BAYESIAN GRADUATION OF FHA/HUD SINGLE FAMILY HOME MORTGAGE INSURANCE CONTRACTS -- SECTION 221, 222, 223(e), and 235

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Mortgage insurance is used to guarantee the mortgage investor (mortgagee) against financial loss due to the inability of the home buyer (mortgagor) to repay the loan in accordance with the terms of the mortgage. In this paper we will describe a procedure used to construct predicted survivorship tables for single-family home mortgage insurance contracts written under four sections of the National Housing Act. These sections are 221, 222, 223(e), and 235. All of the mortgages we will be concerned with have 30-year terms.

In an earlier paper (Herzog [1981]), we constructed a survivorship table for Section 203 single-family home mortgage insurance contracts using a procedure known as "Bayesian graduation." In our earlier paper, the prior distributions for claim and non-claim terminations were based solely on Section 203 data. Here we will show how we used the results of the Section 203 Bayesian graduation procedure together with the observed data for the particular section under consideration to construct the appropriate prior distribution. The completion of the Bayesian graduation process is then largely mechanical. The Section 203 survivorship table constitutes Table l; the two sets of predicted Section 203 central decrement rates are exhibited in Table 2.

To simplify the discussion, we will be concerned primarily with the data of Section 222. After we have described the procedure for the construction of the predicted survivorship table for Section 222, we will indicate how that procedure can be modified to construct survivorship tables for sections 221 , 223(e), and 235.

The paper consists of seven sections. In the first section, we define the two types of terminations considered. Next, we present a brief overview of the Bayesian graduation process and then discuss the tabulation and adjustment procedures. In the fourth and fifth sections we specify the prior vectors and correlation matrices needed to produce the posterior (i.e., predicted) claim and non-claim termination rates, respectively. Next, we show how these predicted rates can be used to produce the desired survivorship table for Section 222. Finally, we discuss how the procedure can be modified to produce survivorship tables for Sections 221, 223(e), and 235.

In our construction process, we only considered those Section 222 policies endorsed between January 1, 1963 and December 31, 1980. Thus, we have made the implicit, but fundamental, assumption that the data (i.e, the policies) under consideration could be used (without adjustment) to predict future experience. The observed data of Section 222 are exhibited in Table 3.

1. CLAIM AND NON-CLAIM TERMINATIONS

For the purpose of constructing a survivorship table, there ace essentially two ways in which a Section 222 mortgage can cease generating premiums:
(1) as the result of a claim termination where the mortgagor defaults on the loan and the mortgagee submits an insurance claim to HUD,
(2) a non-claim termination or withdrawal--usually where the mortgagor prepays the loan in full.

In the first case, claim termination can occur through a foreclosure by the lender (or deed in lieu of foreclosure where HUD acquires title to the property) or an assignment of the loan to HUD. As indicated above, in the second case, the majority of non-claim terminations are due to the full prepayment of the loan by the mortgagor. However, there are a number of other ways this can occur. These include:
(1) voluntary termination of the insurance whereby the mortgagor and mortgagee agree to terminate the insurance without termination of the loan,
(2) default termination where the value of the security (house) is greater than the insurance benefits and the mortgagee does not apply for insurance benefits, or
(3) cancellation due to fraud or other illegal action.

Non-claim terminations other than full prepayment historically constitute less than 3 percent of all non-claim terminations.

Sometimes these loans are partially prepaid. Since the FHA insurance premiums are based on the scheduled outstanding balance at the time of endorsement, we have not considered partial prepayments in our construction process.

For the purpose of constructing a survivorship table, we are concerned with predicting two sets of probabilities. One is a schedule of annual claim termination probabilities, $q_{x}^{c}$; the second is a schedule of annual non-claim termination probabilities, $q_{x}^{n}$, where the time frame, $x$, runs from the origination of the loan to the end of the term of the loan. The unit of exposure is taken to be the individual insurance policy; i.e., each policy is given equal weight in the construction process.
2. AN OVERVIEW OF THE CONSTRUCTION PROCESS

We have used a general procedure known as Bayesian graduation (see Kilmeldorf and Jones [1967] and Hickman and Miller [1977]) to construct the Section 222 survivorship table. Bayesian graduation enables us to incorporate our prior subjective notions (about the salient aspects of our data) in a unified and formal framework, and results in a posterior distribution of the predicted central decrement rates.

Before proceeding, we will review some of the notation of the earlier papers:
u is the (random) vector of observed central mortality rates.
$B$ is the covariance matrix of $\underline{u}$.
$\underline{W}$ is the (random) vector of variables (parameters in the traditional or frequentist sense) for which a
smooth estimate is sought.

A is the covariance matrix of $W_{0}$
$\underline{m}$ is the (prior) mean vector of $W$.
$t(x)=\arcsin \sqrt{x .}$

The procedure is based on the following assumptions:

1. The transformed random vector, $t(\underline{u})$, of observed mortality rates, $\underline{u}$, has a multinormal distribution with covariance matrix $t(B)$.
2. The transformed prior distribution of $W$ is multinormal with mean $t(m)$ and covariance matrix $t(A)$.

As a result of these assumptions, it can be shown that the posterior distribution of $t(\underline{W})$ is multinormal with mean

$$
\text { (*) }\left(t(B)^{-1}+t(A)^{-1}\right)^{-1}\left(t(B)^{-1} t(\underline{u})+t^{(A)^{-1}} t(\underline{m})\right)
$$

and covariance matrix

$$
\left(t(B)^{-1}+t(A)^{-1}\right)^{-1}
$$

The inverse of the function $t$ (i.e., the square of the sine) of expression (*) is used as the set of predicted values.

We have employed the arcsin transformation here in order to "stabilize the variance" in the sense discussed on page 12 of Hickman and Miler [1977]. (This should become clearer later in the discussion.) In addition, the arsin transformation frequently results in a distribution that is closer to the normal distribution than is that of the pre-transformed variable.

In our application, we will perform two Bayesian graduations--one for the claim terminations and one for the nonclaim terminations.

