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Option Pricing Without Tears: Valuing Equity-Linked Death Benefits

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Let T_x denote the time-until-death random variable for a life aged x .

Let $S(t)$ be the time- t price of a stock or mutual fund.

Consider death benefits that depend on the value of $S(T_x)$, i.e., consider $b(S(T_x))$ for some function $b(\cdot)$.

Examples:

$$b(s) = \text{Max}(s, K)$$

$$b(s) = (s - K)_+$$

Problem: Evaluate

$$E[e^{-\delta T_x} b(S(T_x))]$$

where the expectation is taken with respect to an *appropriate* probability distribution and δ is a continuously compounded interest rate.

$$\begin{aligned}
& \mathbf{E}[e^{-\delta T_x} \mathbf{b}(S(T_x))] \\
&= \mathbf{E}[\mathbf{E}[e^{-\delta T_x} \mathbf{b}(S(T_x)) \mid T_x]] \\
&= \int_0^{\infty} \mathbf{E}[e^{-\delta t} \mathbf{b}(S(t)) \mid T_x = t] f_{T_x}(t) dt \\
&= \int_0^{\infty} \mathbf{E}[e^{-\delta t} \mathbf{b}(S(t))] f_{T_x}(t) dt
\end{aligned}$$

if T_x is independent of $\{S(t)\}$.

So we want to calculate

$$\int_0^{\infty} \mathbb{E}[e^{-\delta t} \mathbf{b}(S(t))] f_{T_x}(t) dt.$$

If

$$f_{T_x}(t) = \sum_j c_j f_{\tau_j}(t),$$

then

$$\int_0^{\infty} \mathbb{E}[e^{-\delta t} \mathbf{b}(S(t))] f_{T_x}(t) dt$$

$$= \sum_j c_j \int_0^{\infty} \mathbb{E}[e^{-\delta t} \mathbf{b}(S(t))] f_{\tau_j}(t) dt$$

$$= \sum_j c_j \mathbb{E}[e^{-\delta \tau_j} \mathbf{b}(S(\tau_j))].$$

The time-until-death density function can be *approximated* by linear combinations of *exponential* density functions

$$f_{T_x}(t) \approx \sum_j c_j \times f_{\tau_j}(t) = \sum_j c_j \times \lambda_j e^{-\lambda_j t}.$$

Thus, our valuation problem becomes finding

$$E[e^{-\delta\tau} b(S(\tau))],$$

where τ is an *exponential* random variable *independent* of $\{S(t)\}$. It turns out to be an *elementary calculus exercise* for *geometric Brownian motion* $\{S(t)\}$.

Let $S(t) = S(0)e^{\mu t + \sigma Z(t)}$, $t \geq 0$,

where $\{Z(t)\}$ is a standard Brownian motion.

Let τ be an independent exponential random variable with mean $1/\lambda$. Then,

$$E[e^{-\delta\tau} b(S(\tau), \text{Max}\{S(t); 0 \leq t \leq \tau\})]$$

$$= \frac{2\lambda}{\sigma^2} \int_0^\infty \left[\int_{-\infty}^m b(S(0)e^x, S(0)e^m) e^{-\alpha x} dx \right] e^{-(\beta-\alpha)m} dm$$

where $\alpha < 0$ and $\beta > 0$ are the solutions of

$$\frac{1}{2}\sigma^2 x^2 + \mu x - (\lambda + \delta) = 0.$$

$$\begin{aligned}
& E[e^{-\delta\tau} b(S(\tau), \text{Max}\{S(t); 0 \leq t \leq \tau\})] \\
&= \frac{2\lambda}{\sigma^2} \int_0^\infty \left[\int_{-\infty}^m b(S(0)e^x, S(0)e^m) e^{-\alpha x} dx \right] e^{-(\beta-\alpha)m} dm.
\end{aligned}$$

Examples:

$$b(s, u) = (s - K)_+ \quad \text{call option}$$

$$b(s, u) = (K - s)_+ \quad \text{put option}$$

$$b(s, u) = u \quad \text{high water mark payoff}$$

Barrier Options

Assume $S(0) < \mathbf{B}$, a barrier.

$$b(s, u) = I(u < \mathbf{B}) \times \pi(s)$$

Up-and-out option

$$b(s, u) = I(u \geq \mathbf{B}) \times \pi(s)$$

Up-and-in option

Useful for incorporating lapses or surrenders.

