 7,747 | 76 | 9.81 | 2,934 | 48 | 16.36 | 3,834 | 60 | 15.65 | 0.7411 | 19.95 | 14.79 |
| F | A 1 | 14 | 45 | 10,382 | 101 | 9.73 | 2,303 | 26 | 11.29 | 3,604 | 44 | 12.21 | 0.7964 | 17.32 | 13.79 |
| F | A 1 | 14 | 55 | 8,414 | 106 | 12.60 | 1,937 | 30 | 15.49 | 3,358 | 52 | 15.49 | 0.7284 | 19.60 | 14.27 |
| F | A 1 | 14 | 65 | 2,108 | 25 | 11.86 | 487 | 9 | 18.48 | 528 | 9 | 17.05 | 0.7925 | 27.08 | 21.46 |
| F | A 1 | 30 | 25 | 11.771 | 33 | 2.80 | 4,795 | 14 | 2.92 | 4.949 | 14 | 2.83 | 0.7451 | 6.20 | 4.62 |
| F | A 1 | 30 | 35 | 16,279 | 62 | 3.81 | 5,911 | 38 | 6.43 | 6,187 | 41 | 6.63 | 0.7814 | 7.74 | 6.05 |
| F | A 1 | 30 | 45 | 21,901 | 106 | 4.84 | 4,864 | 35 | 7.20 | 5,122 | 36 | 7.03 | 0.8171 | 7.51 | 6.14 |
| F | A 1 | 30 | 55 | 16,780 | 98 | 5.84 | 3,744 | 23 | 6.14 | 3,867 | 24 | 6.21 | 0.7832 | 9.74 | 7.63 |
| F | A 1 | 30 | 65 | 3.472 | 28 | 8.06 | 632 | 6 | 9.49 | 634 | 6 | 9.46 | 0.9014 | 10.28 | 9.27 |
| F | A 1 | 90 | 25 | 2.375 | 1 | 0.42 | 711 | 1 | 1.41 | 768 | 2 | 2.60 | 1.0867 | 1.78 | 1.93 |
| F | A 1 | 90 | 35 | 3,676 | 8 | 2.18 | 1,314 | 0 | 0.00 | 1,453 | 0 | 0.00 | 0.8300 | 0.58 | 0.48 |
| F | A 1 | 90 | 45 | 5.815 | 8 | 1.38 | 977 | 2 | 2.05 | 1,286 | 2 | 1.56 | 1.0055 | 1.35 | 1.36 |
| F | A 1 | 90 | 55 | 5.445 | 15 | 2.75 | 693 | 3 | 4.33 | 1,089 | 5 | 4.59 | 1.1379 | 2.25 | 2.56 |
| F | A 1 | 90 | 65 | 914 | 3 | 3.28 | 87 | 0 | 0.00 | + 96 | 0 | 0.00 | 1.1225 | 3.60 | 4.04 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company Decennial |  |  | Five Company Original. |  |  | New York Ratios | $\begin{aligned} & \text { SOA } \\ & \hline \text { Two } \\ & \text { Occ. } \end{aligned}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| F | A 2 | 0 | 25 | 2.439 | 113 | 46.33 | 75 | 3 | 40.00 | 93 | 3 | 32.26 | 1.5797 | 31.06 | 49.07 |
| F | A 2 | 0 | 35 | 3,254 | 162 | 49.78 | 74 | 4 | 54.05 | 88 | 5 | 56.82 | 1.6024 | 29.81 | 47.77 |
| F | A 2 | 0 | 45 | 5,396 | 281 | 52.08 | 113 | 7 | 61.95 | 124 | 7 | 56.45 | 1.5048 | 25.32 | 38.10 |
| F | A 2 | 0 | 55 | 5,929 | 312 | 52.62 | 207 | 16 | 77.29 | 212 | 16 | 75.47 | 1.2787 | 29.29 | 37.45 |
| F | A 2 | 0 | 65 | 1,995 | 94 | 47.12 | 172 | 15 | 87.21 | 172 | 15 | 87.21 | 1.3164 | 39.56 | 52.08 |
| F | A 2 | 7 | 25 | 2,557 | 72 | 28.16 | 464 | 8 | 17.24 | 747 | 19 | 25.44 | 1.2039 | 23.72 | 28.55 |
| F | A 2 | 7 | 35 | 3.842 | 117 | 30.45 | 1,527 | 30 | 19.65 | 2,129 | 51 | 23.95 | 1.2324 | 22.70 | 27.98 |
| F | A 2 | 7 | 45 | 5,697 | 165 | 28.96 | 2,506 | 35 | 13.97 | 3,371 | 63 | 18.69 | 1.3208 | 25.19 | 33.27 |
| F | A 2 | 7 | 55 | 6,084 | 181 | 29.75 | 3,713 | 89 | 23.97 | 4,625 | 114 | 24.65 | 1.0767 | 24.18 | 26.04 |
| F | A 2 | 7 | 65 | 675 | 23 | 34.07 | 406 | 8 | 19.70 | 469 | 9 | 19.19 | 1.2969 | 23.42 | 30.38 |
| F | A 2 | 14 | 25 | 2,213 | 42 | 18.98 | 2,239 | 45 | 20.10 | 2,574 | 52 | 20.20 | 1.6154 | 17.67 | 28.54 |
| F | A 2 | 14 | 35 | 2.603 | 61 | 23.43 | 2,419 | 66 | 27.28 | 2,992 | 77 | 25.74 | 1.7704 | 19.95 | 35.32 |
| F | A 2 | 14 | 45 | 3.126 | 64 | 20.47 | 1,999 | 42 | 21.01 | 2,492 | 52 | 20.87 | 1.6761 | 17.32 | 29.03 |
| F | A 2 | 14 | 55 | 2.630 | 85 | 32.32 | 1,446 | 32 | 22.13 | 1,844 | 38 | 20.61 | 1.8688 | 19.60 | 36.62 |
| F | A 2 | 14 | 65 | 565 | 15 | 26.55 | 200 | 5 | 25.00 | 204 | 6 | 29.41 | 1.7741 | 27.08 | 48.04 |
| F | A 2 | 30 | 25 | 2.580 | 21 | 8.14 | 1,772 | 18 | 10.16 | 1,810 | 18 | 9.94 | 2.1632 | 6.20 | 13.42 |
| F | A 2 | 30 | 35 | 3.827 | 36 | 9.41 | 2,849 | 23 | 8.07 | 2,889 | 23 | 7.96 | 1.9299 | 7.74 | 14.94 |
| F | A 2 | 30 | 45 | 4,943 | 53 | 10.72 | 3,041 | 31 | 10.19 | 3,070 | 32 | 10.42 | 1.8102 | 7.51 | 13.60 |
| F | A 2 | 30 | 55 | 4,408 | 60 | 13.61 | 3,101 | 31 | 10.00 | 3,118 | 31 | 9.94 | 1.8253 | 9.74 | 17.78 |
| F | A 2 | 30 | 65 | 552 | 8 | 14.49 | 267 | 0 | 0.00 | 267 | 0 | 0.00 | 1.6200 | 10.28 | 16.66 |
| F | A 2 | 90 | 25 | 206 | 0 | 0.00 | 141 | 0 | 0.00 | 143 | 0 | 0.00 | 0.0000 | 1.78 | 0.00 |
| F | A 2 | 90 | 35 | 519 | 3 | 5.78 | 260 | 0 | 0.00 | 264 | 0 | 0.00 | 2.2044 | 0.58 | 1.28 |
| F | A 2 | 90 | 45 | 763 | 1 | 1.31 | 399 | 0 | 0.00 | 424 | 0 | 0.00 | 0.9579 | 1.35 | 1.29 |
| F | A 2 | 90 | 55 | 751 | 0 | 0.00 | 496 | 3 | 6.05 | 527 | 3 | 5.69 | 0.0000 | 2.25 | 0.00 |
| F | A 2 | 90 | 65 | 112 | 0 | 0.00 | 48 | 0 | 0.00 | 50 | 0 | 0.00 | 0.0000 | 3.60 | 0.00 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)


EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Study |  |  | Four Company Decennial. |  |  | Five Company Original. |  |  | New York Ratios | $\begin{aligned} & \text { SOA } \\ & \text { Two } \\ & \text { Occ. } \end{aligned}$ | $\begin{aligned} & \text { Generated } \\ & \text { Four Occ. } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| F | A 4 | 0 | 25 | 5 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 57.97 | 0.00 |
| $F$ | A 4 | 0 | 35 | 21 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 53.82 | 0.00 |
| F | A 4 | 0 | 45 | 79 | 9 | 113.92 | 4 | 0 | 0.00 | 4 | 0 | 0.00 | 1.6527 | 39.04 | 64.52 |
| F | A 4 | 0 | 55 | 120 | 9 | 75.00 | 8 | 0 | 0.00 | 8 | 0 | 0.00 | 1.2404 | 40.15 | 49.81 |
| F | A 4 | 0 | 65 | 35 | 2 | 57.14 | 4 | 0 | 0.00 | 4 | 0 | 0.00 | 1.0496 | 33.17 | 34.81 |
| F | A 4 | 7 | 25 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 17.42 | 0.00 |
| F | A 4 | 7 | 35 | 5 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 33.41 | 0.00 |
| F | A 4 | 7 | 45 | 0 | 0 | 0.00 | 2 | 1 | 500.00 | 2 | 1 | 500.00 | 0.0000 | 34.60 | 0.00 |
| F | A 4 | 7 | 55 | 0 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 44.32 | 0.00 |
| F | A 4 | 7 | 65 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 58.14 | 0.00 |
| F | A 4 | 14 | 25 | 0 | 0 | 0.00 | 45 | 0 | 0.00 | 45 | 0 | 0.00 | 0.0000 | 29.02 | 0.00 |
| F | A 4 | 14 | 35 | 0 | 0 | 0.00 | 55 | 0 | 0.00 | 55 | 0 | 0.00 | 0.0000 | 30.26 | 0.00 |
| F | A 4 | 14 | 45 | 1 | 0 | 0.00 | 43 | 0 | 0.00 | 44 | 0 | 0.00 | 0.0000 | 39.04 | 0.00 |
| F | A 4 | 14 | 55 | 0 | 0 | 0.00 | 23 | 0 | 0.00 | 23 | 0 | 0.00 | 0.0000 | 32.19 | 0.00 |
| F | A 4 | 14 | 65 | 0 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 18.18 | 0.00 |
| $F$ | A 4 | 30 | 25 | 0 | 0 | 0.00 | 43 | 1 | 23.26 | 43 | 1 | 23.26 | 0.0000 | 13.98 | 0.00 |
| F | A 4 | 30 | 35 | 5 | 0 | 0.00 | 43 | 1 | 23.26 | 44 | 1 | 22.73 | 0.0000 | 19.71 | 0.00 |
| F | A 4 | 30 | 45 | 10 | 1 | 100.00 | 28 | 0 | 0.00 | 28 | 0 | 0.00 | 3.3967 | 20.19 | 68.57 |
| F | A 4 | 30 | 55 | 5 | 0 | 0.00 | 5 | 0 | 0.00 | 5 | 0 | 0.00 | 0.0000 | 18.90 | 0.00 |
| F | A 4 | 30 | 65 | 0 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 25.51 | 0.00 |
| F | A 4 | 90 | 25 | 0 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | A 4 | 90 | 35 | 4 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 5.00 | 0.00 |
| F | A 4 | 90 | 45 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 8.85 | 0.00 |
| F | A 4 | 90 | 55 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | A 4 | 90 | 65 | 2 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)


EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | NEW York Study |  |  | Four Company decennial |  |  | Five Company Orioinal |  |  | New York Ratios | $\begin{aligned} & \text { SOA } \\ & \hline \text { Two } \\ & \text { Occ. } \end{aligned}$ | Generated <br> Four Occ. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claims | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| F | S 2 | 0 | 25 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | S 2 | 0 | 35 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 81.08 | 0.00 |
| F | S 2 | 0 | 45 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 171.05 | 0.00 |
| F | S 2 | 0 | 55 | 0 | 0 | 0.00 | 2 | 1 | 500.00 | 2 | 1 | 500.00 | 0.0000 | 168.61 | 0.00 |
| F | S 2 | 0 | 65 | 0 | 0 | 0.00 | 7 | 0 | 0.00 | 7 | 0 | 0.00 | 0.0000 | 170.48 | 0.00 |
| F | S 2 | 7 | 25 | 4,024 | 384 | 95.43 | 514 | 36 | 70.04 | 813 | 72 | 88.56 | 1.1762 | 65.09 | 76.56 |
| F | S 2 | 7 | 35 | 5,725 | 708 | 123.67 | 1,578 | 158 | 100.13 | 2,193 | 241 | 109.90 | 1.1501 | 86.95 | 100.00 |
| F | S 2 | 7 | 45 | 8,831 | 1,247 | 141.21 | 2,587 | 281 | 108.62 | 3,460 | 405 | 117.05 | 1.2106 | 100.52 | 121.69 |
| F | S 2 | 7 | 55 | 9,553 | 1,254 | 131.27 | 3,878 | 339 | 87.42 | 4,795 | 452 | 94.26 | 1.1742 | 95.27 | 111.87 |
| F | S 2 | 7 | 65 | 1,778 | 229 | 128.80 | 559 | 45 | 80.50 | 622 | 54 | 86.82 | 1.2318 | 102.14 | 125.81 |
| F | S 2 | 14 | 25 | 2.571 | 129 | 50.18 | 2,258 | 113 | 50.04 | 2,594 | 132 | 50.89 | 1.3634 | 44.51 | 60.68 |
| F | S 2 | 14 | 35 | 3,142 | 286 | 91.02 | 2,429 | 211 | 86.87 | 3,002 | 243 | 80.95 | 1.5108 | 63.66 | 96.18 |
| F | S 2 | 14 | 45 | 4,178 | 4.35 | 104.12 | 2,020 | 175 | 86.63 | 2,516 | 217 | 86.25 | 1.3747 | 77.45 | 106.47 |
| F | S 2 | 14 | 55 | 3,776 | 401 | 106.20 | 1,468 | 136 | 92.64 | 1,866 | 175 | 93.78 | 1.3789 | 71.14 | 98.10 |
| F | S 2 | 14 | 65 | 888 | 93 | 104.73 | 204 | 15 | 73.53 | 208 | 15 | 72.12 | 1.4674 | 81.66 | 119.83 |
| F | S 2 | 30 | 25 | 2,881 | 64 | 22.21 | 1,775 | 34 | 19.15 | 1,814 | 34 | 18.74 | 1.9737 | 15.88 | 31.34 |
| F | S 2 | 30 | 35 | 4,284 | 162 | 37.82 | 2,863 | 106 | 37.02 | 2,904 | 106 | 36.50 | 1.5249 | 22.32 | 34.04 |
| F | S 2 | 30 | 45 | 5,76: | 277 | 48.08 | 3,049 | 138 | 45.26 | 3,078 | 141 | 45.81 | 1.4330 | 32.75 | 46.93 |
| F | S 2 | 30 | 55 | 5,258 | 248 | 47.17 | 3,114 | 132 | 42.39 | 3,131 | 133 | 42.48 | 1.3709 | 35.09 | 48.10 |
| F | S 2 | 30 | 65 | 761 | 42 | 55.19 | 272 | 15 | 55.15 | 272 | 15 | 55.15 | 1.5603 | 44.82 | 69.94 |
| F | S 2 | 90 | 25 | 211 | 0 | 0.00 | 145 | 0 | 0.00 | 147 | 0 | 0.00 | 0.0000 | 1.78 | 0.00 |
| F | S 2 | 90 | 35 | 539 |  | 1.86 | 260 | 0 | 0.00 | 264 | 0 | 0.00 | 0.4044 | 3.48 | 1.41 |
| F | S 2 | 90 | 45 | 769 | 8 | 10.40 | 397 | 2 | 5.04 | 422 | 2 | 4.74 | 1.2132 | 4.83 | 5.86 |
| F | S 2 | 90 | 55 | 764 | 12 | 15.71 | 494 | 6 | 12.15 | 525 | 6 | 11.43 | 1.2846 | 11.77 | 15.12 |
| F | S 2 | 90 | 65 | 117 | 4 | 34.19 | 48 | 1 | 20.83 | 50 | 1 | 20.00 | 2.0361 | 20.48 | 41.71 |