Unfortunately, unlike life insurance, mortgage insurance must be considered to be a fundamental $1 /$ (or economic) risk rather than a particular (or personal) risk. This is because the ability of individual homeowners to continue to meet their monthly mortgage obligations is highly dependent on the economy of the United States as well as its effect on local economic conditions. Hence, a group of mortgage insurance policies must be treated as being interdependent rather than independent. As a result, one of the basic assumptions underlying the Bayesian graduation model--that the policies may be assumed to be stochastically independent--is violated. Consequently, Bayesian graduation does not produce here a good model of the posterior distribution of either the claim or non-claim terminations. It is, however, not at all clear how to model the interdependent structure since this largely depends on U. S. economic conditions. We have, therefore, reduced the scope of our problem to that of estimating the average central decrement rates. For this purpose, Bayesian graduation does provide a useful framework and we have used the means of the posterior distribution as our sought-after average central rates of decrement.

[^0]Before specifying the mean vectors and covariance arrays just discussed, we need to discuss the observed data. We begin with the following notation:

$\begin{aligned} C_{x}^{j}= & \text { the number of policies written in year } j \text { which result } \\ & \text { in a claim termination between age } x \text { and } x+1 .\end{aligned}$
$N_{x}^{j}=$ the number of policies written in year $j$ which result in a non-claim termination between age $x$ and $x+1$.

Since not all policies endorsed during calendar year 1980 were reported to and/or processed by HUD on or before December 31, 1980, we had to adjust, the value of $E_{0}^{1980}$.

Because an analysis of recent reporting lags indicated that $\mathrm{E}_{0}^{1980}$ was about 7 percent too low, we decided to replace $E_{0}^{1980}$ by ( $1.07 \times E_{0}^{1980}$ ).

Because, on the average, those policies still in force on December 31,1980 , have only completed about half of their most recent policy year and there is frequently a substantial delay (of up to two years in some cases) in the reporting and/or processing of terminations, we decided (1) to replace $C_{x}^{1980-x}$ and $N_{x}^{1980-x}$ for $x=1, \ldots, 13$ by ( $3.0 x C_{x}^{1980-x}$ ) and ( $3.0 \times \mathrm{N}_{\mathrm{x}}^{1980-\mathrm{x}}$ ), respectively, and (2) to replace
$C_{x}^{1979-x}$ and $N_{x}^{1979-x}$ for $x=1, \ldots, 12$ by
(1.2 $\mathrm{xC}_{\mathrm{x}}^{1979-\mathrm{x}}$ ) and (1.2 $\mathrm{x}_{\mathrm{x}}^{1979-\mathrm{x}}$ ), respectively. Again, these are just ratio adjustments based on reporting lags observed in the recent past.

Having made these adjustments, we may now define more encompassing classes of exposure, claim terminations and nonclaim terminations by combining the data across endorsement years as follows:

$$
\begin{aligned}
& E_{x}=\sum_{j=1963}^{1980-x} E_{x}^{j} \\
& c_{x}=\sum_{j=1963}^{1980-x} c_{x}^{j} \\
& N_{x}=\sum_{j=1963}^{1980-x} N_{x}^{j}
\end{aligned}
$$

where $x=0,1, \ldots, 13$.
Thus, as defined above, $\mathrm{E}_{0}$, for example, represents the total number of policies written between January 1, 1963 and December 31, 1980. Unfortunately, because of a computer system problem, there were no data available on the last four policy years of mortgages written during the period 1957-1969.
4. CLAIM TERMINATIONS

### 4.1 Observed Central Rates

Assuming that terminations are distributed uniformly throughout each policy year, we may approximate the first 14 observed central rates of claim termination by

$$
u_{x}=C_{x} /\left(E_{x}-.5\left(C_{x}+N_{x}\right)\right)
$$

```
for x = 0,1,...,13.
```

We define $t(B)^{-1}$, the inverse of the arcsin transformation of the corresponding covariance matrix, to be the 30 -by- 30 diagonal matrix in which the $i-t h$ diagonal is $4 E_{i}$ for $i=$ $0,1, \ldots ., 13$ and 0 for $i=14, \ldots, 29$. Since the weight assigned to each of the last sixteen policy years is zero, the choice of $u_{x}$ for $x \geq 14$ is irrelevant. Because the components of $\underline{u}$ are assumed to be mutually independent, given $\underline{W}=\underline{w}$, the off-diagonal elements of $B$ (and hence $t(B)^{-1}$ ) are all set equal to zero.

### 4.2 Prior Distribution

Since we want the graduated claim termination rates, $m_{x}$, to equal the observed rates for the first seven policy years, we define

$$
m_{x}=u_{x}
$$

$$
\text { for } x=0,1, \ldots, 6
$$

The remaining prior central claim termination rates were obtained by multiplying the means of the posterior distribution of the central claim termination rates ( $u_{x}^{*}$ ) of Section 203 by the ratio of

$$
\frac{\sum_{x=7}^{13} u_{x}}{\sum_{x=7}^{13} u_{x}^{*}}
$$

This is in the spirit of maximum likelihood estimation and ensures that the average prior central rate of claim termination (for Section 222) for policy years $8 \mathbf{- 1 4}$ will be equal to that observed. Moreover, the mean vector of the prior distribution
(of Section 222 claim terminations) from policy year 8 and beyond is assumed to have the shape of the mean vector of the posterior claim distribution of Section 203.