Assume $S(t) = S(0)e^{X(t)}$, $t \geq 0$, where

$$X(t) = \mu t + \sigma Z(t) + \sum_{j=1}^{N_v(t)} \mathbf{J}_j - \sum_{k=1}^{N_w(t)} \mathbf{K}_k$$

$$f_{\mathbf{J}}(\mathbf{x}) = \sum_{i=1}^m A_i v_i e^{-v_i \mathbf{x}}, \quad \mathbf{x} > 0$$

$$f_{\mathbf{K}}(\mathbf{x}) = \sum_{i=1}^n B_i w_i e^{-w_i \mathbf{x}}, \quad \mathbf{x} > 0$$

$$\sum_{i=1}^m A_i = 1, \quad \sum_{i=1}^n B_i = 1$$

$$S(t) = S(0)e^{X(t)}, \quad t \geq 0$$

Running maximum $M(t) := \text{Max}\{X(u); 0 \leq u \leq t\}$

Running minimum $m(t) := \text{Min}\{X(u); 0 \leq u \leq t\}$

Because $\{X(u)\}$ is a Levy process,

- (i) $M(\tau)$ and $[X(\tau) - M(\tau)]$ are independent random variables,
 - (ii) $[X(\tau) - M(\tau)]$ has the same distribution as $m(\tau)$.
- (i) is hard to prove; (ii) is easy.

In fact, (ii) is true for each fixed t .

$$\begin{aligned} X(t) - M(t) &= X(t) - \text{Max}\{X(s); 0 \leq s \leq t\} \\ &= X(t) + \text{Min}\{-X(s); 0 \leq s \leq t\} \\ &= \text{Min}\{X(t) - X(s); 0 \leq s \leq t\} \\ &= \text{Min}\{X(t - s); 0 \leq s \leq t\} \text{ in distribution} \\ &= \text{Min}\{X(s); 0 \leq s \leq t\} \\ &= m(t) \end{aligned}$$

Running maximum $M(t) := \text{Max}\{X(u); 0 \leq u \leq t\}$

Running minimum $m(t) := \text{Min}\{X(u); 0 \leq u \leq t\}$

(i) $M(\tau)$ and $[X(\tau) - M(\tau)]$ are independent r.v.'s.

(ii) $[X(\tau) - M(\tau)]$ and $m(\tau)$ have the same distribution.

Then,

$$\begin{aligned} E[e^{zX(\tau)}] &= E[e^{z[X(\tau)-M(\tau)+M(\tau)]}] \\ &= E[e^{z[X(\tau)-M(\tau)]}] \times E[e^{zM(\tau)}] \\ &= E[e^{zm(\tau)}] \times E[e^{zM(\tau)}], \end{aligned}$$

which is a version of *Wiener-Hopf factorization*.

Assume

$$X(t) = \mu t + \sigma Z(t) + \sum_{j=1}^{N_v(t)} J_j - \sum_{k=1}^{N_\omega(t)} K_k$$

where

$$f_J(x) = \sum_{i=1}^m A_i v_i e^{-v_i x}, \quad x > 0$$

$$f_K(x) = \sum_{i=1}^n B_i w_i e^{-w_i x}, \quad x > 0$$

Then, $E[e^{zX(t)}] = e^{t\Psi(z)}$ for each $t \geq 0$, with

$$\Psi(z) = \mu z + \frac{1}{2}\sigma^2 z^2 + v \sum_{i=1}^m A_i \frac{z}{v_i - z} - \omega \sum_{i=1}^n B_i \frac{z}{w_i + z}$$

$$\Psi(z) = \mu z + \frac{1}{2}\sigma^2 z^2 + \nu \sum_{i=1}^m A_i \frac{z}{v_i - z} - \omega \sum_{i=1}^n B_i \frac{z}{w_i + z}$$

can be extended by *analytic continuation*.

The moment-generating function of $X(\tau)$ is

$$\begin{aligned} \mathbb{E}[e^{zX(\tau)}] &= \mathbb{E}[\mathbb{E}[e^{zX(\tau)} | \tau]] \\ &= \mathbb{E}[e^{\Psi(z)\tau}] \\ &= \frac{\lambda}{\lambda - \Psi(z)}. \end{aligned}$$

The **zeros** of the RHS are the **poles** of $\Psi(z)$.

The **poles** of the RHS are the **zeros** of $\lambda - \Psi(z)$.

