TRANSACTIONS OF SOCIETY OF ACTUARIES 1963 REPORTS

II. 1959–62 JUVENILE EXTENSION OF THE 1955–60 BASIC TABLES

The 1955-60 Select Basic Tables (TSA, 1962 Reports, 46) did not include rates for issue ages under 10. Data at these ages were first contributed for the year between 1959 and 1960 policy anniversaries, and their volume was insufficient for the development of graduated rates when the 1955-60 Tables were constructed. Since then, the experience of two additional years has become available, and the extension described in this report is based on the experience between 1959 and 1962 policy anniversaries. Data were submitted at issue ages 0, 1, 2-4, and 5-9 for the first fifteen policy years. The names of the seventeen contributing companies and their proportionate contributions to the total exposures are given in Table A at the end of this report.

Separate tables were constructed for male lives, female lives, and both sexes combined. The experience of the fifteen companies which subdivided their data by sex is included in all three tables. The experience of the remaining two companies is included only in the table for both sexes combined. In order to obtain as large a volume of data as possible, medical and nonmedical issues were combined. Almost 80 per cent of the exposure was on nonmedical issues.

The first problem was the choice of an appropriate select period, if any. The fifteen-year select period used in the 1955–60 Select Basic Tables did not seem appropriate for juvenile ages. The results of two tests are presented in Table 1. The upper portion shows the crude mortality rates for attained ages 1–17 at various durations. For the purpose of computing attained ages, ages 3 and 7 were used to represent issue age groups 2–4, and 5–9, respectively.

To the extent that selection by the insurance company is effective, the mortality rate for a given attained age will increase with duration. The upper portion of Table 1 does not indicate a consistent pattern of effective selection. Where there is an increase with duration, it is generally small; in many instances there is actually a decrease.

The lower portion of Table 1 shows mortality rates at attained ages 1-7 with no select period, a one-year select period, and a five-year select period. A dash indicates that data for a meaningful comparison were not available. The maximum difference, which occurs at age 7, is .07 per 1,000 between the rate with no select period and that with a five-year select period.

TABLE 1

TESTS FOR SELECT PERIOD JUVENILE TABLES MALE AND FEMALE LIVES COMBINED CRUDE MORTALITY RATES PER 1,000 COMPARISON OF ATTAINED AGE RATES AT DIFFERENT DURATIONS*

At- tained Age	Issue Age	Dura- tion	Rate	At- tained Age	Issue Age	Dura- tion	Rate	At- tained Age	Issue Age	Dura- tion	Rate
1	1	1	. 66	7	7	1	.27	12	7	6	. 18
• • • • •	Ő	$\frac{1}{2}$.77	••••	73	1 5 7	.25		3	10	.44
	_				1	7	.39		1	12	.41
2	1	$\frac{2}{3}$. 68		0	8	.41		0	13	. 28
	0	3	.66								
				8	7	2	. 26	13 .	7	7	. 31
3	3	1	.47		3	6	.13		3	11	. 37
	1	3	. 59	4	!	8	.40		1	13	33
	0	4	. 54		0	9	.31	İ	0	14	.38
4	3	2	. 50	9	7	3	. 21	14	7	8	.35
	1	4 5	.40		7 3 1	3 7 9	. 28	ļ	3	12	. 38
	0	5	. 79			9	.24		1	14	.35
					0	10	.26		0	15	.37
5	3	3 5	.48		_				_		
	1	5	.56 .41	10	7 3 1	4	. 18	15	7	9	.73
	0	6	.41		3	8	.21		3	13	.43
						10	.21		1	15	. 55
6	3	4	.40 .54 .37		0	11	. 31		-		
		6 7	.54		-			16	7	10	.47
	0	1	.37	11	7 3 1	5 9	.23		3	14	. 79
				1	3	9	.34	17	-		15
	1					11	.32	17	7	11	.65
					0	12	. 20		3	15	. 52

COMPARISON OF ULTIMATE RATES WITH DIFFERENT SELECT PERIODS

SELECT PERIOD						
0 Years	1 Year	5 Years				
.74	. 77					
. 53	. 56					
.46		.41				
. 42		.42				
. 33	. 36	. 40				
	0 Years .74 .53 .46 .42	0 Years 1 Year . 74 . 77 . 53 . 56 . 46				

* "Age" means age nearest birthday.