EXHIBIT C-1-Continued
(FOUR OCC. DATA)


EXHIBIT C-1-Continued
(FOUR OCC. DATA)

| S | COC | EP | X | New York Stuiry |  |  | Forer Company Decennial. |  |  | Five Company Original. |  |  | New York Ratios | $\begin{aligned} & \text { SOA } \\ & \text { Two } \\ & \text { Occ. } \end{aligned}$ | $\begin{aligned} & \text { Generated } \\ & \text { Four Occ. } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Exposure | Claim: | Rate | Exposure | Claims | Rate | Exposure | Claims | Rate |  |  |  |
| F | S 4 | 0 | 25 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | S 4 | 0 | 35 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | S 4 | 0 | 45 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 121.21 | 0.00 |
| F | S 4 | 0 | 55 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 150.38 | 0.00 |
| F | S 4 | 0 | 65 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 311.11 | 0.00 |
| F | S 4 | 7 | 25 | 3 | 1 | 333.33 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 2.8774 | 77.11 | 221.88 |
| F | S 4 | 7 | 35 | 24 | 1 | 41.67 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.2590 | 116.98 | 30.29 |
| F | S 4 | 7 | 45 | 70 | 6 | 85.71 | 5 | 0 | 0.00 | 5 | 0 | 0.00 | 0.4740 | 127.79 | 60.58 |
| F | S 4 | 7 | 55 | 121 | 16 | 132.23 | 10 | 1 | 100.00 | 10 | 1 | 100.00 | 0.7706 | 130.36 | 100.45 |
| F | S 4 | 7 | 65 | 27 | 3 | 111.11 | 4 | 1 | 250.00 | 4 | 1 | 250.00 | 0.6953 | 115.38 | 80.23 |
| $F$ | S 4 | 14 | 25 | 0 | 0 | 0.00 | 46 | 0 | 0.00 | 46 | 0 | 0.00 | 0.0000 | 60.92 | 0.00 |
| F | S 4 | 14 | 35 | 2 | 0 | 0.00 | 55 | 2 | 36.36 | 55 | 2 | 36.36 | 0.0000 | 97.69 | 0.00 |
| F | S 4 | 14 | 45 | 1 | 0 | 0.00 | 43 | 1 | 23.26 | 44 | 2 | 45.45 | 0.0000 | 133.38 | 0.00 |
| F | S 4 | 14 | 55 | 3 | 0 | 0.00 | 23 | 5 | 217.39 | 23 | 5 | 217.39 | 0.0000 | 104.87 | 0.00 |
| F | S 4 | 14 | 65 | 5 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 95.13 | 0.00 |
| F | S 4 | 30 | 25 | 4 | 0 | 0.00 | 43 | 2 | 46.51 | 43 | 2 | 46.51 | 0.0000 | 26.18 | 0.00 |
| F | S 4 | 30 | 35 | 4 | 0 | 0.00 | 43 | 1 | 23.26 | 44 | 1 | 22.73 | 0.0000 | 48.78 | 0.00 |
| F | S 4 | 30 | 45 | 17 | 0 | 0.00 | 29 | 1 | 34.48 | 29 | 1 | 34.48 | 0.0000 | 54.22 | 0.00 |
| F | S 4 | 30 | 55 | 7 | 0 | 0.00 | 5 | 0 | 0.00 | 5 | 0 | 0.00 | 0.0000 | 61.43 | 0.00 |
| F | S 4 | 30 | 65 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 74.77 | 0.00 |
| F | S 4 | 90 | 25 | 0 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | S 4 | 90 | 35 | 4 | 0 | 0.00 | 1 | 0 | 0.00 | 1 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |
| F | S 4 | 90 | 45 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 8.85 | 0.00 |
| F | S 4 | 90 | 55 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 2 | 0 | 0.00 | 0.0000 | 10.31 | 0.00 |
| F | S 4 | 90 | 65 | 2 | 0 | 0.00 | 0 | 0 | 0.00 | 0 | 0 | 0.00 | 0.0000 | 0.00 | 0.00 |

EXHIBIT C-2
RATES PER 1000
ACCIDENT - MALE


RATES PER 1000
ACCIDENT - MALE


RATES PER 1000
ACCIDENT - MALE

INCIDENCE RATES




- One 1
$0 \quad 0 \operatorname{cc} 2$
Occ 3
Oce 4




Ocr

RATES PER 1000
ACCIDENT - MALE

IMPLIED 90 DAY RATES
C. IMPLIED 90 DAY RATES





- Occ 1
- $\operatorname{Occ} 2$
- $\operatorname{Occ} 3$
- Occ 4

RATES PER 1000
ACCIDENT - FEMALE

INCIDENCE RATES




- EPO
- EP 7

O-EP 14

- EP 30
- EP 90








- EPO

O-EP 14

- EP 30
-     - EP 90





Elimination Period 90 Days



- Occ 1
O-Occ 2
- Occ
- Occ 4

RATIO OF INCIDENCE RATES

ccupation Class 1

ation Class 2




3าฟW OL ヨาvWヨy jo OIIVy

- EPO -EP 7

O-EP 14

- EP 30
- EPP 90


# RATIO OF FEMALE TO MALE <br> ACCIDENT 

RATIO OF INCIDENCE RATES






## RATIO OF FEMALE TO MALE ACCIDENT

RATIO OF INCIDENCE RATES


${ }^{17} 14$ Days


GRADUATED Elimination Period 7 Davs CRUDE




Occ 3
$\Delta-\mathrm{Occ} 4$

RATIO OF IMPLIED 90 DAY RATES

## RATIO OF FEMALE TO MALE <br> ACCIDENT

RATIO OF INCIDENCE RATES


GRADUATED





Эา＊W OL ヨา甘Wヨy to OII甘ソ
－Occ 1
O－Occ 2
－Occ 3
$\Delta-\mathrm{Occ} 4$








- Occ 1 O-Occ 2
- Occ 3
- Occ 4


# RATES PER 1000 

SICKNESS - MALE




## RATES PER 1000

SICKNESS - FEMALE



RATIO OF FEMALE TO MALE SICKNESS




# RATIO OF FEMALE TO MALE <br> SICKNESS 

RATIO OF INCIDENCE RATES


Occupation Class 3



RATIO OF IMPLIED 90 DAY RATES




- EP 7

O-EP 14

- EP 30
- EP 90


## RATIO OF FEMALE TO MALE SICKNESS

RATIO OF INCIDENCE RATES




- Occ $1 \quad 0 \quad \mathrm{Occ}_{2}$
$\bullet$
Occ 3
- Occ 4



GRADUATED RATES PER 1000

## TOTAL ACCIDENT \& SICKNESS - MALE

4
8


$\qquad$

GRADUATED RATES PER 1000
TOTAL ACCIDENT \& SICKNESS - MALE


GRADUATED RATES PER 1000
TOTAL ACCIDENT \& SICKNESS - FEMALE
$\qquad$
IMP 90 DAY RATES
INCIDENCE RATES


## TOTAL ACCIDENT \& SICKNESS - FEMALE

$\qquad$
INCIDENCE RATES $\qquad$ incIDENCE RATES $\qquad$
ñ
Occupation Class 3





- EP 90
- Occ 1

O-Occ 2
$-\mathrm{Occ} 3$

- Occ 4

RATIOS OF FEMALE TO MALE
TOTAL ACCIDENT \& SICKNESS
ヨาVW OL ヨาVW3s fo OIIVy


Occupation Class 1
IMP. 90 DAY RATES

2.1


INCIDENCE RATES IMP. 90 DAY RATES

Elimination Period 7 Days



Elimination Period 14 Days



- EPO
$-E P 7$
- EP 14
- EP 30
- EP 90

Occ 1
$0-\operatorname{Occ} 2$

- Occ 3
-Occ 4


## RATIOS OF FEMALE TO MALE

TOTAL ACCIDENT \& SICKNESS








O-EP 14
EP 30

- EP 90
- Occ 1
$0-0 \operatorname{cc} 2$
- Occ 3

Occ 4






Occ 1

- Occ 2Occ 3
- Occ 4


## RATIOS OF FEMALE TO MALE INCIDENCE RATES

 ELIMINATION PERIOD 0 DAYS - ACCIDENT ONLY
GRADUATED

IMPLIED 90 DAY RATES

GRADUATED


CRUDE


- Occ 1
- 0 occ 2
- Occ 3
- $-\operatorname{Occ} 4$


## EXHIBIT C-3

Incidence of Disability
(Rates per 1,000 lives exposed)