Label the parameters (the poles of $\Psi(z)$) such that

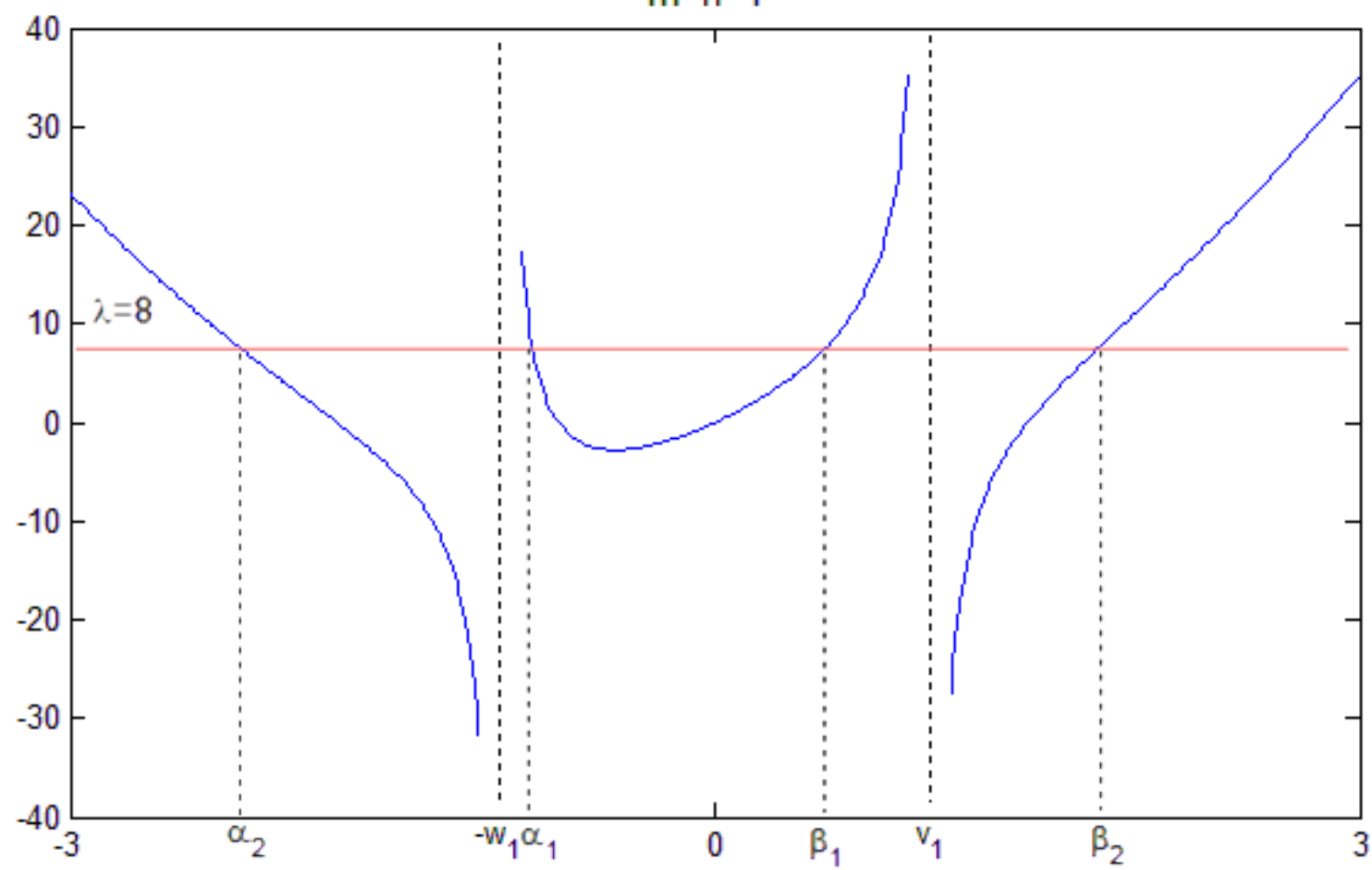
$$v_1 < v_2 < \dots < v_m$$

$$w_1 < w_2 < \dots < w_n$$

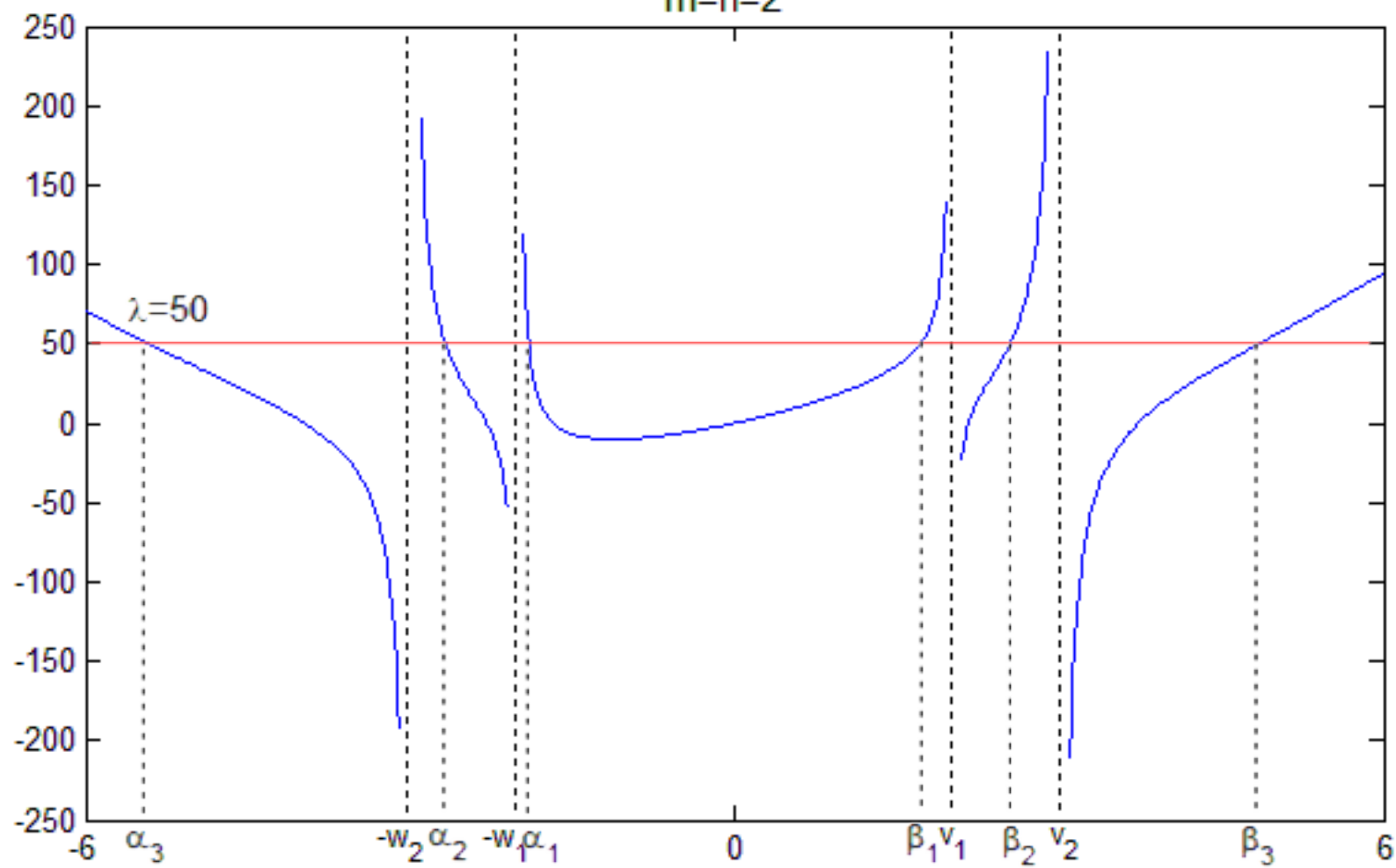
If the weights A 's and B 's are positive, then

$$-\infty < \alpha_{n+1} < -w_n < \dots < -w_1 < \alpha_1 < 0 < \beta_1 < v_1 < \dots < v_m < \beta_{m+1} < \infty$$

$m=n=1$



$m=n=2$



Label the parameters (the poles of $\Psi(z)$) such that

$$v_1 < v_2 < \dots < v_m$$

$$w_1 < w_2 < \dots < w_n$$

If the weights A 's and B 's are positive, then

$$-\infty < \alpha_{n+1} < -w_n < \dots < -w_1 < \alpha_1 < 0 < \beta_1 < v_1 < \dots < v_m < \beta_{m+1} < \infty$$

Wiener-Hopf: $E[e^{zX(\tau)}] = E[e^{zm(\tau)}] \times E[e^{zM(\tau)}]$.

For $z > 0$, $0 < E[e^{zm(\tau)}] \leq 1$. No positive zeros or poles.