The arcsin transformation, $t(A)$, of the corresponding correlation matrix is $t(A)^{*} / 4 N$ where $t(A)^{*}=\left[p^{|i-j|}\right]$ and where $\left[p^{|i-j|}\right]$ is a 30 -by -30 matrix in which the element in the $i-t h$ row and $j$-th column is equal to $p^{|i-j|}=(.99)^{|i-j|}$. The matrix $t(A)$ is constructed in the spirit of Hickman and Miller [1977]. The value of $p$ was chosen to be .99 to equal the observed value of the serial correlation of the central claim termination rates for policy years 2-19 of Section 203 loans. $N$ is chosen to be the median of the last seven non-zero exposure values, i.e., $\mathrm{N}=\mathrm{E}_{10^{\circ}}$

The vector of posterior claim termination values, vc, is the square of the sine of

$$
\left(t(B)^{-1}+t(A)^{-1}\right)^{-1}\left(t(B)^{-1} t(\underline{u})+t(A)^{-1} t(\underline{m})\right)
$$

where $\underline{u}=\left(u_{0}, u_{1}, \ldots, u_{29}\right)$ and $\underline{m}=\left(m_{0}, m_{1}, \ldots, m_{29}\right)$. The observed, prior, and posterior claim termination rates are exhibited in Table 4.
5. NON-CLAIM TERMINATIONS

### 5.1 Observed Central Rates

We approximate the central rate of non-claim termination by

$$
\begin{aligned}
& u_{x}=N_{x} /\left(E_{x}-.5\left(C_{x}+N_{x}\right)\right) \quad \text { for } x=0,1, \ldots, 13 \\
& \text { The matrix } t(B)^{-1} \text { is defined as in section } 4 .
\end{aligned}
$$

### 5.2 Prior Non-Claim Termination Rates

Since we want each graduated non-claim termination rate, $m_{x}$, to equal the corresponding observed rate for the first five policy years, we set

$$
\mathrm{m}_{\mathrm{x}}=\mathrm{u}_{\mathrm{x}} \quad \text { for } \mathrm{x}=0,1, \ldots, 4
$$

The remaining prior central non-claim termination rates were obtained by multiplying the mean vector of the posterior distribution of the central non-claim termination rates ( $u_{x}^{*}$ ) of Section 203 by the ratio of


The reasons for this adjustment are those given in the previous section.

### 5.3 Correlation Matrix

The arcsin transformation, $t(A)$, of the corresponding correlation matrix is $t(A) * / 4 N$ where

$$
t(A) *=\left(\begin{array}{lc}
I & 0_{5,25} \\
0_{25,5} & p^{|i-j|}
\end{array}\right)
$$

and where $I$ is the 5 by 5 identity matrix, $0_{5,25}$ and $0_{25,5}$ are, respectively, 5-by-25 and $25-$ by-5 matrices composed entirely of
zeros, and $p^{|i-j|}$ is a 25-by-25 matrix in which the element in the i-th row and j-th column (of the matrix $t(A)^{*}$ ) is equal to $p^{|i-j|}=(.9)^{|i-j|}$.

Thus, $\mathrm{t}(\mathrm{A})^{*}$ is constructed so that (1) each of the first five graduated values is set equal to the corresponding prior (i.e, observed) value, and (2) each of the graduated values for policy years $6-29$ is influenced to some extent by each of the prior values for policy years 6-29. In the next section, we will show how we determined the non-claim termination rate for the final policy year. Since the observed serial correlation among the initial 19 Section 203 central non-claim termination rates is 0.90 , we set $\mathrm{p}=0.90$; N is as defined in Section 4 .

The vector of graduated non-claim termination values, vn, is the square of the sine of

$$
\left(t(B)^{-1}+t(A)^{-1}\right)^{-1}\left(t(B)^{-1} t(\underline{u})+t(A)^{-1} t(\underline{m})\right)
$$

where

$$
\underline{u}=\left(u_{0}, u_{1}, \ldots, u_{29}\right) \quad \text { and } \quad \underline{m}=\left(m_{0}, m_{1}, \ldots m_{2 g}\right) .
$$

The observed, prior, and posterior (predicted) non-claim termination rates of Section 222 are exhibited in Table 5.
6. CONSTRUCTING THE SURVIVORSHIP SCHEDULE

Using the posterior means (i.e., the predicted values), vc and vn , determined above, we are able to approximate the corresponding vectors of termination probabilities, $q^{c}$ and $q^{n}$, by
$q_{x}^{c}=v c_{x} /\left(1+(.5) v t_{x}\right)$
and for $x=0,1, \ldots, 29$
$q_{x}^{n}=v n_{x} /\left(1+(.5) v t_{x}\right)$
where $v t_{x}=v c_{x}+v n_{x}$. The equations above are derived on pages 273-274 of Jordan [1967].

Before proceeding, we need to introduce the following notation: $S_{x}=$ the number of survivors at the beginning of policy year $x$.

We first set the radix at $100,000-\mathrm{i} . \mathrm{e} ., \mathrm{S}_{0}=100,000$.
We next define

$$
\begin{array}{ll}
c_{x} & =s_{x} q_{x}^{c} \\
N_{x} & =s_{x} q_{x}^{n} \\
s_{x+1} & =s_{x}-\left(N_{x}+c_{x}\right)
\end{array}
$$

for $x=0, \ldots, 28$.

Note that the last three equations represent a recursive procedure in which $S_{x}$ must be computed before $C_{x}$ and $N_{x}$ and in which $S_{x+1}$ is computed after $C_{x}$ and $N_{X}$.

Finally, we define
$C_{29}=S_{29} q_{29}^{c}$ and $\quad N_{29}=S_{29}\left(1-q_{29}^{c}\right)$
so that $\mathbf{S}_{\mathbf{3 0}}=0$.
The resulting survivorship schedule is shown in Table 6 .
7. SECTIONS 221, 223(e), and 235

Because $H U D$ began collecting data for actuarial purposes on both 221 and 222 mortgage insurance contracts in 1963, the procedure applied to the Section 221 data is precisely that applied to the 222 data. Since 1969 was the first year in which substantial amounts of Section $223(e)$ and 235 mortgage insurance contracts were written, we have only considered those 223(e) and 235 mortgages written on or after January 1, 1969. Hence, we must make two minor changes in the procedure applied to those sections. First, for the claim terminations only, we set the first five (rather than seven) prior central rates equal to those observed. Second, since we only had eleven exposure values, we set $N=\left(E_{8}+E_{9}\right) / 2-$ i.e., the median of the last six exposure values.

The observed data for these sections as well as the resulting survivorship tables are shown in Tables 7-18.

