# TRANSACTIONS OF SOCIETY OF ACTUARIES 1995 VOL. 47 

## 1994 GROUP ANNUITY MORTALITY TABLE AND 1994 GROUP ANNUITY RESERVING TABLE

## SOCIETY OF ACTUARIES GROUP ANNUITY VALUATION TABLE TASK FORCE*

## EXECUTIVE SUMMARY

The Society of Actuaries Group Annuity Valuation Table Task Force has completed its research and has developed a table that it recommends as suitable for a new Group Annuity Reserve Valuation Standard.

The proposed new table, recommended as suitable for a new Group Annuity Reserve Valuation Standard, if accepted and adopted by regulators, would incorporate the use of generational mortality into statutory reserving requirements for group annuities for the first time. Generational mortality allows for the recognition of explicit assumptions for future mortality improvement in the calculation of reserve values.

The Task Force strongly believes that the use of generational mortality in group annuity reserving is appropriate given the trends in morality improvement that have been observed in the past and the continued improvement expected to occur in the foreseeable future. Modern systems capabilities are sufficient to allow for the increased refinement and computation intensity that generational mortality requires.

## The 1994 Group Annuity Reserving Table

The 1994 Group Annuity Reserving Table appears in Table 1. This table includes $q_{x}$ values on an age nearest birthday basis for each age in 1994 and projection factors to be used in generating $q_{x}$ values in years beyond 1994.

## Use of the Values in the Table To Produce Projected Mortality Rates

The values in the 1994 Group Annuity Reserving Table are as follows:
$q_{x}^{1994}=$ the mortality rate for a person age $x$ in 1994.

[^0]TABLE 1
1994 Group Annuity Reserving Table

| Age ( $x$ ) | Male |  | Female |  |  | Male |  | Female |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $q_{x}^{1094}$ | $A A_{4}$ | $9^{1944}$ | $A A_{x}$ | Age ( $x$ ) | $9_{1}^{1994}$ | $A A_{1}$ | $q_{1}^{1944}$ | $A A_{1}$ |
| 1 | 0.000592 | 0.020 | 0.000531 | 0.020 | 31 | 0.000821 | 0.005 | 0.000373 | 0.008 |
| 2 | 0.000400 | 0.020 | 0.000346 | 0.020 | 32 | 0.000839 | 0.005 | 0.000397 | 0.008 |
| 3 | 0.000332 | 0.020 | 0.000258 | 0.020 | 33 | 0.000848 | 0.005 | 0.000422 | 0.009 |
| 4 | 0.000259 | 0.020 | 0.000194 | 0.020 | 34 | 0.000849 | 0.005 | 0.000449 | 0.010 |
| 5 | 0.000237 | 0.020 | 0.000175 | 0.020 | 35 | 0.000851 | 0.005 | 0.000478 | 0.011 |
| 6 | 0.000227 | 0.020 | 0.000163 | 0.020 | 36 | 0.000862 | 0.005 | 0.000512 | 0.012 |
| 7 | 0.000217 | 0.020 | 0.000153 | 0.020 | 37 | 0.000891 | 0.005 | 0.000551 | 0.013 |
| 8 | 0.000201 | 0.020 | 0.000137 | 0.020 | 38 | 0.000939 | 0.006 | 0.000598 | 0.014 |
| 9 | 0.000194 | 0.020 | 0.000130 | 0.020 | 39 | 0.000999 | 0.007 | 0.000652 | 0.015 |
| 10 | 0.000197 | 0.020 | 0.000131 | 0.020 | 40 | 0.001072 | 0.008 | 0.000709 | 0.015 |
| 11 | 0.000208 | 0.020 | 0.000138 | 0.020 | 41 | 0.001156 | 0.009 | 0.000768 | 0.015 |
| 12 | 0.000226 | 0.020 | 0.000148 | 0.020 | 42 | 0.001252 | 0.010 | 0.000825 | 0.015 |
| 13 | 0.000255 | 0.020 | 0.000164 | 0.020 | 43 | 0.001352 | 0.011 | 0.000877 | 0.015 |
| 14 | 0.000297 | 0.019 | 0.000189 | 0.018 | 44 | 0.001458 | 0.012 | 0.000923 | 0.015 |
| 15 | 0.000345 | 0.019 | 0.000216 | 0.016 | 45 | 0.001578 | 0.013 | 0.000973 | 0.016 |
| 16 | 0.000391 | 0.019 | 0.000242 | 0.015 | 46 | 0.001722 | 0.014 | 0.001033 | 0.017 |
| 17 | 0.000430 | 0.019 | 0.000262 | 0.014 | 47 | 0.001899 | 0.015 | 0.001112 | 0.018 |
| 18 | 0.000460 | 0.019 | 0.000273 | 0.014 | 48 | 0.002102 | 0.016 | 0.001206 | 0.018 |
| 19 | 0.000484 | 0.019 | 0.000280 | 0.015 | 49 | 0.002326 | 0.017 | 0.001310 | 0.018 |
| 20 | 0.000507 | 0.019 | 0.000284 | 0.016 | 50 | 0.002579 | 0.018 | 0.001428 | 0.017 |
| 21 | 0.000530 | 0.018 | 0.000286 | 0.017 | 51 | 0.002872 | 0.019 | 0.001568 | 0.016 |
| 22 | 0.000556 | 0.017 | 0.000289 | 0.017 | 52 | 0.003213 | 0.020 | 0.001734 | 0.014 |
| 23 | 0.000589 | 0.015 | 0.000292 | 0.016 | 53 | 0.003584 | 0.020 | 0.001907 | 0.012 |
| 24 | 0.000624 | 0.013 | 0.000291 | 0.015 | 54 | 0.003979 | 0.020 | 0.002084 | 0.010 |
| 25 | 0.000661 | 0.010 | 0.000291 | 0.014 | 55 | 10.04425 | 0.019 | 0.002294 | 0.008 |
| 26 | 0.000696 | 0.006 | 0.000294 | 0.012 | 56 | 00.004949 | 0.018 | 0.002563 | 0.006 |
| 27 | 0.000727 | 0.005 | 0.000302 | 0.012 | 57 | 0.005581 | 0.017 | 0.002919 | 0.005 |
| 28 | 0.000754 | 0.005 | 0.000314 | 0.012 | 58 | 0.006300 | 0.016 | 0.003359 | 0.005 |
| 29 | 0.000779 | 0.005 | 0.000331 | 0.012 | 59 | 0.007090 | 0.016 | 0.003863 | 0.005 |
| 30 | 0.000801 | 0.005 | 0.000351 | 0.010 | 60 | 0.007976 | 0.016 | 0.004439 | 0.005 |
| Male q(55) should be 0.004425 |  |  |  |  |  |  |  |  |  |


| Agc (x) | Male |  | Female |  |  | Male |  | Female |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | ${ }_{4}^{104}$ | $A A_{1}$ | $9_{1}^{\text {mex }}$ | AA, | Age ( $x$ ) | 9.194 | $\mathrm{AA}^{\text {a }}$ | $q_{x}^{194}$ | AA, |
| 61 | 0.008986 | 0.015 | 0.005093 | 0.005 | 91 | 0.167260 | 0.004 | 0.128751 | 0.003 |
| 62 | 0.010147 | 0.015 | 0.005832 | 0.005 | 92 | 0.182281 | 0.003 | 0.141973 | 0.003 |
| 63 | 0.011471 | 0.014 | 0.006677 | 0.005 | 93 | 0.198392 | 0.003 | 0.155931 | 0.002 |
| 64 | 0.012940 | 0.014 | 0.007621 | 0.005 | 94 | 0.215700 | 0.003 | 0.170677 | 0.002 |
| 65 | 0.014535 | 0.014 | 0.008636 | 0.005 | 95 | 0.233606 | 0.002 | 0.186213 | 0.002 |
| 66 | 0.016239 | 0.013 | 0.009694 | 0.005 | 96 | 0.251510 | 0.002 | 0.202538 | 0.002 |
| 67 | 0.018034 | 0.013 | 0.010764 | 0.005 | 97 | 0.268815 | 0.002 | 0.219655 | 0.001 |
| 68 | 0.019859 | 0.014 | 0.011763 | 0.005 | 98 | 0.285277 | 0.001 | 0.237713 | 0.001 |
| 69 | 0.021729 | 0.014 | 0.012709 | 0.005 | 99 | 0.301298 | 0.001 | 0.256712 | 0.001 |
| 70 | 0.023730 | 0.015 | 0.013730 | 0.005 | 100 | 0.317238 | 0.001 | 0.276427 | 0.001 |
| 71 | 0.025951 | 0.015 | 0.014953 | 0.006 | 101 | 0.333461 | 0.000 | 0.296629 | 0.000 |
| 72 | 0.028481 | 0.015 | 0.016506 | 0.006 | 102 | 0.350330 | 0.000 | 0.317093 | 0.000 |
| 73 | 0.031201 | 0.015 | 0.018344 | 0.007 | 103 | $\downarrow 0.368542$ | 0.000 | 0.338505 | 0.000 |
| 74 | 0.034051 | 0.015 | 0.020381 | 0.007 | 104 | 0.387885 | 0.000 | 0.361016 | 0.000 |
| 75 | 0.037211 | 0.014 | 0.022686 | 0.008 | 105 | 0.407224 | 0.000 | 0.383597 | 0.000 |
| 76 | 0.040858 | 0.014 | 0.025325 | 0.008 | 106 | 0.425599 | 0.000 | 0.405217 | 0.000 |
| 77 | 0.045171 | 0.013 | 0.028366 | 0.007 | 107 | 0.441935 | 0.000 | 0.424846 | 0.000 |
| 78 | 0.050211 | 0.012 | 0.031727 | 0.007 | 108 | 0.457553 | 0.000 | 0.444368 | 0.000 |
| 79 | 0.055861 | 0.011 | 0.035362 | 0.007 | 109 | 0.473150 | 0.000 | 0.464469 | 0.000 |
| 80 | 0.062027 | 0.010 | 0.039396 | 0.007 | 110 | 0.486745 | 0.000 | 0.482325 | 0.000 |
| 81 | 0.068615 | 0.009 | 0.043952 | 0.007 | 111 | 0.496356 | 0.000 | 0.495110 | 0.000 |
| 82 | 0.075532 | 0.008 | 0.049153 | 0.007 | 112 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 83 | 0.082510 | 0.008 | 0.054857 | 0.007 | 113 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 84 | 0.089613 | 0.007 | 0.060979 | 0.007 | 114 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 85 | 0.097240 | 0.007 | 0.067738 | 0.006 | 115 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 86 | 0.105792 | 0.007 | 0.075347 | 0.005 | 116 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 87 | 0.115671 | 0.006 | 0.084023 | 0.004 | 117 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 88 | 0.126980 | 0.005 | 0.093820 | 0.004 | 118 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 89 | 0.139452 | 0.005 | 0.104594 | 0.003 | 119 | 0.500000 | 0.000 | 0.500000 | 0.000 |
| 90 | 0.152931 | 0.004 | 0.116265 | 0.003 | 120 | 1.000000 | 0.000 | 1.000000 | 0.000 |

$A A_{x}=$ the annual improvement factor in the mortality rate for age $x$.
To produce the mortality rate for a person age $x$ in year $(1994+n)$, the following formula would be used:

$$
q_{x}^{1994+n}=q_{x}^{1994}\left(1-A A_{x}\right)^{n}
$$

The application of generational mortality techniques to produce reserve values is described in this report.

## Standard Table Names

Several tables are presented in this report. To avoid confusion about what each of these tables represents, the following standard table names are used:

1. The 1994 Group Annuity Mortality Basic (or GAM-94 Basic) Table, which is presented as Table 13, is a static mortality table containing unloaded mortality rates for calendar year 1994.
2. The 1994 Group Annuity Mortality Static (or GAM-94 Static) Table, which is presented as Table 18, is a static mortality table containing loaded mortality rates for calendar year 1994.
3. Projection Scale AA (or Scale AA), which is presented as Table 15, represents the annual rates of mortality improvement by age for projecting future mortality rates beyond calendar year 1994.
4. The 1994 Group Annuity Reserving (or GAR-94) Table, which is presented as Table 1, is a combination of the GAM-94 Static Table and Projection Scale AA. Whenever reference is made to the use of this table, it implies that generational mortality derived from static mortality rates and projection scale factors has been used.

## I. INTRODUCTION

## A. Cbarge of The Task Force

The Group Annuity Valuation Table Task Force has been charged by the Society of Actuaries Board of Governors with developing a new Group Annuity Mortality Valuation Standard that would be suitable as a replacement for the current standard, which is based upon the 1983 Group Annuity Mortality Table (GAM-83).

## B. New Standard To Replace GAM-83

The Society of Actuaries committee that published the GAM-83 Table recommended that a new mortality table be developed when credible
annuitant experience became available, since the GAM-83 was only an update of prior data. The Task Force examined the annuitant experience from 1986 through 1990 and found that this was a sufficient basis for a new mortality table.

Further, that experience shows that mortality improvement has resulted in male actual-to-expected mortality ratios near 1.00, as shown in Table 2. Therefore, the margin included in the male rates in the GAM-83 no longer exists, and a new table with a sufficient margin is warranted.

TABLE 2
Retired Experience by Annuity Income Actual-to-Expected Mortality Ratios by Experience Year

|  | Experience Year |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | 1986 | 1987 | 1988 | 1989 | 1990 |
| Males | 1.05 | 1.08 | 1.06 | 1.03 | 1.01 |
| Fernales | 1.21 | 1.26 | 1.22 | 1.18 | 1.14 |

Finally, because the data, especially for female annuitants, are much more extensive than those used in the development of previous tables, the results produced in this report are more representative of current mortality.

For these reasons, the Task Force recommends that the new standard, as described in this report, be adopted as a replacement for the GAM-83.

## C. Intended Form of the New Standard

The Task Force strongly believes that the new standard should accomplish the goals of:

1. Recognizing mortality improvement
2. Serving for at least 15 years.

As shown in this report, while analyzing the data collected through 1990 and comparing them to GAM-83, the Task Force recognized that the trend in mortality improvement had not abated. Consequently, the Task Force decided that the observed mortality improvement trend should be explicitly recognized in this recommended new standard.

This decision to explicitly recognize mortality improvement was discussed in a 1992 position paper [1]. This position paper generated several very worthwhile suggestions and comments. Many of these suggestions were considered in the development of this recommended new standard.

The Task Force further believes that the new standard should be appropriately designed so that it will be useful for a reasonable time and not need as frequent an update as some of the more traditional standards.

To achieve these two results, the Task Force decided that a generational mortality approach, which is more fully discussed later in this report, would be appropriate. Note also that the great majority of input received by the Task Force in response to the position paper supported such a decision. Thus, the Task Force has proceeded with a recommendation that incorporates a generational approach as part of the new standard. This is the first time that projection scales are being recommended as suitable for a new standard for statutory reserving purposes.

The Task Force further recognizes that this approach departs from the traditional one of solely publishing a static table. Prior papers have published projection scales, but these projection scales were not recommended to be part of the statutory reserving standards. While the implementation of this approach is somewhat more complex than that of previous standards, modern systems capabilities facilitate implementation of this new standard. It is also intended that if and when the new standard is adopted for statutory reserving, insurers should be allowed sufficient time to incorporate this generational approach.

The various sections of this report discuss the development and application of this new standard. Note that additional report(s) will discuss how an adaptation of the new standard also serves as an update to the UP-84 Mortality Table and other related issues.

## II. DEVELOPMENT OF 1988 BASE YEAR GROUP ANNUITY MORTALITY TABLE

## A. 1988 Base Year Core Experience

Our objective was to develop a 1994 base year mortality table for males and females on an age-nearest-birthday basis based on credible group annuity mortality experience. The core mortality information for ages 66-95 was derived from group annuity mortality experience for retired lives for the 1986-1990 experience years. These data were obtained from the Society of Actuaries Group Annuity Experience Committee. In turn, their data were based upon the collective experience of annuitants in payment status for insured contracts from 11 large insurance companies. Data from contributors that were excluded in reports published by this committee were also
excluded from our data. The experience we used was examined for data integrity, and where clearly appropriate, data were excluded when determined to be erroneous.