| Male-Accident |  |  |  |  |  |  | Male-Sickness |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | AGE | Elimination Period |  |  |  |  | Age | Elimination Period |  |  |  |  |
|  |  | 0 -day | 7-day | 14-day | 30-day | 90-day |  | 0-day | 7-day | 14-day | 30-day | 90-day |
| Class 1 | 25 | 33.97 | 25.84 | 13.13 | 4.90 | 86 | 25 |  | 32.26 | 18.22 | 5.51 | 1.01 |
|  | 35 | 32.88 | 24.42 | 11.99 | 4.23 | . 51 | 35 | $\cdots$ | 36.11 | 21.55 | 6.48 | 1.13 |
|  | 45 | 30.40 | 20.40 | 9.86 | 4.50 | . 65 | 45 |  | 47.12 | 31.19 | 12.63 | 2.70 |
|  | 55 | 30.19 | 18.32 | 9.63 | 4.71 | . 80 | 55 | $\ldots$ | 69.48 | 52.75 | 25.11 | 7.78 |
|  | 62 | 33.45 | 16.11 | 10.39 | 5.47 | 1.18 | 62 | $\ldots$ | 91.52 | 74.06 | 41.24 | 15.20 |
| Class 2 | 25 | 59.96 | 47.98 | 30.01 | 10.48 | 2.07 | 25 |  | 46.61 | 27.01 | 12.17 | 2.23 |
|  | 35 | 59.96 | 44.62 | 28.83 | 10.14 | 2.09 | 35 |  | 52.79 | 33.37 | 14.47 | 2.56 |
|  | 45 | 56.74 | 38.49 | 25.67 | 9.86 | 2.14 | 45 | $\ldots$ | 65.97 | 46.91 | 25.40 | 6.21 |
|  | 55 | 51.66 | 31.31 | 20.50 | 10.03 | 2.20 | 55 |  | 92.99 | 71.27 | 41.37 | 15.74 |
|  | 62 | 52.84 | 29.85 | 19.86 | 10.92 | 2.57 | 62 |  | 116.81 | 93.05 | 58.54 | 25.94 |
| Class 3 | 25 | 75.80 | 62.68 | 42.87 | 23.69 | 7.04 | 25 |  | 46.83 | 32.22 | 14.75 | 2.99 |
|  | 35 | 74.78 | 58.37 | 39.59 | 22.57 | 6.48 | 35 |  | 52.72 | 38.32 | 18.70 | 3.52 |
|  | 45 | 69.76 | 50.41 | 34.61 | 20.49 | 5.97 | 45 |  | 67.05 | 51.53 | 29.45 | 7.83 |
|  | 55 | 66.37 | 44.27 | 30.51 | 18.49 | 5.46 | 55 |  | 92.60 | 76.39 | 52.66 | 20.07 |
|  | 62 | 65.04 | 39.98 | 27.96 | 18.56 | 5.30 | 62 |  | 116.23 | 98.78 | 78.56 | 36.04 |
| Class 4 | 25 | 89.42 | 77.60 | 52.59 | 27.03 | 8.73 | 25 |  | 48.20 | 33.28 | 15.07 | 3.04 |
|  | 35 | 91.59 | 73.24 | 50.53 | 26.93 | 8.17 | 35 |  | 53.75 | 39.27 | 19.33 | 3.59 |
|  | 45 | 84.64 | 62.13 | 42.61 | 24.78 | 7.68 | 45 |  | 70.03 | 52.71 | 30.13 | 7.97 |
|  | 55 | 79.77 | 52.03 | 37.34 | 22.78 | 7.27 | 55 |  | 95.01 | 77.91 | 55.87 | 20.45 |
|  | 62 | 79.95 | 49.76 | 36.11 | 22.96 | 7.20 | 62 |  | 119.16 | 101.41 | 81.62 | 36.63 |
| Female-Accident |  |  |  |  |  |  | Female Sickness |  |  |  |  |  |
|  | age | Elimination Period |  |  |  |  | Age | Elimination Period |  |  |  |  |
|  |  | 0-day | 7-day | 14-day | 30-day | 90 -day |  | 0-day | 7-day | 14-day | 30-day | 90-day |
| Class 1 | 25 | 23.06 | 19.92 | 12.96 | 6.00 | 1.14 | 25 |  | 61.10 | 39.29 | 14.03 | 2.55 |
|  | 35 | 26.28 | 20.87 | 13.39 | 6.21 | . 91 | 35 |  | 84.38 | 56.89 | 24.75 | 4.37 |
|  | 45 | 32.36 | 22.77 | 13.78 | 6.83 | 1.11 | 45 | $\cdots$ | 94.57 | 68.33 | 34.14 | 7.64 |
|  | 55 | 45.05 | 26.77 | 14.82 | 8.06 | 1.46 | 55 |  | 90.28 | 61.49 | 34.23 | 10.31 |
|  | 62 | 69.00 | 31.56 | 17.54 | 9.91 | 2.25 | 62 |  | 93.06 | 69.44 | 45.30 | 13.85 |
| Class 2 | 25 | 35.05 | 31.48 | 23.39 | 13.40 | 3.22 | 25 |  | 80.97 | 53.57 | 20.03 | 3.75 |
|  | 35 | 39.36 | 32.01 | 23.36 | 14.02 | 3.20 | 35 |  | 116.02 | 80.05 | 35.34 | 6.60 |
|  | 45 | 47.46 | 33.55 | 24.40 | 15.02 | 3.40 | 45 |  | 134.18 | 92.93 | 47.62 | 10.81 |
|  | 55 | 62.53 | 37.10 | 26.13 | 16.11 | 3.75 | 55 |  | 117.29 | 84.93 | 49.00 | 14.95 |
|  | 62 | 88.91 | 44.31 | 29.27 | 17.88 | 4.46 | 62 |  | 120.40 | 87.53 | 63.15 | 18.86 |
| Class 3 | 25 | 41.93 | 38.01 | 27.94 | 17.63 | 6.19 | 25 |  | 86.64 | 57.85 | 24.83 | 5.03 |
|  | 35 | 46.30 | 38.45 | 28.54 | 18.20 | 6.54 | 35 | $\cdots$ | 124.79 | 96.77 | 44.67 | 8.43 |
|  | 45 | 53.01 | 39.08 | 29.09 | 19.24 | 6.75 | 45 |  | 145.58 | 116.19 | 58.44 | 14.43 |
|  | 55 | 66.71 | 41.96 | 30.86 | 20.99 | 7.08 | 55 |  | 122.98 | 99.89 | 59.99 | 17.86 |
|  | 62 | 90.05 | 48.12 | 33.60 | 23.74 | 7.26 | 62 |  | 125.95 | 101.06 | 69.18 | 22.76 |
| Class 4 | 25 | 52.41 | 47.52 | 34.93 | 22.04 | 7.74 | 25 |  | 90.24 | 60.26 | 25.86 | 5.23 |
|  | 35 | 57.87 | 48.07 | 35.67 | 22.75 | 8.17 | 35 |  | 130.00 | 100.81 | 46.53 | 8.79 |
|  | 45 | 66.26 | 48.86 | 36.36 | 24.05 | 8.45 | 45 |  | 151.65 | 121.04 | 60.87 | 15.03 |
|  | 55 | 83.39 | 52.45 | 38.58 | 26.25 | 8.85 | 55 |  | 128.10 | 104.05 | 62.49 | 18.61 |
|  | 62 | 112.57 | 60.16 | 42.00 | 29.67 | 9.08 | 62 |  | 131.20 | 105.27 | 72.07 | 23.71 |

## APPENDIX D

## COMPARATIVE EXHIBITS DETERMINED FROM THE DTS BASIC TABLE

Exhibit D-1 is a comparison of the final results of the DTS Basic Table to material collected and published by the Society of Actuaries on individual loss of time experience. The Society data are published biannually in the Reports, and we refer to that study as the SOA data. The SOA occupation group I data include the same combined occupations as the DTS includes in occupation classes 1 and 2-basically the so-called white collar group. Many of the policies and claims from the DTS have also been included in the SOA data shown here, since there is an overlapping of contributing companies and exposure years.

It was most encouraging for the Committee to find that our final DTS Basic Table compares well with the SOA's somewhat broader classification in both incidence rates and first-year claim costs for both males and females. The DTS class 1 is slightly lower than the SOA I, and the DTS class 2 is correspondingly higher than SOA I in nearly each corresponding cell.

Exhibit D-2 illustrates the high selectivity by elimination period. As an example, a block of policies having 7-day elimination periods, exposed at age 37 , will experience about 2.5 times the amount of claims during the sixth month of disablement as will a corresponding block of policies having 30-day elimination periods (500/207). This implies that net premiums and active life reserves will vary significantly depending upon the elimination period defining the particular table used to calculate them. This is illustrated in Exhibit D-3.

Exhibit D-3 shows the effect of calculating net premiums and active life reserves for policies having 30-day elimination periods using an improper table (7-day elimination period). Values based on the 7-day e.p. table are highly excessive versus the correct values as determined on a 30 -day e.p. table. As you would expect, a table designed for policies with a particular elimination period produces inconsistent results when applied to policies with other elimination periods.

Exhibit D-4 compares values from the DTS Basic Table by class and to the corresponding values from the 1964 CDT for three different policies; 2 year benefits with a 7 -day e.p.; 60 -month benefits with 30 -day e.p.; and benefits to age 65 with 90 -day e.p. This table illustrates that a table designed for any particular occupation class will produce incorrect results when used for any other occupation class.

The exhibit also shows that the 1964 CDT produces very conservative net premiums and reserves for the policies illustrated in occupation class 1 for all of the illustrated ages and elimination periods. It is, correspondingly,
much more representative of class 2 net premiums and reserves. Looking at class 3 and 4, the 1964 CDT produces much lower net premiums, and the active life reserves are generally inadequate.

Exhibit D-5 illustrates that the 1964 CDT generally produces active life reserves that are highly excessive for policies on female lives.

Exhibit D-6 compares claim reserves for class 1 to the corresponding claim reserves for class 2 . The DTS analysis found that the effect of class on termination rate tended to disap pear after three months for both accident and sickness. The effect on claim reserves is minimal. A reading of the exhibit shows claim reserves differing by class in the fourth and ninth month. The difference is not caused by the occupation class effect, but, rather, by the nature of construction. Since accident and sickness continuance tables are constructed separately, then added together, the combined table shows a difference by class until the thirteenth month.

Exhibit D-7 illustrates the general inadequacy of the 1964 table for claim reserves.

Exhibit D-8 compares net premiums, active life reserves, and claim reserves between the standard DTS table and a modified table. The modification employed was to reduce termination rates during the first year by setting them at 95 percent of the standard DTS Basic Table rate, then grading to 100 percent of the DTS termination rates in the eighteenth month. The result is a $9-10$ percent increase in the net premiums for the policies illustrated, with a generally $6-8$ percent higher active life reserve in the early durations, particularly for the lower ages. The very early duration claim reserves are increased by about 10 percent. The effect on claim reserve grades off, of course, during the 18 -month period.

Exhibit D-9 compares values calculated at 3 percent, with those at a more reasonable, but still conservative, rate of 6 percent. Calculations are based upon the DTS Basic Table with 1958 CSO Mortality.

## EXHIBIT D-1

Comparison of Society of Actuaries
SOA* Data to the Disability Table Study Results
(SOA-Occ. Group I: DTS-Classes 1 and 2)

| AGE |  | Elimination Period: 14 Days |  |  |  | Elimination Period: 30 Days |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | DTS | SOA-Grp I |  | DTS |  | SOA-Grp I |  | DTS |  |
| SOA |  | 1976-77 | 1978-79 | CL 1 | CL 2 | 1976-77 | 1978-79 | CL 1 | CL. 2 |
|  |  | Rates of Disablemen-MALES |  |  |  |  |  |  |  |
| <30 | 25 | . 051 | . 049 | . 031 | . 057 | . 015 | . 013 | . 010 | . 023 |
| 30-39 | 35 | . 052 | . 046 | . 034 | . 062 | . 013 | . 012 | . 011 | . 025 |
| 40-49 | 45 | . 056 | . 051 | . 041 | . 073 | . 021 | . 019 | . 017 | . 035 |
| 50-59 | 55 | . 073 | . 066 | . 062 | . 092 | . 035 | . 031 | . 030 | . 051 |
| 60-69 | 62 | . 109 | . 096 | . 084 | . 113 | . 058 | 049 | . 047 | . 069 |
|  |  | Claim Costs per $\$ 100$ per month-MALES ( 12 Month Benefit at $0 \%$ interest) |  |  |  |  |  |  |  |
| <30 | 25 | 12.40 | 11.50 | 6.69 | 12.26 | 4.50 | 3.40 | 2.86 | 6.11 |
| 30-39 | 35 | 12.80 | 11.30 | 7.93 | 14.83 | 4.00 | 3.80 | 3.16 | 7.22 |
| 40-49 | 45 | 14.90 | 14.70 | 10.94 | 19.44 | 7.00 | 6.20 | 5.44 | 11.23 |
| 50-59 | 55 | 21.90 | 20.00 | 19.42 | 28.79 | 13.30 | 12.10 | 10.58 | 18.33 |
| 60-69 | 62 | 42.90 | 33.80 | 30.31 | 41.14 | 24.90 | 19.80 | 18.13 | 27.16 |
|  |  | Rates of Disablement-FEMALES |  |  |  |  |  |  |  |
| $<30$ | 25 | . 067 | . 065 | . 052 | . 077 | . 027 | . 020 | . 020 | . 033 |
| 30-39 | 35 | . 094 | . 080 | . 070 | . 103 | . 033 | . 026 | . 031 | . 049 |
| 40-49 | 45 | . 111 | . 090 | . 082 | . 117 | . 041 | . 040 | . 041 | . 063 |
| 50-59 | 55 | . 096 | . 084 | . 076 | . 111 | . 050 | . 045 | . 042 | . 065 |
| 60-69 | 62 | . 116 | . 099 | . 087 | . 117 | . 054 | . 054 | . 055 | . 081 |
|  |  | Claim Costs per $\$ 100$ per Month-FEMALES <br> ( 12 Month Benefit at $0 \%$ Interest) |  |  |  |  |  |  |  |
| $<30$ | 25 | 15.00 | 13.80 | 11.19 | 16.40 | 6.60 | 5.30 | 5.25 | 8.79 |
| 30-39 | 35 | 22.40 | 19.90 | 16.44 | 24.17 | 10.30 | 7.70 | 8.48 | 13.57 |
| 40-49 | 45 | 31.20 | 26.40 | 21.48 | 30.92 | 11.60 | 12.30 | 12.22 | 18.71 |
| 50-59 | 55 | 28.20 | 25.50 | 22.38 | 33.06 | 21.00 | 18.70 | 14.00 | 21.59 |
| 60-69 | 65 | 40.50 | 32.90 | 27.74 | 38.06 | 20.50 | 23.00 | 19.80 | 29.32 |

*1981 Reports, pages 193-98.

## EXHIBIT D-2

Disability Continuance Tables
100,000 Lives Exposed
Accident and Sickness Combined-Male
DTS Basic Table versus 1964 CDT


EXHIBIT D-3*

|  | Male-Benefit to Age 65-30-Day Elimination Period |  |  |  |  |  | Female-Benctit to Age 65-30-Day Elimination Period |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 7-Day E.P. Tabie |  |  | 30-Day E.P. Table |  |  | 7-Day E.P. Table |  |  | 30-Day E.P. Table |  |  |
| Age | 30 | 40 | 50 | 30 | 40 | 50 | 30 | 40 | 50 | 30 | 40 | 50 |
| Class 1: |  |  |  |  |  |  |  |  |  |  |  |  |
| NLP | 39.34 | 48.64 | 56.90 | 19.80 | 26.72 | 33.78 | 59.82 | 69.44 | 67.16 | 36.75 | 44.13 | 43.35 |
| 5 yr . rsv. | 85 | 67 | -1 | 61 | 54 | 10 | 107 | 13 | -67 | 82 | 12 | $-30$ |
| 10 | 155 | 94 | -34 | 115 | 80 | -9 | 160 | -26 | -97 | 123 | -9 | -43 |
| 15 | 199 | 66 | ... | 152 | 68 | ... | 149 | -86 | $\ldots$ | 116 | -37 | .. |
| 20 | 199 | 3 | $\ldots$ | 159 | 22 | ... | 83 | -107 | ... | 75 | -47 | $\ldots$ |
| Class 2: |  |  |  |  |  |  |  |  |  |  |  |  |
| NLP | 60.56 | 73.33 | 82.66 | 38.24 | 49.09 | 57.14 | 84.83 | 99.08 | 93.66 | 57.07 | 67.41 | 66.00 |
| 5 yr . rsv. | 120 | 82 | -15 | 100 | 71 | -8 | 162 | 7 | $-100$ | 116 | 16 | - 50 |
| 10 | 212 | 106 | -61 | 180 | 91 | -41 | 237 | -62 | $-134$ | 172 | -16 | -72 |
| 15 | 262 | 62 | ... | 224 | 58 | ... | 209 | -145 | ... | 162 | -61 |  |
| 20 | 251 | -19 | $\ldots$ | 215 | -5 | $\ldots$ | 100 | -158 | $\ldots$ | 101 | -78 | $\cdots$ |
| Class 3: |  |  |  |  |  |  |  |  |  |  |  |  |
| NLP | 68.99 | 82.99 | 92.68 | 56.77 |  |  | 94.66 | 111.12 | 105.14 | 71.42 | 83.81 | 81.84 |
| 5 yr . rsv. | 132 | 87 | -21 | 114 | 86 | $-5$ | 187 | 10 | $-110$ | 138 | 20 | $-7{ }^{-84}$ |
| 10 | 233 | 110 | -72 | 205 | 117 | -51 | 274 | -68 | - 145 | 206 | -22 | $-104$ |
| 15 | 285 | 59 | ... | 260 | 79 | ... | 243 | -159 | ... | 195 | -87 |  |
| 20 | 269 | -28 | $\ldots$ | 257 | -4 | $\ldots$ | 119 | -172 | $\ldots$ | 118 | - 113 | $\ldots$ |
| Class 4: |  |  |  |  |  |  |  |  |  |  |  |  |
| NLP | 78.55 | 93.78 | 103.46 | 63.08 | 76.38 | 87.93 | 105.19 | 123.92 | 118.78 | 78.72 | 92.11 | 90.18 |
| 5 yr . rsv. | 146 | 89 | -24 | 122 | 95 | -7 | 209 | 20 | -116 | 149 | 23 | -77 |
| 10 | 253 | 110 | -77 | 221 | 131 | $-60$ | 311 | - 58 | - 156 | 223 | - 22 | - 113 |
| 15 | 305 | 55 | ... | 283 | 87 | 00 | 285 | -158 | ... | 212 | -92 |  |
| 20 | 283 | -34 | ... | 282 | -8 | $\ldots$ | 154 | -179 | $\ldots$ | 130 | -122 | $\ldots$ |

*lllustrating the need for separate tables by elimination period for active life reserves. Net premiums and reserves for a Benefit to Age 65 (2-year minimum) policy with a $30-$ day elimination period. Calculations are based on a 7 -day elimination period table versus a 30 -day elimination period table. DTS Basic Table with 1958 CSO Mortality and 3 percent interest rate.