## TABLE 1

SURVIVORSHIP AND DECREMENT TABLE AS OF DECEMBER 31. 1980 based on aggregate insurance and termination experience FOR HOME MORTGAGES INSURED SINCE 1957

SECTION 203
30 YEAR TERM
U.S. TOTALS

|  | SURVIV ORS | TERMINATIONS |  |
| :---: | :---: | :---: | :---: |
| POLICY | beginning of |  | NON- |
| YEAR | POLICY YEAR | CLAIM | CLAIM |
| 1 | 100000.0 | 413.1 | 706.9 |
| 2 | 98880.1 | 1445.2 | 2262.8 |
| 3 | 95172.1 | 1352.3 | 3729.6 |
| 4 | 90090.2 | 1042.3 | 4082.8 |
| 5 | 84965.0 | 776.6 | 4349.5 |
| 6 | 79838.9 | 567.9 | 4595.4 |
| 7 | 74675.5 | 415.4 | 4475.2 |
| 8 | 69785.0 | 282.4 | 4311.2 |
| 9 | 65191.3 | 183.0 | 3908.4 |
| 10 | 61099.9 | 119.3 | 3494.6 |
| 11 | 57486.0 | 73.7 | 3414.6 |
| 12 | 53997.7 | 44.9 | 3204.7 |
| 13 | 50748.1 | 26.8 | 3006.5 |
| 14 | 47714.7 | 15.7 | 2819.5 |
| 15 | 44879.5 | 9.1 | 2644.2 |
| 16 | 42226.2 | 5.2 | 2681.9 |
| 17 | 39539.1 | 2.9 | 2695.1 |
| 18 | 36841.2 | 1.6 | 2684.8 |
| 19 | 34154.8 | . 8 | 2651.1 |
| 20 | 31502.8 | . 5 | 2593.8 |
| 21 | 28908.5 | . 3 | 2515.9 |
| 22 | 26392.3 | . 2 | 2420.3 |
| 23 | 23971.9 | . 1 | 2309.9 |
| 24 | 21661.9 | . 1 | 2187.7 |
| 25 | 19474.2 | . 0 | 2056.5 |
| 26 | 17417.6 | . 0 | 1919.3 |
| 27 | 15498.2 | . 0 | 1778.7 |
| 28 | 13719.5 | .0 | 1637.0 |
| 29 | 12082.5 | . 0 | 2092.7 |
| 30 | 9989.8 | . 0 | 9989.8 |

[^1]
## TABLE 2

PREDICTED CENTRAL TERMINATION RATES FOR SECTION 203 (MULTIPLIED BY 100.000)

| POLICY <br> YEAR | CLAIMS | NONCLAIMS |
| :---: | ---: | :---: |
| 1 |  |  |
| 2 | 415 | 711 |
| 3 | 1490 | 2332 |
| 4 | 1460 | 4026 |
| 5 | 1191 | 4665 |
| 6 | 942 | 5278 |
| 7 | 735 | 5948 |
| 8 | 575 | 6196 |
| 9 | 419 | 6388 |
| 10 | 290 | 6189 |
|  | 201 | 5894 |
| 11 | 132 | 6126 |
| 12 | 86 | 6119 |
| 13 | 54 | 6107 |
| 14 | 34 | 6090 |
| 15 | 21 | 6071 |
| 16 | 13 | 6560 |
| 17 | 8 | 7057 |
| 18 | 4 | 7563 |
| 19 | 3 | 8076 |
| 20 | 2 | 8587 |
| 21 | 1 |  |
| 22 | 1 | 9099 |
| 23 | 0 | 9611 |
| 24 | 0 | 10124 |
| 25 | 0 | 10636 |
| 26 | 0 | 11149 |
| 27 | 0 | 11662 |
| 28 | 0 | 12175 |
| 29 | 0 | 12689 |
| 30 |  | 18962 |
|  | 6067 |  |
|  |  |  |

TABLE 3

DATA USED TO CONSTUCT SECTION 222 SURVIVORSHIP SCHEDULE (ROUNDED TO NEAREST INTEGER)

| POLICY ANINIVEKSAKY | ${ }_{X}^{E}$ | $c_{X}$ | ${ }_{X}^{N}$ |
| :---: | :---: | :---: | :---: |
| 1 | 151790 | 664 | 964 |
| 2 | 149468 | 2375 | 4028 |
| 3 | 140556 | 2232 | 6187 |
| 4 | 129136 | 1559 | 6264 |
| 5 | 119155 | 928 | 6136 |
| 6 | 109892 | 472 | 6116 |
| 7 | 101240 | 280 | 6151 |
| 8 | 93367 | 168 | 6118 |
| 9 | 80289 | 92 | 5292 |
| 10 | 63978 | 64 | 3942 |
| 11 | 48176 | 27 | 3221 |
| 12 | 32936 | 17 | 2246 |
| 13 | 18837 | 6 | 1304 |
| 14 | 7801 | 3 | 692 |

## TABLE 4

CLALM TERMINATION RATES ROR SECTION 222 (MULTIPLIED BY 100,000)

| $\begin{gathered} \text { POLICY } \\ Y E A R \end{gathered}$ | OBSERVED | PRIOK | PREDICTED |
| :---: | :---: | :---: | :---: |
| 1 | 440 | 440 | 438 |
| 2 | 1623 | 1623 | 1620 |
| 3 | 1637 | 1637 | 1633 |
| 4 | 1245 | 1245 | 1241 |
| 5 | 803 | 803 | 799 |
| 6 | 442 | 442 | 439 |
| 7 | 286 | 286 | 283 |
| 8 | 186 | 204 | 201 |
| 9 | 119 | 141 | 139 |
| 10 | 104 | 98 | 96 |
| 11 | 58 | 65 | 63 |
| 12 | 53 | 42 | 41 |
| 13 | 33 | 27 | 26 |
| 14 | 40 | 17 | 16 |
| 15 |  | 10 | 10 |
| 16 |  | 6 | 6 |
| 17 |  | 4 | 4 |
| 18 |  | 2 | 2 |
| 19 |  | 1 | 1 |
| 20 |  | 1 | 1 |
| 21 |  | 0 | 0 |
| 22 |  | 0 | 0 |
| 23 |  | 0 | 0 |
| 24 |  | 0 | 0 |
| 25 |  | 0 | 0 |
| 26 |  | 0 | 0 |
| 27 |  | 0 | 0 |
| 28 |  | 0 | 0 |
| 29 |  | 0 | 0 |
| 30 |  | 0 | 0 |