For $z < 0$, $0 < E[e^{zM(\tau)}] \leq 1$. No negative zeros or poles.

$$-\infty < \alpha_{n+1} < -w_n < \dots < -w_1 < \alpha_1 < 0 < \beta_1 < v_1 < \dots < v_m < \beta_{m+1} < \infty$$

$$\mathbb{E}[e^{zm(\tau)}] \propto \left(\prod_{j=1}^n (z + w_j) \right) \left(\prod_{j=1}^{n+1} \frac{1}{z - \alpha_j} \right)$$

$$\mathbb{E}[e^{zM(\tau)}] \propto \left(\prod_{j=1}^m (z - v_j) \right) \left(\prod_{j=1}^{m+1} \frac{1}{z - \beta_j} \right)$$

$$\mathbb{E}[e^{zm(\tau)}] = \left(\prod_{j=1}^n \frac{z + w_j}{w_j} \right) \left(\prod_{j=1}^{n+1} \frac{-\alpha_j}{z - \alpha_j} \right)$$

$$\mathbb{E}[e^{zM(\tau)}] = \left(\prod_{j=1}^m \frac{v_j - z}{v_j} \right) \left(\prod_{j=1}^{m+1} \frac{\beta_j}{\beta_j - z} \right)$$

$$\mathbb{E}[e^{\mathbf{z}M(\tau)}]$$

$$= \left(\prod_{j=1}^m \frac{v_j - \mathbf{z}}{v_j} \right) \left(\prod_{j=1}^{m+1} \frac{\beta_j}{\beta_j - \mathbf{z}} \right)$$

$$= \sum_{\mathbf{k}=1}^{m+1} \left(\prod_{j=1}^m \frac{v_j - \beta_{\mathbf{k}}}{v_j} \right) \left(\prod_{j=1, j \neq \mathbf{k}}^{m+1} \frac{\beta_j}{\beta_j - \beta_{\mathbf{k}}} \right) \frac{\beta_{\mathbf{k}}}{\beta_{\mathbf{k}} - \mathbf{z}}.$$

Thus, $f_{M(\tau)}(\mathbf{x}) = \sum_{\mathbf{k}=1}^{m+1} b_{\mathbf{k}} e^{-\beta_{\mathbf{k}} \mathbf{x}}, \quad \mathbf{x} > 0,$

where $b_{\mathbf{k}} = \left(\prod_{j=1}^m \frac{v_j - \beta_{\mathbf{k}}}{v_j} \right) \left(\prod_{j=1, j \neq \mathbf{k}}^{m+1} \frac{\beta_j}{\beta_j - \beta_{\mathbf{k}}} \right) \beta_{\mathbf{k}}$

$$\mathbb{E}[e^{\mathbf{z}m(\tau)}]$$

$$= \left(\prod_{j=1}^n \frac{\mathbf{z} + \mathbf{w}_j}{\mathbf{w}_j} \right) \left(\prod_{j=1}^{n+1} \frac{-\alpha_j}{\mathbf{z} - \alpha_j} \right)$$

$$= \sum_{\mathbf{k}=1}^{n+1} \left(\prod_{j=1}^n \frac{\alpha_{\mathbf{k}} + \mathbf{w}_j}{\mathbf{w}_j} \right) \left(\prod_{j=1, j \neq \mathbf{k}}^{n+1} \frac{-\alpha_j}{\alpha_{\mathbf{k}} - \alpha_j} \right) \frac{-\alpha_{\mathbf{k}}}{\mathbf{z} - \alpha_{\mathbf{k}}}$$

Thus, $f_{m(\tau)}(\mathbf{x}) = \sum_{\mathbf{k}=1}^{n+1} \mathbf{a}_{\mathbf{k}} e^{-\alpha_{\mathbf{k}} \mathbf{x}}, \quad \mathbf{x} < 0,$

where $\mathbf{a}_{\mathbf{k}} = \left(\prod_{j=1}^n \frac{\alpha_{\mathbf{k}} + \mathbf{w}_j}{\mathbf{w}_j} \right) \left(\prod_{j=1, j \neq \mathbf{k}}^{n+1} \frac{-\alpha_j}{\alpha_{\mathbf{k}} - \alpha_j} \right) (-\alpha_{\mathbf{k}})$

For $y \geq \max(x, 0)$,

$$\begin{aligned} & f_{X(\tau), M(\tau)}(x, y) \\ &= f_{M(\tau), X(\tau)-M(\tau)}(y, x-y) \times 1 \\ &= f_{M(\tau)}(y) f_{X(\tau)-M(\tau)}(x-y) \\ &= f_{M(\tau)}(y) f_{m(\tau)}(x-y) \\ &= \left(\sum_{k=1}^{m+1} b_k e^{-\beta_k y} \right) \left(\sum_{j=1}^{n+1} a_j e^{-\alpha_j (x-y)} \right) \\ &= \sum_{k=1}^{m+1} \sum_{j=1}^{n+1} a_j b_k \times e^{-\alpha_j x} e^{-(\beta_k - \alpha_j) y} \end{aligned}$$

Further Work

1. Use SOA CAE research grant to hire graduate students to estimate the parameters and to program the formulas.
2. Binomial tree version.

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