## EXHIBIT D-4

Comparison of Male Active Life Terminal Reserves by Class
PER \$100 MONTHLY INCOME
(DTS Basic Table 1958 CSO Mortality Interest Rate 3\%)


## EXHIBIT D-5

Comparison of Female active Life Terminal Reserves
PER $\$ 100$ Monthly Income
(DTS Basic Tables 1958 CSO Mortality Interest Rate 3\%)

| $\begin{gathered} \text { DURation } \\ \text { (years) } \\ \hline \end{gathered}$ | Issue Age 30 |  |  | Issue Age 40 |  |  | Issue Age 50 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1964 \\ & \text { CDT } \\ & \hline \end{aligned}$ | DTS |  | $\begin{aligned} & 1964 \\ & \text { CDT } \\ & \hline \end{aligned}$ | DTS |  | $\begin{array}{r} 1964 \\ \text { CDT } \\ \hline \end{array}$ | DTS |  |
|  |  | Cl 1 | Cl 2 |  | Cl 1 | Cl 2 |  | CII | Cl 2 |
| EP 7/24 Month           <br> Maximum:           <br> NLP 22.85 30.45 42.99 29.73 35.12 49.75 40.92 38.90 53.92  |  |  |  |  |  |  |  |  |  |
| 5 yr . rsv. | 57 | 46 | 70 | 70.73 | 28.12 | 49.75 32 | 40.92 | 38.90 17 | 53.92 20 |
| 10 | 114 | 78 | 112 | 127 | 43 | 47 | 83 | 22 | 33 |
| 15 | 166 | 94 | 128 | 159 | 48 | 54 | 0 | 0 | 0 |
| 20 | 204 | 96 | 124 | 133 | 39 | 51 | ... | ... |  |
|  |  |  |  |  |  |  |  |  |  |
| ${ }^{\text {NLP }} 5$ | 21.02 | 22.43 | 34.62 91 | 29.64 89 | 29.14. | 74.34 | 43.93 | 36.19 | 54.79 |
| 10 | 143 | 63 112 | 162 | 89 162 | 40 80 | 70 119 | 68 | 41 60 | 59 82 |
| 15 | 210 | 142 | 208 | 196 | 99 | 145 | 0 | 0 | 0 |
| 20 | 259 | 156 | 229 | 133 | 91 | 130 | ... | ... | $\ldots$ |
|  |  |  |  |  |  |  |  |  |  |
| Age 65: |  |  |  |  |  |  |  |  |  |
| NLP | 22.80 | 20.58 | 34.18 | 32.80 | 26.61 | 42.19 | 42.72 | 28.47 | 43.97 |
| 5 yr. rsv. | 84 | 60 | 81 | 91 | 25 | 31 | 36 | (18) | (35) |
| 10 | 166 | 100 | 133 | 112 | 21 | 20 | 11 | (33) | (61) |
| 15 | 232 | 110 | 144 | 117 | (3) | (21) | 0 | 0 | 0 |
| 20 | 225 | 90 | 111 | 56 | (25) | (53) | $\ldots$ | ... | $\ldots$ |

## EXHIBIT D-6

Disability Income Claim Reserves
PER \$100/MONTH BEnefit, 3\% Interest Rate
Comparing Reserves by Occupation Class and Sex
(DTS Basic Table)

| Ades | Months | 60-month Benefit-30-Day Eimmation Preriod |  |  |  | To Age 65 Benefit-30-Day Elimination Pieriod |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Maic |  | Female |  | Maje |  | Female |  |
|  |  | Class 1 | Class 2 | Class 1 | Class 2 | Class 1 | Class 2 | Class 1 | Class 2 |
| 27 | 2 | 545 | 539 | 484 | 486 | 1,052 | 1,039 | 1,018 | 1,023 |
|  | 4 | 975 | 975 | 887 | 891 | 2,165 | 2,166 | 2,214 | 2,223 |
|  | 9 | 1,903 | 1,903 | 1,872 | 1,868 | 5,207 | 5,206 | 5,949 | 5,935 |
|  | 18 | 2.572 | 2,572 | 2,698 | 2,698 | 9,211 | 9,211 | 11,354 | 11,354 |
|  | 42 | 1,661 | 1,661 | 1,711 | 1,711 | 14,187 | 14,187 | 16,990 | 16,990 |
|  | 66 | ... | ... | ... | ... | 16,251 | 16,251 | 18,588 | 18,588 |
| 37 | 2 | 694 | 690 | 607 | 607 | 1,480 | 1,472 | 1,379 | 1,381 |
|  | 4 | 1,234 | 1,235 | 1,121 | 1,125 | 2,969 | 2,971 | 2,934 | 2,946 |
|  | 9 | 2,454 | 2,454 | 2,396 | 2,396 | 6,993 | 6,993 | 7,603 | 7,603 |
|  | 18 | 3,112 | 3.112 | 3,120 | 3,120 | 10.999 | 10,999 | 12,693 | 12,693 |
|  | 42 | 1,732 | 1.732 | 1,764 | 1,764 | 13,319 | 13.319 | 15,226 | 15.226 |
|  | 66 | ... | ... | ... | ... | 13,786 | 13.786 | -15,386 | 15,386 |
| 47 | 2 | 921 | 920 | 823 | 819 | 1,833 | 1.832 | 1,713 | 1.703 |
|  | 4 | 1,625 | 1.623 | 1,512 | 1,505 | 3,521 | 3,516 | 3,466 | 3,449 |
|  | 9 | 3,018 | 3,017 | 2,966 | 2,962 | 7,383 | 7,380 | 7,774 | 7,762 |
|  | 18 | 3.448 | 3,448 | 3,513 | 3,513 | 9,975 | 9,975 | 10,954 | 10,952 |
|  | 42 | 1.766 | 1.766 | 1,789 | 1.789 | 10,220 | 10,220 | 11.209 | 11,209 |
|  | 66 | ... | ... | ... | ... | 9,537 | 9.537 | 10,323 | 10,323 |
| 57 | 2 | 1,243 | 1,242 | 1,108 | 1,105 | 1,614 | 1,613 | 1,451 | 1,446 |
|  | 4 | 2,134 | 2,125 | 2,005 | 1.992 | 2,854 | 2,841 | 2,713 | 2,695 |
|  | 9 | 3.468 | 3.465 | 3,429 | 3,425 | 4,861 | 4,857 | 4,882 | 4,875 |
|  | 18 | 3,576 | 3.576 | 3,633 | 3.633 | 5.419 | 5,419 | 5,603 | 5,603 |
|  | 42 | 1.774 | 1.774 | 1,795 | 1.794 | 4,052 | 4.052 | 4,189 | 4.189 |
|  | 66 | ... | ... | ... | ... | 2,199 | 2,199 | 2,243 | 2,243 |

## EXHIBIT D-7

Comparison of Claim Reserves
PER $\$ 100$ Monthly Benefit
1964 CDT versus Male versus Female
(DTS Basic Table Interest Rate 3\%)

| Duration <br> Since: <br> Disabif. <br> MENT <br> (Months) | 60 Month Benefit |  |  |  |  | To Age 65 Beneft |  |  |  |  | Lifetime Benemt |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1964 \\ & \text { CDT } \end{aligned}$ | DTS Class 1 |  |  |  | $1964$ | DTS Class 1 |  |  |  | $\begin{aligned} & 1964 \\ & \text { CDT } \end{aligned}$ | DTS Class 1 |  |  |  |
|  |  | 1 Month EP |  | 3 Month EP* |  |  | 1 Month EP |  | 3 Month EP |  |  | 1 Mont EP |  | 3 Month EP |  |
|  |  | Male | Female | Male | Female |  | Male | Female | Male | Female |  | Male | Female | Male | Female |
| Age 27: |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 4 | 824 | 975 | 887 | 1,073 | 985 | 1,631 | 2,165 | 2,214 | 2,357 | 2,430 | 1,665 | 2,238 | 2,348 | 2,436 | 2,578 |
| 9 | 2,499 | 1,903 | 1,872 | 1,948 | 1,919 | 5,579 | 5,207 | 5,949 | 5,220 | 5,958 | 5,714 | 5,409 | 6,361 | 5,422 | 6,371 |
| 18 | 2,712 | 2,572 | 2,698 | 2,654 | 2,791 | 7,654 | 9,211 | 11,354 | 9,211 | 11,354 | 7,863 | 9,617 | 12,229 | 9,617 | 12,228 |
| 42 | 1,559 | 1,661 | 1,711 | 1,815 | 1,875 | 10,099 | 14,187 | 16,990 | 14,187 | 16,990 | 10,460 | 14,953 | 18,533 | 14,953 | 18,533 |
| 66 | ... | ... | ... |  | ... | 11,439 | 16,251 | 18,588 | 16,251 | 18,588 | 11,948 | 17.276 | 20,518 | 17,276 | 20,518 |
| Age 37: |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 4 | 903 | 1,234 | 1,121 | 1,362 | 1,245 | 1,908 | 2,969 | 2,934 | 3,219 | 3,203 | 2,023 | 3,194 | 3,294 | 3,463 | 3,598 |
| 9 | 2.735 | 2,454 | 2,396 | 2,519 | 2,463 | 6,433 | 6,993 | 7,603 | 6,994 | 7,604 | 6,873 | 7,582 | 8,638 | 7,583 | 8,639 |
| 18 | 2,885 | 3,112 | 3,120 | 3,225 | 3,322 | 8,490 | 10,999 | 12,693 | 10,999 | 12,693 | 9,127 | 12,022 | 14,580 | 12,022 | 14.580 |
| 42 | 1,598 | 1,732 | 1,764 | 1,899 | 1,937 | 10,529 | 13,319 | 15,226 | 13,319 | 15,226 | 11,542 | 14,822 | 17,902 | 14,822 | 17,902 |
| 66 | ... | ... | ... | ... | ... | 11,251 | 13,786 | 15,386 | 13,786 | 15,386 | 12,599 | 15,640 | 18,545 | 15,640 | 18,545 |
| Age 47: |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 4 | 1,173 | 1,625 | 1,512 | 1,795 | 1,674 | 2,301 | 3,521 | 3,466 | 3,806 | 3,755 | 2,673 | 4,183 | 4,428 | 4,524 | 4,799 |
| 9 | 3,016 | 3,018 | 2,966 | 3,111 | 3,059 | 6,303 | 7,383 | 7,774 | 7,392 | 7,781 | 7,542 | 8,908 | 10,142 | 8,919 | 10,151 |
| 18 | 3,051 | 3,448 | 3,513 | 3,581 | 3,653 | 7.834 | 9,975 | 10,954 | 9,975 | 10,952 | 9,463 | 12,256 | 14,616 | 12,256 | 14,616 |
| 42 | 1,623 | 1,766 | 1,789 | 1,939 | 1,967 | 8,705 | 10,220 | 11,209 | 10,220 | 11,209 | 11,117 | 13,174 | 15,849 | 13,174 | 15,849 |
| 66 | ... | ... | ... | ... | ... | 8,496 | 9,537 | 10,323 | 9,537 | 10,323 | 11,600 | 13,047 | 15,656 | 13,047 | 15,656 |
| Age 57: |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 4 | 1,702 | 2,134 | 2,005 | 2,329 | 2,196 | 2,330 | 2,854 | 2.713 | 3,038 | 2,897 | 3.682 | 4,845 | 5,225 | 5,162 | 5,585 |
| 9 | 3,363 | 3,468 | 3,429 | 3,577 | 3,539 | 4,312 | 4,861 | 4,882 | 4,865 | 4,885 | 7,596 | 8,716 | 10,038 | 8,723 | 10,045 |
| 18 | 3,212 | 3,576 | 3,633 | 3,716 | 3,779 | 4,482 | 5,419 | 5,603 | 5,419 | 5,603 | 8,889 | 10,517 | 12,595 | 10,517 | 12,595 |
| 42 | 1,628 | 1,774 | 1,795 | 1,947 | 1,973 | 3,384 | 4,052 | 4,189 | 4,052 | 4,189 | 9,479 | 10,357 | 12,684 | 10,357 | 12,684 |
| 66 | 1... | ... |  | ... | ... | 1,651 | 2,199 | 2,243 | 2,199 | 2,243 | 9,527 | 9,691 | 12,005 | 9.691 | 12,005 |

*Shown only to compare male to female. Nol comparable to 1964 CDT one month EP because benefit period is two months longer