All experience from the 1986-1990 group annuity mortality studies for the younger ages and for the very old ages were excluded from the experience committee data because of a lack of sufficient exposure at these ages. Mortality rates for these young and old ages were derived using the processes discussed later in this report.

Table 3 presents the crude mortality rates resulting from the income-based experience initially gathered by the committee for these ages. Table 3 forms the core of the initial 1988 base table prior to extensions for younger and older ages.

TABLE 3
Group Annuity Mortality Experience Unadjusted, Ungraduated, Before Margins

Years 1986-1990
1988 Base Year

|  | Values of $q_{x}$ |  |  | Values of $q_{x}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Male | Female | Age | Mate | Female |
| 66 | 0.019269 | 0.011659 | 81 | 0.083702 | 0.050633 |
| 67 | 0.020827 | 0.011558 | 82 | 0.087230 | 0.053618 |
| 68 | 0.021989 | 0.012648 | 83 | 0.100734 | 0.062886 |
| 69 | 0.025223 | 0.014816 | 84 | 0.108259 | 0.067163 |
| 70 | 0.027970 | 0.016470 | 85 | 0.109440 | 0.079880 |
|  |  |  |  |  |  |
| 71 | 0.030305 | 0.018468 | 86 | 0.118562 | 0.083499 |
| 72 | 0.034400 | 0.019646 | 87 | 0.137411 | 0.093969 |
| 73 | 0.037566 | 0.022562 | 88 | 0.151901 | 0.106342 |
| 74 | 0.041715 | 0.022690 | 89 | 0.156454 | 0.112547 |
| 75 | 0.045670 | 0.026181 | 90 | 0.161550 | 0.127477 |
|  |  |  |  |  |  |
| 76 | 0.049899 | 0.031442 | 91 | 0.199729 | 0.144480 |
| 77 | 0.055961 | 0.033878 | 92 | 0.194778 | 0.161609 |
| 78 | 0.060834 | 0.035267 | 93 | 0.234746 | 0.193206 |
| 79 | 0.066465 | 0.040115 | 94 | 0.232451 | 0.178502 |
| 80 | 0.072808 | 0.045878 | 95 | 0.267373 | 0.199738 |

## B. Ages 25-65

The first extension of Table 3 was for ages 25 through 65 . These mortality rates were derived from Civil Service Retirement System (CSRS) mortality experience by lives for the years 1985-1989 for retired annuitants and 1983-1986 (trending to 1985-1989) for active annuitants. Specifically,
experience for active annuitants was used to derive mortality rates for ages 25-50. A blend of experience for active and retired annuitants was used for ages 51-65, based on active/retired distributions of civil service annuitants as shown in Table 4.

TABLE 4
Assumed Active/Retired Split of Civil Service Annuitants
Used to Derive Experience Mortality for Ages 51-65

| Age | Mate Annuitants |  | Female Annuitants |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Active | Retired | Active | Retired |
| 51 | 0.96 | 0.04 | 0.98 | 0.02 |
| 52 | 0.95 | 0.05 | 0.97 | 0.03 |
| 53 | 0.93 | 0.07 | 0.96 | 0.04 |
| 54 | 0.92 | 0.08 | 0.95 | 0.05 |
| 55 | 0.84 | 0.16 | 0.93 | 0.07 |
|  |  |  |  |  |
| 56 | 0.68 | 0.32 | 0.85 | 0.15 |
| 57 | 0.63 | 0.37 | 0.82 | 0.18 |
| 58 | 0.57 | 0.43 | 0.78 | 0.22 |
| 59 | 0.53 | 0.47 | 0.74 | 0.26 |
| 60 | 0.45 | 0.55 | 0.66 | 0.34 |
|  |  |  |  |  |
| 61 | 0.37 | 0.63 | 0.53 | 0.47 |
| 62 | 0.29 | 0.71 | 0.42 | 0.58 |
| 63 | 0.21 | 0.79 | 0.29 | 0.71 |
| 64 | 0.17 | 0.83 | 0.22 | 0.78 |
| 65 | 0.13 | 0.87 | 0.17 | 0.83 |

Because the rates from CSRS based upon number of lives closely matched the rates from 1986-1990 group annuity mortality experience based upon annual income as shown in Table 5, the Task Force concluded that the CSRS experience was a reasonable basis for extension of the initial 1988 base table for ages below 66 .

TABLE 5
Comparison of Mortality Rates for Blended CSRS and 1986-1990 Group Annuity Experience

|  | Blended CSRS <br> Experience |  | Group Annuity <br> Experience |  |  | Ratios |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Female | Male | Female | Male | Female |  |
|  | 0.017188 | 0.009975 | 0.016831 | 0.009770 | 0.97923 | 0.97945 |  |
| 66 | 0.019160 | 0.010456 | 0.019269 | 0.011659 | 1.00569 | 1.11505 |  |
| 67 | 0.021456 | 0.012152 | 0.020827 | 0.011558 | 0.97068 | 0.95112 |  |
| 68 | 0.023483 | 0.012638 | 0.021989 | 0.012648 | 0.93638 | 1.00079 |  |
| 69 | 0.026761 | 0.014862 | 0.025223 | 0.014816 | 0.94253 | 0.99690 |  |
| 70 | 0.029621 | 0.017459 | 0.027970 | 0.016470 | 0.94426 | 0.94335 |  |

The blended mortality rates from the CSRS experience for ages 25-65 were then combined with the Table 3 group annuity experience for ages 66-95. An adjustment of the blended CSRS experience for ages 25-65 to reflect group annuity experience at age 65 was not needed, because the mortality rates for the blended CSRS experience were quite similar to the mortality rates for the group annuity experience at ages following age 64, as shown in Table 5.

Table 6 shows the crude mortality rates derived for ages 25 through 65 from the blended CSRS experience.

TABLE 6
Blended CSRS Experience
Unadiusted, Ungraduated, before Margins
Years 1985-1989

|  | Values of $q_{r}$ |  |  | Values of $q_{x}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Male | Female | Age | Male | Female |
| 25 | 0.000684 | 0.000365 | 46 | 0.002060 | 0.001202 |
| 26 | 0.000804 | 0.000280 | 47 | 0.002124 | 0.001232 |
| 27 | 0.000665 | 0.000369 | 48 | 0.002596 | 0.001387 |
| 28 | 0.000848 | 0.000324 | 49 | 0.002754 | 0.001763 |
| 29 | 0.000867 | 0.000375 | 50 | 0.003070 | 0.001540 |
| 30 | 0.000863 | 0.000414 | 51 | 0.003447 | 0.001766 |
| 31 | 0.000850 | 0.000411 | 52 | 0.003698 | 0.002068 |
| 32 | 0.000821 | 0.000381 | 53 | 0.004081 | 0.002153 |
| 33 | 0.000813 | 0.000438 | 54 | 0.004963 | 0.002313 |
| 34 | 0.000939 | 0.000555 | 55 | 0.004763 | 0.002522 |
|  |  |  |  |  |  |
| 35 | 0.001009 | 0.000539 | 56 | 0.005751 | 0.002669 |
| 36 | 0.000880 | 0.000585 | 57 | 0.007180 | 0.003222 |
| 37 | 0.000976 | 0.000620 | 58 | 0.007569 | 0.003703 |
| 38 | 0.000987 | 0.000568 | 59 | 0.008356 | 0.004186 |
| 39 | 0.001149 | 0.000810 | 60 | 0.009165 | 0.004759 |
| 40 | 0.001219 | 0.000701 | 61 | 0.010456 | 0.004990 |
| 41 | 0.001202 | 0.000991 | 62 | 0.011893 | 0.005865 |
| 42 | 0.001491 | 0.000861 | 63 | 0.013728 | 0.007110 |
| 43 | 0.001683 | 0.001265 | 64 | 0.015347 | 0.008633 |
| 44 | 0.001925 | 0.000993 | 65 | 0.017188 | 0.009975 |
| 45 | 0.001792 | 0.001065 |  |  |  |

## C. Extreme Ages (Ages 1-24 and 96-120)

Mortality rates for ages 1-24 and ages 96-120 were developed based on mortality rates from the Life Tables for calendar year 1990 and published in Actuarial Study No. 107 (SSA 107) [2]. U.S. Census statistics, information compiled by the National Center for Health Statistics and published in the
volumes of Vital Statistics of the United States, and Medicare data are the underlying data sources for SSA 107.

The Life Tables were combined with the group annuity experience and the blended CSRS experience as follows:

1. For ages 1-24, mortality rates from the SSA 107 Life Tables were used with modifications to the rates above age 12 . The mortality rates for ages 12-24 were obtained by adjusting the SSA 107 rates by a formula designed to replicate the SSA 107 age 12 rate and the age 25 rate from the blended CSRS experience. These values are shown in Table 7.

TABLE 7
SSA 107 Life Tables for 1990
Mortality Rates before and after adjustment to Group Annuity Experience Levels

Ages 1-25

|  | Before Adjustment |  | After Adjustment |  |
| :---: | :---: | :---: | :---: | :---: |
| Age | Male | Female | Male | Female |
| 1 | 0.000736 | 0.000647 | 0.000736 | 0.000647 |
| 2 | 0.000497 | 0.000422 | 0.000497 | 0.000422 |
| 3 | 0.000413 | 0.000315 | 0.000413 | 0.000315 |
| 4 | 0.000322 | 0.000236 | 0.000322 | 0.000236 |
| 5 | 0.000295 | 0.000213 | 0.000295 | 0.000213 |
|  |  |  |  |  |
| 6 | 0.000282 | 0.000199 | 0.000282 | 0.000199 |
| 7 | 0.000270 | 0.000187 | 0.000270 | 0.000187 |
| 8 | 0.000249 | 0.000173 | 0.000249 | 0.000173 |
| 9 | 0.000222 | 0.000159 | 0.000222 | 0.000159 |
| 10 | 0.000200 | 0.000148 | 0.000200 | 0.000148 |
|  |  |  |  |  |
| 11 | 0.000209 | 0.000149 | 0.000209 | 0.000149 |
| 12 | 0.000276 | 0.000172 | 0.000276 | 0.000172 |
| 13 | 0.000416 | 0.000221 | 0.000314 | 0.000194 |
| 14 | 0.000608 | 0.000289 | 0.000367 | 0.000226 |
| 15 | 0.000823 | 0.000368 | 0.000426 | 0.000262 |
|  |  |  |  |  |
| 16 | 0.001026 | 0.000441 | 0.000481 | 0.000295 |
| 17 | 0.001203 | 0.000495 | 0.000530 | 0.000320 |
| 18 | 0.001336 | 0.000520 | 0.000566 | 0.000331 |
| 19 | 0.001435 | 0.000524 | 0.000593 | 0.000333 |
| 20 | 0.001533 | 0.000524 | 0.000620 | 0.000333 |
|  |  |  |  |  |
| 21 | 0.001634 | 0.000530 | 0.000648 | 0.000336 |
| 22 | 0.001708 | 0.000539 | 0.000668 | 0.0003340 |
| 23 | 0.001747 | 0.000554 | 0.000679 | 0.0000347 |
| 24 | 0.001764 | 0.000574 | 0.000683 | 0.000356 |
| 25 | 0.001767 | 0.000594 | 0.000684 | 0.000365 |

2. For ages $96-119$, mortality rates from the SSA 107 Life Tables were appended to the experience table. The resulting mortality rates were then set at a maximum rate of 0.5 . No adjustment was required because the age 95 mortality rates in the experience table and the Life Tables were similar. These values are shown in Table 8.

TABLE 8
SSA Life Tables for 1990 Mortality Rates
(Modified above age 107)
Ages 96-119

| Age | Male | Female | Age | Male | Female |
| ---: | :---: | :---: | :---: | :---: | :---: |
| 96 | 0.278505 | 0.237204 | 111 | 0.500000 | 0.500000 |
| 97 | 0.294423 | 0.254388 | 112 | 0.500000 | 0.500000 |
| 98 | 0.310198 | 0.271234 | 113 | 0.500000 | 0.500000 |
| 99 | 0.325708 | 0.287508 | 114 | 0.500000 | 0.500000 |
| 100 | 0.341993 | 0.304758 | 115 | 0.500000 | 0.500000 |
|  |  |  |  |  |  |
| 101 | 0.359093 | 0.323044 | 116 | 0.500000 | 0.500000 |
| 102 | 0.377047 | 0.342426 | 117 | 0.500000 | 0.500000 |
| 103 | 0.395900 | 0.362972 | 118 | 0.500000 | 0.500000 |
| 104 | 0.415695 | 0.384750 | 119 | 0.500000 | 0.500000 |
| 105 | 0.436479 | 0.407835 |  |  |  |
|  |  |  |  |  |  |
| 106 | 0.458303 | 0.432305 |  |  |  |
| 107 | 0.481218 | 0.458243 |  |  |  |
| 108 | 0.500000 | 0.485738 |  |  |  |
| 109 | 0.500000 | 0.500000 |  |  |  |
| 110 | 0.500000 | 0.500000 |  |  |  |

Strong consideration was given to setting an ultimate value equal to 0.5 . Setting the highest mortality rate at a value of 0.5 instead of 1.0 would mean that there is no theoretical end to the mortality table. Such a proposed table would depart from past practice by not setting the mortality rate to 1.0 at some ultimate age. This change from tradition could be proposed for two reasons:

1. A number of studies have shown that the ultimate mortality rate peaks at a rate of less than 500 per 1,000 , so that a rate of 1.0 is not supported by the facts.
2. Current methods of constructing annuity tables do not require an ultimate value of 1.0 .
The mortality curve has long been known to bend upwards during the middle ages, and that is a feature of the proposed new standard table as well as all past tables. Studies of mortality at the very old ages have shown that
the mortality rate has a second bendpoint in the 80 s or 90 s , which reflects a deceleration in the rate of increase. The rate then proceeds to an approximately level ultimate rate after age 100. For example, Bayo and Faber [3] conducted a detailed study of the first OASDI beneficiaries who have now all died. They concluded that the mortality rates began to decelerate at about age 85. Lew and Garfinckel [4] found that the mortality rate first exceeded 0.33 in the late 90 s and fluctuated between 0.28 and 0.44 after that point.

The ungraduated group annuity experience is sparse after age 95 , but the data show the second bendpoint and the peaking of the rate of mortality. The male rates rise to about 0.25 in the mid-90s and then fluctuate around that point. The female rates also seem to peak at about 0.25 at those ages.

The use of such a mortality table without a final value could be implemented as follows:

1. Add an ultimate value to the annuity
2. Stop the table with a value of 1.0 at a certain age
3. Stop the table at a certain age but use 0.5 as the ultimate rate.

While the Task Force strongly believes an ultimate value of 0.5 is appropriate and could be properly programmed, there are some inconsistencies that could result without an assumed actual "end to the table." To avoid these inconsistent practical applications, the ultimate value is set equal to 1.0 at age 120.

Combining Tables 3, 6, 7, 8 and the ultimate rate of 1.0 at age 120 produces Table 9. This represents ungraduated mortality rates (adjusted for CSRS mortality for ages 25-65 and SSA 107 Life Tables for ages 1-24 and 96-119), as limited to a maximum rate of 0.5 , at all ages except the ultimate age of 120, assuming a base year of 1988. Note that Table 9 does not include any margins.

## III. PROJECTION SCALES DECISION-MAKING

The central calendar year of the modified mortality experience shown in Table 9 is 1988. The development of the new standard requires two projections of this 1988 base year mortality experience:

1. To project the mortality experience from the central experience year of 1988 to central year 1994, to produce a 1994 Basic Table
2. To develop the mortality projection scale used to project mortality into the future, after calendar year 1994, for the generational mortality table process.