TABLE 5
NON-CLALM TERMINATION RATES FOR SECTION 222
(MULTIPLIED BY 100,000)

| $\begin{gathered} \text { POLICY } \\ \text { YEAR } \end{gathered}$ | OBSERVED | PRIOR | PREDICTED |
| :---: | :---: | :---: | :---: |
| 1 | 638 | 638 | 638 |
| 2 | 2754 | 2754 | 2754 |
| 3 | 4538 | 4538 | 4538 |
| 4 | 5002 | 5002 | 5002 |
| 5 | 5307 | 5307 | 5307 |
| 6 | 5737 | 6743 | 6112 |
| 7 | 6275 | 7023 | 6451 |
| 8 | 6781 | 7241 | 6787 |
| 9 | 6820 | 7016 | 6691 |
| 10 | 6360 | 6681 | 6447 |
| 11 | 6919 | 6944 | 6807 |
| 12 | 7062 | 6936 | 6883 |
| 13 | 7172 | 6923 | 6933 |
| 14 | 9284 | 6903 | 6963 |
| 15 |  | 6882 | 6936 |
| 16 |  | 7436 | 7486 |
| 17 |  | 7999 | 8046 |
| 18 |  | 8573 | 8617 |
| 19 |  | 9154 | 9194 |
| 20 |  | 9734 | 9771 |
| 21 |  | 10314 | 10348 |
| 22 |  | 10895 | 10926 |
| 23 |  | 11476 | 11505 |
| 24 |  | 12057 | 12083 |
| 25 |  | 12638 | 12662 |
| 26 |  | 13220 | 13242 |
| 27 |  | 13802 | 13822 |
| 28 |  | 14384 | 14402 |
| 29 |  | 21495 | 21515 |
| 30 |  | 6878 | 6889 |

TABLE 6

SUKVIVORSHIP AND DECKENENT TABLE AS OF DECEMBER 31, 1980 BASED ON AGGREGATE INSURANCE AND TEFMINATION EXPERIENCE FOR HOME MORTGAGES INSURED SINCE 1963

SECTION 222
30 YEAK TERM
U.S.TOTALS

|  | SURV IVORS | TERMINATIONS |  |
| :---: | :---: | :---: | :---: |
|  |  | DUR | Y YEAR |
| POLICY | BEGINNING OF |  | NON- |
| YEAK | POLICY YEAR | CLAIM | CLAIM |
| 1 | 100000.0 | 436.0 | 635.0 |
| 2 | 98929.0 | 1568.6 | 2666.1 |
| 3 | 94694.4 | 1500.2 | 4168.5 |
| 4 | 89025.7 | 1071.2 | 4318.3 |
| 5 | 83636.2 | 648.4 | 4307.0 |
| 6 | 78680.8 | 334.5 | 4656.2 |
| 7 | 73690.0 | 201.4 | 4598.8 |
| 8 | 68889.8 | 133.7 | 4517.6 |
| 9 | 64238.5 | 86.1 | 4156.2 |
| 10 | 59996.2 | 56.0 | 3745.3 |
| 11 | 56194.9 | 34.4 | 3698.3 |
| 12 | 52462.2 | 20.9 | 3490.1 |
| 13 | 48951.2 | 12.4 | 3279.6 |
| 14 | 45659.2 | 7.2 | 3071.9 |
| 15 | 42580.1 | 4.1 | 2854.1 |
| 16 | 39721.9 | 2.3 | 2866.2 |
| 17 | 36853.4 | 1.3 | 2850.4 |
| 18 | 34001.7 | . 7 | 2808.7 |
| 19 | 31192.3 | . 4 | 2741.8 |
| 20 | 28450.1 | . 2 | 2650.3 |
| 21 | 25799.6 | . 1 | 2538.5 |
| 22 | 23261.0 | . 1 | 2409.9 |
| 23 | 20851.1 | . 0 | 2268.3 |
| 24 | 18582.7 | . 0 | 2117.5 |
| 25 | 16465.2 | . 0 | 1960.8 |
| 26 | 14504.4 | . 0 | 1801.4 |
| 27 | 12703.0 | . 0 | 1642.3 |
| 28 | 11060.7 | . 0 | 1486.0 |
| 29 | -9574.7 | . 0 | 1859.9 |
| 30 | 7714.8 | . 0 | 7714.8 |

[^2]
## TABLE 7

DATA USED TO CONSTUCT SECTION 221 SURVIVORSHIP SCHEDULE (ROUNDED TO NEAREST INTEGER)

| POLICY |  |  |  |
| :---: | ---: | ---: | ---: |
| ANNIVEKSARY | $X$ | $C$ | $N$ |
|  | $X$ |  |  |
|  |  |  |  |
| 1 | 740140 | 7878 | 2825 |
| 2 | 718083 | 31233 | 11918 |
| 3 | 653868 | 24689 | 21290 |
| 4 | 584763 | 15658 | 25073 |
| 5 | 521396 | 10392 | 26044 |
| 6 | 469123 | 7321 | 26893 |
| 7 | 420537 | 5453 | 25361 |
| 8 | 373672 | 4282 | 22500 |
| 9 | 287392 | 2791 | 16232 |
| 10 | 206883 | 1667 | 10024 |
| 11 | 131803 | 688 | 7139 |
| 12 | 73778 | 193 | 4353 |
| 13 | 39633 | 93 | 2476 |
| 14 | 14889 | 23 | 1067 |