EXHIBIT D-8
Effect of Lowering Termination Rates from Disablement
by 5\% in Year 1, Grading to Standard at End of 18 Months
Male-EP 30 Days- 1958 CSO Mortality- $\mathbf{3 \%}$ Interest Rate

|  | Values per $\$ 100$ per Month |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Class |  |  | Class 2 |  |  |
| Age | 30 | 40 | 50 | 30 | 40 | 50 |
| DTS Basic Table: 60 Month NLP | 13.83 | 19.96 | 30.07 | 25.43 | 35.05 | 49.30 |
| Res. 5 year | 51 | 63 | 60 | 81 | 91 | 77 |
| 10 year | 102 | 115 | 76 | 160 | 162 | 95 |
| To Age 65 NLP | 19.80 | 26.72 | 33.78 | 38.24 | 49.09 | 57.14 |
| Res. 5 year | 61 | 54 | 10 | 100 | 71 | -8 |
| 10 year | 115 | 80 | -9 | 180 | 91 | $-41$ |
| DTS Valuation Table: <br> 60 Manth NLP | 15.07 | 21.7 | 326 | 27.74 | 38.15 | 53.46 |
| Res. 5 year | 55 | 67 | 65 | 88 | 98 | 83 |
| 10 year | 111 | 123 | 81 | 173 | 174 | 101 |
| To Age 65 NLP | 21.89 | 29.38 | 36.83 | 42.36 | 54.07 | 62.33 |
| Res. 5 year | 66 | 57 | 9 | 109 | 74 | -13 |
| 10 year | 125 | 85 | $-12$ | 195 | 94 | -49 |
| Ratio Valuation/Basic: 60 Month NLP | 1.09 | 1.09 |  |  |  |  |
| Res. 5 year | 1.08 | 1.06 | 1.08 | 1.09 | 1.08 | 1.08 |
| 10 year | 1.09 | 1.07 | 1.07 | 1.08 | 1.07 | 1.06 |
| To Age 65 NLP | 1.11 | 1.10 | 1.09 | 1.11 | 1.10 | 1.09 |
| Res. 5 year | 1.08 | 1.06 | . 90 | 1.09 | 1.04 | -- |
| 10 year | 1.08 | 1.06 | - | 1.08 | 1.03 | - |

Claim Reserves
(Class 1)

|  | Age 30 |  | AGE 40 |  | AGE 50 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $60 \text { Month }$ BP | To Age $65$ | $\begin{aligned} & 60 \text { Month } \\ & \mathrm{BP} \end{aligned}$ | To Age 65 | $\begin{aligned} & 60 \text { Month } \\ & \text { BP } \end{aligned}$ | To Age $65$ |
| DTS Basic Table: |  |  |  |  |  |  |
| 2 Month | 587 | 1,186 | 751 | 1,601 | 1,008 | 1,872 |
| 4 Month | 1,046 | 2,425 | 1,334 | 3,178 | 1.771 | 3,530 |
| 9 Month | 2,065 | 5,836 | 2,624 | 7,306 | 3,175 | 7,029 |
| 18 Month | 2.755 | 10,017 | 3,234 | 11,007 | 3,509 | 9,036 |
| DTS Valuation Table: |  |  |  |  |  |  |
| 2 Month | 642 | 1,322 | 819 | 1,771 | 1,091 | 2,040 |
| 4 Month | 1,115 | 2,610 | 1,416 | 3,395 | 1,860 | 3,720 |
| 9 Month | 2,112 | 5,983 | 2,673 | 7,451 | 3,217 | 7,127 |
| 18 Month | 2,755 | 10,017 | 3,244 | 11,007 | 3,509 | 9,036 |
| Ratio Valuation/Basic: |  |  |  |  |  |  |
| 2 Month | 1.09 | 1.11 | 1.09 | 1.11 | 1.08 | 1.09 |
| 4 Month | 1.07 | 1.08 | 1.06 | 1.07 | 1.05 | 1.05 |
| 9 Month | 1.02 | 1.03 | 1.02 | 1.02 | 1.01 | 1.01 |
| 18 Month | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |

EXHIBIT D-9
Comparison of Values for Interest Rates of 3\% and $6 \%$
Male-EP 30 Days- 1958 CSO Mortality-DTS Basic Table

|  | Values fer \$100 pex Month |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Class 1 |  |  | Class 2 |  |  |
| Age | 30 | 40 | 50 | 30 | 40 | 50 |
| 3\% Interest: |  |  |  |  |  |  |
| 60 Month NLP | 13.83 | 19.96 | 30.07 | 25.43 | 35.05 | 49.30 |
| Res. 5 year | 51 | 63 | 60 | 81 | 91 | 77 |
| 10 year | 102 | 115 | 76 | 160 | 162 | 95 |
| To Age 65 NLP | 19.80 | 26.72 | 33.78 | 38.24 | 49.09 | 57.14 |
| Res. 5 year | 61 | 54 | 10 | 100 | 71 | -8 |
| 10 year | 115 | 80 | -9 | 180 | 91 | -41 |
| 6\% Interest: |  |  |  |  |  |  |
| 60 Month NLP | 10.45 | 16.55 | 26.80 | 20.08 | 29.85 | 44.42 |
| Res. 5 year | 36 | 51 | 54 | 60 | 75 | 69 |
| 10 year | 78 | 98 | 72 | 124 | 140 | 89 |
| To Age 65 NLP | 14.77 | 22.09 | 30.43 | 29.34 | 41.29 | 51.81 |
| Res. 5 year | 46 | 49 | 16 | 79 | 69 | 4 |
| 10 year | 93 | 80 | 0 | 152 | 101 | -25 |
| Ratio 6/3: |  |  |  |  |  |  |
| 60 Month NLP | . 76 | 83 | 89 | . 79 | . 85 | . 90 |
| Res. 5 year | . 71 | . 81 | . 90 | . 74 | . 82 | . 90 |
| 10 year | . 76 | 85 | 95 | . 78 | . 86 | . 94 |
| To Age 65 NLP | . 75 | . 83 | 90 | . 77 | . 84 | . 91 |
| Res. 5 year | . 75 | . 91 | 1.60 | . 79 | . 97 | - |
| 10 year | . 81 | 1.00 | - | . 84 | 1.11 | - |

Claim Reserves
(Class 1)

|  | Age 30 |  | Age 40 |  | Age 50 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} 60 \text { Month } \\ \mathrm{BP} \\ \hline \end{gathered}$ | $\begin{gathered} \text { To Age } \\ 65 \\ \hline \end{gathered}$ | $\begin{gathered} 60 \text { Month } \\ \text { BP } \\ \hline \end{gathered}$ | $\begin{gathered} \text { To Age } \\ 65 \\ \hline \end{gathered}$ | $\begin{gathered} 60 \text { Month } \\ \mathrm{BP} \\ \hline \end{gathered}$ | $\begin{gathered} \text { To Age } \\ 65 \\ \hline \end{gathered}$ |
| 3\% Interest: |  |  |  |  |  |  |
| 2 Month | 587 | 1,186 | 751 | 1,601 | 1,008 | 1,872 |
| 4 Month | 1,046 | 2,425 | 1,334 | 3,178 | 1,771 | 3,530 |
| 9 Month | 2,065 | 5,836 | 2,624 | 7,306 | 3,175 | 7,029 |
| 18 Month | 2,755 | 10,017 | 3,234 | 11,007 | 3,509 | 9,036 |
| 6\% Interest: |  |  |  |  |  |  |
| 2 Month | 565 | 961 | 717 | 1,318 | 958 | 1,626 |
| 4 Month | 1,001 | 1,916 | 1,268 | 2,578 | 1,675 | 3,043 |
| 9 Month | 1,968 | 4,502 | 2,490 | 5,855 | 3,005 | 6,038 |
| 18 Month | 2,636 | 7,622 | 3,087 | 8,797 | 3,346 | 7,791 |
| Ratio 6/3: |  |  |  |  |  |  |
| 2 Month | . 96 | . 81 | . 96 | . 82 | . 95 | . 87 |
| 4 Month | . 96 | . 79 | . 95 | . 81 | . 95 | . 86 |
| 9 Month | . 95 | . 77 | . 95 | . 80 | . 95 | . 86 |
| 18 Month | . 96 | . 76 | . 95 | . 80 | . 95 | . 86 |

## APPENDIX E

## DTS BASIC TABLE

The DTS Basic Table has been defined, in pieces, in Appendixes B and C of this study. This appendix illustrates the calculation. The DTS Basic Table includes the incidence rates (probability of becoming disabled), and the termination rates (probability of termination of disability by recovery or death). Incidence rates vary by:

1. Cause: accident and sickness
2. Sex: male and female
3. Class: occupation class $1,2,3$, and 4 where, in a 5 -class manual, class 1 is 4 A and $3 A$, class 2 is $2 A$, class 3 is $A$, and class 4 is $B$
4. E.P.: elimination periods of 0 days, 7 days, 14 days, 30 days, and 90 days.

Each of these 72 cells will produce its own unique continuance table. A table for accident and sickness combined is obtained by adding the accident continuance table and the sickness continuance table cell by cell (i.e., for each age and duration). Values for individual ages were obtained by Lagrange interpolation with adjustment for end values.

Termination from disability rates are for each week during the first 13 weeks of disablement. They are then expressed as monthly factors until the twenty-fourth month, yearly through the tenth year, and by attained age thereafter.

The termination rate for any particular duration since disablement is the product of the factors corresponding to the profile of each claim. Rates for the 10 -year age groups are appropriate for individual ages $25,35,45,55$, and 62. Values for individual ages were determined by Lagrange interpolation as shown in Exhibit E-1.

As an example:

|  | Factors |  |
| :--- | ---: | :---: |
| The probability of terminating from claim in week 2, | .120 duration rate |  |
| for claimants age 35 at disablement, | .961 |  |
| with 7-day elimination periods, | .934 |  |
| in occupation class 3, | .977 |  |
| who are male claimants, | 1.190 |  |
| disabled from accidents, | is |  |
|  | $\underline{.13407}$ |  |

The incidence rate for that same group is .05837 or 5,837 for each 100,000 lives exposed. Terminations during the second week of disablement are 763, leaving 5,074 disabled lives at end of the second week.

The corresponding sickness incidence rate is .05272 , and the corresponding termination rate is .1168 , leaving 4,656 disabled lives at the end of the second week of disablement, out of the 5,272 that entered week 2 , disabled.

The combined accident and sickness table, then, has 11,109 lives disabled at the end of 7 days of whom 9,730 are still disabled at the end of 14 days. See Exhibit E-2.

The one week entries for individual ages may be obtained by multiplying by 100 the appropriate rates per 1,000 shown in Exhibit E-3. These rates for individual ages, as well as the termination rates for individual ages were obtained by the following Lagrange Interpolation Formula, modified for the end points. The aggregate tables are easily constructed for any particular mix of business. The DTS Basic Table is shown in Exhibits E-4a-c and E5. Illustrative continuance tables for combined accident and sickness are shown in Exhibits E-6a and b.

At the time the exposure draft of this report was distributed, a diskette containing a series of programs was made available to perform a variety of calculations. The software functioned on an IBM-PC or IBM-compatible PC.

The program first builds a continuance table for either the experience or the valuation table (margins added). Then, the software can be used to compute any of the following:

1. Claim cost for $\$ 100$ per month
2. Claim cost for $\$ 1,000$ lump sum
3. Disabled life reserves per $\$ 100$ per month
4. Disabled life pure endowment of $\$ 1,000$
5. NLP method net premiums and reserves
6. 1-year FPT net premiums and reserves
7. 2-year FPT net premiums and reserves

Copies of the diskette have been distributed to over 300 individuals. The software is essentially the official working version of the Committee's report and is incorporated into the NAIC recommendation that considers the Committee's DTS Valuation Table to be the " 1985 Commissioners Individual Disability Tables A."

The software originally sent out has not been altered. To obtain a copy of the diskette, contact the Research Department of the Society of Actuaries.

## EXHIBIT E-I

## 5-point Lagrange Interpolation Formula

Used for incidence rates and termination rates. Given points $F(a), F(b), F(c), F(d)$, and $F(e)$, then:

$$
\begin{aligned}
F(x) & =\frac{(x-b)}{(a-b)} \frac{(x-c)}{(a-c)} \frac{(x-d)}{(a-d)} \frac{(x-e)}{(a-e)} F(a) \\
& +\frac{(x-a)}{(b-a)} \frac{(x-c)}{(b-c)} \frac{(x-d)}{(b-d)} \frac{(x-e)}{(b-e)} F(b)
\end{aligned}
$$

$$
+\frac{(x-a)}{(e-a)} \frac{(x-b)}{(e-b)} \frac{(x-c)}{(e-c)} \frac{(x-d)}{(e-d)} F(e)
$$

for $a<x<e$;
$a, b, c, d$, and $e$ are ages $25,35,45,55$, and 62 , respectively.
When $x \leq 25$ :
for incidence rates, $F(x)=F(25)$
for termination rates, $F(x)=F(25)+(25-x)[F(25)-F(26)]$.
When $x \geq 62$ :
$F(x)=F(62)+(x-62)[F(62)-F(61)]$.