TABLE 9
Mortality Experience Ungraduated before Margins
1988 Base YEAR

| Age | Values of $q_{\text {a }}$ |  | Age | Values of 9 . |  | Age | Values of $q$. |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Femate |  | Male | Female |  | Male | Female |
| 1 | 0.000736 | 0.000647 | 21 | 0.000648 | 0.000336 | 41 | 0.001202 | 0.000991 |
| 2 | 0.000497 | 0.000422 | 22 | 0.000668 | 0.000340 | 42 | 0.001491 | 0.000861 |
| 3 | 0.000413 | 0.000315 | 23 | 0.000679 | 0.000347 | 43 | 0.001683 | 0.001265 |
| 4 | 0.000322 | 0.000236 | 24 | 0.000683 | 0.000356 | 44 | 0.001925 | 0.000993 |
| 5 | 0.000295 | 0.000213 | 25 | 0.000684 | 0.000365 | 45 | 0.001792 | 0.001065 |
| 6 | 0.000282 | 0.000199 | 26 | 0.000804 | 0.000280 | 46 | 0.002060 | 0.001202 |
| 7 | 0.000270 | 0.000187 | 27 | 0.000665 | 0.000369 | 47 | 0.002124 | 0.001232 |
| 8 | 0.000249 | 0.000173 | 28 | 0.000848 | 0.000324 | 48 | 0.002596 | 0.001387 |
| 9 | 0.000222 | 0.000159 | 29 | 0.000867 | 0.000375 | 49 | 0.002754 | 0.001763 |
| 10 | 0.000200 | 0.000148 | 30 | 0.000863 | 0.000414 | 50 | 0.003070 | 0.001540 |
| 11 | 0.000209 | 0.000149 | 31 | 0.000850 | 0.000411 | 51 | 0.003447 | 0.001766 |
| 12 | 0.000276 | 0.000172 | 32 | 0.000821 | 0.000381 | 52 | 0.003698 | 0:002068 |
| 13 | 0.000314 | 0.000194 | 33 | 0.000813 | 0.000438 | 53 | 0.004081 | 0.002153 |
| 14 | 0.000367 | 0.000226 | 34 | 0.000939 | 0.000555 | 54 | 0.004963 | 0.002313 |
| 15 | 0.000426 | 0.000262 | 35 | 0.001009 | 0.000539 | 55 | 0.004763 | 0.002522 |
| 16 | 0.000481 | 0.000295 | 36 | 0.000880 | 0.000585 | 56 | 0.005751 | 0.002669 |
| 17 | 0.000530 | 0.000320 | 37 | 0.000976 | 0.000620 | 57 | 0.007180 | 0.003222 |
| 18 | 0.000566 | 0.000331 | 38 | 0.000987 | 0.000568 | 58 | 0.007569 | 0.003703 |
| 19 | 0.000593 | 0.000333 | 39 | 0.001149 | 0.000810 | 59 | 0.008356 | 0.004186 |
| 20 | 0.000620 | 0.000333 | 40 | 0.001219 | 0.000701 | 60 | 0.009165 | 0.004759 |

TABLE 9-Continued

| Age | Values of $q_{\text {. }}$ |  | Age | Values of $q$, |  | Age | Values of $q_{\text {r }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Female |  | Male | Female |  | Male | Female |
| 61 | 0.010456 | 0.004990 | 81 | 0.083702 | 0.050633 | 101 | 0.359093 | 0.323044 |
| 62 | 0.011893 | 0.005865 | 82 | 0.087230 | 0.053618 | 102 | 0.377047 | 0.342426 |
| 63 | 0.013728 | 0.007110 | 83 | 0.100734 | 0.062886 | 103 | 0.395900 | 0.362972 |
| 64 | 0.015347 | 0.008633 | 84 | 0.108259 | 0.067163 | 104 | 0.415695 | 0.384750 |
| 65 | 0.017188 | 0.009975 | 85 | 0.109440 | 0.079880 | 105 | 0.436479 | 0.407835 |
| 66 | 0.019269 | 0.011659 | 86 | 0.118562 | 0.083499 | 106 | 0.458303 | 0.432305 |
| 67 | 0.020827 | 0.011558 | 87 | 0.137411 | 0.093969 | 107 | 0.481218 | 0.458243 |
| 68 | 0.021989 | 0.012648 | 88 | 0.151901 | 0.106342 | 108 | 0.500000 | 0.485738 |
| 69 | 0.025223 | 0.014816 | 89 | 0.156454 | 0.112547 | 109 | 0.500000 | 0.500000 |
| 70 | 0.027970 | 0.016470 | 90 | 0.161550 | 0.127477 | 110 | 0.500000 | 0.500000 |
| 71 | 0.030305 | 0.018468 | 91 | 0.199729 | 0.144480 | 111 | 0.500000 | 0.500000 |
| 72 | 0.034400 | 0.019646 | 92 | 0.194778 | 0.161609 | 112 | 0.500000 | 0.500000 |
| 73 | 0.037566 | 0.022562 | 93 | 0.234746 | 0.193206 | 113 | 0.500000 | 0.500000 |
| 74 | 0.041715 | 0.022690 | 94 | 0.232451 | 0.178502 | 114 | 0.500000 | 0.500000 |
| 75 | 0.045670 | 0.026181 | 95 | 0.267373 | 0.199738 | 115 | 0.500000 | 0.500000 |
| 76 | 0.049899 | 0.031442 | 96 | 0.278505 | 0.237204 | 116 | 0.500000 | 0.500000 |
| 77 | 0.055961 | 0.033878 | 97 | 0.294423 | 0.254388 | 117 | 0.500000 | 0.500000 |
| 78 | 0.060834 | 0.035267 | 98 | 0.310198 | 0.271234 | 118 | 0.500000 | 0.500000 |
| 79 | 0.066465 | 0.040115 | 99 | 0.325708 | 0.287508 | 119 | 0.500000 | 0.500000 |
| 80 | 0.072808 | 0.045878 | 100 | 0.341993 | 0.304758 | 120 | 1.000000 | 1.000000 |

## A. Projection of Mortality Rates to 1994

For the 1988-1994 projection of mortality reduction, the Task Force considered mortality improvements from the following sources:

1. Projections of mortality improvement in the general population presented in SSA 107, with further detail covering the periods 1988-1994 and 1986-1992, from the 1992 Trustees Report Intermediate Alternative II Assumptions, which are consistent with SSA 107
2. CSRS mortality improvement experience
3. Scale H, which was presented with the development of the GAM-83.
4. The Society of Actuaries Group Annuity Mortality Study covering the period 1985-1990.
Comparisons of mortality improvement rates at quinquennial age groups from these studies appear in Table 10.

After much discussion, including interaction with the UP-94 Table Task Force, the Group Annuity Valuation Table Task Force concluded that the CSRS data would provide the most meaningful projection, because they were produced from a large database and also used directly to extend the mortality table for active lives.

This conclusion was arrived at after examining the SSA 107 experience and the age-by-age trends of the CSRS experience without modification. The SSA 107 experience did not include actual experience past calendar year 1988, whereas the CSRS data included experience through 1993. The CSRS data would therefore provide the better projection for all ages, even though some slight modification and smoothing were required.
The scale of mortality improvement factors for projecting the mortality rates shown in Table 9 from 1988 to 1994 was based on the average trends for CSRS over the period 1987-1993. A mortality table based on CSRS experience was constructed for each year over this period and graduated by using a Whittaker-Henderson type B method. Then a mortality improvement rate for each age was determined based on a least-squares best fit trend line through logarithms of the death rates for that age. The resulting scale of mortality improvement trends for each age was then itself graduated using the same method and rounded to the nearest one-tenth of one percentage point. However, the trends for females at ages $60-65$ were changed from negative to zero, because the group annuity trend experience for these ages was slightly positive. A trend of $2 \%$ was used at the younger ages.

Table 11 shows the final mortality improvement factors compared to the actual CSRS 1987-1993 trends.

TABLE 10
Annual Mortality Improvement Rates from Various Studies Rates in Percentage per Year

|  |  |  |  |  |  |  |  |
| :---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: |
| Ages | SSA <br> $88-94$ | SSA <br> $86-92$ | CSRS <br> $86-92$ | Scale H | SOA <br> $85-90$ | CSRS <br> Nondisability |  |
| Male Lives |  |  |  |  |  |  |  |
| $25-29$ | -2.13 | -1.67 | 0.97 | 0.10 |  |  |  |
| $30-34$ | -3.34 | -3.04 | -1.24 | 0.75 |  |  |  |
| $35-39$ | -2.98 | -3.51 | -0.69 | 2.00 |  |  |  |
| $40-44$ | -2.21 | -1.39 | -0.68 | 2.00 |  |  |  |
| $45-49$ | 0.64 | 0.62 | 1.35 | 1.75 |  |  |  |
|  |  |  |  |  |  |  |  |
| $50-54$ | 0.91 | 1.25 | 2.19 | 1.75 |  |  |  |
| $55-59$ | 1.41 | 1.40 | 2.53 | 1.50 | 3.90 | 2.70 |  |
| $60-64$ | 1.60 | 1.57 | 1.78 | 1.50 | 2.30 | 1.77 |  |
| $65-69$ | 1.52 | 1.08 | 1.29 | 1.50 | 3.00 | 1.23 |  |
| $70-74$ | 1.03 | 1.24 | 1.90 | 1.25 | 3.40 | 1.88 |  |
|  |  |  |  |  |  |  |  |
| $75-79$ | 0.79 | 0.89 | 2.18 | 1.25 | 2.90 | 2.18 |  |
| $80-84$ | 0.68 | 0.57 | 1.70 | 1.25 | 1.00 | 1.69 |  |
| $85-89$ | 0.73 | 0.22 | 1.04 | 0.75 | 0.70 | 1.08 |  |
| $90-94$ | 1.00 | -0.23 | 0.16 | 0.75 | 0.30 | 0.17 |  |


| Female Lives |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| $25-29$ | -0.05 | -0.73 | 3.40 | 0.75 |  |  |
| $30-34$ | -1.68 | -1.47 | 0.21 | 1.25 |  |  |
| $35-39$ | 0.00 | -0.68 | 2.13 | 2.25 |  |  |
| $40-44$ | 1.29 | 1.78 | 3.00 | 2.25 |  |  |
| $45-49$ | 1.81 | 1.77 | 0.97 | 2.00 |  |  |
|  |  |  |  |  |  |  |
| $50-54$ | 1.39 | 1.50 | 1.66 | 2.00 |  |  |
| $55-59$ | 1.20 | 0.65 | -0.26 | 1.75 | 0.60 | 3.06 |
| $60-64$ | 0.69 | 0.74 | -0.07 | 1.75 | 1.80 | -0.44 |
| $65-69$ | 0.45 | 0.35 | 0.50 | 1.75 | 2.60 | 0.54 |
| $70-74$ | 0.54 | 0.64 | 1.20 | 1.75 | 3.90 | 1.22 |
|  |  |  |  |  |  |  |
| $75-79$ | 0.87 | 0.74 | 1.20 | 1.50 | 2.50 | 1.19 |
| $80-84$ | 1.22 | 0.86 | 1.16 | 1.50 | 0.70 | 1.17 |
| $85-89$ | 1.19 | 0.66 | 1.15 | 1.00 | 2.00 | 1.14 |
| $90-94$ | 0.93 | 0.18 | 0.85 | 0.50 | 2.10 | 0.81 |

Table 12 shows the ungraduated base year 1994 mortality table rates before margins. The rates in Table 12 were obtained by taking the 1988 base year mortality rates from Table 9 and projecting them to 1994 using the GAM 88-94 mortality improvement factors in Table 11. The following formula was used to project the mortality rates:

$$
\begin{equation*}
q_{x}^{1994}=q_{x}^{1988} \times\left(1-\text { scale }_{x}\right)^{(1994-1988)} \tag{A}
\end{equation*}
$$

TABLE 11
annual Mortality Improvement Factors for Use in Projecting Mortality Rates from 1988 to 1994 (GAM 88-94 Column)

Rates in Percentage per Year

| Age | Male |  | Female |  |  | Mate |  | Female |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & \text { CSRS } \\ & 87-93 \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { GAM } \\ & \text { B8-94 } \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { CSRS } \\ & 87-93 \\ & \hline \end{aligned}$ | $\begin{array}{r} \text { GAM } \\ 88-94 \\ \hline \end{array}$ | Age | $\begin{aligned} & \text { CSRS } \\ & 87-93 \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { GAM } \\ & 88-94 \end{aligned}$ | $\begin{aligned} & \text { CSRS } \\ & 87-93 \\ & \hline \end{aligned}$ | $\begin{array}{r} \text { GAM } \\ 88-94 \\ \hline \end{array}$ |
| 1 | 2.0 | 2.0 | 2.0 | 2.0 | 31 | -1.5 | -1.2 | 0.8 | 0.5 |
| 2 | 2.0 | 2.0 | 2.0 | 2.0 | 32 | -1.3 | -1.0 | 0.9 | 0.5 |
| 3 | 2.0 | 2.0 | 2.0 | 2.0 | 33 | -0.8 | -0.7 | 0.6 | 0.5 |
| 4 | 2.0 | 2.0 | 2.0 | 2.0 | 34 | -0.1 | -0.2 | 0.3 | 0.5 |
| 5 | 2.0 | 2.0 | 2.0 | 2.0 | 35 | 0.6 | 0.1 | 0.1 | 0.6 |
| 6 | 2.0 | 2.0 | 2.0 | 2.0 | 36 | 1.0 | 0.5 | 0.4 | 0.7 |
| 7 | 2.0 | 2.0 | 2.0 | 2.0 | 37 | 0.9 | 0.7 | 0.9 | 0.8 |
| 8 | 2.0 | 2.0 | 2.0 | 2.0 | 38 | 0.7 | 1.0 | 1.3 | 1.0 |
| 9 | 2.0 | 2.0 | 2.0 | 2.0 | 39 | 0.7 | 1.2 | 1.4 | 1.1 |
| 10 | 2.0 | 2.0 | 2.0 | 2.0 | 40 | 1.0 | 1.4 | 1.2 | 1.1 |
| 11 | 2.0 | 2.0 | 2.0 | 2.0 | 41 | 1.6 | 1.7 | 1.0 | 1.2 |
| 12 | 2.0 | 2.0 | 2.0 | 1.9 | 42 | 2.1 | 1.9 | 1.0 | 1.3 |
| 13 | 2.0 | 1.9 | 2.0 | 1.8 | 43 | 2.3 | 2.0 | 1.2 | 1.4 |
| 14 | 1.9 | 1.9 | 1.8 | 1.7 | 44 | 2.4 | 2.0 | 1.5 | 1.6 |
| 15 | 1.9 | 1.9 | 1.6 | 1.6 | 45 | 2.2 | 2.0 | 1.8 | 1.8 |
| 16 | 1.9 | 1.9 | 1.5 | 1.5 | 46 | 1.8 | 1.8 | 2.2 | 1.9 |
| 17 | 1.9 | 1.9 | 1.4 | 1.5 | 47 | 1.4 | 1.7 | 2.3 | 1.9 |
| 18 | 1.9 | 1.9 | 1.4 | 1.5 | 48 | 1.2 | 1.6 | 2.1 | 1.8 |
| 19 | 1.9 | 1.9 | 1.5 | 1.5 | 49 | 1.3 | 1.6 | 1.6 | 1.5 |
| 20 | 1.9 | 1.9 | 1.6 | 1.6 | 50 | 1.5 | 1.6 | 1.0 | 1.2 |
| 21 | 1.8 | 1.8 | 1.7 | 1.6 | 51 | 1.9 | 1.7 | 0.7 | 1.0 |
| 22 | 1.7 | 1.7 | 1.7 | 1.6 | 52 | 2.2 | 1.8 | 0.6 | 0.8 |
| 23 | 1.5 | 1.5 | 1.6 | 1.6 | 53 | 2.2 | 1.9 | 0.6 | 0.8 |
| 24 | 1.3 | 1.2 | 1.5 | 1.5 | 54 | 2.0 | 1.9 | 0.8 | 0.8 |
| 25 | 1.0 | 0.8 | 1.4 | 1.3 | 55 | 1.7 | 1.9 | 0.9 | 0.9 |
| 26 | 0.6 | 0.3 | 1.2 | 1.1 | 56 | 1.6 | 1.9 | 1.0 | 0.9 |
| 27 | -0.4 | -0.1 | 0.9 | 0.9 | 57 | 1.7 | 1.8 | 1.0 | 0.8 |
| 28 | -0.6 | -0.6 | 0.5 | 0.7 | 58 | 1.8 | 1.9 | 0.9 | 0.6 |
| 29 | -0.9 | $-1.0$ | 0.3 | 0.6 | 59 | 1.9 | 1.9 | 0.4 | 0.2 |
| 30 | -1.3 | -1.2 | 0.4 | 0.5 | 60 | 2.0 | 1.9 | -0.1 | 0.0 |