TABLE 8
CLADM TEFMINATION RATES FOR SECTION 221
(MULTIPLIED BY 100,000)

| $\begin{gathered} P O L I C Y \\ Y E A R \end{gathered}$ | OBSERVED | PRIOR | PREDICTED |
| :---: | :---: | :---: | :---: |
| 1 | 1072 | 1072 | 1068 |
| 2 | 4484 | 4484 | 4475 |
| 3 | 3913 | 3913 | 3903 |
| 4 | 2774 | 2774 | 2763 |
| 5 | 2065 | 2065 | 2053 |
| 6 | 1620 | 1620 | 1605 |
| 7 | 1346 | 1346 | 1329 |
| 8 | 1188 | 1457 | 1433 |
| 9 | 1005 | 1009 | 997 |
| 10 | 829 | 700 | 697 |
| 11 | 538 | 461 | 459 |
| 12 | 270 | 298 | 298 |
| 13 | 243 | 190 | 189 |
| 14 | 160 | 118 | 118 |
| 15 |  | 73 | 73 |
| 16 |  | 44 | 44 |
| 17 |  | 26 | 26 |
| 18 |  | 15 | 15 |
| 19 |  | 9 | 9 |
| 20 |  | 5 | 5 |
| 21 |  | 3 | 3 |
| 22 |  | 2 | 2 |
| 23 |  | 1 | 1 |
| 24 |  | 1 | 1 |
| 25 |  | 1 | 1 |
| 26 |  | 1 | 1 |
| 27 |  | 1 | 1 |
| 28 |  | 0 | 0 |
| 29 |  | 0 | 0 |
| 30 |  | 0 | 0 |

## TABLE 9

NON-CLAIM TEFMINATION RATES FOR SECTION 221
(MULTIPLIED BY 100,000)

| POLICY |  |  |  |
| :---: | :---: | :---: | :---: |
| YEAR | OBSERVED | PRIOR | PREDICTED |
| 1 | 384 | 384 | 384 |
| 2 | 1711 | 1711 | 1711 |
| 3 | 3375 | 3375 | 3375 |
| 4 | 4442 | 4442 | 4442 |
| 5 | 5176 | 5176 | 5176 |
| 6 | 5950 | 5926 | 5899 |
| 7 | 6260 | 6172 | 6117 |
| 8 | 6245 | 6364 | 6217 |
| 9 | 5841 | 6166 | 5921 |
| 10 | 4986 | 5871 | 5555 |
| 11 | 5582 | 6102 | 5823 |
| 12 | 6088 | 6096 | 5890 |
| 13 | 6457 | 6084 | 5934 |
| 14 | 7439 | 6067 | 5952 |
| 15 |  | 6048 | 5945 |
| 16 |  | 6535 | 6438 |
| 17 |  | 7030 | 6940 |
| 18 |  | 7535 | 7451 |
| 19 |  | 8045 | 7957 |
| 20 |  | 8554 | 8483 |
| 21 |  | 9064 | 8998 |
| 22 |  | 9575 | 9513 |
| 23 |  | 10085 | 10028 |
| 24 |  | 10596 | 10544 |
| 25 |  | 11107 | 11059 |
| 26 |  | 11618 | 11574 |
| 27 |  | 12129 | 12089 |
| 28 |  | 12641 | 12604 |
| 29 |  | 18890 | 18851 |
| 30 |  | 6044 | 6023 |

SURVIVORSHIP AND DECREMENT TABLE AS OF DECEMBER 31, 1980 based on aggregate insurance and tekmination experience FOR HOME MORTGAGES INSURED SINCE 1963

SECTION 221
30 YEAR TEFM
U.S.TOTALS

|  |  | TERMINATIONS |  |
| :---: | :---: | ---: | ---: |
| POLICY | SURVIVORS <br> BEGINNING OF | DURING POLICY YEAR |  |

ULTIMATE CLAIM TERMINATION RATE
ULTIMATE NON-CLAIM TERMINATION RATE ESTIMATED LIFE EXPECTANCY
17.16 PERCENT
82.84 PERCENT
13.02 YEARS

## TABLE 11

DATA USED TO CONSTUCT SECTION 223(E) SURVIVOFSHIP SCHEDULE (ROUNDED TO NEAKEST INTEGER)

| POLICY ANNIVERSARY | $E_{X}$ | ${ }^{C}$ | ${ }^{N}$ |
| :---: | :---: | :---: | :---: |
| 1 | 153508 | 1791 | 764 |
| 2 | 147411 | 8887 | 2153 |
| 3 | 130882 | 7950 | 3109 |
| 4 | 114632 | 5844 | 2652 |
| 5 | 100292 | 3976 | 2392 |
| 6 | 89425 | 3236 | 2563 |
| 7 | 80033 | 2476 | 2495 |
| 8 | 70242 | 2056 | 2647 |
| 9 | 35692 | 1094 | 1494 |
| 10 | 25109 | 749 | 899 |
| 11 | 13526 | 465 | 897 |

TABLE 12
CLAIM TERMINATION RATES FOR SECTION 223(E) (MULTIPLIED BY 100,000)

| $\begin{gathered} \text { POLICY } \\ \text { YEAR } \end{gathered}$ | OBSEEVED | PRIOR | PREDICTED |
| :---: | :---: | :---: | :---: |
| 1 | 1176 | 1176 | 1078 |
| 2 | 6264 | 6264 | 6026 |
| 3 | 6342 | 6342 | 6074 |
| 4 | 5295 | 5295 | 5011 |
| 5 | 4094 | 4094 | 3799 |
| 6 | 3740 | 6203 | 5775 |
| 7 | 3193 | 4852 | 4480 |
| 8 | 3028 | 3531 | 3256 |
| 9 | 3181 | 2446 | 2257 |
| 10 | 3083 | 1697 | 1565 |
| 11 | 3620 | 1116 | 1020 |
| 12 |  | 723 | 646 |
| 13 |  | 459 | 399 |
| 14 |  | 287 | 240 |
| 15 |  | 176 | 140 |
| 16 |  | 107 | 79 |
| 17 |  | 63 | 43 |
| 18 |  | 37 | 22 |
| 19 |  | 22 | 11 |
| 20 |  | 13 | 5 |
| 21 |  | 8 | 2 |
| 22 |  | 5 | 1 |
| 23 |  | 3 | 0 |
| 24 |  | 2 | 0 |
| 25 |  | 2 | 0 |
| 26 |  | 2 | 0 |
| 27 |  | 1 | 0 |
| 28 |  | 1 | 0 |
| 29 |  | 1 | 0 |
| 30 |  | 1 | 0 |