## EXHIBIT E-2

DTS Continuance Table*
Number of Persons Alive \& Disabled at the End of the
Duration From Date of Disablement
100,000 Lives Exposed to Disablement
Male-Class 3-7-Day Elimination Period-Age 35

| Duration | Accident | Sickness | Combined |
| :---: | :---: | :---: | :---: |
| 1 (Weeks) | 5,837.00 | 5,272.00 | 11,109.00 |
| 2 | 5,073.90 | 4,656.04 | 9,729.94 |
| 3 | 4,384.01 | 4,030.60 | 8,414.61 |
| 4 | 3,786.69 | 3,445.39 | 7.232 .08 |
| 5 | 3,270.06 | 2,918.64 | 6,188.70 |
| 6 | 2,823.74 | 2,461.28 | 5,285.02 |
| 7 | 2,443.84 | 2,073.39 | 4,517.23 |
| 8 | 2,125.76 | 1,754.41 | 3,880.17 |
| 9 | 1,860.35 | 1,495.77 | 3.356 .12 |
| 10 | 1,639.57 | 1,288.70 | 2,928.27 |
| 11 | 1,459.41 | 1,127.31 | 2,586.72 |
| 12 | 1,313.26 | 1,003.72 | 2,316.99 |
| 13 | 1,197.33 | 913.13 | 2,110.46 |
| 4 (Months). | 857.02 | 637.62 | 1,494.64 |
| 5....... | 650.48 | 475.58 | 1,126.07 |
| 6 | 518.37 | 375.00 | 893.37 |
| 7 | 435.15 | 310.51 | 745.66 |
| 8 | 377.81 | 265.49 | 643.29 |
| 9 | 338.48 | 236.05 | 574.52 |
| 10. | 309.83 | 215.44 | 525.27 |
| 11 | 286.09 | 198.92 | 485.01 |
| 12 | 265.21 | 184.81 | 450.02 |
| 13 | 247.50 | 172.47 | 419.97 |
| 14 | 232.54 | 162.04 | 394.58 |
| 15 | 219.86 | 153.21 | 373.07 |
| 16 | 209.22 | 145.79 | 355.01 |
| 17 | 200.63 | 139.81 | 340.44 |
| 18 | 193.43 | 134.79 | 328.22 |
| 19 | 187.46 | 130.63 | 318.10 |
| 20 | 182.40 | 127.11 | 309.51 |
| 21 | 178.02 | 124.05 | 302.06 |
| 22 | 174.24 | 121.42 | 295.66 |
| 23 | 170.82 | 119.03 | 289.85 |
| 24 | 167.72 | 116.87 | 284.60 |
| 3 (Years) | 141.79 | 98.80 | 240.59 |
| 4. | 126.96 | 88.47 | 215.42 |
| 5 | 117.52 | 81.89 | 199.41 |
| 6. | 111.11 | 77.42 | 188.53 |
| 7 | 106.31 | 74.08 | 180.39 |
| 8 | 102.55 | 71.46 | 174.01 |
| 9 | 99.28 | 69.18 | 168.46 |
| 10. | 96.30 | 67.10 | 163.40 |

*illustrating the results of the preceding sample of consinuction.

EXHIBIT E-3
DTS Basic Table
Incidence of Disability
Rates Per 1,000 Lives Exposed

|  | Male-Actident |  |  |  |  |  | Male-Sickness |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Agil | Elmination Period |  |  |  |  | Age | Elimination Period |  |  |  |  |
|  |  | 0-day | 7-day | 14-day | 30-day | 90-day |  | 0-day | 7-day | 14-day | 30-day | 90-day |
| Class 1: | 25 | 33.97 | 25.84 | 13.13 | 4.90 | . 86 | 25 |  | 32.26 | 18.22 | 5.51 | 1.01 |
|  | 35 | 32.88 | 24.42 | 11.99 | 4.23 | . 51 | 35 |  | 36.11 | 21.55 | 6.48 | 1.13 |
|  | 45 | 30.40 | 20.40 | 9.86 | 4.50 | . 65 | 45 |  | 47.12 | 31.19 | 12.63 | 2.70 |
|  | 55 | 30.19 | 18.32 | 9.63 | 4.71 | . 80 | 55 |  | 69.48 | 52.75 | 25.11 | 7.78 |
|  | 62 | 33.45 | 16.11 | 10.39 | 5.47 | 1.18 | 62 | . | 91.52 | 74.06 | 41.24 | 15.20 |
| Class 2: | 25 | 59.96 | 47.98 | 30.01 | 10.48 | 2.07 | 25 |  | 46.61 | 27.01 | 12.17 | 2.23 |
|  | 35 | 59.96 | 44.62 | 28.83 | 10.14 | 2.09 | 35 |  | 52.79 | 33.37 | 14.47 | 2.56 |
|  | 45 | 56.74 | 38.49 | 25.67 | 9.86 | 2.14 | 45 |  | 65.97 | 46.91 | 25.40 | 6.21 |
|  | 55 | 51.66 | 31.31 | 20.50 | 10.03 | 2.20 | 55 |  | 92.99 | 71.27 | 41.37 | 15.74 |
|  | 62 | 52.84 | 29.85 | 19.86 | 10.92 | 2.57 | 62 |  | 116.81 | 93.05 | 58.54 | 25.94 |
| Class 3: | 25 | 75.80 | 62.68 | 42.87 | 23.69 | 7.04 | 25 |  | 46.83 | 32.22 | 14.75 | 2.99 |
|  | 35 | 74.78 | 58.37 | 39.59 | 22.57 | 6.48 | 35 |  | 52.72 | 38.32 | 18.70 | 3.52 |
|  | 45 | 69.76 | 50.41 | 34.61 | 20.49 | 5.97 | 45 |  | 67.05 | 51.53 | 29.45 | 7.83 |
|  | 55 | 66.37 | 44.27 | 30.51 | 18.49 | 5.46 | 55 |  | 92.60 | 76.39 | 52.66 | 20.07 |
|  | 62 | 65.04 | 39.98 | 27.96 | 18.56 | 5.30 | 62 |  | 116.23 | 98.78 | 78.56 | 36.04 |
| Class 4: | 25 | 89.42 | 77.60 | 52.59 | 27.03 | 8.73 | 25 |  | 48.20 | 33.28 | 15.07 | 3.04 |
|  | 35 | 91.59 | 73.24 | 50.53 | 26.93 | 8.17 | 35 |  | 53.75 | 39.27 | 19.33 | 3.59 |
|  | 45 | 84.64 | 62.13 | 42.61 | 24.78 | 7.68 | 45 |  | 70.03 | 52.71 | 30.13 | 7.97 |
|  | 55 | 79.77 | 52.03 | 37.34 | 22.78 | 7.27 | 55 |  | 95.01 | 77.91 | 55.87 | 20.45 |
|  | 62 | 79.95 | 49.76 | 36.11 | 22.96 | 7.20 | 62 |  | 119.16 | 101.41 | 81.62 | 36.63 |

EXHIBIT E-3-Continued

|  | Female-accident |  |  |  |  |  | Female-Sickness |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age | Elimination Period |  |  |  |  | Age | Elimination Period |  |  |  |  |
|  |  | 0 -day | 7-day | 14-day | 30-day | 90-day |  | 0-day | 7-day | 14 -day | 30-day | 90-day |
| Class 1: | 25 | 23.06 | 19.92 | 12.96 | 6.00 | 1.14 | 25 |  | 61.10 | 39.29 | 14.03 | 2.55 |
|  | 35 | 26.28 | 20.87 | 13.39 | 6.21 | . 91 | 35 |  | 84.38 | 56.89 | 24.75 | 4.37 |
|  | 45 | 32.36 | 22.77 | 13.78 | 6.83 | 1.11 | 45 |  | 94.57 | 68.33 | 34.14 | 7.64 |
|  | 55 | 45.05 | 26.77 | 14.82 | 8.06 | 1.46 | 55 |  | 90.28 | 61.49 | 34.23 | 10.31 |
|  | 62 | 69.00 | 31.56 | 17.54 | 9.91 | 2.25 | 62 |  | 93.06 | 69.44 | 45.30 | 13.85 |
| Class 2: | 25 | 35.05 | 31.48 | 23.39 | 13.40 | 3.22 | 25 |  | 80.97 | 53.57 | 20.03 | 3.75 |
|  | 35 | 39.36 | 32.01 | 23.36 | 14.02 | 3.20 | 35 |  | 116.02 | 80.05 | 35.34 | 6.60 |
|  | 45 | 47.46 | 33.55 | 24.40 | 15.02 | 3.40 | 45 |  | 134.18 | 92.93 | 47.62 | 10.81 |
|  | 55 | 62.53 | 37.10 | 26.13 | 16.11 | 3.75 | 55 |  | 117.29 | 84.93 | 49.00 | 14.95 |
|  | 62 | 88.91 | 44.31 | 29.27 | 17.88 | 4.46 | 62 |  | 120.40 | 87.53 | 63.15 | 18.86 |
| Class 3: | 25 |  |  | 27.94 | 17.63 | 6.19 | 25 |  | 86.64 | 57.85 | 24.83 | 5.03 |
|  | 35 | 46.30 | 38.45 | 28.54 | 18.20 | 6.54 | 35 |  | 124.79 | 96.77 | 44.67 | 8.43 |
|  | 45 | 53.01 | 39.08 | 29.09 | 19.24 | 6.75 | 45 |  | 145.58 | 116.19 | 58.44 | 14.43 |
|  | 55 | 66.71 | 41.96 | 30.86 | 20.99 | 7.08 | 55 |  | 122.98 | 99.89 | 59.99 | 17.86 |
|  | 62 | 90.05 | 48.12 | 33.60 | 23.74 | 7.26 | 62 |  | 125.95 | 101.06 | 69.18 | 22.76 |
| Class 4: | 25 | 52.41 | 47.52 | 34.93 | 22.04 | 7.74 | 25 |  | 90.24 | 60.26 |  | 5.23 |
|  | 35 | 57.87 | 48.07 | 35.67 | 22.75 | 8.17 | 35 |  | 130.00 | 100.81 | 46.53 | 8.79 |
|  | 45 | 66.26 | 48.86 | 36.36 | 24.05 | 8.45 | 45 |  | 151.65 | 121.04 | 60.87 | 15.03 |
|  | 55 | 83.39 | 52.45 | 38.58 | 26.25 | 8.85 | 55 |  | 128.10 | 104.05 | 62.49 | 18.61 |
|  | 62 | 112.57 | 60.16 | 42.00 | 29.67 | 9.08 | 62 |  | 131.20 | 105.27 | 72.07 | 23.71 |

## EXHIBIT E-4a

DTS Basic Table
Factors for Calculation of Weekly Termination Rates


Class 1 includes the two lowest premium classes of a 5 -class manual or the lowest premium class of a 4 -class manual.
*Use .084 for 30 -day elimination periods to allow for the short week from 30 to 35 days.

DTS Basic Table
Factors for Determination of Weekly Termination Rates

| Week: | 6 | 7 | 8 |  |  | 9 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | . 123 | . 126 | . 125 |  |  | 122 |  |  |
| Age: 25 | 1.060 | 1.066 | 1.073 |  |  | 1.079 |  |  |
| EP: 0,7,14,30 | 1.0761 .2101 .048 .689 | 1.0181 .1771 .053 .760 | . 9801.1471 .054 |  | . 820 | . $9581.1181 .049 \quad .873$ |  |  |
| Class: 1,2,3,4 | . 9921.0081 .007 .990 | . 9861.0101 .009 .993 | . 9831.0091 .010 . 997 |  |  | . 9781.0071 .0121 .004 |  |  |
| Sex: M,F | 1.036 .965 | 1.022 .978 | 1.012 .988 |  |  | 1.004 .995 |  |  |
| Cause: A.S | . 8781.118 | . 8741.125 | . 8711.129 |  |  | . 8701.131 |  |  |
| Age: 35 | 1.019 | 1.043 | 1.058 |  |  | 1.066 |  |  |
| EP: 0,7,14,30 | $\left\lvert\, \begin{array}{rrrr} 1.164 & 1.153 & .998 & .701 \\ .999 & 1.003 & 1.002 & .994 \end{array}\right.$ | $\begin{array}{r}1.119 \\ \hline .9961 .121 \\ 1.001 \\ \hline 1.006 \\ \hline\end{array}$ | $1.0821 .0991 .013 .807$ |  |  | $1.05!1.0821 .017$ |  |  |
| Class: 1,2,3,4 |  |  | $9931.0001 .0041 .003$ |  |  | $990 \quad .9991 .0051 .006$ |  |  |
| Sex: M,F | 1.019 .981 | .9961 .0011 .003 .998 .9941 .005 | . 9781.022 |  |  | . 9671.033 |  |  |
| Cause: A,S | . 9251.062 | . 9161.073 | . 9121.078 |  |  | . 9131.078 |  |  |
| Age: 45 | . 988 | 1.007 | 1.019 |  |  | 1.024 |  |  |
| EP: 0,7,14,30 | $1.2061 .096 \quad .962 .738$ | 1.1721 .073 . 974.783 | 1.1431 .057 | . 983 | . 818 | 1.1131 .046 | . 990 | . 851 |
| Class: 1,2,3,4 | 1.0151 .006 .995 .983 | 1.0101 .003 . 996.990 | 1.0061 .000 | . 997 | . 995 | 1.004 .999 | . 998 | . 998 |
| Sex: M,F | 1.005 .995 | . 9841.016 | . 9691.031 |  |  | . 9591.042 |  |  |
| Cause: A,S | 1.002 . 981 | . 989.994 | .9821 .001 |  |  | . 9811.003 |  |  |
| Age: 55 | . 969 | . 964 | . 961 |  |  | . 957 |  |  |
| EP: 0,7,14,30 | $1.2201 .052 \quad .930 .786$ | $1.1961 .041 \quad .946 .814$ | 1.1711 .031 | . 957 | . 841 | 1.1471 .021 | . 964 | . 869 |
| Class: 1,2,3,4 | 1.0411 .018 .985 .956 | 1.0301 .013 .989 .968 | 1.0231 .009 | . 991 | . 976 | 1.0201 .007 | . 993 | . 981 |
| Sex: M,F | . 9951.005 | . 9901.010 | . 9841.016 |  |  | . 9761.024 |  |  |
| Cause: A,S | 1.111 .884 | 1.098 . 895 | $1.089 \quad .902$ |  |  | 1.084 .908 |  |  |
| Age: 62 | . 965 | . 920 | . 890 |  |  | . 874 |  |  |
| EP: 0,7,14,30 | 1.1961 .031 .896 .849 | $1.1911 .031 \quad .910 .857$ | 1.1801 .024 | . 917 | . 876 | 1.1661 .010 | . 919 | . 907 |
| Class: 1,2,3,4 | $1.0901 .040 \quad .971 .906$ | $1.0711 .037 \quad .977 .921$ | 1.0581 .033 | 980 | . 933 | 1.0481 .028 | . 982 | . 944 |
| Sex: M,F | . 9881.012 | 1.011 .988 | $1.025 \quad .975$ |  |  | 1.024 .976 |  |  |
| Cause: A,S | 1.260 .780 | $1.253 \quad .785$ | 1.245 . 790 |  |  | $1.236 \quad .796$ |  |  |