TABLE 11-Continued

| Age | Malc |  | Female |  |  | Male |  | Femate |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & \text { CSRS } \\ & 87-93 \end{aligned}$ | $\begin{aligned} & \text { GAM } \\ & 88-94 \end{aligned}$ | $\begin{aligned} & \hline \text { CSRS } \\ & 87-93 \end{aligned}$ | $\begin{gathered} \text { GAM } \\ 88-94 \end{gathered}$ | Age | $\begin{aligned} & \hline \text { CSRS } \\ & 87-93 \end{aligned}$ | $\begin{aligned} & \text { GAM } \\ & 88-94 \end{aligned}$ | $\begin{aligned} & \hline \text { CSRS } \\ & 87-93 \end{aligned}$ | $\begin{aligned} & \text { GAM } \\ & 88-94 \end{aligned}$ |
| 61 | 1.9 | 1.9 | -0.6 | 0.0 | 91 | 0.9 | 0.9 | 0.8 | 0.8 |
| 62 | 1.9 | 1.8 | -0.8 | 0.0 | 92 | 0.7 | 0.7 | 0.8 | 0.9 |
| 63 | 1.8 | 1.7 | -0.8 | 0.0 | 93 | 0.6 | 0.6 | 0.9 | 0.9 |
| 64 | 1.6 | 1.5 | -0.6 | 0.0 | 94 | 0.5 | 0.5 | 1.0 | 1.0 |
| 65 | 1.3 | 1.3 | -0.2 | 0.0 | 95 | 0.4 | 0.5 | 1.1 | 1.1 |
| 66 | 1.1 | 1.2 | 0.2 | 0.2 | 96 | 0.4 | 0.4 | 1.1 | 1.1 |
| 67 | 1.0 | 1.2 | 0.8 | 0.7 | 97 | 0.4 | 0.4 | 1.1 | 1.0 |
| 68 | 1.0 | 1.2 | 1.3 | 1.2 | 98 | 0.4 | 0.3 | 1.1 | 0.9 |
| 69 | 1.3 | 1.3 | 1.7 | 1.5 | 99 | 0.5 | 0.2 | 1.0 | 0.7 |
| 70 | 1.6 | 1.5 | 1.9 | 1.8 | 100 | 0.0 | 0.2 | 0.0 | 0.4 |
| 71 | 1.9 | 1.7 | 2.0 | 2.0 | 101 | 0.0 | 0.1 | 0.0 | 0.2 |
| 72 | 2.0 | 1.9 | 2.0 | 2.0 | 102 | 0.0 | 0.0 | 0.0 | 0.1 |
| 73 | 2.0 | 2.1 | 1.9 | 2.0 | 103 | 0.0 | 0.0 | 0.0 | 0.0 |
| 74 | 2.1 | 2.2 | 1.8 | 1.9 | 104 | 0.0 | 0.0 | 0.0 | 0.0 |
| 75 | 2.2 | 2.3 | 1.8 | 1.8 | 105 | 0.0 | 0.0 | 0.0 | 0.0 |
| 76 | 2.3 | 2.3 | 1.7 | 1.7 | 106 | 0.0 | 0.0 | 0.0 | 0.0 |
| 77 | 2.4 | 2.3 | 1.6 | 1.5 | 107 | 0.0 | 0.0 | 0.0 | 0.0 |
| 78 | 2.4 | 2.3 | 1.4 | 1.4 | 108 | 0.0 | 0.0 | 0.0 | 0.0 |
| 79 | 2.3 | 2.2 | 1.2 | 1.2 | 109 | 0.0 | 0.0 | 0.0 | 0.0 |
| 80 | 2.2 | 2.1 | 1.1 | 1.1 | 110 | 0.0 | 0.0 | 0.0 | 0.0 |
| 81 | 1.9 | 1.9 | 1.0 | 1.0 | 111 | 0.0 | 0.0 | 0.0 | 0.0 |
| 82 | 1.7 | 1.8 | 0.9 | 1.0 | 112 | 0.0 | 0.0 | 0.0 | 0.0 |
| 83 | 1.6 | 1.7 | 0.9 | 0.9 | 113 | 0.0 | 0.0 | 0.0 | 0.0 |
| 84 | 1.5 | 1.6 | 1.0 | 0.9 | 114 | 0.0 | 0.0 | 0.0 | 0.0 |
| 85 | 1.4 | 1.5 | 1.0 | 0.9 | 115 | 0.0 | 0.0 | 0.0 | 0.0 |
| 86 | 1.4 | 1.4 | 1.0 | 0.9 | 116 | 0.0 | 0.0 | 0.0 | 0.0 |
| 87 | 1.4 | 1.3 | 1.0 | 0.9 | 117 | 0.0 | 0.0 | 0.0 | 0.0 |
| 88 | 1.3 | 1.3 | 0.9 | 0.9 | 118 | 0.0 | 0.0 | 0.0 | 0.0 |
| 89 | 1.2 | 1.2 | 0.8 | 0.8 | 119 | 0.0 | 0.0 | 0.0 | 0.0 |
| 90 | 1.1 | 1.0 | 0.8 | 0.8 | 120 | 0.0 | 0.0 | 0.0 | 0.0 |

TABLE 12
Group annuity Mortality Rates Ungraduated-No Margin
1994 Base Year

| Age | Male | Female | Age | Male | Female | Age | Male | Female |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.000652 | 0.000573 | 21 | 0.000581 | 0.000305 | 41 | 0.001084 | 0.000922 |
| 2 | 0.000440 | 0.000374 | 22 | 0.000603 | 0.000309 | 42 | 0.001329 | 0.000796 |
| 3 | 0.000366 | 0.000279 | 23 | 0.000620 | 0.000315 | 43 | 0.001491 | 0.001162 |
| 4 | 0.000285 | 0.000209 | 24 | 0.000635 | 0.000325 | 44 | 0.001705 | 0.000901 |
| 5 | 0.000261 | 0.000189 | 25 | 0.000652 | 0.000337 | 45 | 0.001587 | 0.000955 |
| 6 | 0.000250 | 0.000176 | 26 | 0.000790 | 0.000262 | 46 | 0.001847 | 0.001071 |
| 7 | 0.000239 | 0.000166 | 27 | 0.000669 | 0.000350 | 47 | 0.001916 | 0.001098 |
| 8 | 0.000221 | 0.000153 | 28 | 0.000879 | 0.000311 | 48 | 0.002357 | 0.001244 |
| 9 | 0.000197 | 0.000141 | 29 | 0.000920 | 0.000362 | 49 | 0.002500 | 0.001610 |
| 10 | 0.000177 | 0.000131 | 30 | 0.000927 | 0.000402 | 50 | 0.002787 | 0.001432 |
| 11 | 0.000185 | 0.000132 | 31 | 0.000913 | 0.000399 | 51 | 0.003110 | 0.001663 |
| 12 | 0.000244 | 0.000153 | 32 | 0.000872 | 0.000370 | 52 | 0.003316 | 0.001971 |
| 13 | 0.000280 | 0.000174 | 33 | 0.000848 | 0.000425 | 53 | 0.003637 | 0.002052 |
| 14 | 0.000327 | 0.000204 | 34 | 0.000950 | 0.000539 | 54 | 0.004423 | 0.002204 |
| 15 | 0.000380 | 0.000238 | 35 | 0.001003 | 0.000520 | 55 | 0.004245 | 0.002389 |
| 16 | 0.000429 | 0.000269 | 36 | 0.000854 | 0.000561 | 56 | 0.005126 | 0.002528 |
| 17 | 0.000472 | 0.000292 | 37 | 0.000936 | 0.000591 | 57 | 0.006439 | 0.003070 |
| 18 | 0.000504 | 0.000302 | 38 | 0.000929 | 0.000535 | 58 | 0.006746 | 0.003572 |
| 19 | 0.000529 | 0.000304 | 39 | 0.001069 | 0.000758 | 59 | 0.007448 | 0.004136 |
| 20 | 0.000553 | 0.000302 | 40 | 0.001120 | 0.000656 | 60 | 0.008169 | 0.004759 |

TABLE 12-Continued

| Age | Male | Female | Age | Male | Female | Age | Male | Female |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 61 | 0.009319 | 0.004900 | 81 | 0.074602 | 0.047670 | 101 | 0.356944 | 0.319187 |
| 62 | 0.010665 | 0.005865 | 82 | 0.078223 | 0.050480 | 102 | 0.377047 | 0.340376 |
| 63 | 0.012386 | 0.007110 | 83 | 0.090886 | 0.059566 | 103 | 0.395900 | 0.362972 |
| 64 | 0.014017 | 0.008633 | 84 | 0.098273 | 0.063617 | 104 | 0.415695 | 0.384750 |
| 65 | 0.015890 | 0.009975 | 85 | 0.099952 | 0.075662 | 105 | 0.436479 | 0.407835 |
| 66 | 0.017923 | 0.011520 | 86 | 0.108945 | 0.079090 | 106 | 0.458303 | 0.432305 |
| 67 | 0.019372 | 0.011081 | 87 | 0.127035 | 0.089007 | 107 | 0.481218 | 0.458243 |
| 68 | 0.020453 | 0.011764 | 88 | 0.140431 | 0.100727 | 108 | 0.500000 | 0.485738 |
| 69 | 0.023318 | 0.013532 | 89 | 0.145522 | 0.107252 | 109 | 0.500000 | 0.500000 |
| 70 | 0.025545 | 0.014769 | 90 | 0.152096 | 0.121479 | 110 | 0.500000 | 0.500000 |
| 71 | 0.027342 | 0.016360 | 91 | 0.189183 | 0.137682 | 111 | 0.500000 | 0.500000 |
| 72 | 0.030660 | 0.017403 | 92 | 0.186739 | 0.153076 | 112 | 0.500000 | 0.500000 |
| 73 | 0.033074 | 0.019986 | 93 | 0.226421 | 0.183005 | 113 | 0.500000 | 0.500000 |
| 74 | 0.036503 | 0.020223 | 94 | 0.225564 | 0.168056 | 114 | 0.500000 | 0.500000 |
| 75 | 0.039719 | 0.023478 | 95 | 0.259451 | 0.186913 | 115 | 0.500000 | 0.500000 |
| 76 | 0.043397 | 0.028368 | 96 | 0.271887 | 0.221973 | 116 | 0.500000 | 0.500000 |
| 77 | 0.048669 | 0.030941 | 97 | 0.287427 | 0.239501 | 117 | 0.500000 | 0.500000 |
| 78 | 0.052907 | 0.032406 | 98 | 0.304656 | 0.256913 | 118 | 0.500000 | 0.500000 |
| 79 | 0.058160 | 0.037312 | 99 | 0.321819 | 0.275642 | 119 | 0.500000 | 0.500000 |
| 80 | 0.064102 | 0.042932 | 100 | 0.337910 | 0.297517 | 120 | 1.000000 | 1.000000 |

where
$q_{x}^{y}=$ mortality rate in calendar year $y$ at attained age $x$
scale $_{x}=$ mortality improvement factor for attained age $x$.
The resulting rates are an ungraduated set of mortality rates for ages $1-120$, by sex, with a base experience year of 1994.

## B. Graduated Mortality Rates

The resulting set of mortality rates in Table 12 was then graduated by using the Karup-King four point graduation formula, as follows. Mortality rates were averaged by quinquennial age groups $q_{n}, q_{n}, q_{n}, q_{n}, \ldots$. Graduated mortality rates $q_{n}$, were derived based on the following formula:

$$
\begin{equation*}
q_{n+t}=A_{1} \times q_{n-5}+A_{2} \times q_{n}+A_{3} \times q_{n+5}+A_{4} \times q_{n+10} \tag{B}
\end{equation*}
$$

where

$$
\begin{aligned}
& A_{1}=-0.5 \times S_{1} \times\left(1-S_{1}\right)^{2} \\
& A_{2}=1.5 \times S_{1}^{3}-2.5 \times S_{1}^{2}+1 \\
& A_{3}=-1.5 \times S_{1}^{3}+2 \times S_{1}^{2}+0.5 \times S_{1} \\
& A_{4}=0.5 \times S_{1}^{2} \times\left(S_{1}-1\right) \\
& S_{1}=t / 5
\end{aligned}
$$

At the extreme ages (under age 7 and over age 102), minor adjustments were made.

The adjusted mortality rates with a base year of 1994 of Table 12 after graduation are shown in Table 13. Table 13 is the 1994 Group Annuity Mortality Basic (GAM-94 Basic) Table.

## C. Projection of Mortality Rates beyond 1994

For the projection of mortality reduction beyond 1994, the Task Force decided to use a blend of the CSRS and SSA 107 mortality reduction trends based upon experience between years 1977 through 1993, with adjustments. A mortality improvement scale based entirely on CSRS data over the period 1977-1993 was constructed. The starting point was a mortality table for each year 1977 through 1993, graduated by Whittaker-Henderson type B. Then a mortality improvement trend for each age was determined based on

TABLE 13
1994 Group Annuity Mortality Table Graduated-No Margin

1994 Base Year

| Age | Male | Female | Age | Mate | Female | Age | Male | Female |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.000637 | 0.000571 | 21 | 0.000570 | 0.000308 | 41 | 0.001243 | 0.000826 |
| 2 | 0.000430 | 0.000372 | 22 | 0.000598 | 0.000311 | 42 | 0.001346 | 0.000888 |
| 3 | 0.000357 | 0.000278 | 23 | 0.000633 | 0.000313 | 43 | 0.001454 | 0.000943 |
| 4 | 0.000278 | 0.000208 | 24 | 0.000671 | 0.000313 | 44 | 0.001568 | 0.000992 |
| 5 | 0.000255 | 0.000188 | 25 | 0.000711 | 0.000313 | 45 | 0.001697 | 0.001046 |
| 6 | 0.000244 | 0.000176 | 26 | 0.000749 | 0.000316 | 46 | 0.001852 | 0.001111 |
| 7 | 0.000234 | 0.000165 | 27 | 0.000782 | 0.000324 | 47 | 0.002042 | 0.001196 |
| 8 | 0.000216 | 0.000147 | 28 | 0.000811 | 0.000338 | 48 | 0.002260 | 0.001297 |
| 9 | 0.000209 | 0.000140 | 29 | 0.000838 | 0.000356 | 49 | 0.002501 | 0.001408 |
| 10 | 0.000212 | 0.000141 | 30 | 0.000862 | 0.000377 | 50 | 0.002773 | 0.001536 |
| 11 | 0.000223 | 0.000148 | 31 | 0.000883 | 0.000401 | 51 | 0.003088 | 0.001686 |
| 12 | 0.000243 | 0.000159 | 32 | 0.000902 | 0.000427 | 52 | 0.003455 | 0.001864 |
| 13 | 0.000275 | 0.000177 | 33 | 0.000912 | 0.000454 | 53 | 0.003854 | 0.002051 |
| 14 | 0.000320 | 0.000203 | 34 | 0.000913 | 0.000482 | 54 | 0.004278 | 0.002241 |
| 15 | 0.000371 | 0.000233 | 35 | 0.000915 | 0.000514 | 55 | 0.004758 | 0.002466 |
| 16 | 0.000421 | 0.000261 | 36 | 0.000927 | 0.000550 | 56 | 0.005322 | 0.002755 |
| 17 | 0.000463 | 0.000281 | 37 | 0.000958 | 0.000593 | 57 | 0.006001 | 0.003139 |
| 18 | 0.000495 | 0.000293 | 38 | 0.001010 | 0.000643 | 58 | 0.006774 | 0.003612 |
| 19 | 0.000521 | 0.000301 | 39 | 0.001075 | 0.000701 | 59 | 0.007623 | 0.004154 |
| 20 | 0.000545 | 0.000305 | 40 | 0.001153 | 0.000763 | 60 | 0.008576 | 0.004773 |