## TABLE 13

NON-CLAIM TERMINATION RATES FOR SECTION 223(E) (MULTIPLIED BY 100,000)

| POLICY |  |  |  |
| :---: | :---: | :---: | :---: |
| YEAR | OBSERVED | PRIOR | PREDICTED |
| 1 | 502 | 502 | 502 |
| 2 | 1517 | 1517 | 1517 |
| 3 | 2480 | 2480 | 2480 |
| 4 | 2402 | 2402 | 2402 |
| 5 | 2464 | 2464 | 2464 |
| 6 | 2962 | 4065 | 3353 |
| 7 | 3217 | 4234 | 3577 |
| 8 | 3899 | 4365 | 3885 |
| 9 | 4344 | 4229 | 3935 |
| 10 | 3701 | 4027 | 3858 |
| 11 | 6983 | 4186 | 4150 |
| 12 |  | 4181 | 4149 |
| 13 |  | 4173 | 4144 |
| 14 |  | 4161 | 4135 |
| 15 |  | 4149 | 4125 |
| 16 |  | 4483 | 4461 |
| 17 |  | 4822 | 4802 |
| 18 |  | 5168 | 5149 |
| 19 |  | 5518 | 5501 |
| 20 |  | 5868 | 5852 |
| 21 |  | 6218 | 6203 |
| 22 |  | 6568 | 6554 |
| 23 |  | 6918 | 6905 |
| 24 |  | 7268 | 7256 |
| 25 |  | 7619 | 7608 |
| 26 |  | 7969 | 7959 |
| 27 |  | 8320 | 8311 |
| 28 |  | 8671 | 8662 |
| 29 |  | 12958 | 12949 |
| 30 |  | 4146 | 4141 |

SURVIVORSHIP AND DECREMENT TPABLE AS OF DECEMBER 31, 1980 based on aggregate insurance and termination experience FOR HOME MORTGAGES INSURED SINCE 1969

SECTION 223(E)<br>30 YEAR TERM<br>U.S.totals

|  | SURVIVORS | TERMINATIONS |  |
| :---: | :---: | :---: | :---: |
| POLICY | beginning of |  | NON- |
| YEAR | POLICY YEAR | CLAIM | CLAIM |
| 1 | 100000.0 | 1070.0 | 497.9 |
| 2 | 98432.0 | 5715.6 | 1439.0 |
| 3 | 91277.4 | 5316.9 | 2170.7 |
| 4 | 83789.8 | 4048.5 | 1941.0 |
| 5 | 77800.3 | 2866.1 | 1858.5 |
| 6 | 73075.7 | 4036.2 | 2343.3 |
| 7 | 66696.2 | 2872.5 | 2293.5 |
| 8 | 61530.1 | 1934.3 | 2307.9 |
| 9 | 57287.9 | 1254.3 | 2186.5 |
| 10 | 53847.0 | 820.5 | 2022.6 |
| 11 | 51004.0 | 506.9 | 2063.4 |
| 12 | 48433.7 | 305.6 | 1962.5 |
| 13 | 46165.6 | 180.1 | 1870.7 |
| 14 | 44114.8 | 103.6 | 1785.3 |
| 15 | 42225.9 | 57.9 | 1705.6 |
| 16 | 40462.4 | 31.4 | 1764.9 |
| 17 | 38666.1 | 16.2 | 1812.8 |
| 18 | 36837.1 | 8.0 | 1849.1 |
| 19 | 34980.1 | 3.7 | 1872.6 |
| 20 | 33103.8 | 1.6 | 1882.0 |
| 21 | 31220.2 | . 6 | 1878.2 |
| 22 | 29341.4 | . 2 | 1861.9 |
| 23 | 27479.2 | . 1 | 1834.1 |
| 24 | 25645.1 | . 0 | 1795.7 |
| 25 | 23849.3 | . 0 | 1747.9 |
| 26 | 22101.5 | . 0 | 1691.8 |
| 27 | 20409.7 | . 0 | 1628.5 |
| 28 | 18781.2 | . 0 | 1559.3 |
| 29 | 17221.8 | . 0 | 2094.4 |
| 30 | 15127.4 | . 0 | 15127.4 |

ultimate claim termination rate
ULTIMATE NON-CLAIM TERMINATION RATE ESTIMATED LIFE EXPECTANCY

## TABLE 15

data used to constuct section 235 SURVIVORSHIP SChedule (ROUNDED TO NEAREST INTEGER)

| POLICY |  |  |  |
| ---: | ---: | ---: | ---: |
| ANNIVERSARY | $\boldsymbol{E}$ | $\boldsymbol{C}$ | $\boldsymbol{N}$ |
|  |  | $\boldsymbol{X}$ | $\boldsymbol{X}$ |
|  |  |  |  |
| 2 | 456598 | 11329 | 672 |
| 3 | 444591 | 25298 | 2848 |
| 4 | 416415 | 20843 | 7370 |
| 5 | 388193 | 12992 | 11822 |
| 6 | 363367 | 7913 | 17776 |
| 7 | 337604 | 4932 | 22566 |
| 8 | 306949 | 3177 | 23622 |
| 9 | 272920 | 2294 | 20387 |
| 10 | 204334 | 1274 | 13210 |
| 11 | 126686 | 627 | 6002 |
|  | 48813 | 282 | 2862 |