## EXHIBIT E-4a-(Cominued)

DTS Basic Table
Factors for Determination of Weekly Termination Rates

| Week: |  |  | 11 |  |  | 12 |  |  | 13 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | $\qquad$ |  | . 109 |  |  | . 099 |  |  | . 086 |  |  |
| Age: 25 | 1.086 |  | 1.096 |  |  | 1.110 |  |  | 1.133 |  |  |
| EP: 0,7,14.30 | $\begin{array}{rrrr}.951 & 1.087 & 1.038 & .921 \\ .972 & 1.002 & 1.013 & 1.013\end{array}$ |  | $.9631 .051 \quad 1.018 \quad .964$ $966 \quad 9941.0151026$ |  |  | $.9961 .008 \quad .9851 .007$ |  |  | $1.059 \quad .949 \quad .9351 .050$ |  |  |
| Class: 1,2,3,4 |  |  | . 9751.018 1.021.074 |  |  |  |  |  |
| Sex: M,F | . 9971.002 |  |  |  |  | . 9901.008 |  |  | . 9841.013 |  |  |
| Cause: A,S | . 8711.131 |  | . 8761.127 |  |  | . 8841.118 |  |  | . 9871.104 |  |  |
| Age: 35 | 1.068 |  | 1.062 |  |  | 1.049 |  |  | 1.027 |  |  |
| EP: 0,7,14,30 | $\begin{array}{rrrr}1.025 & 1.069 & 1.019 & .885 \\ .986 & .997 & 1.006 & 1.010\end{array}$ |  | $1.0031 .058 \quad 1.017 \quad .920$ |  |  | . 9851.0491 .008 . 955 |  |  | $\begin{array}{llll}.971 & 1.038 & .989 & .992\end{array}$ |  |  |
| Class: 1,2,3,4 |  |  | . $981 . .9961 .0071 .015$ |  |  | . 974 . 9941.0091 .002 |  |  | . 962.9931 .0121 .032 |  |  |
| Sex: M,F | .986 .997 1.0061 .010 <br> .961 1.040  |  | . 9301.060 |  |  | . 9591.039 |  |  | . 9671.026 |  |  |
| Cause: A,S | . 9191.072 |  |  |  |  | $.9501 .040$ |  |  | . 9841.006 |  |  |
| Age: 45 | 1.022 |  | 1.012 |  |  | . 993 |  |  | . 962 |  |  |
| EP: 0,7,14,30 | 1.0831 .040 | . 995.882 | $\begin{array}{rrrr}1.048 & 1.039 & .998 & .914 \\ 1.001 & 1.000 & 1.000 & .999\end{array}$ |  |  | $\left\lvert\, \begin{array}{llr} 1.007 & 1.043 & .997 \\ 1.000 & 1.003 & 1.000 \end{array}\right.$ |  | .951.995 | $\begin{array}{rrrr}.952 & 1.054 & .989 & .995 \\ 1.000 & 1.008 & 1.001 & .989\end{array}$ |  |  |
| Class: 1,2,3,4 | $1.002 \quad .999$ | .9991 .000 |  |  |  |  |  |  |  |  |  |
| Sex: M,F | . 9511.050 |  | . 9461.055 |  |  | $\begin{array}{r}.9431 .057 \\ 1.020 \\ \hline\end{array}$ |  |  | . 9421.053 |  |  |
| Cause: A,S | . 986 . 999 |  | . 998.989 |  |  |  |  |  | $1.058 \quad .935$ |  |  |
| Age: 55 | . 953 |  | . 948 |  |  | . 941 |  |  | . 932 |  |  |
| EP: 0,7,14,30 | 1.1211 .013 | . 967 . 900 | 1.090 1.005 .966 .938 <br> 1.022 1.006 .992 .980 |  |  | $\begin{array}{ll} 1.052 & .997 \\ 1.031 & 1.009 \end{array}$ |  | $\begin{aligned} & .989 \\ & .971 \end{aligned}$ | $\begin{array}{rr} .999 & .988 \\ 1.048 & 1.015 \end{array}$ | . 9431.062 |  |
| Class: 1,2,3,4 | 1.0191 .005 | . 993.982 |  |  |  | $\begin{array}{\|rrrr} 1.048 & 1.015 & .984 & .953 \\ .908 & 1.092 & & \end{array}$ |  |  |  |
| Sex: M,F | . 9661.034 |  | 1.022 1.006 .992 .980 <br> .953 1.048   |  |  |  |  |  |  |  | $\begin{array}{\|rrrr} 1.031 & 1.009 & .989 & .971 \\ .935 & 1.066 & & \end{array}$ |  |  |  |
| Cause: A,S | $1.082 \quad .910$ |  | $1.086 \quad .909$ |  |  | $\begin{array}{r} .9351 .066 \\ 1.094 \quad .904 \\ \hline \end{array}$ |  |  | $1.110 \quad 891$ |  |  |
| Age: 62 | . 871 |  | . 881 |  |  | . 907 |  |  | . 946 |  |  |
| EP: 0,7,14,30 | 1.147 . 987 | . 917 . 951 | $\begin{array}{rrrr} 1.119 & .956 & .913 & 1.017 \\ 1.041 & 1.016 & .984 & .961 \end{array}$ |  |  | $\left[\begin{array}{lrrr} 1.079 & .914 & .906 & 1.114 \\ 1.043 & 1.009 & .982 & .967 \end{array}\right.$ |  |  | 1.024 .853 .894 1.265 <br> 1.052 .998 .978 .972 |  |  |
| Class: 1,2,3,4 | 1.0431 .022 | . 984.953 |  |  |  |  |  |  |  |  |  |  |  |
| Sex: M,F | 1.008 . 991 |  | $\begin{array}{r} 1.0411 .016 \\ .9751 .024 \end{array}$ | $.984$ |  | $\begin{array}{r} 1.0431 .009 \\ .9201 .083 \end{array}$ | $.982$ | $.967$ | $\begin{array}{rr} 1.052 & .998 \\ .844 & 1.175 \end{array}$ | . $978 \quad .972$ |  |
| Cause: A,S | $1.223 \quad .806$ |  | $1.210 \quad .816$ |  |  | $\begin{array}{r} .9201 .083 \\ 1.193 \quad .829 \\ \hline \end{array}$ |  |  | $1.166 \quad .849$ |  |  |

## EXHIBIT E-4b

DTS Basic Table
Factors for Calculation of Monthly Termination Rates

| Month: | 4 |  | 5 |  | 6 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | 236 |  | . 208 |  | . 182 |  |
| $\begin{aligned} & \text { <90d EP } \\ & 90 \mathrm{~d} \mathrm{EP} \end{aligned}$ | $\begin{array}{r} 1.172 \\ \hline .828 \end{array}$ |  | $1.109$ |  | 1.051 |  |
| Male: Female: | $\begin{array}{r} .989 \\ 1.011 \end{array}$ |  | .981 1.019 |  | .9751.025 |  |
| Age: 25 A,S | 1.082 | 1.186 | 1.103 | 1.182 | 1.149 | 1.173 |
| 35 A,S | 1.039 | 1.103 | 1.065 | 1.123 | 1.089 | 1.134 |
| 45 A,S | 1.012 | . 989 | 1.045 | . 993 | 1.061 | . 989 |
| 55 A,S | 1.017 | . 857 | . 980 | . 837 | . 970 | . 809 |
| $62 \mathrm{~A}, \mathrm{~S}$ | 981 | . 732 | . 971 | . 701 | 963 | 663 |
| Month: | 7 |  | 8 |  | 9 |  |
| Duration Rate: | . 153 |  | . 124 |  | . 095 |  |
| Male: | $\begin{array}{r} .947 \\ 1.053 \end{array}$ |  | $\begin{array}{r} .943 \\ 1.057 \end{array}$ |  | $\begin{array}{r} .939 \\ 1.061 \end{array}$ |  |
| Female: |  |  |  |  |  |  |
| Age: 25 A,S | 1.2041 .218 |  | 1.2591 .262 |  | 1.351 1.289 |  |
| 35 A,S | 1.1081 .187 |  | 1.1271 .240 |  | 1.1671.031 | 1.243 |
| 45 A,S | $1.040 \quad 1.019$ |  | 1.0191 .048 |  |  | $1.031 \quad 1.021$ |
| 55 A,S | $\begin{array}{ll} .920 & .815 \\ .835 & .657 \end{array}$ |  | .869 .820 <br> .706 .651 |  | $\begin{array}{r}.856 \\ .671 \\ \hline\end{array}$ | .772.600 |
| $62 \mathrm{~A}, \mathrm{~S}$ |  |  |  |  |  |  |
| Month: | 10 |  | 11 |  | 12 |  |
| Duration Rate: | . 075 |  | . 066 |  | . 060 |  |
| Male: | . 935 |  | . 931 |  | . 945 |  |
| Female: | 1.065 |  | 1.069 |  | 1.055 |  |
| Age: 25 A,S | 1.442 | 1.317 | 1.534 | 1.344 | 1.626 | 1.371 |
| 35 A,S | 1.207 | 1.245 | 1.247 | 1.248 | 1.287 | 1.251 |
| 45 A,S | 1.042 | . 993 | 1.054 | 966 | 1.066 | . 939 |
| 55 A,S | . 844 | 724 | . 831 | . 676 | . 818 | . 628 |
| 62 A.S | . 637 | . 550 | . 602 | 499 | . 567 | 448 |

## EXHIBIT E-4b-Continued

DTS Basic Table
Factors for Calculation of Monthly Termination Rates
Second Year of Disablement

| Month: | 13 | 14 | 15 | 16 | 17 | 18 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration Rate: | .054 | .048 | .043 | .038 | .032 | .028 |
| Male: | .960 | .975 | .978 | .981 | .984 | .988 |
| Female: | 1.040 | 1.025 | 1.022 | 1.019 | 1.016 | 1.012 |
| Age: 25 | 1.558 | 1.625 | 1.692 | 1.758 | 1.825 | 1.897 |
| 35 | 1.288 | 1.292 | 1.296 | 1.299 | 1.303 | 1.298 |
| 45 | .971 | .937 | .903 | .869 | .835 | .797 |
| 55 | .658 | .629 | .600 | .571 | .542 | .516 |
| 62 | .524 | .517 | .510 | .503 | .496 | .493 |
| Month: | 19 | 20 | 21 | 22 | 23 | 24 |
| Duration Rate: | .024 | .021 | .019 | .017 | .016 | .015 |
| Male: | .993 | .997 | 1.001 | 1.005 | 1.009 | 1.013 |
| Female: | 1.007 | 1.003 | .999 | .995 | .991 | .987 |
| Age: 25 | 1.970 | 2.042 | 2.061 | 2.079 | 2.098 | 2.117 |
| 35 | 1.294 | 1.289 | 1.265 | 1.241 | 1.217 | 1.193 |
| 45 | .758 | .720 | .706 | .693 | .679 | .665 |
| 55 | .489 | .463 | .471 | .479 | .487 | .495 |
| 62 | .489 | .486 | .497 | .508 | .519 | .530 |

## EXHIBIT E-4c

DTS Basic Table
Factors for Calculation of
Annual Termination Rates
Years 3 Through 10

| Year: | 3 | 4 | 5 | 0 |
| :--- | :---: | :---: | :---: | :---: |
| Duration Rate: | .123 | .084 | .062 | .050 |
| Male: | 1.080 | 1.129 | 1.179 | 1.200 |
| Female: | .920 | .871 | .821 | .800 |
| Age: 25 | 2.085 | 1.832 | 1.554 | 1.262 |
| 35 | 1.164 | 1.103 | 1.017 | .909 |
| 45 | .727 | .757 | .767 | .754 |
| 55 | .536 | .616 | .697 | .832 |
| 62 | .489 | .691 | .965 | 1.244 |
| Year: | 7 | 8 | 9 | 10 |
| Duration Rate: | .045 | .042 | .042 | .043 |
| Male: | 1.212 | 1.210 | 1.204 | 1.200 |
| Female: | .788 | .790 | .796 | .800 |
| Age: 25 | .994 | .776 | .617 | .524 |
| 35 | .792 | .696 | .631 | .582 |
| 45 | .741 | .737 | .739 | .751 |
| 55 | 1.489 | 1.103 | 1.182 | 1.226 |
| 62 |  | 1.688 | 1.830 | 1.918 |

EXHIBIT E-5
DTS Basic Table
Ultimate Termination Rates for Duration 11 Years and Over
by Attained Age

| Attained Age | Mate | Female | Attained <br> AgE | Male | Female |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 30 | . 0238 | . 0160 | 65 | . 0665 | . 0446 |
| 31 | . 0240 | . 0161 | 66 | . 0707 | . 0474 |
| 32 | . 0242 | . 0162 | 67 | . 0753 | . 0504 |
| 33 | . 0244 | . 0163 | 68 | . 0802 | . 0538 |
| 34 | . 0246 | . 0165 | 69 | . 0857 | . 0574 |
| 35 | . 0249 | . 0167 | 70 | . 0916 | . 0614 |
| 36 | 0251 | . 0168 | 71 | . 0986 | . 0657 |
| 37 | . 0254 | . 0170 | 72 | . 1051 | . 0704 |
| 38 | . 0258 | . 0173 | 73 | . 1127 | . 0755 |
| 39 | . 0261 | . 0175 | 74 | . 1210 | . 0811 |
| 40 | . 0265 | . 0178 | 75 | . 1301 | . 0871 |
| 41 | . 0270 | . 0181 | 76 | . 1398 | . 0937 |
| 42 | . 0275 | . 0184 | 77 | . 1504 | 1008 |
| 43. | . 0280 | . 0188 | 78 | . 1619 | . 1085 |
| 44. | . 0286 | . 0192 | 79 | . 1743 | . 1168 |
| 45 | . 0292 | . 0196 | 80 | . 1878 | . 1258 |
| 46 | . 0299 | . 0200 | 81 | . 2022 | . 1355 |
| 47 | . 0306 | . 0205 | 82 | . 2178 | . 1459 |
| 48 | . 0315 | . 0211 | 83 | . 2345 | . 1571 |
| 49 | . 0324 | . 0217 | 84 | . 2525 | . 1691 |
| 50 | . 0334 | . 0224 | 85 | . 2717 | . 1820 |
| 51 | . 0345 | . 0231 | 86 | . 2922 | . 1958 |
| 52 | . 0357 | . 0239 | 87 | . 3140 | . 2104 |
| 53. | . 0370 | . 0248 | 88 | . 3372 | . 2259 |
| 54 | . 0384 | . 0257 | 89 | . 3618 | . 2424 |
| 55. | . 0400 | . 0268 | 90 | . 3877 | . 2598 |
| 56. | . 0417 | . 0279 | 91 | . 4149 | . 2780 |
| 57 | . 0436 | . 0292 | 92 | . 4435 | . 2971 |
| 58 | . 0456 | . 0306 | 93 | . 4732 | . 3171 |
| 59 | . 0479 | . 0321 | 94 | . 5041 | . 3378 |
| 60 | . 0503 | . 0337 | 95 | . 5360 | 3591 |
| 61 | . 0530 | . 0355 | 96 | . 5686 | . 3801 |
| 62. | . 0559 | . 0375 | 97 | . 6020 | . 4033 |
| 63. | . 0592 | . 0397 | 98 | . 6357 | . 4259 |
| 64. | . 0627 | . 0420 | 99. | . 6695 | 4486 |