TABLE 13-Continued

| Age | Male | Female | Age | Male | Female | Age | Male | Female |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 61 | 0.009663 | 0.005476 | 81 | 0.073780 | 0.047260 | 101 | 0.358560 | 0.318956 |
| 62 | 0.010911 | 0.006271 | 82 | 0.081217 | 0.052853 | 102 | 0.376699 | 0.340960 |
| 63 | 0.012335 | 0.007179 | 83 | 0.088721 | 0.058986 | 103 | 0.396884 | 0.364586 |
| 64 | 0.013914 | 0.008194 | 84 | 0.096358 | 0.065569 | 104 | 0.418855 | 0.389996 |
| 65 | 0.015629 | 0.009286 | 85 | 0.104559 | 0.072836 | 105 | 0.440585 | 0.415180 |
| 66 | 0.017462 | 0.010423 | 86 | 0.113755 | 0.081018 | 106 | 0.460043 | 0.438126 |
| 67 | 0.019391 | 0.011574 | 87 | 0.124377 | 0.090348 | 107 | 0.475200 | 0.456824 |
| 68 | 0.021354 | 0.012648 | 88 | 0.136537 | 0.100882 | 108 | 0.485670 | 0.471493 |
| 69 | 0.023364 | 0.013665 | 89 | 0.149949 | 0.112467 | 109 | 0.492807 | 0.483473 |
| 70 | 0.025516 | 0.014763 | 90 | 0.164442 | 0.125016 | 110 | 0.497189 | 0.492436 |
| 71 | 0.027905 | 0.016079 | 91 | 0.179849 | 0.138442 | 111 | 0.499394 | 0.498054 |
| 72 | 0.030625 | 0.017748 | 92 | 0.196001 | 0.152660 | 112 | 0.500000 | 0.500000 |
| 73 | 0.033549 | 0.019724 | 93 | 0.213325 | 0.167668 | 113 | 0.500000 | 0.500000 |
| 74 | 0.036614 | 0.021915 | 94 | 0.231936 | 0.183524 | 114 | 0.500000 | 0.500000 |
| 75 | 0.040012 | 0.024393 | 95 | 0.251189 | 0.200229 | 115 | 0.500000 | 0.500000 |
| 76 | 0.043933 | 0.027231 | 96 | 0.270441 | 0.217783 | 116 | 0.500000 | 0.500000 |
| 77 | 0.048570 | 0.030501 | 97 | 0.289048 | 0.236188 | 117 | 0.500000 | 0.500000 |
| 78 | 0.053991 | 0.034115 | 98 | 0.306750 | 0.255605 | 118 | 0.500000 | 0.500000 |
| 79 | 0.060066 | 0.038024 | 99 | 0.323976 | 0.276035 | 119 | 0.500000 | 0.500000 |
| 80 | 0.066696 | 0.042361 | 100 | 0.341116 | 0.297233 | 120 | 1.000000 | 1.000000 |

a least-squares best-fit trend line through the logarithms of the rates for that age. The opening year of 1977 was chosen because it provided a reasonable representation of anticipated trends in the future and, properly, did not reflect more rapid mortality improvement rates found in the experience of prior periods.

The trends for Social Security are based on data from SSA 107 along with additional data used in this study, which were provided by the Office of the Actuary at the Social Security Administration. These additional data included central death rates for five-year age groups for each calendar year over the period 1960-1988. Before the Social Security trends for 1977-1988 could be blended with the CSRS trends for 1977-1993, it was necessary to extend the Social Security trends up through 1993. This extension was based on mortality improvement trends for the CSRS from 1988 through 1993. The SSA 107 extended central death rates for each year 1989 through 1993 were obtained by multiplying the SSA central death rate for 1988 by the ratio of the CSRS central death rate for the corresponding year to the CSRS central death rate for 1988 . Then the average trend for each central age over the entire 1977-1993 period was determined based on a least-squares bestfit trend line through the logarithms of these central death rates. The Social Security data did not cover central ages beyond age 92, and the CSRS data at these older ages were limited. The mortality improvement trends for individual ages were interpolated from the trends for the central ages by using the Karup-King four-point interpolation formula.

The trends at ages $1-25$ were based on Social Security data and on the Social Security assumptions for future trends listed in SSA 107 and start out at a rate of improvement of $2 \%$ per year. Then the CSRS mortality improvement trend for each age was averaged with the corresponding trend for Social Security. These average trends were then rounded to the nearest onetenth of one percentage point. The resulting mortality improvement factors are shown in Table 14.

To obtain the mortality improvement factors for projecting mortality beyond 1994, the following modifications were made in this scale:

1. Any mortality improvement factors that were less than $0.5 \%$ for ages under 85 were changed to $0.5 \%$, because the Task Force thought that the use of lower factors would result in excessive mortality rates in the future.
2. A maximum mortality improvement rate of $2.0 \%$ was set for ages under 60 . This reduced the highest rate of $2.3 \%$ to $2.0 \%$ at male ages $52-54$ and provided a smoother progression of rates around these ages.

TABLE 14
Annual Mortality lmprovement Factors from the SSA 107 and CSRS Studies
Based upon 1977-1993 Experience Rates in Percentage per Year

| Age | Male |  |  | Female |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \hline \text { SS } \\ 77-93 \end{gathered}$ | $\begin{aligned} & \text { CSRS } \\ & 77-93 \\ & \hline \end{aligned}$ | Average | $\begin{gathered} \text { SS } \\ 77-93 \end{gathered}$ | $\begin{aligned} & \text { CSRS } \\ & 77-93 \\ & \hline \end{aligned}$ | Average |
| 1 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 2 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 3 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 4 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 5 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 6 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 7 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 8 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 9 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 10 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 11 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 12 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 13 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 | 2.0 |
| 14 | 1.9 | 1.9 | 1.9 | 1.8 | 1.8 | 1.8 |
| 15 | 1.9 | 1.9 | 1.9 | 1.6 | 1.6 | 1.6 |
| 16 | 1.9 | 1.9 | 1.9 | 1.5 | 1.5 | 1.5 |
| 17 | 1.9 | 1.9 | 1.9 | 1.4 | 1.4 | 1.4 |
| 18 | 1.9 | 1.9 | 1.9 | 1.4 | 1.4 | 1.4 |
| 19 | 1.9 | 1.9 | 1.9 | 1.5 | 1.5 | 1.5 |
| 20 | 1.9 | 1.9 | 1.9 | 1.6 | 1.6 | 1.6 |
| 21 | 1.8 | 1.8 | 1.8 | 1.7 | 1.7 | 1.7 |
| 22 | 1.7 | 1.7 | 1.7 | 1.7 | 1.7 | 1.7 |
| 23 | 1.5 | 1.5 | 1.6 | 1.6 | 1.6 | 1.6 |
| 24 | 1.3 | 1.3 | 1.3 | 1.5 | 1.5 | 1.5 |
| 25 | 1.0 | 1.0 | 1.0 | 1.4 | 1.4 | 1.4 |
| 26 | 0.6 | 0.6 | 0.6 | 1.2 | 1.2 | 1.2 |
| 27 | 0.3 | 0.4 | 0.3 | 1.1 | 0.9 | 1.0 |
| 28 | 0.0 | -0.4 | -0.2 | 0.9 | 1.4 | 1.2 |
| 29 | -0.5 | -1.0 | -0.8 | 0.8 | 1.6 | 1.2 |
| 30 | $-1.0$ | -1.4 | $-1.2$ | 0.6 | 1.5 | 1.0 |
| 31 | $-1.4$ | $-1.6$ | $-1.5$ | 0.5 | 1.1 | 0.8 |
| 32 | -1.7 | -1.5 | -1.6 | 0.5 | 0.6 | 0.6 |
| 33 | -1.7 | -1.2 | $-1.4$ | 0.7 | 0.2 | 0.4 |
| 34 | $-1.5$ | -0.9 | $-1.2$ | 1.0 | 0.0 | 0.5 |
| 35 | -1.2 | -0.5 | -0.9 | 1.3 | 0.1 | 0.7 |
| 36 | -0.9 | -0.3 | -0.6 | 1.6 | 0.4 | 1.0 |
| 37 | -0.6 | -0.1 | -0.4 | 1.8 | 0.8 | 1.3 |
| 38 | $-0.3$ | 0.0 | -0.2 | 1.9 | 1.0 | 1.5 |
| 39 | 0.0 | 0.0 | 0.0 | 2.0 | 1.1 | 1.5 |
| 40 | 0.3 | 0.2 | 0.2 | 2.0 | 1.0 | 1.5 |
| 41 | 0.6 | 0.4 | 0.5 | 2.0 | 0.8 | 1.4 |
| 42 | 0.9 | 0.7 | 0.8 | 2.0 | 0.7 | 1.3 |
| 43 | 1.1 | 1.0 | 1.1 | 2.0 | 0.7 | 1.4 |
| 44 | 1.3 | 1.3 | 1.3 | 2.0 | 0.9 | 1.4 |
| 45 | 1.5 | 1.6 | 1.5 | 2.0 | 1.2 | 1.6 |

TABLE 14-Continued

| Age | Male |  |  | Female |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { SS } \\ 77-93 \\ \hline \end{gathered}$ | $\begin{gathered} \text { CSRS } \\ 77-93 \\ \hline \end{gathered}$ | Average | $\begin{gathered} \text { SS } \\ 77-93 \\ \hline \end{gathered}$ | $\begin{aligned} & \text { CSRS } \\ & 77-93 \\ & \hline \end{aligned}$ | Average |
| 46 | 1.6 | 1.8 | 1.7 | 1.9 | 1.4 | 1.7 |
| 47 | 1.7 | 2.1 | 1.9 | 1.9 | 1.6 | 1.8 |
| 48 | 1.8 | 2.3 | 2.0 | 1.9 | 1.6 | 1.8 |
| 49 | 1.8 | 2.5 | 2.2 | 1.9 | 1.6 | 1.8 |
| 50 | 1.9 | 2.7 | 2.3 | 1.9 | 1.5 | 1.7 |
| 51 | 1.9 | 2.8 | 2.3 | 1.9 | 1.3 | 1.6 |
| 52 | 1.9 | 2.7 | 2.3 | 1.9 | 1.0 | 1.4 |
| 53 | 1.8 | 2.6 | 2.2 | 1.7 | 0.8 | 1.2 |
| 54 | 1.8 | 2.4 | 2.1 | 1.5 | 0.6 | 1.0 |
| 55 | 1.7 | 2.2 | 1.9 | 1.3 | 0.4 | 0.8 |
| 56 | 1.6 | 2.0 | 1.8 | 1.0 | 0.2 | 0.6 |
| 57 | 1.6 | 1.8 | 1.7 | 0.8 | 0.1 | 0.5 |
| 58 | 1.6 | 1.7 | 1.6 | 0.6 | 0.0 | 0.3 |
| 59 | 1.6 | 1.5 | 1.6 | 0.4 | -0.2 | 0.1 |
| 60 | 1.7 | 1.4 | 1.6 | 0.2 | -0.4 | 0.0 |
| 61 | 1.8 | 1.3 | 1.5 | 0.0 | -0.5 | $-0.2$ |
| 62 | 1.8 | 1.2 | 1.5 | 0.0 | -0.5 | -0.2 |
| 63 | 1.7 | 1.2 | 1.4 | 0.0 | -0.5 | -0.2 |
| 64 | 1.6 | 1.2 | 1.4 | 0.0 | -0.4 | -0.1 |
| 65 | 1.5 | 1.2 | 1.4 | 0.1 | -0.2 | 0.0 |
| 66 | 1.4 | 1.3 | 1.3 | 0.2 | 0.0 | 0.1 |
| 67 | 1.3 | 1.4 | 1.3 | 0.3 | 0.1 | 0.2 |
| 68 | 1.3 | 1.5 | 1.4 | 0.4 | 0.3 | 0.3 |
| 69 | 1.3 | 1.5 | 1.4 | 0.5 | 0.3 | 0.4 |
| 70 | 1.3 | 1.6 | 1.5 | 0.7 | 0.4 | 0.5 |
| 71 | 1.4 | 1.6 | 1.5 | 0.8 | 0.4 | 0.6 |
| 72 | 1.4 | 1.6 | 1.5 | 0.9 | 0.4 | 0.6 |
| 73 | 1.3 | 1.6 | 1.5 | 0.9 | 0.5 | 0.7 |
| 74 | 1.3 | 1.6 | 1.5 | 0.9 | 0.5 | 0.7 |
| 75 | 1.3 | 1.6 | 1.4 | 1.0 | 0.6 | 0.8 |
| 76 | 1.2 | 1.5 | 1.4 | 0.9 | 0.6 | 0.8 |
| 77 | 1.2 | 1.4 | 1.3 | 0.9 | 0.5 | 0.7 |
| 78 | 1.1 | 1.3 | 1.2 | 0.9 | 0.5 | 0.7 |
| 79 | 1.0 | 1.2 | 1.1 | 0.9 | 0.5 | 0.7 |
| 80 | 0.9 | 1.0 | 1.0 | 0.9 | 0.5 | 0.7 |
| 81 | 0.9 | 0.9 | 0.9 | 0.9 | 0.5 | 0.7 |
| 82 | 0.8 | 0.8 | 0.8 | 0.9 | 0.6 | 0.7 |
| 83 | 0.8 | 0.8 | 0.8 | 0.9 | 0.5 | 0.7 |
| 84 | 0.8 | 0.7 | 0.7 | 0.8 | 0.5 | 0.7 |
| 85 | 0.8 | 0.6 | 0.7 | 0.8 | 0.4 | 0.6 |
| 86 | 0.8 | 0.6 | 0.7 | 0.8 | 0.3 | 0.5 |
| 87 | 0.7 | 0.5 | 0.6 | 0.7 | 0.2 | 0.4 |
| 88 | 0.6 | 0.5 | 0.5 | 0.7 | 0.1 | 0.4 |
| 89 | 0.5 | 0.4 | 0.5 | 0.7 | 0.0 | 0.3 |
| 90 | 0.4 | 0.3 | 0.4 | 0.7 | 0.0 | 0.3 |
| 91 | 0.3 | 0.2 | 0.3 | 0.6 | 0.0 | 0.3 |
| 92 | 0.2 | 0.2 | 0.2 | 0.6 | 0.0 | 0.3 |

3. At the higher ages the mortality improvement rates were graded to a value of $0.1 \%$ at age 100 and set to 0 for all ages greater than 100 .
4. Other minor adjustments were made as described below.

After age 37, the factors for males start to increase fairly rapidly from one age to the next, going from a factor of $0.2 \%$ (before change) at age 38 to $2.3 \%$ at age 50 .

When there are large mortality improvement factor increases from one age to the next like this, it is possible that, after the mortality improvement scale has been applied for a number of years, the mortality rate for a particular age could become lower than the rate for an age one year younger. To minimize this possibility, it was decided to limit the increase in the factor from one age to the next to one-tenth of one percentage point. As a result, the mortality improvement factors for males were modified so that they increase from $0.5 \%$ at age 37 to $2.0 \%$ at age 52 . The factors for some ages were increased by this process, and factors for other ages were reduced.

There are also significant age-to-age increases for females in the factors from ages 33 through 38 . The factors for females for ages 32 through 38 were therefore also modified, as were the factors at female ages 41 to 44 .