In view of these results, it was decided that mortality rates on an attained-age basis, disregarding duration since issue, would be satisfactory for issue ages under 10. Accordingly, all the data were combined on this basis, the attained ages for issue age groups 2-4 and 5-9 being calculated as indicated above. A previous report on juvenile mortality (TASA, XXXVIII, 205-7 and 587-89) also concluded that the effect of selection could be disregarded at juvenile ages.

Data contributed on an age last birthday basis were excluded at attained ages 0 and 1 (this was also done in Table 1); at higher attained ages, their inclusion had no effect on the crude rates.

It should be noted that issues at age 0 represent an average age somewhat greater than 0. When the X17 Mortality Table (which was later developed into the 1958 CSO Table) was constructed, sampling in one large company gave an average age of 55 days, and the rate was adjusted to a true age 0 basis in that Table (TSA, IX, 8–9). A similar adjustment was not considered necessary or desirable for the present purpose as the Basic Tables are intended to facilitate the analysis of mortality trends and to provide companies with a tool for mortality comparisons.

At ages 1–21, the crude rates were graduated by use of a Whittaker-Henderson second difference A Formula, with a = .5 for males and both sexes combined, and a = 2 for females. At age 0 the following procedure was used:

- 1. In the male table, the graduated rate was taken as equal to the crude rate.
- 2. The crude rate for females was higher than that for males. This was believed to be due to large claims because the average female claim at age 0 (\$2,210) was higher than at any other age; it was also higher in relation to the average male claim at age 0 (\$2,385) than at the other juvenile ages. Accordingly, the average female claim was adjusted to an amount bearing the same ratio to the average male claim at age 0 as the ratio (82.2 per cent) of the average female to the average male claim for all juvenile ages combined. The female claims reduced in this manner were then used to calculate the graduated mortality rate for females at age 0.
- 3. In the table for both sexes combined, female claims as adjusted in (2) were combined with the unadjusted male claims to produce the graduated rate for age 0.

At age 4, one company whose contribution was not subdivided by sex reported a claim for \$200,000. This claim caused a relatively high crude rate of .63 for both sexes combined; the graduated rate produced by the Whittaker-Henderson formula was .56. Both the crude and the graduated rate were higher than the corresponding rates for male lives. Accordingly, the graduated rate at age 4 for both sexes combined was adjusted to .53. Table 2 shows the crude and graduated rates; U.S. population rates are shown for comparison. For convenience in calculating expected deaths for the annual study of Standard Ordinary mortality, which calls for data on a fifteen-year select and ultimate basis, Table 3 shows the graduated rates in select format. Table 4 shows a test of the graduation.

In order to maintain fidelity to the underlying data, no attempt was made to blend the graduated rates for ages 15-21 with the 1955-60 Ultimate Basic Tables, which have somewhat higher rates at these ages.

TABLE 2

1959-62 JUVENILE AGGREGATE BASIC TABLE CRUDE AND GRADUATED RATES PER 1,0001961 U.S. POPULATION RATES SHOWN FOR COMPARISON