## EXHIBIT E-6a

DTS Continuance Table (Basic Table)
Number of Persons Alive \& Disabled at the End of the Duration from Date of Disablement
100.000 Lives Exposed to Disablement

Sex: Male Cause: Combined Class: 1 EP: 30-day

| Dukation | Ages at Disablement |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 27 | 37 | 47 | 57 |
| 4 Week (30-day) | 970.30 | 1,162.16 | 1,901.25 | 3,372.00 |
| 5 | 916.80 | 1,099.34 | 1,796.17 | 3,172.94 |
| 6 | 833.30 | 1,000.06 | 1,634.45 | 2,889.74 |
| 7 | 748.66 | 900.40 | 1,475.48 | 2,623.05 |
| 8 | 667.91 | 806.50 | 1,327.83 | 2,379.27 |
| 9 | 593.92 | 721.02 | 1,194.02 | 2,159.40 |
| 10 | 528.07 | 645.49 | 1,076.22 | 1,963.71 |
| 11 | 471.57 | 581.25 | 975.77 | 1,792.87 |
| 12 | 424.20 | 528.07 | 892.28 | 1,645.22 |
| 3 Month | 385.98 | 486.00 | 825.95 | 1,520.86 |
| 4 | 267.37 | 344.90 | 605.51 | 1.168 .27 |
| 5 | 199.06 | 260.13 | 471.26 | 951.46 |
| 6 | 156.47 | 206.70 | 385.54 | 811.37 |
| 7 | 129.27 | 172.72 | 330.32 | 719.17 |
| 8 | 110.35 | 149.13 | 291.65 | 653.67 |
| 9 | 97.50 | 133.45 | 266.22 | 610.99 |
| 10 | 88.21 | 122.31 | 248.43 | 581.63 |
| 11 | 80.55 | 113.29 | 234.24 | 558.82 |
| 12 | 73.85 | 105.49 | 222.22 | 540.10 |
| 13 | 68.08 | 98.78 | 211.80 | 523.07 |
| 14 | 63.11 | 93.13 | 203.19 | 508.76 |
| 15 | 58.82 | 88.37 | 196.09 | 496.75 |
| 16 | 55.17 | 84.39 | 190.27 | 486.80 |
| 17 | 52.17 | 81.20 | 185.73 | 478.93 |
| 18 | 49.61 | 78.54 | 182.03 | 472.39 |
| 19 | 47.44 | 76.35 | 179.08 | 467.06 |
| 20 | 45.57 | 74.50 | 176.67 | 462.63 |
| 21 | 43.93 | 72.90 | 174.55 | 458.56 |
| 22 | 42.51 | 71.52 | 172.69 | 454.84 |
| 23 | 41.21 | 70.27 | 170.98 | 451.28 |
| 24 | 40.02 | 69.15 | 169.40 | 447.89 |

EXHIBIT E-6a-Continued

| Dukation | ages at Disablement |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 27 | 37 | 47 | 57 |
| 3 Year | 30.20 | 59.53 | 154.23 | 417.19 |
| 4 | 25.51 | 53.81 | 143.80 | 392.75 |
| 5 | 22.90 | 50.05 | 136.10 | 371.70 |
| 6 | 21.30 | 47.45 | 130.04 | 351.47 |
| 7 | 20.21 | 45.46 | 124.66 | 330.53 |
| 8 | 19.43 | 43.87 | 119.74 | 309.76 |
| 9 | 18.82 | 42.45 | 114.94 | 288.87 |
| 10 | 18.30 | 41.13 | 110.12 | 268.22 |
| 11 | 17.84 | 39.87 | 105.32 | 248.03 |
| 12 | 17.38 | 38.62 | 100.51 | 228.13 |
| 13 | 16.92 | 37.37 | 95.70 | 208.58 |
| 14 | 16.47 | 36.12 | 90.88 | 189.48 |
| 15 | 16.03 | 34.87 | 86.07 | 170.79 |
| 16 | 15.59 | 33.63 | 81.26 | 152.84 |
| 17 | 15.15 | 32.38 | 76.45 | 135.62 |
| 18 | 14.72 | 31.14 | 71.65 | 119.21 |
| 19 | 14.29 | 29.90 | 66.89 | 103.70 |
| 20 | 13.86 | 28.65 | 62.16 | 89.20 |
| 21 | 13.44 | 27.40 | 57.48 | 75.79 |
| 22 | 13.01 | 26.15 | 52.87 | 63.57 |
| 23 | 12.59 | 24.90 | 48.34 | 52.49 |
| 24 | 12.17 | 23.65 | 43.91 | 42.63 |
| 25 | 11.75 | 22.39 | 39.58 | 34.01 |
| 26 | 11.33 | 21.14 | 35.42 | 26.60 |
| 27 | 10.91 | 19.89 | 31.43 | 20.37 |
| 28 | 10.49 | 18.64 | 27.63 | 15.22 |
| 29 | 10.07 | 17.40 | 24.03 | 11.09 |
| 30 | 9.65 | 16.17 | 20.67 | 7.85 |
| 31 | 9.23 | 14.95 | 17.56 | 5.38 |
| 32 | 8.81 | 13.75 | 14.73 | 3.57 |
| 33 | 8.39 | 12.58 | 12.16 | 2.28 |
| 34 | 7.97 | 11.42 | 9.88 | 1.39 |
| 35 | 7.55 | 10.30 | 7.88 | 0.82 |
| 36 | 7.12 | 9.22 | 6.17 | 0.45 |
| 37 | 6.70 | 8.18 | 4.72 | 0.24 |
| 38 | 6.28 | 7.19 | 3.53 | 0.12 |
| 39 | 5.86 | 6.25 | 2.57 | 0.06 |
| 40 | 5.45 | 5.38 | 1.82 | 0.02 |
| 41 | 5.04 | 4.57 | 1.25 | 0.01 |
| 42 | 4.64 | 3.83 | 0.83 | 0.00 |

## EXHIBIT E-6b

DTS Continuance Table (Basic Table)
Number of Persons alive \& Disabled at the End of the
Duration from Date of Disablement
100.000 Lives Exposed to Disablement

Sex: Male Cause: Combined Class: 3 EP:7-day

| Durathon | Ages at Disablement |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 27 | 37 | 47 | 57 |
| I Week | 11,027.45 | 11,153.01 | 12,026.30 | 14.212.98 |
| 2 | 9,476.60 | 9,810.86 | 10,815.26 | 13,005.77 |
| 3 | 8,064.23 | 8,517.64 | 9,585.30 | 11.751.55 |
| 4 | 6,824.12 | 7.348.04 | 8,428.65 | 10,526.79 |
| 5 | 5,756.71 | 6,310.26 | 7,369.39 | 9,373.49 |
| 6 | 4,852.04 | 5,406.22 | 6,419.79 | 8,309.58 |
| 7 | 4,097.55 | 4,634.51 | 5,588.29 | 7,357.03 |
| 8 | 3,483.66 | 3,991.30 | 4,880.64 | 6.534 .52 |
| 9 | 2,987.65 | 3,460.05 | 4,285.85 | 5,836.00 |
| 10 | 2,589.83 | 3,025.01 | 3,792.95 | 5,252.59 |
| 11 | 2,277.14 | 2,676.58 | 3,393.46 | 4,779.26 |
| 12 | 2,034.06 | 2,400.87 | 3,076.46 | 4,405.68 |
| 13 | 1,851.62 | 2,189.41 | 2,834.39 | 4,126.64 |
| 4 Month | 1,289.26 | 1.558.31 | 2,072.62 | 3,146.38 |
| 5 | 963.66 | 1,177.19 | 1,608.39 | 2,549.78 |
| 6 | 758.90 | 936.10 | 1,312.01 | 2,165.26 |
| 7 | 627.82 | 783.46 | 1,123.01 | 1,915.43 |
| 8 | 536.45 | 677.89 | 991.70 | 1,739.76 |
| 9 | 473.98 | 607.19 | 904.84 | 1,624.86 |
| 10 | 428.55 | 556.72 | 843.74 | 1,545.47 |
| 11 | 390.96 | 515.59 | 794.77 | 1,483.49 |
| 12 | 357.97 | 479.94 | 753.03 | 1.432.31 |
| 13 | 330.02 | 449.38 | 717.72 | 1.387.15 |
| 14 | 305.91 | 423.69 | 688.57 | 1.349 .20 |
| 15 | 285.13 | 402.04 | 664.49 | 1,317.37 |
| 16 | 267.40 | 383.94 | 644.79 | 1,290.98 |
| 17 | 252.89 | 369.42 | 629.38 | 1,270.10 |
| 18 | 240.46 | 357.31 | 616.84 | 1,252.76 |
| 19 | 229.96 | 347.34 | 606.84 | 1.238.64 |
| 20 | 220.88 | 338.94 | 598.69 | 1.226 .89 |
| 21 | 212.94 | 331.66 | 591.50 | 1,216.07 |
| 22 | 206.05 | 325.39 | 585.19 | 1,206.22 |
| 23 | 199.74 | 319.72 | 579.39 | 1.196 .79 |
| 24 | 193.97 | 314.59 | 574.06 | 1,187.78 |

EXHIBIT E-6b-Contimued

| Duration | Ages at disablement |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 27 | 37 | 47 | 57 |
| 3 Year | 146.40 | 270.81 | 522.64 | 1,106.36 |
| 4 | 123.67 | 244.79 | 487.31 | $1,041.56$ |
| 5 | 111.01 | 227.70 | 461.21 | 985.74 |
| 6 | 103.26 | 215.86 | 440.66 | 932.09 |
| 7 | 97.95 | 206.82 | 422.43 | 876.56 |
| 8 | 94.18 | 199.58 | 405.76 | 821.46 |
| 9 | 91.21 | 193.15 | 389.49 | 766.08 |
| 10 | 88.71 | 187.14 | 373.16 | 711.32 |
| 11 | 86.46 | 181.41 | 356.89 | 657.76 |
| 12 | 84.23 | 175.70 | 340.61 | 605.00 |
| 13 | 82.03 | 170.00 | 324.30 | 553.15 |
| 14 | 79.85 | 164.33 | 307.99 | 502.49 |
| 15 | 77.70 | 158.66 | 291.66 | 452.94 |
| 16 | 75.56 | 152.99 | 275.36 | 405.34 |
| 17 | 73.45 | 147.33 | 259.06 | 359.66 |
| 18 | 71.35 | 141.68 | 242.81 | 316.14 |
| 19 | 69.26 | 136.01 | 226.67 | 275.01 |
| 20 | 67.19 | 130.34 | 210.64 | 236.56 |
| 21 | 65.13 | 124.65 | 194.78 | 200.98 |
| 22 | 63.08 | 118.97 | 179.16 | 168.58 |
| 23 | 61.04 | 113.27 | 163.80 | 139.20 |
| 24 | 59.00 | 107.57 | 148.80 | 113.06 |
| 25 | 56.97 | 101.87 | 134.13 | 90.20 |
| 26 | 54.93 | 96.18 | 120.03 | 70.55 |
| 27 | 52.90 | 90.48 | 106.50 | 54.01 |
| 28 | 50.87 | 84.81 | 93.62 | 40.37 |
| 29 | 48.83 | 79.17 | 81.44 | 29.40 |
| 30 | 46.80 | 73.57 | 70.05 | 20.81 |
| 31 | 44.76 | 68.03 | 59.52 | 14.28 |
| 32 | 42.72 | 62.58 | 49.92 | 9.46 |
| 33 | 40.67 | 57.21 | 41.22 | 6.04 |
| 34 | 38.62 | 51.97 | 33.48 | 3.70 |
| 35 | 36.58 | 46.85 | 26.71 | 2.16 |
| 36 | 34.53 | 41.92 | 20.89 | 1.20 |
| 37 | 32.49 | 37.20 | 15.99 | 0.63 |
| 38 | 30.45 | 32.70 | 11.96 | 0.31 |
| 39 | 28.43 | 28.44 | 8.71 | 0.15 |
| 40 | 26.42 | 24.47 | 6.16 | 0.06 |
| 41 | 24.43 | 20.79 | 4.23 | 0.03 |
| 42 | 22.47 | 17.44 | 2.80 | 0.01 |


[^0]:    * W. Duane Kidwell, Co-chairman, William J. Taylor, Co-chairman, David S. Cox, William Daniels, Bryant O. Gamble, Frank E. Knorr, Roland E. Nelson, James J. Olsen, Richard Ostuw. Donald M. Pearsall, Edward J. Seligman, Robert B. Shapland, and John Haynes Miller (Special Consultant to the Committee).

[^1]:    *Rates of incidence and termination vary by age, sex, cause, elimination period, and class for incidence and for each claim duration as shown. Claim termination data were analyzed as to the significance by duration for each of 12 reported variables. The variables found to be significant and the durations for which they are significant are shown. (Note: The period of observation provided too little data to determine the significance of the "his own occ." clause.)
    **Incidence rates for accident only.
    ***Small volume of data.

[^2]:    ${ }^{1}$ The recommendation was adopted by the NAIC at its December 1985 meeting. The DTS Valuation Table is now known officially as the "Commissioners 1985 Individual Disability Tables A."

[^3]:    

[^4]:    ${ }^{2}$ Transactions of the Society of Actuaries, 1980 Reports, page 163.