Mortality improvement factors to be used in the new Group Annuity Mortality Table when projecting mortality rates beyond 1994 are shown in Table 15 and are referred to as the Projection Scale AA. Figure 1 displays a graph of the Projection Scale AA factors for males. Figure 2 displays a graph of the Projection Scale AA factors for females.

## IV. MARGINS

Consistent with accepted actuarial practice and precedent set in the development of existing mortality tables used in reserving, the Task Force deemed it necessary and appropriate to add margins to the $q_{x}$ values of the 1994 Group Annuity Mortality Basic Table. The overall margin comprises two components:

1. Margins for random variation in mortality rates
2. Margins for other contingencies.

## A. Margins for Random Variation in Mortality Rates

The unloaded 1994 Group Annuity Mortality Basic Table $q_{x}$ values shown in Table 13 represent expected values. Considering current reserving theory, the Task Force decided to incorporate margins to produce annuity reserve

TABLE 15
Projection Scale AA
Mortality Improvement Factors To Be Used in the New Table
When Projecting Mortality Rates beyond 1994
Factors are Shown as Percentage per Year

| Altained Age | Male Factor | Female Factor | Atained Age | Male Factor | Female Factor | Allained Age | Male Factor | Female Factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| I | 2.0 | 2.0 | 21 | 1.8 | 1.7 | 41 | 0.9 | 1.5 |
| 2 | 2.0 | 2.0 | 22 | 1.7 | 1.7 | 42 | 1.0 | 1.5 |
| 3 | 2.0 | 2.0 | 23 | 1.5 | 1.6 | 43 | 1.1 | 1.5 |
| 4 | 2.0 | 2.0 | 24 | 1.3 | 1.5 | 44 | 1.2 | 1.5 |
| 5 | 2.0 | 2.0 | 25 | 1.0 | 1.4 | 45 | 1.3 | 1.6 |
| 6 | 2.0 | 2.0 | 26 | 0.6 | 1.2 | 46 | 1.4 | 1.7 |
| 7 | 2.0 | 2.0 | 27 | 0.5 | 1.2 | 47 | 1.5 | 1.8 |
| 8 | 2.0 | 2.0 | 28 | 0.5 | 1.2 | 48 | 1.6 | 1.8 |
| 9 | 2.0 | 2.0 | 29 | 0.5 | 1.2 | 49 | 1.7 | 1.8 |
| 10 | 2.0 | 2.0 | 30 | 0.5 | 1.0 | 50 | 1.8 | 1.7 |
| 11 | 2.0 | 2.0 | 31 | 0.5 | 0.8 | 51 | 1.9 | 1.6 |
| 12 | 2.0 | 2.0 | 32 | 0.5 | 0.8 | 52 | 2.0 | 1.4 |
| 13 | 2.0 | 2.0 | 33 | 0.5 | 0.9 | 53 | 2.0 | 1.2 |
| 14 | 1.9 | 1.8 | 34 | 0.5 | 1.0 | 54 | 2.0 | 1.0 |
| 15 | 1.9 | 1.6 | 35 | 0.5 | 1.1 | 55 | 1.9 | 0.8 |
| 16 | 1.9 | 1.5 | 36 | 0.5 | 1.2 | 56 | 1.8 | 0.6 |
| 17 | 1.9 | 1.4 | 37 | 0.5 | 1.3 | 57 | 1.7 | 0.5 |
| 18 | 1.9 | 1.4 | 38 | 0.6 | 1.4 | 58 | 1.6 | 0.5 |
| 19 | 1.9 | 1.5 | 39 | 0.7 | 1.5 | 59 | 1.6 | 0.5 |
| 20 | 1.9 | 1.6 | 40 | 0.8 | 1.5 | 60 | 1.6 | 0.5 |

TABLE 15--Continued

| Attained Age | Male Factor | Female Factor | Alained Age | Male Factor | Female Factor | Altained Age | Male Factor | Female Factor |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 61 | 1.5 | 0.5 | 81 | 0.9 | 0.7 | 101 | 0.0 | 0.0 |
| 62 | 1.5 | 0.5 | 82 | 0.8 | 0.7 | 102 | 0.0 | 0.0 |
| 63 | 1.4 | 0.5 | 83 | 0.8 | 0.7 | 103 | 0.0 | 0.0 |
| 64 | 1.4 | 0.5 | 84 | 0.7 | 0.7 | 104 | 0.0 | 0.0 |
| 65 | 1.4 | 0.5 | 85 | 0.7 | 0.6 | 105 | 0.0 | 0.0 |
| 66 | 1.3 | 0.5 | 86 | 0.7 | 0.5 | 106 | 0.0 | 0.0 |
| 67 | 1.3 | 0.5 | 87 | 0.6 | 0.4 | 107 | 0.0 | 0.0 |
| 68 | 1.4 | 0.5 | 88 | 0.5 | 0.4 | 108 | 0.0 | 0.0 |
| 69 | 1.4 | 0.5 | 89 | 0.5 | 0.3 | 109 | 0.0 | 0.0 |
| 70 | 1.5 | 0.5 | 90 | 0.4 | 0.3 | 110 | 0.0 | 0.0 |
| 71 | 1.5 | 0.6 | 91 | 0.4 | 0.3 | 111 | 0.0 | 0.0 |
| 72 | 1.5 | 0.6 | 92 | 0.3 | 0.3 | 112 | 0.0 | 0.0 |
| 73 | 1.5 | 0.7 | 93 | 0.3 | 0.2 | 113 | 0.0 | 0.0 |
| 74 | 1.5 | 0.7 | 94 | 0.3 | 0.2 | 114 | 0.0 | 0.0 |
| 75 | 1.4 | 0.8 | 95 | 0.2 | 0.2 | 115 | 0.0 | 0.0 |
| 76 | 1.4 | 0.8 | 96 | 0.2 | 0.2 | 116 | 0.0 | 0.0 |
| 77 | 1.3 | 0.7 | 97 | 0.2 | 0.1 | 117 | 0.0 | 0.0 |
| 78 | 1.2 | 0.7 | 98 | 0.1 | 0.1 | 118 | 0.0 | 0.0 |
| 79 | 1.1 | 0.7 | 99 | 0.1 | 0.1 | 119 | 0.0 | 0.0 |
| 80 | 1.0 | 0.7 | 100 | 0.1 | 0.1 | 120 | 0.0 | 0.0 |

FIGURE 1
Mortality Improvement Factors-Male
Scale AA


FIGURE 2
Mortality Improvement Factors-Female
Scale AA

values that would be adequate for random variation of two standard deviations from expected mortality.

Probability theory was used to develop variances of distributions of annuity values as indicated below.

For a single life age $x$, assume $Y$ is a random variable representing the present value of annuity payments received. $Y$ would have the following probability distribution:

| $\underline{Y}$ | $\operatorname{Pr}(Y=y)$ |
| :---: | :---: |
| 0 | $1-p$ |
| $a_{1}$ | ${ }_{1} q_{y}\left[=p_{y}\left(1-p_{y+1}\right)\right]$ |
| $a_{2}$ | ${ }_{21} q_{y}$ |
| $a_{3}$ | ${ }_{3}{ }^{9} y$ |
| - | - |
| - | - |

The mean, variance, and standard deviation of this distribution would be determined as follows:

$$
\begin{align*}
\mu=E[Y] & =\sum_{t=0}^{\infty} a_{\eta} \times \operatorname{Pr}\left(Y=a_{i}\right)  \tag{C}\\
E\left[Y^{2}\right] & =\sum_{i=0}^{\infty} a_{i}^{2} \times \operatorname{Pr}\left(Y=a_{\eta}\right)  \tag{D}\\
\sigma^{2} & =E\left[Y^{2}\right]-(E[Y])^{2}  \tag{E}\\
\sigma & =\sqrt{\sigma^{2}} \tag{F}
\end{align*}
$$

For a distribution of annuity values for $N$ lives age $x$, assumed to be independent, the mean, variance and standard deviations would be calculated as follows:

$$
\begin{align*}
\boldsymbol{\mu}_{N} & =N \times \boldsymbol{\mu}  \tag{G}\\
\boldsymbol{\sigma}_{N}^{2} & =N \times \boldsymbol{\sigma}^{2}  \tag{H}\\
\boldsymbol{\sigma}_{N} & =\sqrt{N} \times \boldsymbol{\sigma} \tag{I}
\end{align*}
$$

As the size of a company's group annuity block of business increases, the required margins for random variations decrease. The Task Force reviewed
recent statutory annual statement data on group annuity business to determine an appropriate company block of business volume assumption to use in calculating the random variation margin component. To ensure that the new standard would provide at least a two-standard-deviation margin for the vast majority of companies (more than $95 \%$ ) having insured group annuity business, the Task Force decided that a 3,000 -life block of business would be appropriate for computing margins for random variation.

Tables 16 and 17 show the results of applying these concepts and the determination of required margins to be built into the GAM-94 Basic Table $q_{x}$ values shown in Table 13. Expected values and standard deviations were calculated by using the formulas presented in this section with a value of $N$ 3,000 . The interest rate used in the analysis was $6 \%$. Note that use of other interest assumptions and forms of annuity did not significantly change the level of required margins.

TABLE 16
Random Variation Analysis of Required Margins for Males
GAM-94 Basic Table Expected Mortality
3,000-Life Group, Interest at 6\%

| Annuity Type | Age | Expected Value | Standard <br> Deviation | Required Margins |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 1 Standard Deviation | 2 Standard Deviations |
| Immediate Life Annuities | 45 | 41,400 | 128 | 3.0\% | 6.0\% |
|  | 50 | 38,987 | 148 | 2.8 | 5.6 |
|  | 55 | 36,043 | 167 | 2.6 | 5.2 |
|  | 60 | 32,574 | 184 | 2.5 | 4.9 |
|  | 65 | 28,724 | 194 | 2.3 | 4.6 |
|  | 70 | 24,697 | 195 | 2.2 | 4.3 |
|  | 75 | 20.459 | 188 | 2.0 | 4.0 |
|  | 80 | 16,180 | 175 | 1.9 | 3.8 |
| Deferred to Age 65 Life Annuity | 30 | 3,281 | 33 | 2.3\% | 4.6\% |
|  | 35 | 4,410 | 43 | 2.3 | 4.6 |
|  | 40 | 5,931 | 58 | 2.3 | 4.6 |
|  | 45 | 7,990 | 76 | 2.3 | 4.6 |
|  | 50 | 10,804 | 101 | 2.3 | 4.7 |
|  | 55 | 14,713 | 132 | 2.4 | 4.7 |
|  | 60 | 20,301 | 167 | 2.4 | 4.7 |

Based on these results, the Task Force concluded that a $5 \%$ margin would make adequate provision for random variations in mortality for reserving purposes.

TABLE 17
Random Variation Analysis of Required Margins for Females GAM-94 Basic Table Expected Mortality

3,000-Life Group, Interest at 6\%

|  |  |  |  |  | Required Margins |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Annuity Type | Expected <br> Value | Standard <br> Deviation | 1 Standard <br> Deviation | 2 Standard <br> Deviations |  |  |
| Immediate Life | 45 | 43,301 | 107 | $3.3 \%$ | $6.5 \%$ |  |
| Annuities | 50 | 41,342 | 125 | 3.0 | 6.0 |  |
|  | 55 | 38,861 | 145 | 2.8 | 5.6 |  |
|  | 60 | 35,810 | 165 | 2.7 | 5.3 |  |
|  | 65 | 32,307 | 180 | 2.5 | 4.9 |  |
|  | 70 | 28,439 | 186 | 2.3 | 4.6 |  |
|  | 75 | 24,039 | 187 | 2.1 | 4.3 |  |
|  | 80 | 19,406 | 180 | 2.0 | 4.0 |  |
| Deferred to Age | 30 | 3,907 | 30 | $2.6 \%$ | $5.1 \%$ |  |
| 65 Life Annuity | 35 | 5,239 | 40 | 2.6 | 5.1 |  |
|  | 40 | 7,032 | 53 | 2.6 | 5.1 |  |
|  | 45 | 9,453 | 70 | 2.6 | 5.1 |  |
|  | 50 | 12,727 | 92 | 2.6 | 5.1 |  |
|  | 55 | 17,192 | 120 | 2.6 | 5.1 |  |

## B. Margins for Otber Contingencies

The Task Force thought that the $5 \%$ margin was adequate for random variation for most insurance companies. However, blocks of business of less than 3,000 lives would have a greater standard deviation than shown above. Also, variations in the mix of business that companies write may cause the underlying mortality for a given company to differ from the underlying mortality in the valuation standard. Examples of business characteristics that could affect the underlying mortality averages include:

1. The mix of white-collar and blue-collar workers
2. The mix of higher-income and lower-income annuitants
3. Degree of concentration by geographic area.

For these reasons, the Task Force decided to recommend a specific margin to be added to the $5 \%$ statistical margin. The conclusion was to add $2 \%$ to the $5 \%$ statistical margin to produce a total $7 \%$ margin. It is anticipated that this margin produces reserves that are adequate to cover various business characteristics and random variations.

The resulting $q_{x}$ values, including the $7 \%$ margin, comprise the 1994 Group Annuity Mortality Static Table and are presented in Table 18. Table 18 is calculated as $93 \%$ of the corresponding Table 13 values, with modification after age 102 . No margin was applied to the mortality rates of 0.5

TABLE 18
1994 Group annuity Mortality Static Table
1994 Base Year

| Age | Values of $q_{x}$ |  | Age | Values of $q_{\text {, }}$ |  | Age | Values of $q_{\text {, }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Female |  | Male | Female |  | Male | Femate |
| 1 | 0.000592 | 0.000531 | 21 | 0.000530 | 0.000286 | 41 | 0.001156 | 0.000768 |
| 2 | 0.000400 | 0.000346 | 22 | 0.000556 | 0.000289 | 42 | 0.001252 | 0.000825 |
| 3 | 0.000332 | 0.000258 | 23 | 0.000589 | 0.000292 | 43 | 0.001352 | 0.000877 |
| 4 | 0.000259 | 0.000194 | 24 | 0.000624 | 0.000291 | 44 | 0.001458 | 0.000923 |
| 5 | 0.000237 | 0.000175 | 25 | 0.000661 | 0.000291 | 45 | 0.001578 | 0.000973 |
| 6 | 0.000227 | 0.000163 | 26 | 0.000696 | 0.000294 | 46 | 0.001722 | 0.001033 |
| 7 | 0.000217 | 0.000153 | 27 | 0.000727 | 0.000302 | 47 | 0.001899 | 0.001112 |
| 8 | 0.000201 | 0.000137 | 28 | 0.000754 | 0.000314 | 48 | 0.002102 | 0.001206 |
| 9 | 0.000194 | 0.000130 | 29 | 0.000779 | 0.000331 | 49 | 0.002326 | 0.001310 |
| 10 | 0.000197 | 0.000131 | 30 | 0.000801 | 0.000351 | 50 | 0.002579 | 0.001428 |
|  | 0.000208 | 0.000138 | 31 | 0.000821 | 0.000373 | 51 | 0.002872 | 0.001568 |
| 12 | 0.000226 | 0.000148 | 32 | 0.000839 | 0.000397 | 52 | 0.003213 | 0.001734 |
| 13 | 0.000255 | 0.000164 | 33 | 0.000848 | 0.000422 | 53 | 0.003584 | 0.001907 |
| 14 | 0.000297 | 0.000189 | 34 | 0.000849 | 0.000449 | 54 | 0.003979 | 0.002084 |
| 15 | 0.000345 | 0.000216 | 35 | 0.000851 | 0.000478 | 55 | 0.004425 | 0.002294 |
| 16 | 0.000391 | 0.000242 | 36 | 0.000862 | 0.000512 | 56 | 0.004949 | 0.002563 |
| 17 | 0.000430 | 0.000262 | 37 | 0.000891 | 0.000551 | 57 | 0.005581 | 0.002919 |
| 18 | 0.000460 | 0.000273 | 38 | 0.000939 | 0.000598 | 58 | 0.006300 | 0.003359 |
| 19 | 0.000484 | 0.000280 | 39 | 0.000999 | 0.000652 | 59 | 0.007090 | 0.003863 |
| 20 | 0.000507 | 0.000284 | 40 | 0.001072 | 0.000709 | 60 | 0.007976 | 0.004439 |