Attained Age Nearest	Male and Fe- male Lives Combined		MALE LIVES		FEMALE LIVES		U.S. 1961 Official Life Tables	
BIRTHDAY	Crude	Grad- uated	Crude	Grad- uated	Crude	Grad uated	White Males	White Females
0 1 2 3 4	1.83 .74 .67 .53 .63	1.78 .74 .65 .59 .53	1.85 .85 .79 .59 .52	1.85 .86 .74 .63 .54	1.89 .54 .44 .47 .38	1.68 .52 .48 .43 .39	25.4 1.5 1.0 .8 .7	19.3 1.3 .8 .6 .5
5 6 7 8 9	.46 .42 .33 .27 .24	. 48 . 41 . 34 . 28 . 24	.50 .47 .38 .29 .26	.49 .44 .37 .31 .27	.41 .29 .27 .25 .23	.35 .31 .28 .25 .23	.7 .6 .5 .4 .4	.4 .4 .3 .3
10 11 12 13 14	.23 .25 .27 .35 .36	. 23 . 24 . 28 . 34 . 42	.28 .30 .32 .37 .42	.27 .29 .32 .38 .46	.15 .17 .20 .29 .26	.22 .23 .25 .27 .30	.4 .4 .5 .6 .7	.3 .3 .3 .3 .3
15 16 17 18 19	.63 .57 .60 .57 .92	. 53 . 58 . 61 . 68 . 81	. 66 . 58 . 69 . 80 1.08	.56 .63 .72 .85 .99	.57 .46 .19 .26 .21	.32 .32 .32 .32 .32 .32	.9 1.1 1.2 1.3 1.4	.4 .5 .5 .5 .5
20 21	.86 .89	.87 .91	1.05 .95	1.02 1.00	. 33 . 43	.33 .35	1.5 1.6	.5

TABLE 3

1959-62 JUVENILE BASIC TABLES GRADUATED MORTALITY RATES PER 1,000

	ISSUE AGE GROUP (NEAREST BIRTHDAY)							
POLICY YEAR	0	1	2-4	5-9				
	MALE AND FEMALE LIVES COMBINED							
1 2 3 4 5 5 6 7 7 8 9 10 11 11 12 13 14 15	1.78.74.65.59.53.48.41.34.24.23.24.23.24.28.34.42	$\begin{array}{c} .74\\ .65\\ .59\\ .53\\ .48\\ .41\\ .34\\ .28\\ .24\\ .23\\ .24\\ .23\\ .24\\ .28\\ .34\\ .42\\ .53\end{array}$	$\begin{array}{c} .59\\ .53\\ .48\\ .41\\ .34\\ .28\\ .24\\ .23\\ .24\\ .23\\ .24\\ .28\\ .34\\ .42\\ .53\\ .58\\ .61\end{array}$.34 .28 .24 .23 .24 .28 .34 .42 .53 .58 .61 .68 .81 .87 .91				
		Male	LIVES	·				
1	1.85.86.74.63.54.49.44.37.31.27.27.29.32.38.46	.86 .74 .63 .54 .49 .44 .37 .31 .27 .27 .27 .29 .32 .38 .46 .56	.63 .54 .49 .44 .37 .27 .27 .29 .32 .38 .46 .56 .63 .72	.37 .21 .27 .29 .32 .38 .46 .56 .63 .72 .85 .99 1.02 1.00				
		Female	Lives					
1	$1.68 \\ .52 \\ .48 \\ .43 \\ .39 \\ .35 \\ .31 \\ .28 \\ .25 \\ .23 \\ .22 \\ .23 \\ .25 \\ .27 \\ .30$.52 .48 .43 .39 .35 .31 .28 .25 .23 .22 .23 .22 .23 .25 .27 .30 .32	.43 .39 .35 .31 .28 .25 .23 .22 .23 .25 .27 .30 .32 .32 .32	.28 .25 .23 .22 .23 .25 .25 .27 .30 .32 .32 .32 .32 .32 .32 .33 .35				