TABLE 16

CLAIM TENMINATION RATES FOR SECTION 235 (MULTIPLIED BY 100,000)

| $\begin{gathered} \text { POLICY } \\ \text { YEAR } \end{gathered}$ | OBSERVED | PRIOR | PREDICTED |
| :---: | :---: | :---: | :---: |
| 1 | 2514 | 2514 | 2491 |
| 2 | 5876 | 5876 | 5840 |
| 3 | 5181 | 5181 | 5144 |
| 4 | 3457 | 3457 | 3423 |
| 5 | 2257 | 2257 | 2226 |
| 6 | 1523 | 1636 | 1605 |
| 7 | 1082 | 1280 | 1251 |
| 8 | 877 | 931 | 909 |
| 9 | 646 | 645 | 630 |
| 10 | 508 | 448 | 437 |
| 11 | 597 | 294 | 287 |
| 12 |  | 191 | 185 |
| 13 |  | 121 | 116 |
| 14 |  | 76 | 72 |
| 15 |  | 46 | 44 |
| 16 |  | 28 | 26 |
| 17 |  | 17 | 15 |
| 18 |  | 10 | 9 |
| 19 |  | 6 | 5 |
| 20 |  | 3 | 3 |
| 21 |  | 2 | 2 |
| 22 |  | 1 | 1 |
| 23 |  | 1 | 1 |
| 24 |  | 1 | 0 |
| 25 |  | 0 | 0 |
| 26 |  | 0 | 0 |
| 27 |  | 0 | 0 |
| 28 |  | 0 | 0 |
| 29 |  | 0 | 0 |
| 30 |  | 0 | 0 |

TABLE 17
NON-CLAIM TEFMINATION RATES FOR SECTION 235
(MULTIPLIED BY 100,000)

| POLICY |  |  |  |
| :---: | :---: | :---: | :---: |
| YEAR | OBSERVED | PRIOR | PREDICTED |
| 1 | 149 | 149 | 149 |
| 2 | 662 | 662 | 662 |
| 3 | 1832 | 1832 | 1832 |
| 4 | 3146 | 3146 | 3146 |
| 5 | 5071 | 5071 | 5071 |
| 6 | 6968 | 6546 | 6967 |
| 7 | 8047 | 6819 | 7296 |
| 8 | 7794 | 7030 | 7367 |
| 9 | 6703 | 6812 | 6905 |
| 10 | 4865 | 6486 | 6384 |
| 11 | 6058 | 6741 | 6625 |
| 12 |  | 6734 | 6630 |
| 13 |  | 6721 | 6627 |
| 14 |  | 6702 | 6618 |
| 15 |  | 6682 | 6606 |
| 16 |  | 7219 | 7149 |
| 17 |  | 7766 | 7700 |
| 18 |  | 8324 | 8262 |
| 19 |  | 8887 | 8831 |
| 20 |  | 9450 | 9398 |
| 21 |  | 10014 | 9965 |
| 22 |  | 10577 | 10532 |
| 23 |  | 11141 | 11100 |
| 24 |  | 11705 | 11667 |
| 25 |  | 12270 | 12235 |
| 26 |  | 12834 | 12802 |
| 27 |  | 13399 | 13370 |
| 28 |  | 13964 | 13938 |
| 29 |  | 20868 | 20840 |
| 30 |  | 6677 | 6662 |

SURVIVORSHIP AND DECREMENT TABLE AS OF DECENBER 31. 1980 bASED ON AGGREGATE INSURANCE AND TERMINATION EXPERIENCE FOR HOME MORIGAGES INSURED SINCE 1969

> SECTION 235
> 30 YEAR TERM
> U.S.TOTALS

|  |  | TERMINATIONS |  |
| :---: | :---: | :---: | :---: |
|  | SURVIVORS | DURING POLICY YEAR |  |
| POLICY | BEGINNING OF | NON- |  |
| YEAR | POLICY YEAR | CLAIM |  |


| 1 | 100000.0 | 2458.6 | 147.2 |
| ---: | ---: | ---: | ---: |
| 2 | 97394.2 | 5508.5 | 624.0 |
| 3 | 91261.7 | 4535.9 | 1615.4 |
| 4 | 85110.3 | 2820.9 | 2592.4 |
| 5 | 79697.1 | 1711.7 | 3899.5 |
| 6 | 74086.0 | 1140.2 | 4949.7 |
| 7 | 67996.1 | 815.7 | 4757.6 |
| 8 | 62422.8 | 545.1 | 4416.1 |
| 9 | 57461.6 | 348.6 | 3823.7 |
| 10 | 53289.3 | 225.1 | 3289.9 |
| 11 | 49774.2 | 138.0 | 3187.6 |
| 12 | 46448.6 | 43.0 | 2977.9 |
| 13 | 43387.7 | 28.9 | 2781.5 |
| 14 | 40557.4 | 16.0 | 2597.1 |
| 15 | 37932.1 | 8.9 | 2425.1 |
| 16 | 35491.0 | 4.8 | 2449.3 |
| 17 | 33032.9 | 2.5 | 2449.2 |
| 18 | 30578.9 | 1.3 | 2380.2 |
| 19 | 28150.2 | .7 | 2312.6 |
| 20 | 25768.2 | .3 | 2226.3 |
| 21 | 23454.7 | .2 | 2124.0 |
| 22 | 21228.0 | .1 | 2009.0 |
| 23 | 19103.8 | .1 | 1884.6 |
| 24 | 17094.7 | .0 | 1753.7 |
| 25 | 15210.1 | .0 | 1619.1 |
| 26 | 13456.4 | .0 | 1483.5 |
| 27 | 11837.3 | .0 | 1349.1 |
| 28 | 10353.8 | .0 | 1699.5 |
| 29 | -9004.7 | .0 | 7305.2 |

[^3]20.44 PERCENT
79.56 PERCENT
12.38 YEARS

## REFERENCES

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Kulp, C. A. and Hall, J. L., [1963] Casualty Insurance, Fourth Edition, The Ronald Press Co.


[^0]:    1/ The terminology employed here appears in Kulp and Hall
    [1968, Chapter 1]. Also, there is a brief discussion of mortgage insurance on page 69 of Bickelhaupt [1974].

[^1]:    ULTIMATE CLAIM TERMINATION RATE
    ULTIMATE NON-CLAIM TERMINATION RATE
    ESTIMATED LIfE EXPECTANCY

[^2]:    ULTIMATE CLAIM TERMINATION RATE
    ULTIMATE NON-CLAIM TERMINATION RATE ESTIMAI'ED LIFE EXPECTANCY

    > 6.12 PERCENT
    > 93.88 PERCENT
    > 13.38 YEARS

[^3]:    ULTIMATE CLAIM TERMINATION RATE
    ULTIMATE NON-CLAIM TERMINATION RATE
    ESTIMATED LIFE EXPECTANCY