| Age | Values of $q_{x}$ |  | Age | Values of $q$, |  | Age | Values of $q_{x}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Male | Female |  | Male | Female |  | Male | Female |
| 61 | 0.008986 | 0.005093 | 81 | 0.068615 | 0.043952 | 101 | 0.333461 | 0.296629 |
| 62 | 0.010147 | 0.005832 | 82 | 0.075532 | $\checkmark 0.049153$ | 102 | 0.350330 | 0.317093 |
| 63 | 0.011471 | 0.006677 | 83 | 0.082510 | 0.054847 | 103 | 0.368542 | 0.338505 |
| 64 | 0.012940 | 0.007621 | 84 | 0.089613 | 0.060979 | 104 | 0.387855 | 0.361016 |
| 65 | 0.014535 | 0.008636 | 85 | 0.097240 | 0.067738 | 105 | 0.407224 | 0.383597 |
| 66 | 0.016239 | 0.009694 | 86 | 0.105792 | 0.075347 | 106 | 0.425599 | 0.405217 |
| 67 | 0.018034 | 0.010764 | 87 | 0.115671 | 0.084023 | 107 | 0.441935 | 0.424846 |
| 68 | 0.019859 | 0.011763 | 88 | 0.126980 | 0.093820 | 108 | 0.457553 | 0.444368 |
| 69 | 0.021729 | 0.012709 | 89 | 0.139452 | 0.104594 | 109 | 0.473150 | 0.464469 |
| 70 | 0.023730 | 0.013730 | 90 | 0.152931 | 0.116265 | 110 | 0.486745 | 0.482325 |
| 71 | 0.025951 | 0.014953 | 91 | 0.167260 | 0.128751 | 111 | 0.496356 | 0.495110 |
| 72 | 0.028481 | 0.016506 | 92 | 0.182281 | 0.141973 | 112 | 0.500000 | 0.500000 |
| 73 | 0.031201 | 0.018344 | 93 | 0.198392 | 0.155931 | 113 | 0.500000 | 0.500000 |
| 74 | 0.034051 | 0.020381 | 94 | 0.215700 | 0.170677 | 114 | 0.500000 | 0.500000 |
| 75 | 0.037211 | 0.022686 | 95 | 0.233606 | 0.186213 | 115 | 0.500000 | 0.500000 |
| 76 | 0.040858 | 0.025325 | 96 | 0.251510 | 0.202538 | 116 | 0.500000 | 0.500000 |
| 77 | 0.045171 | 0.028366 | 97 | 0.268815 | 0.219655 | 117 | 0.500000 | 0.500000 |
| 78 | 0.050211 | 0.031727 | 98 | 0.285277 | 0.237713 | 118 | 0.500000 | 0.500000 |
| 79 | 0.055861 | 0.035362 | 99 | 0.301298 | 0.256712 | 119 | 0.500000 | 0.500000 |
| 80 | 0.062027 | 0.039396 | 100 | 0.317238 | 0.276427 | 120 | 1.000000 | 1.000000 |

at ages 112 and older. A modified Karup-King graduation process was used to obtain a smooth transition from the rates under age 103 to the rates at age 112 and above. Figure 3 displays a graph of the mortality rates for male and female ages 1-119 shown in Table 18. Figures 4,5 and 6 display those rates by the age categories of $1-40,40-70$, and $70-119$, respectively.

FIGURE 3
1994 Group annuity Mortality Static Table Rates
1994 Base YEar
Ages 1-119


## V. THE GENERATION MORTALITY TABLE

## A. Development of Generation Mortality Tables

The Task Force was now in a position to produce the generation mortality tables for males and females.

Prior mortality table generation methodologies included mortality tables produced from projection scales. Thus, if we have a static mortality table that is appropriate for 1994, together with mortality improvement factors that are assumed to apply in the calendar years 1995 and later, we can produce a static mortality table for each calendar year 1995 and later.

FIGURE 4
1994 Group Annutty Mortality Static Table Rates
1994 Base Year
Ages 1-40


FIGURE 5
1994 Group Annuity Mortality Static Table Rates 1994 Base Year Ages 40-70


FIGURE 6
1994 Group annuity Mortality Static Table Rates
1994 Base Year
AGES 70-119


For example, assume a set of generation mortality rates is required to calculate annuity values for issue age 65 in calendar year 1997. The attained age $65 q_{x}$ value would be taken from the 1997 static table. The attained age $66 q_{x}$ value would be taken from the 1998 static table. This process would be continued until the ultimate age $q_{x}$ value is taken from the appropriate final static table. Table 19 illustrates this process if we understand that the columnar rates come from the individual static mortality tables.

$$
\begin{equation*}
q_{x}^{y}=q_{x}^{1994}\left(1-A A_{x}\right)^{y-1994} \tag{J}
\end{equation*}
$$

An abbreviated example illustrates the principles involved in determining the $q_{x}$ values needed to calculate an annuity value using generational mortality techniques.

As a specific example, assume one wishes to calculate, in 1994, a fiveyear temporary life annuity for a male age 63, using the GAM-94 Static Table from Table 18 and Projection Scale AA from Table 15. This requires determination of mortality rates for male ages 63-67 that would be applicable in 1994-1998. This example requires only five abbreviated "static" tables. However, a life annuity calculation would require the generation table

TABLE 19
Illustration of the Development of a Generation Mortality Table from Static Mortality Tables

| Age | 1994 | 1995 | 1996 | 1997 | 1998 | 1999 | 2000 | 2001 | .- | 2052 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 65 | 965 | $9_{65}^{1995}$ | 9659 | $q_{65}^{1997}$ | $q_{65}^{1998}$ | 9859 | $q_{65}^{2000}$ | $q_{65}^{2001}$ | ... | $q_{65}^{2052}$ |
| 66 | $9_{66}^{1994}$ | $q_{60}^{1995}$ | $q_{66}^{1996}$ | $q_{66}^{1997}$ | 966 | $9{ }_{60}^{1999}$ | $\boldsymbol{q}_{66}^{2000}$ | $9^{2001}$ | $\ldots$ | $q_{00}^{2052}$ |
| 67 | $q_{67}^{1994}$ | $q_{67}^{1995}$ | $q_{67}^{1996}$ | $9{ }_{67}^{1997}$ | $q_{67}^{1998}$ | $9{ }^{1979}$ | $q_{67}^{2000}$ | $q^{207}$ | $\cdots$ | $q_{67}^{2092}$ |
| 68 | $9_{68}^{1994}$ | $99_{68}^{1995}$ | $9_{68}^{1996}$ | $9_{68}^{199}$ | $q_{68}^{1998}$ | $9_{68}^{1999}$ | $\mathcal{G}^{2080}$ | $q_{88}^{2001}$ | ... | $q_{68}^{2052}$ |
| 69 | 969 | $q_{69}^{1995}$ | $q_{69}^{1996}$ | $q_{69}^{1997}$ | $q_{69}^{1998}$ | $9{ }_{69}^{1999}$ | $\boldsymbol{q}_{69}^{2000}$ | $9_{669}^{8001}$ | ... | $q_{69}^{2052}$ |
| - | - | $\cdot$ | - | - | - | - | - | - | ..* | - |
| - | - | - | $\cdot$ | $\cdot$ | $\cdot$ | - | - | - | $\cdots$ | - |
| - | - | - | - | - | - | - | - | - | ... | - |
| 120 | $q_{120}^{1994}$ | $q_{120}^{1093}$ | $q_{120}^{19 \%}$ | $q_{120}^{1997}$ | $q_{120}^{1998}$ | $q_{120}^{1909}$ | $q_{120}^{2000}$ | $q_{120}^{2001}$ | $\cdots$ |  |

to comprise 57 "static" tables, using age 120 as the last attained age in the calculations.

Table 20 shows our assumptions of mortality improvement factors and $1000 q_{x}$ values in columns (1) and (2), respectively, and the resulting calculated values for future years in columns (3) to (6). Column (1) shows the final male mortality improvement factors from Table 15 by attained age, ages 63-67 in our example. Column (2) shows the GAM-94 Static Table of death rates in 1994 for attained ages 63-67 from Table 18. Columns (3) to (6) show calculated generation table death rates during the calendar years 1995-1998.

TABLE 20
Generation Mortality Table for the Years 1994 - 1998 Based on Gam-94 Static Table for Males with Full Generation Using Projection Scale aA Specimen $1000 q_{x}$ Mortality Rates for Issue Year 1994

| Attained Age | (1) <br> Montality Improvement Factor | Values of $1000 q_{\text {, }}$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | (2) | (3) | (4) | (5) | (6) |
|  |  | 1994 | 1995 | 1996 | 1997 | 1998 |
| 63 | 1.4\% | 11.471 | 11.310 | 11.152 | 10.996 | 10.842 |
| 64 | 1.4 | 12.940 | 12.759 | 12.580 | 12.404 | 12.230 |
| 65 | 1.4 | 14.535 | 14.332 | 14.131 | 13.933 | 13.738 |
| 66 | 1.3 | 16.239 | 16.028 | 15.820 | 15.614 | 15.411 |
| 67 | 1.3 | 18.034 | 17.800 | 17.568 | 17.340 | 17.114 |

The values in columns (3) to (6) for age 63 are calculated as 11.471 multiplied successively by ( $1-0.014$ ). For age 65 values under columns (3) to ( 6 ), 14.535 would be multiplied successively by ( $1-0.014$ ).

Our required mortality rates are therefore found along the diagonal beginning with 11.471 , followed by $12.759,14.131,15.614$, and 17.114.

Generation mortality rates from the GAM-94 Static Table for males and females at issue age 65 until attained age 120 are shown in Table 21 and Table 22, respectively. These tables compare the rates of mortality for issues of 1994, 1999, 2004, and 2009. A similar set of tabular rates applies to each issue age, for each issue year. Note that the mortality rates by issue year are the same for attained ages 101 and older because no mortality improvement is assumed at these advanced ages.

Note that the generation tables shown for each issue year in Table 21 (male) and Table 22 (female) reflect projected mortality using the general formula on page 909:

## TABLE 21

Generation Mortality Rates per 1,000 for Issues of 1994, 1999, 2004, and 2009 at Male Issue Age 65 in the Indicated Year
Based upon GaM-94 Static Table with Full Generation and Projection Scale AA

| Age | Mate Issue Agc 65 in the Year |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 1994 | 1999 | 2004 | 2009 |
| 65 | 14.535 | 13.546 | 12.624 | 11.764 |
| 66 | 16.028 | 15.013 | 14.062 | 13.171 |
| 67 | 17.568 | 16.456 | 15.413 | 14.437 |
| 68 | 19.037 | 17.741 | 16.533 | 15.408 |
| 69 | 20.537 | 19.140 | 17.837 | 16.623 |
| 70 | 22.003 | 20.401 | 18.917 | 17.540 |
| 71 | 23.701 | 21.976 | 20.377 | 18.894 |
| 72 | 25.622 | 23.757 | 22.028 | 20.425 |
| 73 | 27.648 | 25.635 | 23.770 | 22.040 |
| 74 | 29.720 | 27.557 | 25.552 | 23.692 |
| 75 | 32.318 | 30.118 | 28.068 | 26.157 |
| 76 | 34.988 | 32.607 | 30.387 | 28.319 |
| 77 | 38.607 | 36.162 | 33.872 | 31.727 |
| 78 | 42.918 | 40.404 | 38.037 | 35.809 |
| 79 | 47.847 | 45.273 | 42.837 | 40.532 |
| 80 | 53.347 | 50.732 | 48.246 | 45.881 |
| 81 | 59.374 | 56.750 | 54.242 | 51.844 |
| 82 | 65.892 | 63.298 | 60.806 | 58.412 |
| 83 | 71.403 | 68.592 | 65.892 | 63.298 |
| 84 | 78.416 | 75.710 | 73.097 | 70.574 |
| 85 | 84.495 | 81.579 | 78.763 | 76.045 |
| 86 | 91.282 | 88.132 | 85.090 | 82.153 |
| 87 | 101.327 | 98.323 | 95.409 | 92.581 |
| 88 | 113.153 | 110.352 | 107.621 | 104.957 |
| 89 | 123.646 | 120.585 | 117.601 | 114.690 |
| 90 | 138.350 | 135.605 | 132.915 | 130.277 |

TABLE 21-Continued

| Age | Male Issue Age 65 in the Year |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 1994 | 1999 | 2004 | 2009 |
| 91 | 150.708 | 147.717 | 144.787 | 141.914 |
| 92 | 168.078 | 165.572 | 163.103 | 160.671 |
| 93 | 182.385 | 179.665 | 176.987 | 174.348 |
| 94 | 197.701 | 194.754 | 191.850 | 188.989 |
| 95 | 219.989 | 217.798 | 215.628 | 213.481 |
| 96 | 236.375 | 234.021 | 231.690 | 229.382 |
| 97 | 252.134 | 249.622 | 247.136 | 244.675 |
| 98 | 276.012 | 274.635 | 273.265 | 271.901 |
| 99 | 291.221 | 289.768 | 288.322 | 286.883 |
| 100 | 306.321 | 304.793 | 303.272 | 301.759 |
| 101 | 333.461 | 333.461 | 333.461 | 333.461 |
| 102 | 350.330 | 350.330 | 350.330 | 350.330 |
| 103 | 368.542 | 368.542 | 368.542 | 368.542 |
| 104 | 387.855 | 387.855 | 387.855 | 387.855 |
| 105 | 407.224 | 407.224 | 407.224 | 407.224 |
| 106 | 425.599 | 425.599 | 425.599 | 425.599 |
| 107 | 441.935 | 441.935 | 441.935 | 441.935 |
| 108 | 457.553 | 457.553 | 457.553 | 457.553 |
| 109 | 473.150 | 473.150 | 473.150 | 473.150 |
| 110 | 486.745 | 486.745 | 486.745 | 486.745 |
| 111 | 496.356 | 496.356 | 496.356 | 496.356 |
| 112 | 500.000 | 500.000 | 500.000 | 500.000 |
| 113 | 500.000 | 500.000 | 500.000 | 500.000 |
| 114 | 500.000 | 500.000 | 500.000 | 500.000 |
| 115 | 500.000 | 500.000 | 500.000 | 500.000 |
| 116 | 500.000 | 500.000 | 500.000 | 500.000 |
| 117 | 500.000 | 500.000 | 500.000 | 500.000 |
| 118 | 500.000 | 500.000 | 500.000 | 500.000 |
| 119 | 500.000 | 500.000 | 500.000 | 500.000 |
| 120 | 1000.000 | 1000.000 | 1000.000 | 1000.000 |