TABLE 4

A	Male and Lives Co		Male I	IVES	Female Lives		
Attained Age Nearest Birthday	Actual Claims	Ratio of Actual to Tabular Claims	Actual Claims	Ratio of Actual to Tabular Claims	Actual Claims	Ratio of Actual to Tabular Claims	
$\begin{array}{c} 0, \\ 1, \\ 2, \\ 3, \\ 4, \\ 5, \\ 6, \\ 7, \\ 8, \\ 9, \\ 10, \\ 11, \\ 12, \\ 13, \\ 14, \\ 15, \\ 16, \\ 17, \\ 18, \\ 19, \\ 20, \\ 21, \\ 0-5, \\ 6-10, \\ 11-15, \\ 16-21 \\ \end{array}$		$\begin{array}{c} 102.7\%\\ 99.4\\ 102.8\\ 90.2\\ 119.1\\ 95.8\\ 101.4\\ 97.3\\ 100.9\\ 97.3\\ 100.9\\ 101.4\\ 103.2\\ 97.6\\ 101.6\\ 86.3\\ 118.8\\ 98.9\\ 98.2\\ 83.4\\ 113.1\\ 98.6\\ 97.5\\ 100.7\\ 99.3\\ 100.2\\ 98.6 \end{array}$	\$ 543,751 393,739 372,670 394,788 343,650 325,641 280,178 332,249 218,500 161,861 151,675 150,974 154,338 169,265 174,878 169,265 174,878 169,265 174,878 169,265 174,878 169,265 174,878 169,265 174,878 169,265 174,878 169,500 73,900 72,099 68,530 2,374,239 1,144,463 809,955 461,793	$\begin{array}{c} 100.0\%\\ 98.9\\ 106.1\\ 93.3\\ 95.5\\ 103.0\\ 106.6\\ 102.1\\ 93.6\\ 94.8\\ 105.0\\ 104.3\\ 99.9\\ 97.2\\ 91.0\\ 117.4\\ 92.7\\ 96.5\\ 93.9\\ 109.6\\ 103.2\\ 94.7\\ 99.2\\ 100.7\\ 100.9\\ 97.9\end{array}$	$\begin{array}{c} 331,550\\ 152,371\\ 132,500\\ 190,816\\ 157,559\\ 167,189\\ 114,600\\ 130,250\\ 107,000\\ 82,600\\ 46,000\\ 47,568\\ 55,200\\ 78,407\\ 67,000\\ 74,000\\ 74,000\\ 74,000\\ 74,000\\ 10,527\\ 14,000\\ 9,000\\ 7,000\\ 11,000\\ 15,000\\ 11,31,985\\ 480,450\\ 322,175\\ 90,527\\ \end{array}$	$\begin{array}{c} 112.4\%\\ 104.6\\ 92.2\\ 108.8\\ 98.0\\ 116.3\\ 94.3\\ 94.7\\ 100.8\\ 99.3\\ 68.3\\ 73.3\\ 80.6\\ 107.7\\ 87.6\\ 177.5\\ 142.8\\ 60.3\\ 79.8\\ 65.4\\ 99.7\\ 124.3\\ 106.4\\ 99.3\\ 106.4\\ 99.3\\ 99.8\\ \end{array}$	
	\$7,792,893		\$4,790,450		\$2,031,137	101.5%	

Test of Graduation of 1959-62 Juvenile Aggregate Tables Amounts to Nearer Dollar

TABLE A

CONTRIBUTING COMPANIES PROPORTION OF TOTAL EXPOSURES CONTRIBUTED BY EACH

	Firs Policy		First 15 Policy Years by Sex				
Company	Medical Issues	Non- medical	Med Issu		Nonmedical Issues		
	Issues	Issues	Male	Female	Male	Female	
Northwestern Mutual New York Life Mutual Benefit	22.5% 21.8 11.7	10.9%	18.8% 18.2	8.9% 8.6	7.4%	4.7%	
Metropolitan Penn Mutual	8.0 6.2	28.8 1.4	1.9 5.3	.8 2.4	12.8 1.0	9.0	
Prudential	5.8	29.2	5.4	1.7	20.9	11.4	
John Hancock	4.5	13.2	3.6	1.9	8.4	6.1	
Massachusetts Mutual	4.3	2.2	3.7	1.6	1.4	.9	
Sun Life, Canada	2.8	2.9	2.4	1.0	2.0	1.2	
Equitable, N.Y.*	2.6 2.3	1.8	2.4 1.9	.8	1.3	.7 .2 .2 .2 2.3	
Connecticut General	2.3	.5	1.9	.9	.3 .5	. 4	
New England Life*	1.4	.6 .5	1.0	.9 .7 .5 .5	.5	. 2	
Mutual Life, N.Y.	1.4	5.4	1.1	.5	.4 3.7	2 3	
Provident Mutual	1.2	.2	1.1		0.7	2.0	
Aetna	.9	1.4		4	1.0	.6	
Lincoln National	.7	.8	.6	.4 .3	.5	.3	
Total	100.0%	100.0%	69.0%	31.0%	61.6%	38.4%	

* Only data for issue age group 5-9 used.