TABLE 22
Generation Mortality Rates per 1,000 for Issues of 1994, 1999, 2004, and 2009 at Female Issue Age 65 in the Indicated Year
Based upon Gam-94 Static Table with Full Generation and Projection Scale aa

| Age | Fermale Issue Age 65 in the Year |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 1994 | 1999 | 2004 | 2009 |
| 65 | 8.636 | 8.422 | 8.214 | 8.010 |
| 66 | 9.646 | 9.407 | 9.174 | 8.947 |
| 67 | 10.657 | 10.393 | 10.136 | 9.885 |
| 68 | 11.587 | 11.301 | 11.021 | 10.748 |
| 69 | 12.457 | 12.148 | 11.848 | 11.554 |
| 70 | 13.390 | 13.059 | 12.736 | 12.420 |
| 71 | 14.423 | 13.995 | 13.580 | 13.178 |
| 72 | 15.825 | 15.356 | 14.901 | 14.459 |
| 73 | 17.342 | 16.743 | 16.165 | 15.607 |
| 74 | 19.132 | 18.472 | 17.835 | 17.219 |
| 75 | 20.935 | 20.111 | 19.319 | 18.559 |
| 76 | 23.183 | 22.271 | 21.394 | 20.552 |
| 77 | 26.073 | 25.173 | 24.304 | 23.465 |
| 78 | 28.958 | 27.959 | 26.994 | 26.062 |
| 79 | 32.050 | 30.944 | 29.876 | 28.845 |
| 80 | 35.456 | 34.232 | 33.051 | 31.910 |
| 81 | 39.280 | 37.924 | 36.615 | 35.351 |
| 82 | 43.620 | 42.115 | 40.661 | 39.258 |
| 83 | 48.341 | 46.673 | 45.062 | 43.507 |
| 84 | 53.360 | 51.518 | 49.740 | 48.024 |
| 85 | 60.057 | 58.276 | 56.549 | 54.873 |
| 86 | 67.819 | 66.140 | 64.503 | 62.907 |
| 87 | 76.931 | 75.405 | 73.909 | 72.443 |
| 88 | 85.558 | 83.860 | 82.197 | 80.566 |
| 89 | 97.317 | 95.866 | 94.437 | 93.029 |
| 90 | 107.852 | 106.244 | 104.660 | 103.099 |

TABLE 22-Continued

| Age | Female Issue Age 65 in the Year |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 1994 | 1999 | 2004 | 2009 |
| 91 | 119.076 | 117.301 | 115.552 | 113.829 |
| 92 | 130.911 | 128.959 | 127.036 | 125.142 |
| 93 | 147.431 | 145.962 | 144.508 | 143.069 |
| 94 | 161.050 | 159.446 | 157.858 | 156.286 |
| 95 | 175.358 | 173.612 | 171.882 | 170.171 |
| 96 | 190.350 | 188.454 | 186.577 | 184.719 |
| 97 | 212.734 | 211.672 | 210.616 | 209.565 |
| 98 | 229.993 | 223.845 | 227.703 | 226.567 |
| 99 | 248.126 | 246.888 | 245.656 | 244.430 |
| 100 | 266.915 | 265.583 | 264.258 | 262.339 |
| 101 | 296.629 | 296.629 | 296.629 | 296.629 |
| 102 | 317.093 | 317.093 | 317.093 | 317.093 |
| 103 | 338.505 | 338.505 | 338.505 | 338.505 |
| 104 | 361.016 | 361.016 | 361.016 | 361.016 |
| 105 | 383.597 | 383.597 | 383.597 | 383.597 |
| 106 | 405.217 | 405.217 | 405.217 | 405.217 |
| 107 | 424.846 | 424.846 | 424.846 | 424.846 |
| 108 | 444.368 | 444.368 | 444.368 | 444.368 |
| 109 | 464.469 | 464.469 | 464.469 | 464.469 |
| 110 | 482.325 | 482.325 | 482.325 | 482.325 |
| 111 | 495.110 | 495.110 | 495.110 | 495.110 |
| 112 | 500.000 | 500.000 | 500.000 | 500.000 |
| 113 | 500.000 | 500.000 | 500.000 | 500.000 |
| 114 | 500.000 | 500.000 | 500.000 | 500.000 |
| 115 | 500.000 | 500.000 | 500.000 | 500.000 |
| 116 | 500.000 | 500.000 | 500.000 | 500.000 |
| 117 | 500.000 | 500.000 | 500.000 | 500.000 |
| 118 | 500.000 | 500.000 | 500.000 | 500.000 |
| 119 | 500.000 | 500.000 | 500.000 | 500.000 |
| 120 | 1000.000 | 1000.000 | 1000.000 | 1000.000 |

$$
\begin{equation*}
q_{65+n}^{1994+n+t}=q_{65+n}^{1994} \times\left(1-A A_{65+n}\right)^{n+t} \tag{K}
\end{equation*}
$$

where
$n$ attained age less 65
$t$ issue year less 1994 .

## B. The 1994 Group Annuity Reserving Table

As initially indicated, the Task Force was charged with recommending a new Group Annuity Mortality Valuation Standard that would be suitable for calculating group annuity valuation reserves. By definition, this new standard shall be known as the 1994 Group Annuity Reserving (GAR-94) Table.

The GAR-94 Table combines three components:

1. Projection Scale AA, whose mortality improvement factors are shown in Table 15, for projecting mortality beyond the year 1994
2. The GAM-94 Static Table, whose $q_{x}$ values are shown in Table 18
3. All the generation tables produced by multiplying the Projection Scale AA mortality improvement factors by the respective GAM-94 Static Table $q_{x}$ values (of which examples for issue age 65 and certain issue years are shown in Table 21 and Table 22).
The complete GAR-94 Table appears as Table 1 in the Executive Summary.

Note that this approach implies that a different set of mortality rates should be used for each different issue year for a specific issue age. However, it also implies that the same mortality rate should be used when the attained age and issue year offsets are the same. Thus, the mortality rate for issue age 65 in 1994 five years after issue is the same as that for issue age 70 in 1999 (and issue age 67 in 1996 two years after issue).

## C. Financial Values Using the GAR-94 Tables

Table 23 shows and compares the life annuity net single premiums for an annuity due of $\$ 1$ per year, payable monthly, for various issue ages based upon GAM-83 mortality and $7 \%$ level interest and for various issue ages and issue years based upon GAR-94 mortality and the same interest rate, In this table, on the GAR-94 basis, the net single premiums are significantly greater (at least $3 \%$ ) for male issue ages $40-90$ in 1994, 1999 and 2004, and $35-90$ in 2009. At male issue age 65 , for these issue years, the percentages are $6.2 \%, 7.7 \%, 9.1 \%$, and $10.4 \%$, respectively. Female issue ages show no significantly greater net single premiums in 1994 and only for issue age 75 in 1999. Because of the improving mortality, issues in 2004 show

TABLE 23
Life Annuity Net Single Premiums Assuming 7\% Level Interest Rate and Mortality from GaR-94 Table versus Mortality from GaM-83 Table


TABLE 23-Continued

significantly greater net single premiums for ages 70-80. At female issue age 65 , for these issue years, the percentages are $0.8 \%, 1.3 \%, 1.8 \%$, and $2.3 \%$, respectively. The progression of ratios of GAR-94 net single premiums for males to those of the GAM-83 is relatively smooth and increasing until issue age 80 and then proceeds to decrease. Such ratios for females begin to materially decrease starting at issue age 80 but then show a sharp increase at issue age 95 .

The mortality rate equivalence under GAR-94 mortality outlined above implies the same equivalence between net single premiums and reserves. Thus the net single premium for issue age 70 in 1999 is the same as the fifth-year reserve for issue age 65 in 1994, while the net single premium for issue age 75 in 2004 is the same as the tenth year reserve for issue age 65 in 1994 (and the fifth-year reserve for issue age 70 in 1999). Thus, for male issue age 65 in 1994, the percentage increases of the initial and fifth-, tenth-, and fifteenth-year reserves relative to the GAM-83 values are 6.2, $9.4,12.3$, and 13.9 , respectively. For female issue age 65 , the values are 0.8 , $2.5,4.2$, and 4.9 .

This analysis further confirms the need for a new reserve valuation standard to replace the GAM-83.

## 7. CONCLUSION

Present-day mortality levels have eroded the margins built into the 1983 Group Annuity Mortality Tables. They are no longer adequate for valuation purposes. Therefore, the Task Force has developed the 1994 Group Annuity Reserving Table presented in this report. The Task Force recommends that this new table replace the 1983 Group Annuity Mortality Table for use as a Valuation Mortality Standard.

## A. Potential Uses of the New Standard

This report does not preclude other uses of the new standard, as long as the user clearly understands the development and coinciding limitations (margins, annual mortality improvement, and so on) of this new standard. Other reports that will be released will discuss additional uses of the tables presented in this report.

## B. Acknowledgment

The Chair would like to thank each Task Force member as well as their employers for the time and unceasing effort devoted to this endeavor. This
report and incorporated recommendation would not be as complete or as well-defined without the effort extended by each of these Task Force members. In addition, the Society of Actuaries staff and especially our assigned actuarial liaison's support have been invaluable throughout the process.

The Task Force thanks the following individuals for their written comments on the Exposure Draft of this report: Robert L. Brown and Shaun Wang, William H. Crosson, Harvey Fishman and Zachary Granovetter, G. Thomas Mitchell, Michael Mudry, Bruce E. Nickerson, Owen A. Reed, Robert Stalzer, David A. Wiener, William S. Wright, and especially Walter J. McLaughlin. Their comments only served to improve the final report.

Special thanks go to Charles F. Brown and Marian Rivera from Bankers Security Life for their tireless assistance in developing the Exposure Draft and presenting it in its final format. To try and name each and every individual who helped the Task Force would be to forget someone who should not be forgotten. However, we want to acknowledge the many helpful suggestions we received during the development of the tables. Thus, our appreciation to all, even those not named, is total.

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## DISCUSSION OF PRECEDING PAPER

## JACQUES F. CARRIERE:

The purpose of this discussion is to present a parametric model or mathematical formula that will explain the pattern of mortality for the male and female GAM-94 tables. In my opinion, parametric formulas for mortality rates are always preferable to tabular rates, if the formula gives a good fit. Generally, the parametric approach always yields very smooth rates.

Before proceeding, it is instructive to plot the crude and graduated rates and examine the pattern of mortality in the GAM-94 tables. Consider the function

$$
y_{x}=\log _{e}\left\{-\log _{e}\left(1-q_{x}\right)\right\}
$$

where $q_{x}$ is the mortality rate for a life aged $x$. Remember that $\mu_{x+0.5} \approx \exp \left(y_{x}\right)$, where $\mu_{x+0.5}$ is a force of mortality. Let $y_{x}^{\text {crude }}$ be the value based on the crude or ungraduated rates found in Table 12 and let $y_{x}^{\text {grad }}$ be the value based on the graduated rates found in Table 13. Figure 1 presents plots of $y_{x}^{\text {crude }}, y_{x}^{\text {grad }}$, and $y_{x}^{\text {crude }}-y_{x}^{\text {grad }}$ versus $x=1,2, \ldots, 108$ for both the female and male rates. The graduated values $y_{x}^{\text {grad }}$ are based on the classical Karup-King method, which did an excellent job.
The parametric models for the male and female rates have nine parameters, denoted as $\theta=\left(\psi_{1}, m_{1}, \sigma_{1}, m_{2}, \sigma_{2}, \psi_{3}, m_{3}, \sigma_{3}\right)$. The parametric formula for the male rates is

$$
y_{x}^{f o r m}(\theta)=\psi_{1} \times r_{x}\left(m_{1}, \sigma_{1}\right)+\psi_{2}+s_{x}\left(m_{2}, \sigma_{2}\right)+\psi_{3} \times t_{x}\left(m_{3}, \sigma_{3}\right),
$$

while the formula for the female rates is

$$
y_{x}^{f o r m}(\theta)=\psi_{1} \times r_{x}\left(m_{1}, \sigma_{1}\right)+\psi_{2} \times r_{x}\left(m_{2}, \sigma_{2}\right)+\psi_{3} \times t_{x}\left(m_{3}, \sigma_{3}\right),
$$

where

$$
\begin{aligned}
& r_{x}(m, \sigma)=\exp \left\{-\left(\frac{x}{m}\right)^{m / \sigma}\right\} \\
& s_{x}(m, \sigma)=1-\exp \left\{-\left(\frac{x}{m}\right)^{-m / \sigma}\right\} \\
& t_{x}(m, \sigma)=\frac{1}{\sigma}\left(\frac{x}{m}\right)^{(m / \sigma)-1} \exp \left\{-\left(\frac{x}{m}\right)^{m / \sigma}\right\} .
\end{aligned}
$$

FIGURE 1
a Comparison of the Crude and Graduated Rates from the GaM-94.
The Horizontal Axis on All Graphs is the attained Age $x$.
The First Column of Graphs Gives $y_{x}=\log _{f}\left(-\log _{e}\left(1-q_{x}\right)\right)$ Versus $x$. The Second Column Shows $y_{x}^{\text {crude }-y_{s}^{\text {grad }} \text {, the Difference in Crude and Graduated Rates. }}$





Notice that the male formula is a function of $r_{x}, s_{x}$, and $t_{x}$, while the female formula is a function of $r_{x}, r_{x}$, and $t_{x}$, so the male and female formulas are different. Both the male and female formulas were specifically created for the GAM-94 tables. Thus, applying these formulas on other tables may or may not yield good results. To estimate the parameters $\theta$, we minimized the loss function

$$
L(\theta)=\sum_{x=1}^{108}\left[y_{x}^{\text {form }}(\theta)-y_{x}^{\text {crude }}\right]^{2} .
$$

The parameter estimates are shown in Table 1. Figure 2 presents plots of $y_{x}^{\text {crude }}, y_{x}^{\text {form }}$, and $y_{x}^{\text {crude }}-y_{x}^{\text {form }}$ versus $x=1,2, \ldots, 108$ for both the female and male rates. The graduated values, $y_{x}^{\text {form, }}$, based on our formulas did an excellent job. Comparison of Figure 1 and Figure 2 shows $y_{x}^{\text {fom }}$ is smoother than $y_{x}^{\text {grad }}$, but the Karup-King values, $y_{x}^{\text {grad }}$, are slightly better fitting than $y_{x}^{\text {form. }}$.

TABLE 1
Parameter Estimates for the Parametric Models

| Model | $\psi_{1}$ | $m_{1}$ | $\sigma_{1}$ | $\psi_{2}$ | $m_{2}$ | $\sigma_{2}$ | $\psi_{3}$ | $m_{3}$ | $\sigma_{3}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Male | 9.800 | 82.09 | 22.15 | -2.428 | 75.88 | 11.67 | 13.82 | 11.20 | 4.882 |
| Female | 188.3 | 96.59 | 59.35 | -180.9 | 97.84 | 62.85 | 12.10 | 9.973 | 4.459 |

FIGURE 2
A Comparison of the Crude and Graduated Rates from the Formulas.
The Horizontal Axis on All Graphs is the Attained Age $x$.
The First Column of Graphs Gives $y_{x}=\log _{e}\left(-\log _{e}\left(1-q_{x}\right)\right\}$ Versus $x$.
The Second Column Shows $y_{x}^{\text {ciude }}-y_{x}^{\text {form }}$, the Difference in Crude and Formula Rates.





## (AUTHOR'S REVIEW OF DISCUSSION)

## LINDSAY J. MALKIEWICH:

The Task Force thanks Dr. Carriere for his thought-provoking discussion on a possible parametric model for explaining mortality patterns. While it is interesting to note that such a model can help to explain some of the GAR-94 results, it seems that an exploration of the model's computational viability would enhance the discussion. In other words, whereas the solution to a one-parameter model is readily developed, additional parameters will surely increase the difficulty of finding such a solution.

In addition to the increased difficulty of solving a multiple-parameter problem, the issue of parameter sensitivity should be addressed. Small changes in the provided formula's coefficients can cause drastic changes in the estimates of the given parameters. Therefore, an easy solution for a model making use of nine parameters would, logically, be quite difficult to find. It would be an interesting addition if the discussion explored various methods that would be used to discover such solutions.

In a related observation, a range of choices was shown for these parameter estimates in Table 1. Some of the differences between a given male versus female parameter are quite large. Therefore, it follows that small changes in the formula's coefficients may yield very different results, given the same parameters or even slightly different ones. Is this situation intended? Based on the discussion alone, it is unclear.

While the above concerns suggest areas in which additional demonstrations and more research could enhance the model, this discussion does serve as a worthwhile start. As stated, the model as presented does help to explain some of the GAR-94 results. The Task Force is, of course, pleased that such a model could do so. Perhaps, additional research could be that much more enlightening.


[^0]:    *Lindsay J. Malkiewich, Chairperson, David B. Berg, Neil J. Broderick, John B. Gould, Edwin C. Hustead, Naftali Teitelbaum, Charles N. Vest, Michael R. Virga, and John A. Luff, SOA Staff Liaison.

